bigDM

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Introduction

Disease mapping is the field of spatial epidemiology that studies the link between geographic locations and the occurrence of diseases, focusing on the estimation of the spatial and/or spatio-temporal distribution of disease incidence or mortality patterns. The great variability inherent to classical risk estimation measures, such as standardized mortality/incidence ratios or crude rates, makes necessary the use of statistical models to estimate smooth spatial risk surfaces borrowing information from neighbouring regions. The information adquired from these analyses is invaluable for health researchers and policy-makers as it helps to formulate hypothesis about the etiology of a disease, to look for risk factors and and also to allocate funds efficiently in hot spot areas.

Spatial models in disease mapping

Let us assume that the spatial domain of interest is divided into n contiguous small areas labeled as i = 1, ..., n. For a given area i, let O_i and E_i denote the observed and expected number of disease cases, respectively. Using these quantities, the *standardized mortality/incidence ratio* (SMR or SIR) is defined as the ratio of observed and expected cases for the corresponding areal unit. Although its interpretation is very simple, these measures are extremely variable when analyzing rare diseases or low-populated areas, as it is the case of high-dimensional data. To cope with this situation, it is necessary to use statistical models that stabilize the risks (rates) borrowing information from neighbouring regions.

Generalized linear mixed models (GLMM) are typically used for the analysis of count data within a hierarchical Bayesian framework. Conditional to the relative risk r_i , the number of observed cases in the *i*th area is assumed to be Poisson distributed with mean $\mu_i = E_i r_i$. That is,

$$O_i|r_i \sim Poisson(\mu_i = E_i r_i), i = 1, ..., n$$

 $\log \mu_i = \log E_i + \log r_i,$

% where log E_i is an offset. Depending on the specification of the log-risks different models are defined. Here we assume that %

$$\log r_i = \alpha + \xi_i,\tag{1}$$

% where α is an intercept representing the overall log-risk and $\xi = (\xi_1, \dots, \xi_n)'$ is a spatial random effect for which a conditional autoregressive (CAR) prior is usually assumed. The spatial correlation between CAR random effects is determined by the neighbouring structure (represented as an undirected graph) of the areal units. Let $\mathbf{W} = (w_{ij})$ be a binary $n \times n$ adjacency matrix, whose ijth element is equal to one if areas j and k are defined as neighbours (usually if they share a common border), and it is zero otherwise. Here, the prior distribution proposed by Leroux et. al.(2000) has been considered, which is given by

$$\xi \sim N(0, Q_{\xi}^{-1}), \text{ where } Q_{\xi} = \tau_{\xi}[\lambda_{\xi}(D_W - W) + (1 - \lambda_{\xi})I_n]$$

% and $\tau_{\xi} = 1/\sigma_{\xi}^2$ is the precision parameter, $\lambda_{\xi} \in [0,1)$ is a spatial smoothing parameter, $D_W = diag(w_{1+},\ldots,w_{n+})$ and $w_{i+} = \sum_j w_{ij}$ is the *i*th row sum of W, and I_n is the $n \times n$ identity matrix. % We will refer to this model as the *Global model*.

A initial step: Merge cartography file with datasets

This package has as input files of the class sf or spatialpolygonsdataframe, but not always the user has this type of file. For example: in the case of Spain, we have the following cartography files:

- MGN MPIO POLITICO.shp
- MGN MPIO POLITICO.shx
- MGN MPIO POLITICO.prj
- MGN MPIO POLITICO.dbf

To convert this cartography in SpatialPolygonsDataFrame the following is required:

Alternatively, an sf can be used:

To merge the observed and expected data with the imported cartography it is necessary to have the data organized by a common identifier. Using function as merge() or logical relations it is possible to have in a unique file cartography and data.

```
#### data with expected and observed cases in csv file ####
library(readr)
data <- read_csv(system.file("csv/data.csv", package = "bigDM"),</pre>
                 col_types = cols(ID = col_character(),
                                  exp = col_number(),
                                  obs = col_number()))
zones_sf<- merge (zones_sf, data, by.x="ID", by.y="ID")</pre>
#> Simple feature collection with 6 features and 9 fields
#> geometry type: MULTIPOLYGON
#> dimension:
#> bbox:
                 xmin: 485318 ymin: 4727428 xmax: 543317 ymax: 4779153
#> proj4string: +proj=utm +zone=30 +ellps=GRS80 +units=m +no_defs
                                              name lat
                                                              long
                                                                        area
#> 1 01001
                                  Alegria-Dulantzi 4742974 4742974 19913794
#> 2 01002
                                           Amurrio 4765886 4765886 96145595
#> 3 01003
                                           Aramaio 4766225 4766225 73338806
#> 4 01004
                                        Artziniega 4775453 4775453 27506468
```

```
#> 5 01006
                                          Arminon 4729884 4729884 10559721
#> 6 01008 Arrazua-Ubarrundia (San Martin de Ania) 4752714 4752714 57502811
    perimeter
                  region obs
                                    exp
                                                              geometry
#> 1 34372.11 Pais Vasco
                           4 3.0237149 MULTIPOLYGON (((538259 4737...
#> 2 63352.32 Pais Vasco 27 20.8456682 MULTIPOLYGON (((503520 4760...
#> 3 41430.46 Pais Vasco
                           6 3.7527301 MULTIPOLYGON (((533286 4759...
#> 4 22605.22 Pais Vasco
                           5 3.2093191 MULTIPOLYGON (((491260 4776...
#> 5 17847.35 Pais Vasco
                           0 0.4817391 MULTIPOLYGON (((509851 4727...
                           1 1.9643891 MULTIPOLYGON (((534678 4746...
#> 6 64968.81 Pais Vasco
```

Disjoint model

This model is characterized by the area of interest D is formed by a partition of sub-areas d_i such that $\bigcup_{i=1}^n d_i = D$ and $\bigcap_{i=1}^n d_i = \emptyset$. This partition can be determined by an administrative or natural region (provinces, municipalities,...), although it can also be determined randomly. \mathbf{E}_d and \mathbf{O}_d are expected and observed cases, respectively in each sub-area $d_i \in D$. Then we will have the log-risk vector $\log \mathbf{r}_d$, overall log-risk vector α_d and spatial random effect vector ξ_d . The steps for fitting a model whose small area is large are illustrated below, for colorectal cancer in Spain with simulated data.

Stage 1: Build disjoint partition.

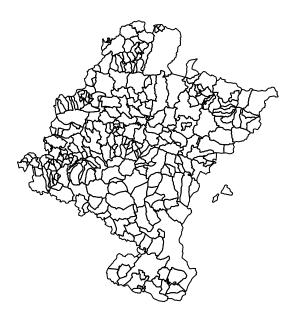
The disjointed partition is carried out taking as a reference the autonomous regions of Spain. However, a random partition is possible and is explained below.

```
#### libraries and data####
library(bigDM)
data("Carto_SpainMUN")
head(Carto_SpainMUN)
#> Simple feature collection with 6 features and 9 fields
#> geometry type: MULTIPOLYGON
#> dimension:
                   XY
#> bbox:
                   xmin: 485318 ymin: 4727428 xmax: 543317 ymax: 4779153
#> CRS:
                   +proj=utm +zone=30 +ellps=GRS80 +units=m +no_defs
#>
        ID
                                               name
                                                        lat
                                                               long
                                                                        area
#> 1 01001
                                  Alegria-Dulantzi 4742974 4742974 19913794
#> 2 01002
                                            Amurrio 4765886 4765886 96145595
#> 3 01003
                                           Aramaio 4766225 4766225 73338806
#> 4 01004
                                        Artziniega 4775453 4775453 27506468
#> 5 01006
                                            Arminon 4729884 4729884 10559721
#> 6 01008 Arrazua-Ubarrundia (San Martin de Ania) 4752714 4752714 57502811
    perimeter obs
                                                                geometry
                 4 3.0237149 Pais Vasco MULTIPOLYGON (((538259 4737...
#> 1 34372.11
#> 2 63352.32
                27 20.8456682 Pais Vasco MULTIPOLYGON (((503520 4760...
#> 3 41430.46
                 6 3.7527301 Pais Vasco MULTIPOLYGON (((533286 4759...
#> 4 22605.22
                 5 3.2093191 Pais Vasco MULTIPOLYGON (((491260 4776...
#> 5 17847.35
                 0 0.4817391 Pais Vasco MULTIPOLYGON (((509851 4727...
#> 6 64968.81
                 1 1.9643891 Pais Vasco MULTIPOLYGON (((534678 4746...
#### disjoint partition ####
carto.d <- divide_carto(carto=Carto_SpainMUN, ID.group="region", k=0)</pre>
data.d <- lapply(carto.d, function(x) sf::st_set_geometry(x, NULL))</pre>
```

The returned object is list with 15 elements composed with the autonomus regions except insular zone, Ceuta and Melilla. Plotting Navarre region:

```
plot(carto.d[[14]]$geometry, main = "Navarre")
```

Navarre



Stage 2: Object's neighbours and spatial neighbourhood matrix

In this step, we obtain a neighbours list of class ${\tt nb}$ and its associated spatial adjacency matrix computed by adding to isolated areas its nearest neighbour for each region. LCAR prior has been considered, $\xi_a \sim N({\bf 0},{\bf Q}_{\xi_d}^{-1})$, with ${\bf Q}_{\xi_d} = \tau_{\xi_d}[\lambda_{\xi_d}(D_W-W) + (1-\lambda_{\xi_d})I_{n_d}]$. The matrix ${\bf Q}_{\xi_d}$ is computed as follows:

```
#### list of object's neighbours and spatial neighbourhood matrix ####
library(Matrix)
aux <- lapply(carto.d, function(x) connect_subgraphs(x, ID.area = "ID"))
Wd <- lapply(aux, function(x) x$W)
nd <- lapply(Wd, function(x) nrow(x))
Rd <- mapply(function(x,y){Diagonal(x,colSums(y))-y}, x=nd, y=Wd)
Rd.Leroux <- mapply(function(x,y){Diagonal(x)-y}, x=nd, y=Rd)</pre>
```

Stage 3: Fit of the models

The datasets for fit the models are as follows:

```
SIMPLIFY = FALSE)
D <- length(data.INLA)</pre>
```

The hyperparameter prior assumed sdunif and lunif, but, you can select others using the CAR_INLA function explained below.

We fit the model selection criteria in a classic way,

```
#### Fit of the models ####
library(INLA)
Models <- vector("list",D)</pre>
for(i in 1:D) {
  cat(sprintf("+ Model %d of %d", i, D), "\n")
  Rs <- Rd[[i]]
  Rs.Leroux <- Rd.Leroux[[i]]</pre>
  formula \leftarrow 0 \sim f(
    ID.area,
    model = "generic1",
    Cmatrix = Rs.Leroux,
    constr = TRUE,
    hyper = list(
      prec = list(prior = sdunif),
      beta = list(prior = lunif, initial = 0)
    )
  )
  Models[[i]] <-
    inla(
      formula,
      family = "poisson",
      data = data.INLA[[i]],
      E = E,
      control.predictor = list(compute = TRUE,
                                 cdf = c(log(1))),
      control.compute = list(
        dic = TRUE,
        cpo = TRUE,
        waic = TRUE,
        config = TRUE
      ),
      control.inla = list(strategy = "gaussian")
    )
}
```

Stage 4: Merging disjoint models

The merging of disjoint models is carried out with the function mergeINLA. We have merge the models in three ways:

Computing only the models selection criteria By default, the function extracts the model selection criteria with compute.DIC = TRUE.

```
#### Merge models ####
disjoint <-
  mergeINLA(
    inla.models = Models,
    k=0,
    ID.area = "Area",
    0 = "0",
    E = "E"
    seed=1234,
    compute.fixed = FALSE,
    compute.DIC = TRUE
  )
#### DIC ####
disjoint$dic$dic
#> [1] 28205.59
#### WAIC ####
disjoint$waic$waic
#> [1] 28181.93
```

Aditionaly computing the overall fixed effect. To compute the overall fixed effect compute.fixed=TRUE.

```
#### Merge models with \alpha ####
disjoint <- mergeINLA(
  inla.models = Models,
  k = 0,
  ID.area = "muni",
  0 = "0",
  E = "E",
  seed = 1234,
  n.sample = 1000,
  compute.fixed = TRUE,
  compute.DIC = TRUE
)
summary(disjoint)</pre>
```

Immediately without steps by steps with Car_INLA. The Car_INLA function allows without the previous steps to obtain a merged model with several variants in the parameters. It is possible to retrieve elements from the previous steps such as save the sub-models with save.models = TRUE. The prior = "Leroux" specifies the prior distribution considered for the spatial random effect in this case Leroux et. al. (2000), but we can choose another: "intrinsic", "BYM" or "BYM2". The PCpriors is a functionality is developed in Simpson et. al. (2017) for scaling intrinsic Gaussian Markov Random Fields (GMRFs). If PC priors are used for the precision of a Gaussian random effect, the parameters U and α must be specified so that $P(\sigma > U) = \alpha$ or $P(\lambda_{\xi} > U) = \alpha$, as appropriate. If PCpriors = TRUE, the spatial random effects must be "intrinsic" or "BYM2", with values $(U, \alpha) = (1, 0.01)$ and $(U, \alpha) = (0.5, 0.5)$.

```
disjoint_1 <- CAR_INLA(
  carto = Carto_SpainMUN,
  ID.area = "ID",
  ID.group = "region",
  0 = "obs",</pre>
```

```
E = "exp",
  prior = "Leroux",
 model = "partition",
 k = 0,
  seed = 1234,
  strategy = "gaussian",
 PCpriors = FALSE,
  compute.fixed = FALSE,
 compute.DIC = TRUE,
  save.models = FALSE
#> STEP 1: Pre-processing data
\#> STEP 2: Fitting partition (k=0) model with INLA
#> + Model 1 of 15
#> + Model 2 of 15
#> + Model 3 of 15
#> + Model 4 of 15
#> + Model 5 of 15
#> + Model 6 of 15
#> + Model 7 of 15
#> + Model 8 of 15
#> + Model 9 of 15
#> + Model 10 of 15
#> + Model 11 of 15
#> + Model 12 of 15
#> + Model 13 of 15
#> + Model 14 of 15
#> + Model 15 of 15
#> STEP 3: Merging the results
#### DIC ####
disjoint_1$dic$dic
#> [1] 28205.6
#### WAIC ####
disjoint_1$waic$waic
#> [1] 28181.97
```

K-order neighbours model

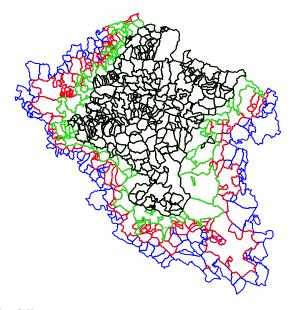
This model is characterized by the area of interest D is formed by a partition of sub-areas d_i such that $\bigcup_{i=1}^n d_i = D$ and $\bigcap_{i=1}^n d_i \neq \emptyset$. This implies that there are regions with multiple estimates. Below is how to extract the cartography list.

```
#### 1st order neighbours ####
a <-
divide_carto(
    carto = Carto_SpainMUN,
    ID.group = "region",
    k = 1,
    plot = FALSE
)

#### 2nd order neighbours ####</pre>
```

```
b <-
  divide_carto(
    carto = Carto_SpainMUN,
    ID.group = "region",
   k = 2
    plot = FALSE
#### 3rd order neighbours ####
c <-
  divide_carto(
    carto = Carto_SpainMUN,
    ID.group = "region",
   k = 3
   plot = FALSE
  )
#### Navarre with 1st, 2nd and 3rd order neighbours ####
plot(c[[14]]$geometry, border = "blue",main="Navarre with 1st, 2nd and 3rd order neighbours")
plot(b[[14]]$geometry, border = "red", add = TRUE)
plot(a[[14]]$geometry, border = "green", add = TRUE)
plot(carto.d[[14]]$geometry, add = TRUE)
```

Navarre with 1st, 2nd and 3rd order neighbours



To fit the models, we have the following:

```
#### 1st order neighbours model ####
order1<-CAR_INLA(
carto = Carto_SpainMUN,
ID.area = "ID",
ID.group = "region",
O = "obs",
E = "exp",</pre>
```

```
prior = "Leroux",
model = "partition",
k = 1,
strategy = "gaussian")
#### 2nd order neighbours model ####
order2<-CAR_INLA(</pre>
carto = Carto_SpainMUN,
ID.area = "ID",
ID.group = "region",
0 = "obs",
E = "exp",
prior = "Leroux",
model = "partition",
k = 2
strategy = "gaussian")
#### 3rd order neighbours model ####
order3<-CAR_INLA(
carto = Carto_SpainMUN,
ID.area = "ID",
ID.group = "region",
0 = "obs",
E = "exp",
prior = "Leroux",
model = "partition",
k = 3,
strategy = "gaussian")
```

Random partition

Many times, it is not possible to divide by a natural or administrative region. In this case, a random grid partition could be generated by specifying the desired rows and columns.

```
library(tmap)

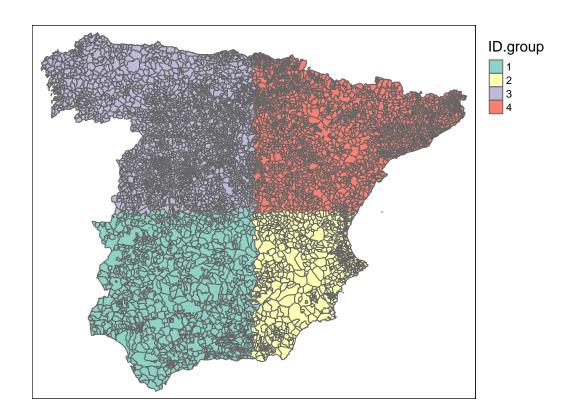
## define a random partition based on a 2x2 regular grid ##
carto_random <- random_partition(carto=Carto_SpainMUN, rows=2, columns=2)

## plot of the grouping variable 'ID.group' ##

tm_shape(carto_random) +

tm_polygons(col="ID.group") +

tm_layout(legend.outside=TRUE)</pre>
```



References

Leroux, Brian G., Xingye Lei, and Norman Breslow. 2000. "Estimation of Disease Rates in Small Areas: A New Mixed Model for Spatial Dependence." In *Statistical Models in Epidemiology, the Environment, and Clinical Trials*, 179–91. Springer New York. https://doi.org/10.1007/978-1-4612-1284-34.

Simpson, Daniel, Håvard Rue, Andrea Riebler, Thiago G. Martins, and Sigrunn H. Sørbye. 2017. "Penalising Model Component Complexity: A Principled, Practical Approach to Constructing Priors." *Statist. Sci.* 32 (1): 1–28. https://doi.org/10.1214/16-STS576.