

Lecture 13: February 26

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13.1 Preliminaries

Definition 13.1 (Operator Norm of a Symmetric Matrix) Let $\mathbf{A} \in S^{d \times d}$ be a symmetric $d \times d$ matrix. Then the operator norm of \mathbf{A} , denoted $\|\mathbf{A}\|_{op}$ is defined as

$$\|\mathbf{A}\|_{op} := \max_{x \in \mathbb{S}^{d-1}} |x^T \mathbf{A} x| \quad (13.1)$$

Definition 13.2 (Sub-Gaussian Random Vector) A vector $\mathbf{x} \in \mathbb{R}^d$ is vector sub-gaussian ($\mathbf{x} \in SG_d(\sigma^2)$) with parameter σ if for all $\mathbf{v} \in \mathbb{S}^{d-1}$

$$\mathbb{E}[\exp(\mathbf{v}^T (\mathbf{x} - \mu))] \leq \exp(\lambda^2 \sigma^2 / 2) \quad (13.2)$$

13.2 Concentration Inequalities for Covariance Matrices

Theorem 13.3 Let $\mathbf{x}_1, \dots, \mathbf{x}_n$ be an i.i.d sequence of σ sub-gaussian random vectors such that $\mathbb{V}[\mathbf{x}_1] = \Sigma$ and let $\hat{\Sigma}_n := \frac{1}{n} \sum_{i=1}^n x_i x_i^T$ be the empirical covariance matrix. Then there exists a universal constant $C > 0$ such that, for $\delta \in (0, 1)$, with probability at least $1 - \delta$

$$\frac{\|\hat{\Sigma}_n - \Sigma\|_{op}}{\sigma^2} \leq C \max \left\{ \sqrt{\frac{d + \log(2/\delta)}{n}}, \frac{d + \log(2/\delta)}{n} \right\} \quad (13.3)$$

If $d/n \rightarrow 0$ then the confidence interval goes to 0 at a rate of $\sqrt{d/n}$ which is the minimax rate

Proof: We break the proof up into two steps:

1. Use a discretization argument to reduce the problem to task of computing the maximum of finitely many random variables
2. Use standard concentration inequalities

Step 1:

Lemma 13.4 Let $\mathbf{A} \in S^{d \times d}$ and let N_ϵ be an ϵ -net of \mathbb{S}^{d-1} . Then

$$\|\mathbf{A}\|_{op} \leq \frac{1}{1 - 2\epsilon} \max_{y \in N_\epsilon} |y^T \mathbf{A} y| \quad (13.4)$$

Proof of Lemma 13.4:

Let $y \in N_\epsilon$ satisfy $\|x - y\| \leq \epsilon$. Then

$$|x^T \mathbf{A} x - y^T \mathbf{A} y| = |x^T \mathbf{A} (x - y) + y^T \mathbf{A} (x - y)| \quad (13.5)$$

$$\leq |x^T \mathbf{A} (x - y)| + |y^T \mathbf{A} (x - y)| \quad (13.6)$$

Looking at $|x^T \mathbf{A} (x - y)|$ we have

$$|x^T \mathbf{A} (x - y)| \leq \|\mathbf{A} (x - y)\| \|x\| \quad (13.7)$$

$$\leq \|\mathbf{A}\|_{op} \underbrace{\|x - y\|}_{\leq \epsilon} \underbrace{\|x\|}_{=1} \quad (13.8)$$

$$\leq \|\mathbf{A}\|_{op} \epsilon \quad (13.9)$$

Applying the same argument to $|y^T \mathbf{A} (x - y)|$ gives us $|x^T \mathbf{A} x - y^T \mathbf{A} y| \leq 2\epsilon \|\mathbf{A}\|_{op}$. To complete the proof, we see that $\|\mathbf{A}\|_{op} = \max_{x \in \mathbb{S}^{d-1}} x^T \mathbf{A} x \leq 2\epsilon \|\mathbf{A}\|_{op} + \max_{y \in N_\epsilon} y^T \mathbf{A} y$. Rearranging the equation gives $\|\mathbf{A}\|_{op} \leq \frac{1}{1-2\epsilon} \max_{y \in N_\epsilon} y^T \mathbf{A} y$ as desired.

Step 2:

Applying Lemma 13.4 on $\hat{\Sigma}_n - \Sigma$ with $\epsilon = 1/4$ we have

$$\|\hat{\Sigma}_n - \Sigma\|_{op} \leq 2 \max_{v \in N_{1/4}} |v^T (\hat{\Sigma}_n - \Sigma) v| \quad (13.10)$$

Additionally, we know that $N_{1/4} \leq 9^d$. From here, we can apply standard concentration tools as follows:

$$\mathbb{P}(\|\hat{\Sigma}_n - \Sigma\|_{op} \geq t) \leq \mathbb{P}(\max_{v \in N_{1/4}} |v^T (\hat{\Sigma}_n - \Sigma) v| \geq t/2) \quad (13.11)$$

$$\leq |N_{1/4}| \mathbb{P}(|v_i^T (\hat{\Sigma}_n - \Sigma) v_i| \geq t/2) \quad (13.12)$$

We rewrite $v_i^T (\hat{\Sigma}_n - \Sigma) v_i$ as follows:

$$v_i^T (\hat{\Sigma}_n - \Sigma) v_i = \frac{1}{n} \sum_{j=1}^n (v_i^T x_j)^2 - \mathbb{E}[(v_i^T x_j)^2] \quad (13.13)$$

$$= \frac{1}{n} \sum_{j=1}^n z_j - \mathbb{E}[z_j] \quad (13.14)$$

where z_j 's are independent and by assumption $v_i^T x_j \in SG(\sigma^2)$ so that $z_j - \mathbb{E}[z_j] \in SE((16\sigma^2)^2, 16\sigma^2)$. Applying the sub-exponential tail bound gives us

$$\mathbb{P}(|v_i^T (\hat{\Sigma}_n - \Sigma) v_i| \geq t/2) \leq 2 \exp \left\{ -\frac{n}{2} \min \left\{ \left(\frac{t}{32\sigma^2} \right)^2, \frac{t}{32\sigma^2} \right\} \right\} \quad (13.15)$$

so that

$$\mathbb{P}(\|\hat{\Sigma}_n - \Sigma\|_{op} \geq t) \leq 2 \cdot 9^d \exp \left\{ -\frac{n}{2} \min \left\{ \left(\frac{t}{32\sigma^2} \right)^2, \frac{t}{32\sigma^2} \right\} \right\} \quad (13.16)$$

Inverting the bound gives the desired result ■

13.3 Matrix Concentration Inequalities

Theorem 13.5 (Matrix Bernstein) *Let X_1, \dots, X_n be independent mean 0 symmetric $d \times d$ random matrices such that $\|X_i\|_{op} \leq C$ almost surely. Then for any $t \geq 0$*

$$P(\|\sum X_i\| > t) \leq 2d \exp(-\frac{t^2}{2(\sigma^2 + Ct/3)}) \quad (13.17)$$

where $\sigma^2 = \|\sum \mathbb{E}[X_i]\|_{op}$

Some applications of matrix concentration inequalities include:

- Solving Linear Systems
- Matrix Multiplication
- Sub Sampling
- Sparsification methods for spectral clustering
- Dimensionality Reduction
- Compressed Sensing
- Network Models

For a more in-depth discussion on these topics, refer to [tropp2012user, tropp2015introduction]

References

[tropp2012user] J. TROPP, “User-friendly tail bounds for sums of random matrices,”

[tropp2015introduction] J. TROPP, “An introduction to matrix concentration inequalities,”