

# Sub-Gaussian random variable and its properties

Shengming Luo

6 Sep 2017

## 0.1 Recap

Last time we've defined  $X \in SG(\sigma^2)$  if  $\mathbb{E}[e^{\lambda(X - \mathbb{E}(X))}] \leq e^{\lambda^2 \sigma^2 / 2}$ , where the RHS is mgf of  $N(0, \sigma^2)$ .

We're also given the tail bound as  $\mathbb{P}(|X - \mu| \geq t) \leq 2 \exp\{-\frac{t^2}{2\sigma^2}\}$ . Some other properties of  $X \in SG(\sigma^2)$ .

- $V[X] \leq \sigma^2$
- $a \leq X - \mu \leq b$ , then  $X \in SG((\frac{b-a}{2})^2)$
- $X \in SG(\sigma^2) \implies aX \in SG(a^2 \sigma^2)$
- $X \in SG(\sigma^2)$  and  $Y \in SG(\tau^2)$ , then  $X + Y \in SG((\sigma + \tau)^2)$ . Additionally, if  $X$  and  $Y$  are independent,  $X + Y \in SG(\sigma^2 + \tau^2)$ , which is a sharper bound.

Question: let  $\sigma(X) = \inf\{\sigma > 0 : \mathbb{E}[e^{\lambda(X - \mathbb{E}X)}]\} \leq e^{\frac{\lambda^2 \sigma^2}{2}}\}$ , is  $V[X] = \sigma^2$ ? Not in general. Let

$$X = \begin{cases} 1, & \text{with prob } \frac{1-p}{2}, \\ 0, & \text{with prob } p, \\ -1, & \text{with prob } \frac{1-p}{2}, \end{cases}$$

Then,  $\mathbb{E}[e^{\lambda X}] > \exp\{\frac{\lambda^2 \mathbb{E}X^2}{2}\}$  for all  $\lambda$  small enough.

Proof of property (4): W.L.O.G.,  $\mathbb{E}[X] = \mathbb{E}[Y] = 0$ . Then,  $\forall \lambda \in \mathbb{R}$ ,

$$\mathbb{E}[e^{\lambda(X+Y)}] \leq (\mathbb{E}[e^{\lambda p X}])^{\frac{1}{p}} (\mathbb{E}[e^{\lambda q Y}])^{\frac{1}{q}} = \exp\left\{\frac{\lambda^2 \sigma^2 p^2}{2} \frac{1}{p} + \frac{\lambda^2 \tau^2 q^2}{2} \frac{1}{q}\right\},$$

where the inequality is Holder's inequality,  $\frac{1}{p} + \frac{1}{q} = 1$ .

Minimize RHS over all pairs of  $p, q$ , we get, by Cauchy inequality,

$$\frac{\lambda^2 \sigma^2 p}{2} + \frac{\lambda^2 \sigma^2 q}{2} = \left(\frac{\lambda^2 \sigma^2 p}{2} + \frac{\lambda^2 \tau^2 q}{2}\right) \left(\frac{1}{p} + \frac{1}{q}\right) \geq \frac{\lambda^2 (\sigma + \tau)^2}{2}.$$

## 1 Hoeffding inequality

Let  $X_1, \dots, X_n$  be independent with  $\mu_i = \mathbb{E}[X_i]$ ,  $X_i \in SG(\sigma_i^2)$  for all  $1 \leq i \leq n$ . Then,

$$\mathbb{P}\left(\left|\frac{\sum (X_i - \mu_i)}{n}\right| \geq t\right) \leq 2 \exp\left\{-\frac{n^2 t^2}{2 \sum_{i=1}^n \sigma_i^2}\right\}.$$

If  $\sigma_i^2 = \sigma^2$ , then RHS becomes  $\leq 2 \exp\{-\frac{nt^2}{2\sigma^2}\}$ .

Example:

Let  $X_1, \dots, X_n$  be independent *Bernoulli*( $p_i$ ), where  $\pi_i \in (0, 1)$ . Then  $X_i \in SG(\frac{1}{4})$ .  
 As a result,  $\mathbb{P}(|\bar{X} - \bar{p}_n| \geq t) \leq 2 \exp\{-2t^2n\}$ . Let  $\text{RHS}=\delta$  and solve for  $t$ , we get  $|\bar{X} - \bar{p}_n| \leq \sqrt{\frac{1}{2n} \log(\frac{1}{\delta})}$  with prob at least  $1 - \delta$ .  
 Let  $\delta = n^{-c}$ , then  $|\bar{X} - \bar{p}_n| = O(\sqrt{\frac{\log(n)}{n}})$  with prob at least  $1 - n^{-c}$ .

### 1.1 Is Hoeffding the sharpest bound?

The answer is No! For Bernoulli, Chernoff bound gives sharper result.

Consider following multiplicative bounds,

$$\mathbb{P}(\sum X_i \geq (1 + \epsilon) \sum p_i) \leq \begin{cases} \exp\left\{-\frac{\epsilon^2 \mu}{3}\right\}, & \epsilon \in (0, 1), \\ \exp\left\{-\frac{\epsilon^2 \mu}{2+\mu}\right\}, & \epsilon > 1, \end{cases}$$

Hoeffding vs Multiplicative bound:

Suppose  $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bern}(p)$ , then Hoeffding bound would give us:

$$\mathbb{P}(p - \frac{\sum X_i}{n} \geq t) \leq \exp\{-2nt^2\} \implies p - \bar{X}_n \leq \sqrt{\frac{1}{2n} \log(1/\delta)}, \text{ with probs } \geq 1 - \delta.$$

However, using multiplicative bound,

$$\mathbb{P}(p - \frac{\sum X_i}{n} \geq t) \leq \exp\{-np\epsilon^2/2\} \implies p - \bar{X}_n \leq \sqrt{\frac{2p}{n} \log(1/\delta)}, \text{ with probs } \geq 1 - \delta.$$

If  $p \leq 1/4$ , multiplicative bound is sharper. If  $p \rightarrow 0$ , multiplicative bounds is much better (e.g. See paper minimax classification with optimal decision tree.)

## 2 Equivalent characterizations of SubGaussian random variables

Assuming  $\mathbb{E}X = 0$ , we've told the definition of Sub gaussian random variable is 1)  $\mathbb{E}[e^{\lambda X}] \leq e^{\frac{\lambda^2 \sigma^2}{2}}, \forall \lambda \in \mathbb{R}$ .

Equivalent definitions: 2) there exists  $c > 0$  and  $Z \in N(0, \tau^2)$ , s.t.

$$\mathbb{P}(|X| > t) \leq c\mathbb{P}(|Z| \geq \tau).$$

$$3) \mathbb{E}[e^{\frac{\lambda X^2}{2\sigma^2}}] \leq \frac{1}{\sqrt{1-\lambda}}, \text{ where } \lambda \in (0, 1).$$

$$4) \mathbb{E}[e^{aX^2}] \leq 2 \text{ for some } a > 0.$$

$$5) \text{ Finally, if } X \in SG(\sigma^2) \text{ and } \forall p > 0,$$

$$\mathbb{E}[|X|^p] \leq p2^{p/2}\sigma^p\Gamma(p/2) \leq c\sigma\sqrt{p},$$

where  $\Gamma(x) = \int_0^\infty e^{-\mu}\mu^{x-1}d\mu$ .

Proof of (5):

$$\begin{aligned} \mathbb{E}[|X|^p] &= \int_0^\infty \mathbb{P}(|X|^p \geq t) dt \\ &\leq 2 \int_0^\infty \exp\left\{-\frac{t^{2/p}}{2\sigma^2}\right\} dt = (2\sigma^2)^{p/2} p\Gamma(p/2). \end{aligned}$$

### 3 Sub-Exponential random variables

$X \sim \text{Laplacian}(b), b > 0$ . Pdf of  $X$  is  $p(x) = \frac{b}{2}e^{-|x|b}$ , i.e. the tail probability does not decay as fast as sub-gaussian random variables (thicker tails). In fact, the MGF of  $X$  equals  $\frac{1}{1-\lambda^2 b^2}$ , where  $|\lambda| < \frac{1}{|b|}$ .

Thus, we might need consider a another class of random variables with slower tail probability decay rate. That motivates us to the definition of sub exponential random variables.

Definition: Suppose  $v, \alpha > 0$ , then  $X \in SE(v^2, \alpha)$  when  $\mathbb{E}[e^{\lambda(X - \mathbb{E}[X])}] \leq e^{\frac{\lambda^2 v^2}{2}}$  if  $|\lambda| < \frac{1}{\alpha}$ .

If  $X \in SG(\sigma^2)$ , we have  $X \in SE(v^2, 0)$ .

Example: If  $Z \sim N(0, 1)$ ,  $X = Z^2 \sim \mathcal{X}_1^2$ .  $\mathbb{E}[X] = 1$ .

$$\begin{aligned}\mathbb{E}[e^{\lambda(X-1)}] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{+\infty} e^{\lambda(z^2-1)} e^{-\frac{z^2}{2}} dz \\ &= \frac{1}{\sqrt{2\pi}} e^{-\lambda} \int_{-\infty}^{+\infty} e^{\frac{z^2}{2}(1-2\lambda)} dz \\ &= \frac{e^{-\lambda}}{\sqrt{1-2\lambda}} \leq \exp\{\frac{\lambda^2}{1-2\lambda}\},\end{aligned}$$

where the last inequality is due to  $-\log(1-u) - u \leq \frac{u^2}{1-u}$

Hence,  $\mathbb{E}[e^{\lambda(X-1)}] \leq \exp\{\frac{\lambda^2}{1-2\lambda}\} \leq \exp\{\frac{4\lambda^2}{2}\} \implies Z^2 \sim SE(4, 4) \implies \mathbb{P}(Z^2 - 1 \geq 2t + 2\sqrt{t}) \leq e^{-t}, \forall t > 0$ .

Properties:

- 1) If  $X \in SG(\sigma^2)$ , then  $X^2 \in SE(256\sigma^2, 16\sigma)$ .
- 2) If  $V[X] = \sigma^2$  and  $|X - \mathbb{E}[X]| \leq b, a.s.$  Then  $X \in SE(2\sigma^2, 2b)$ .