

**36-755, Fall 2017**  
**Homework 5 Solution**

Due Wed Nov 15 by 5:00pm in Jisu's mailbox

**Points:** 100+1 pts total for the assignment.

We first review some basic relations with norms and the singular value decomposition on matrices.

**Lemma 0.1** For any matrix  $A \in \mathbb{R}^{m \times n}$ , let  $A = U\Sigma V^\top$  be its singular value decomposition with  $U^\top U = UU^\top = I_m$  and  $V^\top V = VV^\top = I_n$ , and  $\Sigma$  being  $m \times n$  diagonal matrix with nonnegative diagonal values  $\sigma_1, \dots, \sigma_{\min(m,n)}$ .

(i) Frobenius norm can be calculated as  $\|A\|_F = \sqrt{\sum_{i=1}^{\min(m,n)} \sigma_i^2}$ .

(ii) Operator 2-norm  $\|A\|_{op}$  can be calculated as  $\|A\|_{op} = \max_i \sigma_i$ .

**Proof** (i)

Note that Frobenius norm equals  $\|A\|_F = \sqrt{\sum_{i=1}^m \sum_{j=1}^n A_{ij}^2} = \sqrt{\text{tr}(A^\top A)}$ . And hence from the singular value decomposition of  $A$ ,

$$\begin{aligned} \|A\|_F &= \sqrt{\text{tr}(A^\top A)} = \sqrt{\text{tr}(U\Sigma V^\top V\Sigma^\top U^\top)} \\ &= \sqrt{\text{tr}(U\Sigma\Sigma^\top U^\top)} = \sqrt{\text{tr}(\Sigma\Sigma^\top U^\top U)} \\ &= \sqrt{\sum_{i=1}^{\min(m,n)} \sigma_i^2}. \end{aligned}$$

(ii)

Let  $r = \min(m, n)$ , and let columns of  $U$  and  $V$  be  $u_1, \dots, u_m$  and  $v_1, \dots, v_n$ , respectively. Expand any  $w \in \mathbb{R}^n$  using  $v_1, \dots, v_n$  as basis, so that  $w = \sum_{i=1}^n a_i v_i$  with  $a_1, \dots, a_n \in \mathbb{R}^n$ . Then

$$Aw = \sum_{i=1}^n a_i Av_i = \sum_{i=1}^r a_i \sigma_i u_i,$$

and hence

$$\|Aw\|_2 = \sqrt{\sum_{i=1}^r \sigma_i^2 a_i^2} \leq \max_{1 \leq i \leq r} \sigma_i \sqrt{\sum_{i=1}^r a_i^2} = \left( \max_{1 \leq i \leq r} \sigma_i \right) \|w\|_2,$$

and sufficient condition for the equality is when  $a_1 = 1$  and  $a_2 = \dots = a_n = 0$ , i.e.  $w = v_1$ . Hence

$$\|A\|_{op} = \sup_{w \neq 0} \frac{\|Aw\|_2}{\|w\|_2} = \max_{1 \leq i \leq r} \sigma_i.$$

1. In this exercise you will fill in some of the details from the proof of the upper bound for sparse PCA under a spike covariance model.

- (a) For  $p \geq 1$  the Schatten  $p$ -norm of a  $n \times m$  matrix  $A$  is the  $\ell_p$  norm of its singular values:

$$\|A\|_p = \left( \sum_{i=1}^r \sigma_i^p \right)^{1/p},$$

where  $\sigma_1 \geq \sigma_2 \geq \dots \sigma_r \geq 0$  are the singular values of  $A$  and  $r = \min\{m, n\}$ . Prove the non-commutative Hölder inequality for conformal matrices  $A$  and  $B$ :

$$|\operatorname{tr}(A^\top B)| \leq \|A\|_p \|B\|_q,$$

for all  $p, q \geq 1$  such that  $\frac{1}{p} + \frac{1}{q} = 1$ .

- (b) Let  $u$  and  $v$  be two unit norm vectors in  $\mathbb{R}^d$ . Show that

$$\|uu^\top - vv^\top\|_{\text{op}} = \|uu^\top - vv^\top\|_F = \sqrt{2 - 2(u^\top v)^2} = \sqrt{2 \sin^2(\angle(u, v))}.$$

where  $\angle(u, v) = \cos^{-1}(|u^\top v|)$

**Points:** 15 pts = 5 + 10.

**Solution.**

(a)

Let  $A$  be  $n \times m_1$  matrix and let  $B$  be  $n \times m_2$  matrix, and let  $r = \min\{n, m_1\}$ ,  $s = \min\{n, m_2\}$ . Let  $\sigma_1 \geq \dots \geq \sigma_r \geq 0$  be singular values of  $A$ , then from singular value decomposition, there exists an orthonormal set of vectors  $\{u_1, \dots, u_r\} \subset \mathbb{R}^n$  and an orthonormal set of vectors  $\{v_1, \dots, v_r\} \subset \mathbb{R}^{m_1}$  such that  $A = \sum_{i=1}^r \sigma_i u_i v_i^\top$ . These  $\{u_i\}$ 's are first  $r$  columns of  $U$  and  $\{v_i\}$ 's are first  $r$  columns of  $V$ , where  $A = U \Sigma V^\top$  is the singular value decomposition of  $A$ . Similarly, let  $\lambda_1 \geq \dots \geq \lambda_s \geq 0$  be singular values of  $B$ , then there exists an orthonormal set of vectors  $\{w_1, \dots, w_s\} \subset \mathbb{R}^n$  and an orthonormal set of vectors  $\{t_1, \dots, t_s\} \subset \mathbb{R}^{m_2}$  such that  $B = \sum_{j=1}^s \lambda_j w_j t_j^\top$ . Note that  $\|A\|_p = (\sum_{i=1}^r \sigma_i^p)^{1/p}$  and  $\|B\|_q = (\sum_{j=1}^s \lambda_j^q)^{1/q}$ . Meanwhile, LHS can be expanded as

$$\begin{aligned} |\operatorname{tr}(A^\top B)| &= \left| \operatorname{tr} \left( \left( \sum_{i=1}^r \sigma_i v_i u_i^\top \right) \left( \sum_{j=1}^s \lambda_j w_j t_j^\top \right) \right) \right| \\ &= \left| \sum_{i=1}^r \sum_{j=1}^s \sigma_i \lambda_j \operatorname{tr} (v_i (u_i^\top w_j) t_j^\top) \right| \\ &\leq \sum_{i=1}^r \sum_{j=1}^s \sigma_i \lambda_j |u_i^\top w_j| |\operatorname{tr} (v_i t_j^\top)|. \end{aligned}$$

Now, note that for any  $u, w \in \mathbb{R}^n$  with  $\|u\| = \|w\| = 1$ , Cauchy-Schwarz inequality yields  $|u^\top w| \leq \|u\| \|w\| = 1$ . Also, for any  $v \in \mathbb{R}^{m_1}$  and  $t \in \mathbb{R}^{m_2}$  with  $\|v\| = \|t\| = 1$ , Cauchy-Schwarz inequality yields

$$|\operatorname{tr} (v_i t_j^\top)| = \sum_{i=1}^{\min\{m_1, m_2\}} v_i t_i \leq \sqrt{\sum_{i=1}^{\min\{m_1, m_2\}} v_i^2} \sqrt{\sum_{i=1}^{\min\{m_1, m_2\}} t_i^2} \leq \|v\| \|t\| = 1.$$

Hence applying this and further applying Hölder's inequality gives

$$\begin{aligned} \left| \operatorname{tr}(A^\top B) \right| &\leq \sum_{i=1}^r \sum_{j=1}^s \sigma_i \lambda_j \\ &\leq \left( \sum_{i=1}^r \sigma_i^p \right)^{\frac{1}{p}} \left( \sum_{j=1}^s \lambda_j^q \right)^{\frac{1}{q}} = \|A\|_p \|B\|_q. \end{aligned}$$

Note that above inequality holds for  $p = 1, q = \infty$  or  $p = \infty, q = 1$  as well, where  $\|A\|_\infty = \max_{1 \leq i \leq r} \sigma_i$ .

(b)

The rest of the inequality follows from HW4 Problem 4(c), hence we are only left to show the first inequality  $\sqrt{2}\|uu^\top - vv^\top\|_{\text{op}} = \|uu^\top - vv^\top\|_F$ .

Note that the rank of  $uu^\top - vv^\top$  is at most 2, so nonzero eigenvalues can be at most 2. Hence  $uu^\top - vv^\top$  can be expressed as

$$uu^\top - vv^\top = \lambda_1 w_1 w_1^\top + \lambda_2 w_2 w_2^\top,$$

where  $\lambda_1, \lambda_2 \in \mathbb{R}$  and  $w_1, w_2 \in \mathbb{R}^d$  with  $w_1 \perp w_2$ . Then from

$$\lambda_1 + \lambda_2 = \operatorname{tr}(uu^\top - vv^\top) = u^\top u - v^\top v = 0,$$

$\lambda_2 = -\lambda_1$ , i.e. there exists  $\lambda > 0$  such that

$$uu^\top - vv^\top = \lambda w_1 w_1^\top - \lambda w_2 w_2^\top.$$

Then since  $uu^\top - vv^\top$  is real symmetric, all the singular values are absolute values of eigenvalues. Hence singular values of  $uu^\top - vv^\top$  are  $\lambda, \lambda, 0, \dots, 0$ . Hence from Lemma 0.1,

$$\|uu^\top - vv^\top\|_{\text{op}} = \lambda \quad \text{and} \quad \|uu^\top - vv^\top\|_F = \sqrt{\lambda^2 + \lambda^2} = \sqrt{2}\lambda,$$

and hence  $\sqrt{2}\|uu^\top - vv^\top\|_{\text{op}} = \|uu^\top - vv^\top\|_F$  holds.

2. This is a result that I cited when discussing spectra clustering for stochastic block models. A random matrix  $A$  of dimension  $n \times m$  is sub-Gaussian with parameter  $\sigma^2$ , written as  $A \in SG_{m,n}(\sigma^2)$ , when  $y^\top A x$  is  $SG(\sigma^2)$  for any  $y \in \mathbb{S}^{n-1}$  and  $x \in \mathbb{S}^{m-1}$ . You may assume that  $\mathbb{E}[A] = 0$  (or otherwise replace  $A$  by  $A - \mathbb{E}[A]$ ).

(a) Suppose that the entries of  $A$  are independent variables that are  $SG(\sigma^2)$ . Show that  $A \in SG_{m,n}(\sigma^2)$ .

(b) Let  $A \in SG_{n,m}(\sigma^2)$  and recall that the operator norm of  $A$  is

$$\|A\|_{\text{op}} = \max_{x \in \mathbb{R}^m, x \neq 0} \frac{\|Ax\|}{\|x\|} = \max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top A x.$$

Show that, for some  $C > 0$ ,

$$\mathbb{E}[\|A\|_{\text{op}}] \leq C(\sqrt{n} + \sqrt{m}).$$

(c) Find a concentration inequality for  $\|A\|_{\text{op}}$ .

Hint: work with a  $1/4$  net for  $\mathbb{S}^{n-1}$  and a  $1/4$  net for  $\mathbb{S}^{m-1}$ .

**Points:** 25 pts = 5 + 10 + 10.

**Solution.**

Assume that  $\mathbb{E}[A] = 0$ .

(a)

For all  $y \in \mathbb{S}^{n-1}$ ,  $x \in \mathbb{S}^{m-1}$ , and  $\lambda \in \mathbb{R}$ , by using the entries of  $A$  being independent and  $SG(\sigma^2)$ , the mgf of  $y^\top Ax$  is bounded as

$$\begin{aligned} \mathbb{E} \left[ \exp \left( \lambda y^\top Ax \right) \right] &= \mathbb{E} \left[ \exp \left( \lambda \sum_{i=1}^n \sum_{j=1}^m y_i A_{ij} x_j \right) \right] \\ &= \prod_{i=1}^n \prod_{j=1}^m \mathbb{E} [\exp ((\lambda y_i x_j) A_{ij})] \quad (\text{using independence of } A_{ij}) \\ &\leq \prod_{i=1}^n \prod_{j=1}^m \exp \left( \frac{1}{2} \sigma^2 (\lambda y_i x_j)^2 \right) \\ &= \exp \left( \frac{1}{2} \sigma^2 \lambda^2 \left( \sum_{i=1}^n y_i^2 \right) \left( \sum_{j=1}^m x_j^2 \right) \right) \\ &= \exp \left( \frac{1}{2} \sigma^2 \lambda^2 \right), \quad (y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}) \end{aligned}$$

hence  $A \in SG_{m,n}(\sigma^2)$ .

(b)

Let  $\mathcal{B}_n \subset \mathbb{S}^{n-1}$  and  $\mathcal{B}_m \subset \mathbb{S}^{m-1}$  be the  $\frac{1}{4}$ -net for  $\mathbb{S}^{n-1}$  and  $\mathbb{S}^{m-1}$  with respect to  $l_2$  distances. Then from Lecture note 5(Sep 18),

$$|\mathcal{B}_n| \leq \frac{\text{Vol} \left( \frac{2}{1/4} \mathcal{B}(0, 1) + \mathcal{B}(0, 1) \right)}{\text{Vol}(\mathcal{B}(0, 1))} = \frac{\text{Vol}(9\mathcal{B}(0, 1))}{\text{Vol}(\mathcal{B}(0, 1))} = 9^n,$$

and similarly  $|\mathcal{B}_m| \leq 9^m$  holds. Also for all  $y \in \mathbb{S}^{n-1}$  and  $x \in \mathbb{S}^{m-1}$ , there exists  $y' \in \frac{1}{4}\mathbb{S}^{n-1}$ ,  $x' \in \frac{1}{4}\mathbb{S}^{m-1}$ ,  $w \in \mathcal{B}_n$ ,  $z \in \mathcal{B}_m$  with  $y = w + y'$  and  $x = z + x'$ . Hence

$$\begin{aligned} &\max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top Ax \\ &\leq \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az + \max_{w \in \mathcal{B}_n, x \in \frac{1}{4}\mathbb{S}^{m-1}} w^\top Ax' + \max_{y' \in \frac{1}{4}\mathbb{S}^{n-1}, z \in \mathcal{B}_m} y'^\top Az + \max_{y' \in \frac{1}{4}\mathbb{S}^{n-1}, x' \in \frac{1}{4}\mathbb{S}^{m-1}} y'^\top Ax' \\ &= \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az + \frac{1}{4} \max_{w \in \mathcal{B}_n, x \in \mathbb{S}^{m-1}} w^\top Ax + \frac{1}{4} \max_{y \in \mathbb{S}^{n-1}, z \in \mathcal{B}_m} y^\top Az + \frac{1}{16} \max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top Ax. \end{aligned} \tag{1}$$

Now, note that

$$\begin{aligned} \max_{w \in \mathcal{B}_n, x \in \mathbb{S}^{m-1}} w^\top Ax &\leq \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az + \max_{w \in \mathcal{B}_n, x' \in \frac{1}{4}\mathbb{S}^{m-1}} w^\top Ax' \\ &\quad \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az + \frac{1}{4} \max_{w \in \mathcal{B}_n, x \in \mathbb{S}^{m-1}} w^\top Ax, \end{aligned}$$

so  $\max_{w \in \mathcal{B}_n, x \in \mathbb{S}^{m-1}} w^\top Ax \leq \frac{4}{3} \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az$ , and similarly  $\max_{y \in \mathbb{S}^{n-1}, z \in \mathcal{B}_m} y^\top Az \leq \frac{4}{3} \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az$ . And applying these to (1) yields

$$\max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top Ax \leq \frac{16}{9} \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az. \quad (2)$$

Hence these yields the upper bound of  $\mathbb{E}[\|A\|_{\text{op}}]$  as

$$\begin{aligned} \mathbb{E}[\|A\|_{\text{op}}] &= \mathbb{E} \left[ \max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top Ax \right] \\ &\leq \frac{16}{9} \mathbb{E} \left[ \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az \right] \quad (\text{using (2)}) \\ &\leq \frac{16\sigma}{9} \sqrt{2 \log(|\mathcal{B}_m| |\mathcal{B}_n|)} \\ &\leq \frac{16\sigma}{9} \sqrt{2(m+n) \log 9} \\ &< \frac{32\sigma \sqrt{\log 3}}{9} (\sqrt{m} + \sqrt{n}) \quad (\text{using } \sqrt{m+n} < \sqrt{m} + \sqrt{n} \text{ for } m, n > 0). \end{aligned}$$

(c)

From (2), the tail probability  $\mathbb{P}(\|A\|_{\text{op}} \geq t)$  can be bounded as

$$\begin{aligned} \mathbb{P}(\|A\|_{\text{op}} \geq t) &= \mathbb{P} \left( \max_{y \in \mathbb{S}^{n-1}, x \in \mathbb{S}^{m-1}} y^\top Ax \geq t \right) \\ &\leq \mathbb{P} \left( \frac{16}{9} \max_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} w^\top Az \geq t \right) \quad (\text{using (2)}) \\ &\leq \sum_{w \in \mathcal{B}_n, z \in \mathcal{B}_m} \mathbb{P} \left( w^\top Az \geq \frac{9}{16} t \right). \end{aligned} \quad (3)$$

Then for each  $w \in \mathcal{B}_n$  and  $z \in \mathcal{B}_m$ , since  $w^\top Az \in SG(\sigma^2)$  with  $\mathbb{E}[w^\top Az] = 0$ , so by applying Hoeffding's inequality,

$$\mathbb{P} \left( w^\top Az \geq \frac{9}{16} t \right) \leq \exp \left( -\frac{81t^2}{512\sigma^2} \right).$$

Hence applying this to (3) gives the bound for the tail probability  $\mathbb{P}(\|A\|_{\text{op}} \geq t)$  as

$$\begin{aligned} \mathbb{P}(\|A\|_{\text{op}} \geq t) &\leq |\mathcal{B}_m| |\mathcal{B}_n| \exp \left( -\frac{81t^2}{512\sigma^2} \right) \\ &\leq 9^{m+n} \exp \left( -\frac{81t^2}{512\sigma^2} \right). \end{aligned}$$

### 3. Exercise 8.4:

Show that the orthogonal matrix  $V \in \mathbb{R}^{d \times r}$  maximizing the criterion (8.2) has columns formed by the top  $r$  eigenvectors of  $\Sigma = \text{cov}(X)$ :

$$\mathbb{E} \left\| V^\top X \right\|_2^2 = \sum_{j=1}^r \mathbb{E} \left[ \langle v_j, X \rangle^2 \right].$$

**Points:** 15 pts.

**Solution.**

$$\begin{aligned}
\mathbb{E} \left\| V^\top X \right\|_2^2 &= \mathbb{E} \left[ X^\top V V^\top X \right] = \mathbb{E} \left[ \text{tr} \left( X^\top V V^\top X \right) \right] \\
&= \mathbb{E} \left[ \text{tr} \left( V V^\top X X^\top \right) \right] = \text{tr} \left( V V^\top \mathbb{E} \left[ X X^\top \right] \right) \\
&= \text{tr} \left( V V^\top \Sigma \right).
\end{aligned}$$

Let eigenvalues and corresponding eigenvectors of  $\Sigma$  be  $\lambda_1 \geq \dots \geq \lambda_d$  and  $u_1, \dots, u_d$ . Then

$$\begin{aligned}
\mathbb{E} \left\| V^\top X \right\|_2^2 &= \text{tr} \left( V V^\top \sum_{i=1}^d \lambda_i u_i u_i^\top \right) = \sum_{i=1}^d \lambda_i \text{tr} \left( V V^\top u_i u_i^\top \right) \\
&= \sum_{i=1}^d \lambda_i \text{tr} \left( u_i^\top V V^\top u_i \right) = \sum_{i=1}^d \lambda_i \left\| V^\top u_i \right\|_2^2.
\end{aligned}$$

Now, note that  $(V V^\top) V = V (V^\top V) = V I_r = V$ , and hence for any  $w \in \text{Col}(V)$ ,  $(V V^\top) w = w$ , where  $\text{Col}(V)$  is the column space of  $V$ . Also, for any  $w^\perp \in \text{Col}(V)^\perp$ ,  $V^\top w^\perp = 0$ , which implies  $V V^\top w^\perp = 0$ . Hence for any  $u \in \mathbb{R}^d$  with  $\|u\|_2 = 1$ , let  $u = w + w^\perp$  with  $w \in \text{Col}(V)$  and  $w^\perp \in \text{Col}(V)^\perp$ , then

$$\begin{aligned}
\left\| V^\top u \right\|_2^2 &= (w + w^\perp)^\top V V^\top (w + w^\perp) \\
&= (w + w^\perp)^\top w = \|w\|_2^2 \leq \|u\|_2^2 = 1,
\end{aligned}$$

hence  $V^\top u_i$  satisfy

$$0 \leq \left\| V^\top u_i \right\|_2^2 \leq 1. \quad (4)$$

Also,  $\left\| V^\top u_i \right\|_2^2 = 1$  holds if and only if  $u_i \in \text{Col}(V)$  (i.e.  $\text{Col}(V)^\perp$  part is 0) and  $\left\| V^\top u_i \right\|_2^2 = 0$  if and only if  $u_i \in \text{Col}(V)^\perp$  (i.e.  $\text{Col}(V)$  part is 0). Also, note that

$$\begin{aligned}
\sum_{i=1}^d \left\| V^\top u_i \right\|_2^2 &= \sum_{i=1}^d \text{tr} \left( V V^\top u_i u_i^\top \right) = \text{tr} \left( V V^\top \sum_{i=1}^d u_i u_i^\top \right) \\
&= \text{tr} \left( V V^\top I_d \right) = \text{tr} \left( V^\top V \right) = \text{tr} \left( I_r \right) = r.
\end{aligned} \quad (5)$$

Now, consider the following optimization problem:

$$\begin{aligned}
&\text{maximize} \quad \sum_{i=1}^d \lambda_i x_i, \\
&\text{subject to} \quad \forall i, 0 \leq x_i \leq 1, \sum_{i=1}^d x_i = r.
\end{aligned} \quad (6)$$

Then, the constraints imply that  $1 \leq \forall i \leq d, \sum_{j=1}^i x_j \leq \min\{i, r\}$ . Let  $\lambda_{d+1} = 0$  for conve-

nience, and note that

$$\begin{aligned}
\sum_{i=1}^d \lambda_i x_i &= \sum_{i=1}^d (\lambda_i - \lambda_{i+1}) \left( \sum_{j=1}^i x_j \right) \\
&\leq \sum_{i=1}^r (\lambda_i - \lambda_{i+1}) i + \sum_{i=r+1}^d (\lambda_i - \lambda_{i+1}) r \\
&= \sum_{i=1}^r \lambda_i,
\end{aligned}$$

and the equality holds if and only if  $\sum_{j=1}^i x_j = \min\{i, r\}$  whenever  $\lambda_i > \lambda_{i+1}$ . Hence

$$x_1 = \cdots = x_r = 1, x_{r+1} = \cdots = x_d = 0 \quad (7)$$

maximizes the objective function  $\sum_{i=1}^d \lambda_i x_i$ , and if we further have  $\lambda_r > \lambda_{r+1}$ , it is the unique maximizer.

Hence, applying (4) and (5) to (6) implies that

$$\mathbb{E} \|V^\top X\|_2^2 = \sum_{i=1}^d \lambda_i \|V^\top u_i\|_2^2 \leq \sum_{i=1}^r \lambda_i.$$

And from (7), the equality holds if

$$\|V^\top u_1\|_2^2 = \cdots = \|V^\top u_r\|_2^2 = 1 \text{ and } \|V^\top u_{r+1}\|_2^2 = \cdots = \|V^\top u_d\|_2^2 = 0, \quad (8)$$

and only if holds under  $\lambda_r > \lambda_{r+1}$ . Then from equality conditions from (4), (8) is equivalent to

$$\text{Col}(V) = \langle u_1, \dots, u_r \rangle,$$

where  $\langle u_1, \dots, u_r \rangle \subset \mathbb{R}^d$  is the linear subspace spanned by  $u_1, \dots, u_r$ . In conclusion,

$$\mathbb{E} \|V^\top X\|_2^2 \leq \sum_{i=1}^r \lambda_i,$$

and the equality holds if  $\text{Col}(V) = \langle u_1, \dots, u_r \rangle$ , and under  $\lambda_r > \lambda_{r+1}$ , it is necessary as well. In particular, when top  $r$  eigenvectors  $\{u_1, \dots, u_r\}$  form the columns of  $V$ ,  $\mathbb{E} \|V^\top X\|_2^2$  is maximized.

4. Suppose we observe an i.i.d. sample  $X_1, \dots, X_n$  from the mixture distribution

$$\frac{1}{2}P_1 + \frac{1}{2}P_2,$$

where  $P_1 = N_d(\mu, I_d)$  and  $P_2 = N_d(-\mu, I_d)$ , with  $\mu \in \mathbb{R}^d$  a non-zero vector. Ours task is to cluster the sample points into two groups, where points in the same group originated from the same component of the mixture (i.e. Either  $P_1$  or  $P_2$ )

We will use spectral clustering: compute the leading eigenvector  $\hat{v}$  of the empirical covariance matrix and cluster the points depending on the sign of  $X_i^\top \hat{v}$ ,  $i = 1, \dots, n$ .

Use the Davis-Kahan theorem to derive an upper bound on the proportion of misclustered nodes.

**Points:** 20 + 1 pts.

**Solution.**

$X$  has the same distribution as  $W\mu + Y$ , where  $W \sim \text{Rademacher}$ ,  $Y \sim N_d(0, I_d)$ . Then  $\mathbb{E}[X] = 0$  and

$$\Sigma := \text{Var}[X] = \mathbb{E}[XX^\top] = \mu\mu^\top + I_d.$$

Hence by letting  $v := \frac{\mu}{\|\mu\|}$ , then  $v$  is the leading eigenvector of  $\Sigma$  with eigenvalue  $1 + \|\mu\|^2$ . Hence from Davis-Kahan theorem,

$$\sin \angle(v, \hat{v}) \leq \frac{2 \|\hat{\Sigma} - \Sigma\|_{op}}{\|\mu\|^2}. \quad (9)$$

Then note that the number of misclustered nodes is upper bounded by

$$\sum_{i=1}^n I \left( \left\{ X_i^\top \hat{v} \leq 0, W_i = 1 \right\} \cup \left\{ X_i^\top \hat{v} > 0, W_i = -1 \right\} \right).$$

Hence the misclustering proportion  $R$  is bounded as

$$\begin{aligned} R &\leq \mathbb{P} \left( X_i^\top \hat{v} \leq 0, W_i = 1 \right) + \mathbb{P} \left( X_i^\top \hat{v} > 0, W_i = -1 \right) \\ &\leq \mathbb{P} \left( (\mu + Y_i)^\top \hat{v} \leq 0 \right) + \mathbb{P} \left( (-\mu + Y_i)^\top \hat{v} > 0 \right). \end{aligned}$$

Now, consider the first event  $\{(\mu + Y_i)^\top \hat{v} \leq 0\}$ . Note that

$$\begin{aligned} \langle \mu + Y_i, \hat{v} \rangle &= \langle \mu, \hat{v} \rangle + \langle Y_i, \hat{v} \rangle \\ &\geq \langle \mu, \hat{v} \rangle - \|Y_i\|, \end{aligned}$$

Hence  $\langle \mu + Y_i, \hat{v} \rangle \leq 0$  implies

$$\|Y_i\| \geq \langle \mu, \hat{v} \rangle - \langle \mu + Y_i, \hat{v} \rangle \geq \langle \mu, \hat{v} \rangle.$$

Its geometric interpretation is in Figure 1. And applying  $-Y_i$  instead of  $Y_i$  gives that  $\langle -\mu + Y_i, \hat{v} \rangle > 0$  implies  $\|Y_i\| > \langle \mu, \hat{v} \rangle$  as well. Since  $\hat{v}$  is aligned with  $\mu$  so that  $\langle \mu, \hat{v} \rangle \geq 0$ , and hence the misclustering proportion  $R$  is bounded as

$$R \leq 2\mathbb{P} \left( \|Y_i\|^2 \geq (\langle \mu, \hat{v} \rangle)^2 \right).$$

Then applying the result of Davis-Kahan Theorem in (9) implies

$$\begin{aligned} (\langle \mu, \hat{v} \rangle)^2 &= \|\mu\|^2 \cos^2 \angle(v, \hat{v}) = \|\mu\|^2 (1 - \sin^2 \angle(v, \hat{v})) \\ &\geq \|\mu\|^2 \left( 1 - \frac{4 \|\hat{\Sigma} - \Sigma\|_{op}^2}{\|\mu\|^4} \right). \end{aligned}$$

Hence the misclustering proportion  $R$  is bounded as

$$R \leq 2\mathbb{P} \left( \|Y_i\|^2 \geq \|\mu\|^2 \left( 1 - \frac{4 \|\hat{\Sigma} - \Sigma\|_{op}^2}{\|\mu\|^4} \right) \right).$$



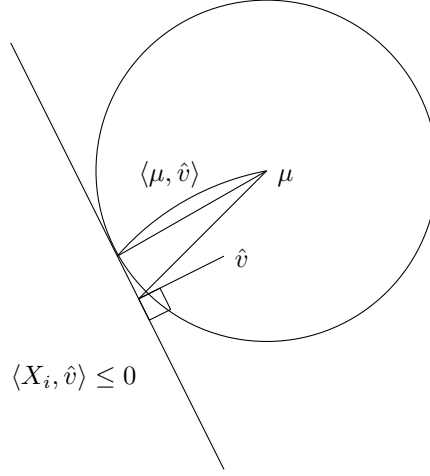


Figure 1: Geometrical interpretation of  $\langle \mu + Y_i, \hat{v} \rangle \leq 0$  implying  $\|Y_i\| \geq \langle \mu, \hat{v} \rangle$ .  $\langle \mu + Y_i, \hat{v} \rangle \leq 0$  is equivalent to  $X_i = \mu + Y_i$  lying on the bottom-left region of the hyperplane which is perpendicular to  $\hat{v}$ . Then  $X_i$  lies outside of a sphere centered at  $\mu$  and of radius  $\langle \mu, \hat{v} \rangle$ . Hence  $\|X_i - \mu\| = \|Y_i\| \geq \langle \mu, \hat{v} \rangle$ .

Now, note that for any  $v \in \mathbb{R}^d$  with  $\|v\| = 1$ ,

$$\begin{aligned} \mathbb{E} \left[ \exp \left( \lambda v^\top X_i \right) \right] &= \frac{1}{2} \mathbb{E} \left[ \exp \left( \lambda v^\top (\mu + Y_i) \right) \right] + \frac{1}{2} \mathbb{E} \left[ \exp \left( \lambda v^\top (-\mu + Y_i) \right) \right] \\ &= \frac{1}{2} \exp \left( \lambda v^\top \mu + \frac{1}{2} \lambda^2 \right) + \frac{1}{2} \exp \left( -\lambda v^\top \mu + \frac{1}{2} \lambda^2 \right) \\ &= \exp \left( \frac{1}{2} \lambda^2 \right) \cosh \left( \lambda v^\top \mu \right) \\ &\leq \exp \left( \frac{1}{2} \lambda^2 (1 + (v^\top \mu)^2) \right) \quad (\because \cosh x \leq \exp(x^2/2)) \\ &\leq \exp \left( \frac{1}{2} \lambda^2 (1 + \|\mu\|^2) \right), \quad (\because \text{Cauchy-Schwarz}) \end{aligned}$$

Hence  $X_i \in SG(1 + \|\mu\|^2)$ . Hence from Lecture note 6 on Sep 20, for all  $\delta \in (0, 1)$ , there exists some  $C > 0$  such that

$$\mathbb{P} \left( \left\| \hat{\Sigma} - \Sigma \right\|_{op} \leq C(1 + \|\mu\|^2) \max \left\{ \sqrt{\frac{d + \log(2/\delta)}{n}}, \frac{d + \log(2/\delta)}{n} \right\} \right) \geq 1 - \delta. \quad (10)$$

Let  $E$  be the event that  $\left\| \hat{\Sigma} - \Sigma \right\|_{op} \leq C(1 + \|\mu\|^2) \max \left\{ \sqrt{\frac{d + \log(2/\delta)}{n}}, \frac{d + \log(2/\delta)}{n} \right\}$  happens. Then

$$\|Y_i\|^2 \geq \|\mu\|^2 \left( 1 - \frac{4 \left\| \hat{\Sigma} - \Sigma \right\|_{op}^2}{\|\mu\|^4} \right) \text{ and } E \text{ implies}$$

$$\begin{aligned} \|Y_i\|^2 &\geq \|\mu\|^2 \left( 1 - \frac{4C^2 (1 + \|\mu\|^2)^2}{\|\mu\|^4} \max \left\{ \frac{d + \log(2/\delta)}{n}, \left( \frac{d + \log(2/\delta)}{n} \right)^2 \right\} \right) \\ &\geq \|\mu\|^2 \left( 1 - 8C^2 \left( 1 + \frac{1}{\|\mu\|^4} \right) \max \left\{ \frac{d + \log(2/\delta)}{n}, \left( \frac{d + \log(2/\delta)}{n} \right)^2 \right\} \right) \\ &:= K_{n,d,\delta,\|\mu\|,C}, \end{aligned}$$

where  $K_{n,d,\delta,\|\mu\|,C} = \|\mu\|^2 \left(1 - 8C^2 \left(1 + \frac{1}{\|\mu\|^4}\right) \max \left\{ \frac{d+\log(2/\delta)}{n}, \left(\frac{d+\log(2/\delta)}{n}\right)^2 \right\} \right)$ . And hence

$$\begin{aligned} R &\leq 2\mathbb{P} \left( \left\{ \|Y_i\|^2 \geq \|\mu\|^2 \left(1 - \frac{4 \left\| \hat{\Sigma} - \Sigma \right\|_{op}^2}{\|\mu\|^4} \right) \right\} \cap E \right) \\ &\quad + 2\mathbb{P} \left( \left\{ \|Y_i\|^2 \geq \|\mu\|^2 \left(1 - \frac{4 \left\| \hat{\Sigma} - \Sigma \right\|_{op}^2}{\|\mu\|^4} \right) \right\} \cap E^c \right) \\ &\leq 2\mathbb{P} (\|Y_i\|^2 \geq K_{n,d,\delta,\|\mu\|,C}) + 2\mathbb{P}(E^c) \\ &\leq 2\mathbb{P} (\|Y_i\|^2 \geq K_{n,d,\delta,\|\mu\|,C}) + 2\delta. \end{aligned}$$

Now, since  $Y_i \sim N_d(0, I_d)$ ,  $\|Y_i\|^2 \sim \chi_d^2$ , chi-square distribution with degree  $d$ . Hence

$$\mathbb{P} (\|Y_i\|^2 \geq K_{n,d,\delta,\|\mu\|,C}) = 1 - F_{\chi_d^2}(K_{n,d,\delta,\|\mu\|,C}),$$

where  $F_{\chi_d^2}$  is the CDF function for  $\chi_d^2$ . To conclude, the misclustering proportion  $R$  is bounded as

$$R \leq 2 \inf_{\delta \in (0,1)} \left\{ 1 - F_{\chi_d^2}(K_{n,d,\delta,\|\mu\|,C}) + \delta \right\},$$

where  $K_{n,d,\delta,\|\mu\|,C} = \|\mu\|^2 \left(1 - 8C^2 \left(1 + \frac{1}{\|\mu\|^4}\right) \max \left\{ \frac{d+\log(2/\delta)}{n}, \left(\frac{d+\log(2/\delta)}{n}\right)^2 \right\} \right)$ .

Note that when getting the bound (10), matrix Bernstein inequality is a bit inconvenient to use since  $\|X_i\|$  is not bounded. There is 1 bonus point for using the correct covariance bound for getting (10).

#### 5. Exercise 4.3:

(a) Recall the re-centered function class  $\bar{\mathcal{F}} = \{f - \mathbb{E}[f] \mid f \in \mathcal{F}\}$ . Show that

$$\mathbb{E}_{X,\epsilon} [\|R_n\|_{\bar{\mathcal{F}}}] \geq \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] - \frac{\sup_{f \in \mathcal{F}} |\mathbb{E}[f]|}{\sqrt{n}}.$$

(b) Use concentration results to complete the proof of Proposition 4.2.

**Points:** 15 pts = 10 + 5.

#### Solution.

(a)

Note that for any  $\bar{f} \in \bar{\mathcal{F}}$ , there exists  $f \in \mathcal{F}$  such that  $\bar{f}(x) = f(x) + \mathbb{E}[f]$ . Hence

$$\begin{aligned} \|R_n\|_{\bar{\mathcal{F}}} &= \sup_{\bar{f} \in \bar{\mathcal{F}}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i \bar{f}(X_i) \right| \\ &= \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i f(X_i) - \frac{1}{n} \sum_{i=1}^n \epsilon_i \mathbb{E}[f] \right| \\ &\geq \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i f(X_i) \right| - \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i \mathbb{E}[f] \right|. \end{aligned} \tag{11}$$

Now, note that by applying Cauchy Schwarz,  $\mathbb{E} \left[ \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i \right| \right]$  can be bounded as

$$\begin{aligned} \mathbb{E} \left[ \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i \right| \right] &\leq \frac{1}{n} \sqrt{\mathbb{E} \left[ \left( \sum_{i=1}^n \epsilon_i \right)^2 \right]} = \frac{1}{n} \sqrt{\mathbb{E} \left[ \sum_{i=1}^n \epsilon_i^2 + \sum_{i \neq j} \epsilon_i \epsilon_j \right]} \\ &= \frac{1}{n} \sqrt{n} = \frac{1}{\sqrt{n}}. \end{aligned}$$

By plugging this to expectation of (11), we get

$$\begin{aligned} \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] &\geq \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] - \sup_{f \in \mathcal{F}} |\mathbb{E}[f]| \mathbb{E} \left[ \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i \right| \right] \\ &\geq \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] - \frac{\sup_{f \in \mathcal{F}} |\mathbb{E}[f]|}{\sqrt{n}}. \end{aligned}$$

(b)

Note that from (a) and Proposition 4.1 (a), we have

$$\begin{aligned} \mathbb{E}_X [\|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}}] &\geq \mathbb{E}_{X,\epsilon} \left[ \frac{1}{2} \|R_n\|_{\mathcal{F}} \right] \\ &\geq \frac{1}{2} \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] - \frac{\sup_{f \in \mathcal{F}} |\mathbb{E}[f]|}{2\sqrt{n}}. \end{aligned} \tag{12}$$

Now, for each  $f \in \mathcal{F}$  let  $\bar{f}(x) = f(x) - \mathbb{E}[f]$  as in (a), and let  $G(x_1, \dots, x_n) = \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \bar{f}(x_i) \right|$  so that  $\|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}} = G(X_1, \dots, X_n)$ . Note that if  $x, y \in \mathcal{X}^n$  satisfies  $x_j = y_j$  for all  $j \neq i$ , then for any  $f \in \mathcal{F}$ ,

$$\begin{aligned} \left| \frac{1}{n} \sum_{j=1}^n \bar{f}(x_j) \right| - \sup_{h \in \mathcal{F}} \left| \frac{1}{n} \sum_{j=1}^n \bar{h}(y_j) \right| &\leq \left| \frac{1}{n} \sum_{j=1}^n \bar{f}(x_j) \right| - \left| \frac{1}{n} \sum_{j=1}^n \bar{f}(y_j) \right| \\ &\leq \left| \frac{1}{n} \sum_{j=1}^n \bar{f}(x_j) - \frac{1}{n} \sum_{j=1}^n \bar{f}(y_j) \right| \\ &= \frac{1}{n} |\bar{f}(x_i) - \bar{f}(y_i)| \\ &\leq \frac{2}{n} \|f\|_{\infty} \leq \frac{2b}{n}. \end{aligned}$$

Hence taking supremum over  $f \in \mathcal{F}$  yields  $G(x) - G(y) \leq \frac{2b}{n}$ . Since the same argument can be applied to get  $G(y) - G(x) \leq \frac{2b}{n}$ , so  $|G(x) - G(y)| \leq \frac{2b}{n}$ . Hence from Bounded differences inequality,

$$\mathbb{P}(G(X) - \mathbb{E}[G(X)] \geq \delta) \leq \exp \left( -\frac{n\delta^2}{2b^2} \right).$$

Then applying  $G(X) = \|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}}$  and (12) to this gives

$$\begin{aligned} \mathbb{P} \left( \|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}} - \left( \frac{1}{2} \mathbb{E}_{X,\epsilon} [\|R_n\|_{\mathcal{F}}] - \frac{\sup_{f \in \mathcal{F}} |\mathbb{E}[f]|}{2\sqrt{n}} \right) \geq \delta \right) \\ \leq \mathbb{P} (\|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}} - \mathbb{E} [\|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}}] \geq \delta) \leq \exp \left( -\frac{n\delta^2}{2b^2} \right). \end{aligned}$$

Hence  $\|\mathbb{P}_n - \mathbb{P}\|_{\mathcal{F}} \leq \frac{1}{2} \mathbb{E}_{X, \epsilon} [\|R_n\|_{\mathcal{F}}] - \frac{\sup_{f \in \mathcal{F}} |\mathbb{E}[f]|}{2\sqrt{n}} + \delta$  with  $\mathbb{P}$ -probability at least  $1 - e^{-\frac{n\delta^2}{2b^2}}$ .

6. **Massart's finite class Lemma** Let  $\mathcal{A}$  be a finite subset of  $\mathbb{R}^n$  and let  $\epsilon = (\epsilon_1, \dots, \epsilon_n)$  be a vector of i.i.d. Rademacher variables. Show that

$$\mathbb{E} \left[ \frac{1}{n} \sup_{a \in \mathcal{A}} a^\top \epsilon \right] \leq D \frac{\sqrt{2 \log |\mathcal{A}|}}{n}$$

where  $D = \max_{a \in \mathcal{A}} \|a\|$ . Use this result to prove Lemma 4.1 of Chapter 4. (In proving both claims it is OK if you get different constants; know however that the constants in Lemma 4.1 are sub-optimal).

**Points:** 10 pts.

**Solution.**

Note that  $\mathbb{E}[e^{\lambda \epsilon_i}]$  can be bounded as

$$\begin{aligned} \mathbb{E}[\exp(\lambda \epsilon_i)] &= \frac{1}{2}(e^\lambda + e^{-\lambda}) = \sum_{k=0}^{\infty} \frac{\lambda^{2k}}{(2k)!} \\ &\leq \sum_{k=0}^{\infty} \frac{(\lambda^2)^k}{k!} \times \frac{k!}{(2k)!} \leq \sum_{k=0}^{\infty} \frac{(\lambda^2)^k}{k!} \times \left(\frac{1}{2}\right)^k = e^{\frac{1}{2}\lambda^2}, \end{aligned}$$

hence  $\epsilon_i \in SG(1)$ . Since  $\epsilon_i$ 's are independent, from HW1 Problem 6 Details,

$$a^\top \epsilon \in SG(\|a\|_2^2) \subset SG(D^2).$$

Hence from maximum inequality for subgaussian random variables,

$$\mathbb{E} \left[ \frac{1}{n} \sup_{a \in \mathcal{A}} a^\top \epsilon \right] \leq D \frac{\sqrt{2 \log |\mathcal{A}|}}{n},$$

and from HW2 Problem 1, we also have a bound for absolute value version,

$$\mathbb{E} \left[ \frac{1}{n} \sup_{a \in \mathcal{A}} |a^\top \epsilon| \right] \leq D \frac{\sqrt{2 \log(2|\mathcal{A}|)}}{n}. \quad (13)$$

For proving Lemma 4.1 of Chapter 4, let  $\mathcal{F}(x_1^n) := \{(f(x_1), \dots, f(x_n)) \in \mathbb{R}^n \mid f \in \mathcal{F}\}$  as in (4.11) in Wainwright, then  $\max_{a \in \mathcal{F}(x_1^n)} \|a\| = \sqrt{n}D(x_1^n)$ . Hence by applying (13),

$$\mathbb{E} \left[ \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i f(x_i) \right| \right] \leq D(x_1^n) \sqrt{\frac{2(\log |\mathcal{F}(x_1^n)| + \log 2)}{n}}.$$

Then  $\mathcal{F}$  being polynomial discrimination of order  $\nu$  implies  $|\mathcal{F}(x_1^n)| \leq (n+1)^\nu$ , hence

$$\begin{aligned} \mathbb{E} \left[ \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n \epsilon_i f(x_i) \right| \right] &\leq D(x_1^n) \sqrt{\frac{2(\nu \log(n+1) + \log 2)}{n}} \\ &\leq 2D(x_1^n) \sqrt{\frac{\nu \log(n+1)}{n}}. \end{aligned}$$