

SDS 387 Linear Models

Fall 2025

Lecture 16 - Tue, Oct 23, 2025

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- HW3 is out. For the delta method question you need to invoke Lemma 2.12 in van der Vaart's book.

Let $R: \mathbb{R}^d \rightarrow \mathbb{R}$ s.t. $R(0) = 0$. Let $\{X_n\} \subset \mathbb{R}^d$ be a sequence of random vectors s.t. $X_n \xrightarrow{P} 0$.

Then, $\forall p > 0$

- $\forall \{h_n\}$
- i) if $R(h_n) = o(\|h_n\|^p)$ then $R(X_n) = o_p(\|X_n\|^p)$
 - ii) if $R(h_n) = O(\|h_n\|^p)$ then $R(X_n) = O_p(\|X_n\|^p)$

- Let's finish the proof of L_2 projection.

Thm 11.1 (of Van der Vaart) \hat{S} is the projection of T onto S

iff i) $\hat{S} \in S$ and ii) $\mathbb{E}[(T - \hat{S})S] = 0$

orthogonality $\rightarrow \langle T - \hat{S}, S \rangle = 0 \quad \forall S \in S$ ①

This projection is unique (in the sense if \hat{S}' is another projection then $\mathbb{P}(\hat{S} \neq \hat{S}') = 0$)

If S contains the constant functions then

$$\mathbb{E}[\hat{S}] = \mathbb{E}[T] \quad \text{and} \quad \text{cov}(T - \hat{S}, S) = 0 \quad \forall S \in S$$

where T and S are r.v.'s in L_2 and S is a vector space

PP/ Last time we show that if orthogonality holds then \hat{S} is the unique projection.

Suppose \hat{S} is a projection (that is

$$\hat{S} \in \arg \min_{S \in S} \mathbb{E}[(S - T)^2])$$

Then $\forall \alpha \in \mathbb{R} \quad \forall S \in S$

$$0 \leq \mathbb{E}[(T - \underbrace{\hat{S} - \alpha S}_{\in S})^2] = \mathbb{E}[(T - \hat{S})^2]$$

$$= \alpha^2 \mathbb{E}[S^2] - 2\alpha \mathbb{E}[(T - \hat{S})S]$$

This is a parabola in α that has to stay above the α -axis. The zeros of this parabola are

$$\alpha = 0 \quad \text{and} \quad \alpha = 2 \frac{\mathbb{E}[(T - \hat{S})S]}{\mathbb{E}[S^2]}$$

\hookrightarrow therefore $\mathbb{E}[(T - \hat{S})S] = 0$ and

(2)

since S is generic, the orthogonality condition

is satisfied.

If S contains the constant functions then, by orthogonality

$$\mathbb{E}[(T - \hat{S}) \cdot c] = 0 \quad \forall c \in \mathbb{R}$$

$$\hookrightarrow \mathbb{E}[T] = \mathbb{E}[\hat{S}]$$

Corollary Pythagora theorem for r.v.'s:

$$\mathbb{E}[T^2] = \mathbb{E}[\hat{S}^2] + \mathbb{E}[(T - \hat{S})^2]$$

Writing $\|X\|_2^2 = \mathbb{E}[X^2]$, this gives us

$$\|T\|^2 = \|\hat{S}\|^2 + \|T - \hat{S}\|^2$$

direct sum decomposition

- The most important type of L_2 projection is the conditional expectation. Suppose Y and X are L_2 random variables and let

$$S = \left\{ f(X), f \text{ is arbitrary st. } \mathbb{E}[f(X)^2] < \infty \right\}$$

I want to find the function g st. $\mathbb{E}[g(X)^2] < \infty$

$$\text{and } \mathbb{E}[(Y - g(X))^2] \leq \mathbb{E}[(Y - f(X))^2]$$

for all $f(X) \in \mathcal{S}$.

Then $g(X) = E[Y|X]$. This is true because
 $X \mapsto E[Y|X]$ satisfies the orthogonality condition

$$E[(Y - E[Y|X]) f(X)] = E[Y f(X)] - E[E[Y|X] f(X)]$$

\downarrow
any $f(\cdot)$ s.t. $E[f^2(X)] < \infty$ by law of iterated expectation

$$\left(E[E[Y|X]^2] \leq E[E[Y^2|X]] = E[Y^2] < \infty \right)$$

by conditional Jensen

- Remark: Conditional expectation is well-defined even without a second moment. In this case the defining condition is

$$E[Y 1_{\{X \in A\}}] = E[E[Y|X] 1_{\{X \in A\}}]$$

all set A .

LINEAR REGRESSION

General regression settings:

- response or dependent variable Y univariate r.v. of interest
- covariates X random vector in \mathbb{R}^d
- independent variable
- features
- explanatory variables
- Our goal is to "model" or "learn" Y using X

- Assuming Y and X have finite second moments ($\mathbb{E}[Y^2] < \infty$ and

$$\begin{aligned} \Sigma &= \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T] \\ &= \mathbb{E}[XX^T] - \mathbb{E}[X](\mathbb{E}[X])^T \end{aligned}$$

exist) this can be cast as the

problem of finding or modeling a function

$f: \mathbb{R}^d \rightarrow \mathbb{R}$ s.t.

$$\mathbb{E}[(Y - f(X))^2] \text{ is "small"}$$

- Hierarchy of modeling assumption

Agnostic or model free approach:

$$Y = \underbrace{\mathbb{E}[Y|X]}_{\substack{\downarrow \\ \text{best approximation} \\ \text{of } Y \text{ using } X}} + \underbrace{(Y - \mathbb{E}[Y|X])}_{\substack{\varepsilon \quad \text{error}}}$$

The function $x \in \mathbb{R}^d \mapsto \mathbb{E}[Y|X=x]$ is the regression function. So letting $f^{\text{opt}}(x) = \mathbb{E}[Y|X]$

$$Y = \underbrace{f^{\text{opt}}(x)}_{\text{signal}} + \underbrace{\varepsilon}_{\text{error}}$$

Properties: i) $\mathbb{E}[\varepsilon] = \mathbb{E}[(Y - \mathbb{E}[Y|X])] = 0$

\downarrow
natural desirable property of ε

ii) $\mathbb{E}[f^{\text{opt}}(x) \varepsilon] = \text{cov}[f^{\text{opt}}(x) \varepsilon] = 0$

by orthogonality

\hookrightarrow Remark This does not mean $\varepsilon \perp f^{\text{opt}}(x)$

- Signal + independent noise assumption non-parametric regression

$$Y = f(X) + \varepsilon \quad \text{where} \quad \mathbb{E}[\varepsilon] = 0 \\ \varepsilon \perp X$$

where f belongs to (large) class of functions satisfying certain properties. Typically these functions are assumed to be "smooth".

- Parametric regression

$$Y = f(X) + \varepsilon \quad \mathbb{E}[\varepsilon] = 0 \\ X \perp \varepsilon = 0$$

where f is a function belonging to a parametric class, i.e. each f in this class is

parametrized by a parameter $\theta \in \Theta \subseteq \mathbb{R}^n$

so regression function is f_θ some $\theta \in \Theta$

- Linear regression:

$$Y = f_\theta^*(X) + \varepsilon \quad \mathbb{E}[\varepsilon] = 0 \quad X \perp \varepsilon$$

where $f_{\theta^*}(x) = x^T \theta^*$, with

θ^* unknown but in a set $\Theta \subseteq \mathbb{R}^d$

Remark: to allow for an intercept term, we assume that the first coordinate of x is a constant, say 1. In this case

$$f_{\theta^*}(x) = \theta_1^* + \sum_{j=2}^d \theta_j^* x_j$$

- We usually need to have an intercept in your model, to make sense of ANOVA decomposition and R^2 statistic.

- Mis-specified modeling: You are in a agnostic setting

Fit a parametric model and carry out inference on some well-defined "projection parameters"

- The fixed- x setting: the covariates are deterministic

- 2 tasks:
 - 1) Prediction: we want to predict a new instance of response variable, say y_{new} , using a new instance

of the covariates X_{new} .

We observe X_{new} and would
like to predict Y_{new}

2) statistical inference on β^* the true parameter
indexing the regression function, or the
projection parameter if the model is
mis-specified