## 36-755, Fall 2017 Homework 4

Due Wed Nov 1 by 5:00pm in Jisu's mailbox

1. In earlier works on the lasso, people have used a even stronger assumptions than the restricted eigenvalue property. Here is one. Suppose that the design matrix X is such that, for some integer k > 0,

$$\max_{i,j} \left| \frac{X_i^{\top} X_j}{n} - 1(i=j) \right| \le \frac{1}{23k} \tag{1}$$

where  $X_i$  is the *i*th column of X, i = 1, ..., d. Think about what that means.

(a) Show that this condition implies that, for any subset S of  $\{1, \ldots, d\}$  of cardinality no larger than k < d and any  $\Delta \in \mathbb{R}^d$  with  $\|\Delta_{S^c}\|_1 \leq 3\|\Delta_S\|_1$ ,

$$\|\Delta\|^2 \le \frac{2}{n} \|X\Delta\|^2.$$

That is, show that this condition implies the RE(3,1/2) condition given in class for all nonempty subsets S of  $\{1,\ldots,d\}$  of size no larger than k. Instead of the constant 23 you may take a larger one if it simplifies your calculations.

(b) Suppose that the entries of X are now populated by independent Rademacher variables (a Ra detacher variable is one that takes the values +1 and -1 with equal probability). Show that, for any  $\delta \in (0,1)$ , if

$$n \ge Ck^2(\log(d) + \log(1/\delta)),$$

for some constant C > 0, then X satisfies the condition (1), with probability at least  $1 - \delta$ . Again, instead of 23 feel free to show the result for a different constant if it helps with the calculations.

- 2. Read the paper "p-Values for High-Dimensional Regression" by Nicolai Meinshausen, Lukas Meier and Peter Bühlmann, JASA 2009, volume 104, issue 448, pages 1671-1681. Reproduce the proof of Theorem 3.2.
- 3. **Performance of the best selection procedure.** Consider the regression framework  $Y = X\beta^* + \epsilon$ , where X is a  $n \times d$  deterministic design matrix, and  $\epsilon$  a vector in  $\mathbb{R}^n$  of independent  $SG(\sigma^2)$  error variables. Assume that the true regression coefficient  $\beta^*$  belongs to the set  $S_0(k) = \{x \in \mathbb{R}^d : , ||x||_0 = k\}$  of k-sparse vectors, where  $k \leq d$ . Consider the estimator

$$\hat{\beta} = \operatorname{argmin}_{\beta \in B_0(k)} ||Y - X\beta||^2.$$

This the best least squares solution computed over all subsets of the coordinates of size k. Computationally, it requires evaluating  $\binom{d}{k}$  least squares. Analyze the performance of  $\hat{\beta}$  by showing that, with probability at lest  $1 - \delta$ 

$$||X(\hat{\beta} - \beta^*)||^2 \le C(\delta) \frac{\sigma^2 k}{n} \log\left(\frac{ed}{2k}\right),$$

where  $C(\delta)$  is a constant that depend on  $\delta$ . Notice that, up to a logarithmic term, this is the (optimal) performance of the least squares estimator if the support of  $\beta^*$  were known. This is something that

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is quite typical: the statistical price for not knowing the support of of  $\beta^*$  is only logarithmic (and therefore rather minimal). However, at least for the estimator  $\hat{\beta}$ , the computational price is huge. The trade-off between computational and statistical guarantees in a very important topic in the theoretical literature on high-dimensional statistics.

Hint: follow the proof of the performance of the least squares estimator. You may want to use the fact that  $\binom{d}{k} \leq \left(\frac{ed}{k}\right)^k$ .

## 4. Matrix Algebra Problems.

- (a) Problem 8.3 (You may assume the result of Problem 8.1 as given).
- (b) Recall the spiked covariance model:  $\Sigma = \theta v v^{\top} + I_d$ , where  $\theta > 0$  and  $v \in \mathbb{S}^{d-1}$ . Let  $\hat{v}$  be another unit vector in  $\mathbb{S}^{d-1}$ . Show that

$$v^{\mathsf{T}} \Sigma v - \hat{v}^{\mathsf{T}} \Sigma \hat{v} = \theta \sin^2(\angle(v, \hat{v}))$$

where  $\angle(v, \hat{v}) = \cos^{-1}(|v^{\top}\hat{v}|)$ 

(c) Show that

$$\left\| \hat{v}\hat{v}^{\top} - vv^{\top} \right\|_F^2 = 2\sin^2(\angle(v, \hat{v})),$$

where, for a matrix  $A = (A_{i,j}), \|A\|_F = \sqrt{\sum_{i,j} A_{i,j}^2}$ .