

1. bilinear regression

we have a dataset $z[\omega]$ that has noise $\eta[\omega]$, $z[\omega] = z_0[\omega] + \eta[\omega]$,
 where $\eta[\omega]$ has power spectrum $\pi[\omega] = (1/\Omega) \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle$,
 where Ω is the total time which sets the coarse grain of frequency $\Delta\omega = \frac{2\pi}{\Omega}$ (similar to crystal unit cell). We fit this to a function $Z_{\lambda_1, \lambda_2, \dots, \lambda_N}[\omega]$ where $\lambda_1, \lambda_2, \dots, \lambda_N$ are fitting parameters. The best fit values $\Lambda_1, \Lambda_2, \dots, \Lambda_N$ of the parameters are determined by minimizing the goodness function

$$(1) \quad G = \int d\omega W[\omega] (z[\omega] - Z_{\{\lambda_i\}}[\omega])^* (z[\omega] - Z_{\{\lambda_i\}}[\omega]) + \text{c.c.} \rightarrow \min$$

where $W[\omega]$ is weight function. The best fit values $\Lambda_1, \Lambda_2, \dots, \Lambda_N$ of the parameters depend on the noise $\eta[\omega]$ in the data and therefore the values of $\Lambda_1, \Lambda_2, \dots, \Lambda_N$ will be spread over a finite range when averaged over all configuration of the noise.

Question: given the power spectrum $\pi[\omega]$ of the noise $\eta[\omega]$, what is the variance of each parameter, $\langle \langle \delta\Lambda_1^2 \rangle \rangle$, $\langle \langle \delta\Lambda_2^2 \rangle \rangle$, $\dots \langle \langle \delta\Lambda_N^2 \rangle \rangle$ averaged over the noise $\eta[\omega]$. What is the distribution function of the $\Lambda_1, \Lambda_2, \dots, \Lambda_N$ given the distribution function of the noise?

Answer :

Noise with power spectrum

(2 a) $\pi[\omega] = (1/\Omega) \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle$ has a probability distribution function

$$(2) \quad \mathbb{P}\{\eta[\omega]\} = \text{const} \times e^{-\int d\omega \frac{\eta[\omega]^* \eta[\omega]}{\pi[\omega]}}$$

the probability distribution function of the values of the best fit parameters $\Lambda_i + \delta\Lambda_i$ for small deviations $\delta\Lambda_i$ away from their average Λ_i is given by the same function where $\eta[\omega]$ is replaced by the difference of the fit $Z_{\{\lambda_i\}}[\omega]$ and the data $z[\omega]$, $\eta[\omega] \rightarrow z[\omega] - Z_{\{\lambda_i + \delta\Lambda_i\}}[\omega]$

$$(3) \quad \mathbb{P}\{\delta\Lambda_i\} = \text{const} \times e^{-\int d\omega \frac{(z[\omega] - Z_{\{\lambda_i + \delta\Lambda_i\}}[\omega])^* (z[\omega] - Z_{\{\lambda_i + \delta\Lambda_i\}}[\omega])}{\pi[\omega]}}$$

For small $\delta\Lambda_i$ we can expand the exponent

$$(4) \quad G_{\eta\{\lambda_i\}}\{z[\omega]\} = \int d\omega \frac{(z[\omega] - Z_{\{\lambda_i\}}[\omega])^* (z[\omega] - Z_{\{\lambda_i\}}[\omega])}{\pi[\omega]}$$

as

$$(5) \quad G_{\eta} \{ \Lambda_i + \delta \Lambda_i \} = A_{\eta} + (B_{\eta})_i \delta \Lambda_j + \delta \Lambda_i (C_{\eta})_{ij} \delta \Lambda_j,$$

where $(C_{\eta})_{ij}$ is the matrix of second derivatives

$$(6) \quad (C_{\eta})_{ij} = \left(\frac{d^2 G_{\eta}}{d\lambda_i d\lambda_j} \right)_{\lambda_i = \Lambda_i},$$

the variances $\langle \delta \Lambda_i \delta \Lambda_j \rangle =$

$\langle \delta \Lambda_i \delta \Lambda_j \rangle - \langle \delta \Lambda_i \rangle \langle \delta \Lambda_j \rangle$ of the best fit parameters are given by

$$(7) \quad \langle \delta \Lambda_i \delta \Lambda_j \rangle = (C_{\eta}^{-1})_{ij}$$

In particular,

$$(8) \quad \begin{aligned} \langle \delta \Lambda_1^2 \rangle &= (C_{\eta}^{-1})_{11} \\ \langle \delta \Lambda_2^2 \rangle &= (C_{\eta}^{-1})_{22} \\ &\dots \\ \langle \delta \Lambda_N^2 \rangle &= (C_{\eta}^{-1})_{NN} \end{aligned}$$

2. noise in the calorimeter response function $R[\omega]$

the calorimeter response $R[\omega]$ is

$$\begin{aligned} (2.1) \quad R[\omega] &= \frac{\Delta T_{\text{Therm}}[\omega]}{\Delta P_{\text{heater}}[\omega]} \\ &= \frac{\Delta T_{\text{Therm}}[\omega]}{\Delta V_{\text{heater}}[\omega/2] \Delta I_{\text{heater}}[\omega/2]} \\ &= \left(\frac{d \log T_{\text{therm}}}{d \log R_{\text{therm}}} \right) \frac{\Delta V_{\text{Therm}}[\omega + \omega_1] / \Delta I_{\text{Therm}}[\omega_1]}{\Delta V_{\text{heater}}[\omega/2] \Delta I_{\text{heater}}[\omega/2]} \end{aligned}$$

where $\alpha = \left(\frac{d \log T_{\text{therm}}}{d \log R_{\text{therm}}} \right)$ is a thermometer calibration factor

each of the four quantities, $\Delta V_{\text{Therm}}[\omega + \omega_1]$, $\Delta I_{\text{Therm}}[\omega_1]$, $\Delta V_{\text{heater}}[\omega/2]$,

$\Delta I_{\text{heater}}[\omega/2]$ are independently measured by the lockin -- therefore,

each accumulating its noise-- and the calorimeter response =

Eq. (2.1) is calculated digitally using results of these four measurements.

We now consider noise in $R[\omega]$ as a result of the noise in $\Delta V_{\text{Therm}}[\omega + \omega_1]$, $\Delta I_{\text{Therm}}[\omega_1]$, $\Delta V_{\text{heater}}[\omega/2]$.

The noise in the 4 quantities is not uncorrelated because of the coupling on the calorimeter and in the wiring between calorimeter and the lockin. First,

assume that only one of the four , $\Delta V_{\text{Therm}}[\omega + \omega_1]$, is noisy. Then the noise $\eta[\omega]$ in $R[\omega]$ is

$$(2.2) \quad \eta[\omega] = \delta R[\omega] = \delta V_{\text{Therm}}[\omega + \omega_1] \frac{\alpha / \Delta I_{\text{Therm}}[\omega_1]}{\Delta V_{\text{heater}}[\omega / 2] \Delta I_{\text{heater}}[\omega / 2]}$$

the noise $\delta V_{\text{Therm}}[\omega + \omega_1]$ includes

- (1) the thermal noise on the thermometer resistor (probably small at cryo temperatures)
- (2) the vibration - in - field and antenna noise in the < calorimeter → lockin > wires
- (2) amplifier noise on the channels processing $\Delta V_{\text{Therm}}[\omega + \omega_1]$

we need to estimate the power spectrum $\pi_{V_{\text{Therm}}}[\omega] = \frac{1}{\Omega} \langle \langle \delta V_{\text{Therm}}[\omega + \omega_1]^* \delta V_{\text{Therm}}[\omega + \omega_1] \rangle \rangle$,

of $\delta V_{\text{Therm}}[\omega + \omega_1]$. The assumption of the white noise is not satisfactory as in the calorimeter spectroscopy the frequency changes by 10^5 over the measured frequency range.

On top of that, $\eta[\omega]$ has an additional ω - dependent factor, $\eta[\omega] = \phi_{\text{Therm}}[\omega] \delta V_{\text{Therm}}[\omega]$, where

$$(2.3) \quad \phi_{\text{Therm}}[\omega] = \frac{\alpha / \Delta I_{\text{Therm}}[\omega_1]}{\Delta V_{\text{heater}}[\omega / 2] \Delta I_{\text{heater}}[\omega / 2]}$$

which is not constant in frequency because the amplitudes of the AC components of $\Delta I_{\text{heater}}[\omega / 2]$ and $\Delta I_{\text{Therm}}[\omega_1]$ are continuously adjusted in the experiment as frequency is swept to maintain comfortable S / N conditions

Overall, the power spectrum $\pi[\omega] =$

$(1 / \Omega) \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle$ in the calorimeter response $R[\omega]$ is given by

$$(2.4) \quad \pi[\omega] = (1 / \Omega) \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle = \left| \phi_{\text{Therm}}[\omega] \right|^2 \frac{1}{\Omega} \langle \langle \delta V_{\text{Therm}}[\omega + \omega_1]^* \delta V_{\text{Therm}}[\omega + \omega_1] \rangle \rangle$$

If all four voltages are noisy then (2.4) will have a corresponding contribution from noise in each of them as well as terms coming from the cross - noise between different channels.

3. question -- what is the power spectrum of the noise in the thermometer voltage

$$\frac{1}{\Omega} \langle \langle \delta V_{\text{Therm}}[\omega + \omega_1]^* \delta V_{\text{Therm}}[\omega + \omega_1] \rangle \rangle$$

after lockin amplification ?

$$? \frac{\text{nV}}{\sqrt{\text{Hz}}} \text{ or equivalent temperature ? What is its frequency}$$

dependence in the range between 0.01 Hz and 4 kHz ?

4. Derivation of Eq.(3)

For a given data $z[\omega]$, the best fit value $\Lambda_1, \Lambda_1, \dots, \Lambda_N$, of parameters $\lambda_1, \lambda_2, \dots, \lambda_N$ can be calculated as an average value over the probability distribution

$$(10) \quad P[\{\lambda_1, \dots, \lambda_N\}] = e^{\frac{G\theta_T\{z[\omega]\}}{T}} e^{-\frac{G_{\{\lambda_i\}}\{z[\omega]\}}{T}}$$

$$e^{-\frac{G\theta_T\{z[\omega]\}}{T}} = \int d\lambda_1 \dots d\lambda_N e^{-\frac{G_{\lambda_1, \lambda_2, \dots, \lambda_N}\{z[\omega]\}}{T}}$$

where T is a parameter. The average value of each parameter $\langle \lambda_i \rangle_T$ at the temperature T is determined from

$$(11) \quad \langle \lambda_i \rangle_T = e^{\frac{G\theta_T\{z[\omega]\}}{T}} \int d\lambda_1 \dots d\lambda_N \lambda_i e^{-\frac{1}{T} G_{\{\lambda_i\}}\{z[\omega]\}}, \text{ etc.}$$

the best fit values $\Lambda_1, \Lambda_1, \dots, \Lambda_N$, are equal to the average value of parameters $\lambda_1, \lambda_2, \dots$ in the limit $T \rightarrow 0$.

$$(12) \quad \Lambda_i = \lim_{T \rightarrow 0} \langle \lambda_i \rangle_T$$

The averages $\langle \lambda_i \rangle_T$ -- as well as the best fit Λ_i -- depend on the data and noise $z[\omega] + \eta[\omega]$ and therefore will be distributed over the finite range even in the limit of $T \rightarrow 0$. For a given noise configuration $\eta[\omega]$,

$$(13) \quad P[\{\lambda_1, \dots, \lambda_{1N}\}; T, \eta[\omega]] = e^{\frac{G\theta_T \{z[\omega] + \eta[\omega]\}}{T}} e^{-\frac{G_{\{\lambda_i\}} \{z[\omega] + \eta[\omega]\}}{T}}$$

Because noise is random,

the values of the Λ_i are themselves distributed over a finite range when we consider all configurations of noise. The probability distribution of the noise is

$$(14) \quad \mathbb{P}\{\eta[\omega]\} = \text{const} e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}}$$

where $\pi[\omega] = \langle \eta[\omega]^* \times \eta[\omega] \rangle$ is proportional to the power spectrum. The probability distribution $\mathbb{P}\{\Lambda_i\}$ for the best fit values Λ_i is

$$(15) \quad \mathbb{P}\{\Lambda_i\} = \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* \mathbb{P}\{\eta[\omega]\} \times P[\{\lambda_i\}; T, \eta[\omega]] \\ = \text{const} \times \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{\frac{G\theta_T \{z[\omega] + \eta[\omega]\}}{T}} e^{-\frac{G_{\{\lambda_i\}} \{z[\omega] + \eta[\omega]\}}{T}}$$

in the limit of $T \rightarrow 0$ the function $G\theta_T \{z[\omega] + \eta[\omega]\}$ depends very weakly on $\eta[\omega]$ (see SI 1) and therefore we can absorb it into the definition of the const. Therefore, at small T the probability distribution $\mathbb{P}\{\Lambda_i\}$ is

$$(16) \quad \mathbb{P}\{\Lambda_i\} = e^{\frac{F_T}{T}} \times \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{-\frac{G_{\{\lambda_i\}} \{z[\omega] + \eta[\omega]\}}{T}} \\ e^{-\frac{F_T}{T}} = \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* \int d\lambda_1 \dots d\lambda_{1N} e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{-\frac{G_{\{\lambda_i\}} \{\eta[\omega]\}}{T}}$$

and all averages are done with

$$(17) \quad \langle \dots \rangle = \int d\lambda_1 \dots d\lambda_{1N} \mathbb{P}\{\Lambda_i\} (\dots)$$

To evaluate $\mathbb{P}\{\Lambda_i\}$ we need to integrate over all noise configurations

$$(18) \quad \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{-\frac{G_{\{\lambda_i\}} \{z[\omega] + \eta[\omega]\}}{T}} = \\ \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{-\int d\omega W[\omega] \frac{(\eta[\omega] - (Z[\omega] - z[\omega]))^* (\eta[\omega] - (Z[\omega] - z[\omega]))}{T}} =$$

The integral is a product of simple gaussian integrals at each ω , which are evaluated using an identity

$$(19) \quad \int d\eta d\eta^* e^{-\frac{\eta^* \eta}{T_2} - \frac{(\eta - a)^* (\eta - a)}{T_1}} = \int d\eta d\eta^* e^{-\frac{\frac{T_1}{T_2} \eta^* \eta - (\eta - a)^* (\eta - a)}{T_1}} \\ = \int dx dy e^{-\frac{\frac{T_1}{T_2} (\eta_1^2 + \eta_2^2) + (\eta_1 - a_1 - i\eta_2 - a_1 + ia_2) (\eta_1 + ia_2 - a_1 - ia_2)}{T_1}} \\ = \int dx dy e^{-\frac{\frac{T_1}{T_2} (\eta_1^2 + \eta_2^2) + 2 \{ (a_1 - \eta_1)^2 + (a_2 - \eta_2)^2 \}}{T_1}}$$

$$\begin{aligned}
&= \int dx dy e^{-\frac{\frac{2}{T_1+2T_2} a_1^2 + \left(2 + \frac{T_1}{T_2}\right) \left(\eta_1 - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2 + (\eta_1 \rightarrow \eta_2, a_1 \rightarrow a_2)}{T_1}} \\
&= \frac{\pi T_1}{\left(2 + \frac{T_1}{T_2}\right)} e^{-\frac{2(a_1^2 + a_2^2)}{T_1+2T_2}}
\end{aligned}$$

Note that the parameters T_1 and T_2 enter as a sum only in the exponent

(the prefactor is not important). This reflects the fact that when T_1 goes to 0, the value of a is very close to η , however, distribution of a is still broad because η is distributed broadly. In other words, the temperature T rather than being zero, is set by the power spectrum $\pi[\omega]$.

We have

$$\begin{aligned}
(20) \quad \lim_{T \rightarrow 0} \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\eta[\omega]^* \cdot \eta[\omega]}{\pi[\omega]}} e^{-\frac{G_{\{\lambda_i\}}(\eta[\omega])}{T}} &= \\
\lim_{T \rightarrow 0} \int \mathcal{D}\eta[\omega] \times \mathcal{D}\eta[\omega]^* e^{-\int d\omega \frac{\frac{T/W[\omega]}{\pi[\omega]} \eta[\omega]^* \cdot \eta[\omega] + \left(\eta[\omega] - (Z_{\{\lambda_i\}}[\omega] - z[\omega])\right)^* \left(\eta[\omega] - (Z_{\{\lambda_i\}}[\omega] - z[\omega])\right)}{T/W[\omega]}} &= \\
\text{const } e^{-\int d\omega \frac{(Z_{\{\lambda_i\}}[\omega] - z[\omega])^* (Z_{\{\lambda_i\}}[\omega] - z[\omega])}{\pi[\omega]}} &
\end{aligned}$$

Finally, the probability distribution $\mathbb{P}\{\Lambda_i\}$ for the parameters λ_i when $z[\omega]$ is characterized by the noise power spectrum $\pi[\omega]$ is

$$\begin{aligned}
(21) \quad \mathbb{P}\{\Lambda_i\} &= e^F e^{-G_{\eta\{\lambda_i\}}\{z[\omega]\}}, \\
e^{-F} &= \int d\lambda_1 \dots d\lambda_N e^{-G_{\eta\{\lambda_i\}}\{z[\omega]\}},
\end{aligned}$$

where the "noise-goodness" $G_{\eta\{\lambda_i\}}\{z[\omega]\}$ is

determined by the properties of the noise alone,

$$(22) \quad G_{\eta\{\lambda_i\}}\{z[\omega]\} = \int d\omega \frac{1}{\pi[\omega]} (Z_{\{\lambda_i\}}[\omega] - z[\omega])^* (Z_{\{\lambda_i\}}[\omega] - z[\omega])$$

Note that $G_{\eta\{\lambda_i\}}\{z[\omega]\}$ is equivalent to a specific choice of Weight

function in the goodness function discussed above. The noise spectrum $\langle \eta[\omega]^* \times \eta[\omega] \rangle = \pi[\omega]$ simply sets the (frequency-- dependent) value of T in the representation of the best fit value of parameters.

Note that one could obtain $\mathbb{P}\{\Lambda_i\}$ by simply replacing $\eta[\omega]$ with $Z_{\{\lambda_i\}}[\omega] - z[\omega]$ in the noise distribution function $\mathbb{P}\{\eta[\omega]\}$,

$$(21a) \quad \mathbb{P}\{\Lambda_i\} = \mathbb{P}\{\eta[\omega] \rightarrow Z_{\{\lambda_i\}}[\omega] - z[\omega]\}$$

The averages $\langle \langle \Lambda_i \rangle \rangle$ and the variances $\langle \langle \delta \Lambda_i^2 \rangle \rangle$ of the fit parameters

are determined in a standard way by their probability distribution $\mathbb{P}\{\Lambda_i\}$

$$\begin{aligned}
(23) \quad \langle \langle \delta \Lambda_i^2 \rangle \rangle &= \langle \Lambda_i^2 \rangle - \langle \Lambda_i \rangle^2, \\
\langle \Lambda_i^2 \rangle &= \int d\lambda_1 \dots d\lambda_{1N} (\lambda_i)^2 \wp \{ \Lambda_i \}, \\
\langle \Lambda_i \rangle &= \int d\lambda_1 \dots d\lambda_{1N} (\lambda_i) \wp \{ \Lambda_i \}
\end{aligned}$$

We will now calculate the variance $\langle \langle \delta \Lambda_i^2 \rangle \rangle$. Near the minimum we expand $G_\eta \{ \Lambda_i + \delta \Lambda_i \}$ as

$$(24) \quad G_\eta \{ \Lambda_i + \delta \Lambda_i \} = A_\eta + (B_\eta)_i \delta \Lambda_j + \delta \Lambda_i (C_\eta)_{ij} \delta \Lambda_j,$$

where $(C_\eta)_{ij}$ is the matrix of second derivatives

$$(25) \quad (C_\eta)_{ij} = \frac{d^2 G_\eta}{d\delta \Lambda_i d\delta \Lambda_j}$$

of G_η near its minimum. We note that if we used a different

goodness function G_W to find the best fit values Λ_i for parameters λ_i ,

$$(26) \quad \Lambda_i \leftarrow \left\{ \frac{dG_W}{d\Lambda_i} = 0 \right\}$$

then these values are not at the minimum of $G_\eta \{ \Lambda_i \}$ and therefore $(B_\eta)_i$ are nonzero.

To calculate the averages $\langle \langle \delta \Lambda_i^2 \rangle \rangle$ we use a characteristic function $X \{ \xi_i \}$ defined via

$$(27) \quad e^{-X \{ \xi_i \}} = \int d\delta \Lambda_1 \dots d\delta \Lambda_N e^{-G_\eta \{ \Lambda_i + \delta \Lambda_i \} + \sum \delta \Lambda_i \xi_i}$$

$$\begin{aligned}
(28) \quad X \{ \xi_i \} &= A_\eta \\
&= -\log \int d\delta \Lambda_1 \dots d\delta \Lambda_N e^{-(B_\eta)_i \delta \Lambda_j - \delta \Lambda_i (C_\eta)_{ij} \delta \Lambda_j + \sum \delta \Lambda_i \xi_i} \\
&= -\log \int d\delta \Lambda_1 \dots d\delta \Lambda_N e^{-\delta \Lambda_i (C_\eta)_{ij} \delta \Lambda_j + (\sum \xi_i - (B_\eta)_i) \delta \Lambda_j} \\
&= -\log \int d\delta \Lambda_1 \dots d\delta \Lambda_N e^{-\delta \Lambda_i (C_\eta)_{ij} \delta \Lambda_j + (\sum \xi_i - (B_\eta)_i) \delta \Lambda_j} \\
&= -\log \left(\sqrt{\frac{\pi}{\det (C_\eta)}} e^{-\frac{1}{4} (\sum \xi_i - (B_\eta)_i) ((C_\eta)^{-1})_{ij} (\sum \xi_j - (B_\eta)_j)} \right) \\
&= \frac{1}{4} (\sum \xi_i - (B_\eta)_i) ((C_\eta)^{-1})_{ij} (\sum \xi_j - (B_\eta)_j) + \frac{1}{2} \log \det (C_\eta)
\end{aligned}$$

The averages are calculated as derivatives of the characteristic function $X \{ \xi_i \}$,

$$(29) \quad \langle \delta \Lambda_i \rangle = \frac{1}{\sum} \left(e^{X \{ \xi_i \}} \frac{d}{d\xi_j} e^{-X \{ \xi_i \}} \right)_{\xi \rightarrow 0} = -\frac{1}{\sum} \left(\frac{dX}{d\xi_j} \right)_{\xi \rightarrow 0}$$

$$\begin{aligned}
(30) \quad & \langle \delta\Lambda_i \delta\Lambda_j \rangle \\
&= \left(\frac{1}{\hbar} \right)^2 \left(e^{x\{\xi_i\}} \frac{d}{d\xi_i} \frac{d}{d\xi_j} e^{-x\{\xi_i\}} \right)_{\xi \rightarrow 0} \\
&= \left(\frac{1}{\hbar} \right)^2 \left(e^{x\{\xi_i\}} \frac{d}{d\xi_i} \left(-\frac{dx\{\xi_i\}}{d\xi_j} e^{-x\{\xi_i\}} \right) \right)_{\xi \rightarrow 0} \\
&= \left(\frac{1}{\hbar} \right)^2 \left(e^{x\{\xi_i\}} \left(-\frac{d^2 x\{\xi_i\}}{d\xi_i d\xi_j} + \frac{dx\{\xi_i\}}{d\xi_i} \frac{dx\{\xi_i\}}{d\xi_j} \right) e^{-x\{\xi_i\}} \right)_{\xi \rightarrow 0} \\
&= \left(\frac{1}{\hbar} \right)^2 \left(-\frac{d^2 x}{d\xi_i d\xi_j} + \frac{dx}{d\xi_i} \frac{dx}{d\xi_j} \right)_{\xi \rightarrow 0} \\
&= \left(\frac{1}{\hbar} \right)^2 \left(-\frac{d^2 x\{\xi_i\}}{d\xi_i d\xi_j} \right)_{\xi \rightarrow 0} + \left(\frac{1}{\hbar} \right)^2 \langle \delta\Lambda_i \rangle \langle \delta\Lambda_j \rangle
\end{aligned}$$

$$\begin{aligned}
(31) \quad & \langle \langle \delta\Lambda_i \delta\Lambda_j \rangle \rangle \\
&= \langle (\delta\Lambda_i - \langle \delta\Lambda_i \rangle) (\delta\Lambda_j - \langle \delta\Lambda_j \rangle) \rangle \\
&= \langle \delta\Lambda_i \delta\Lambda_j \rangle - \langle \delta\Lambda_i \rangle \langle \delta\Lambda_j \rangle \\
&= \left(-\frac{d^2 x}{d\xi_i d\xi_j} + \frac{dx}{d\xi_i} \frac{dx}{d\xi_j} \right)_{\xi \rightarrow 0} - \left(-\left(\frac{dx}{d\xi_j} \right)_{\xi \rightarrow 0} \right) \left(-\left(\frac{dx}{d\xi_i} \right)_{\xi \rightarrow 0} \right) \\
&= \left(\frac{1}{\hbar} \right)^2 \left(-\frac{d^2 x}{d\xi_i d\xi_j} \right)_{\xi \rightarrow 0} \\
&= \frac{1}{4} \left((C_\eta)^{-1} \right)_{ij}
\end{aligned}$$

SI0

$$\delta G = \delta\Lambda A \delta\Lambda + B \delta\Lambda$$

$$e^{-\frac{x\{\xi_i\}}{T}} = \int d\delta\Lambda_1 d\delta\Lambda_2 \dots d\delta\Lambda_N e^{-\frac{\delta\Lambda A \delta\Lambda + B \delta\Lambda}{T} + \hbar \delta\Lambda_i \xi_i}$$

$$\delta\Lambda A \delta\Lambda + B \delta\Lambda + \hbar \delta\Lambda \xi =$$

$$\left(\delta\Lambda + \frac{1}{2} A^{-1} B + \frac{\hbar}{2} A^{-1} \xi \right) A \left(\delta\Lambda + \frac{1}{2} A^{-1} B + \frac{\hbar}{2} A^{-1} \xi \right) - \frac{1}{4} (B + \hbar \xi) A^{-1} (B + \hbar \xi) =$$

$$X = \frac{1}{4} (B + \hbar \xi) A^{-1} (B + \hbar \xi)$$

5. dependence of $e^{-\frac{G_0 T \{z[\omega] + \eta[\omega]\}}{T}}$ on $\eta[\omega]$?

$$\begin{aligned}
& (S1) \quad e^{-\frac{G\theta_T \{z[\omega] + \eta[\omega]\}}{T}} \\
& = \int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\frac{G_{\lambda_1, \lambda_2, \dots, \lambda_N} \{z[\omega] + \eta[\omega]\}}{T}} \\
& = \int d\lambda_1 d\lambda_2 \dots d\lambda_N \\
& \quad e^{-\frac{1}{T} \int d\omega W[\omega] \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] - \eta[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)^* \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] - \eta[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)} \\
& = \int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\frac{1}{T} \int d\omega W[\omega] \times (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j)^* \times (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j)}
\end{aligned}$$

we define Λ_i as the best fit values of noiseless data ,

$\eta[\omega] = \theta$, defined by N equations for the extremum

$$\begin{aligned}
\frac{d G_{\lambda_1, \lambda_2, \dots, \lambda_N} \{z[\omega] + \eta[\omega]\}}{d \lambda_i} \Big|_{\Lambda_i} &= \int d\omega W[\omega] \left(\frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} \right)^* (Z_{\{\Lambda_i\}}[\omega] - z[\omega]) + c.c. \Big|_{\Lambda_i} \\
&= \int d\omega W[\omega] (B_i[\omega]^* A[\omega] + B_i[\omega] A[\omega]^*) \Big|_{\Lambda_i} \\
&= \theta
\end{aligned}$$

$$\text{where } A[\omega] = Z_{\{\Lambda_i\}}[\omega] \text{ and } B_i[\omega] = \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} .$$

We write

$$\begin{aligned}
& \int d\omega W[\omega] \times (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j)^* \times \\
& \quad (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j) \\
& = \int d\omega W[\omega] \left(\begin{aligned} & (A[\omega] - \eta[\omega])^* \times (A[\omega] - \eta[\omega]) \\ & + (A[\omega] - \eta[\omega])^* (B_i[\omega] d\Lambda_i) + c.c. \\ & + (B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega]) d\Lambda_i d\Lambda_j + c.c. \end{aligned} \right) \\
& = \int d\omega W[\omega] \left(\begin{aligned} & (A[\omega] - \eta[\omega])^* \times (A[\omega] - \eta[\omega]) \\ & - (\eta[\omega]^* B_i[\omega] + c.c.) d\Lambda_i \\ & + (B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c.) d\Lambda_i d\Lambda_j \end{aligned} \right) \\
& = \int d\omega W[\omega] \left(\begin{aligned} & (A[\omega] - \eta[\omega])^* \times (A[\omega] - \eta[\omega]) \\ & + (B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c.) \\ & \left(d\Lambda_i - \frac{1}{2} (B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c.)^{-1} (\eta[\omega]^* B_i[\omega] + c.c.) \right) \end{aligned} \right)
\end{aligned}$$

$$\begin{aligned}
& \left(d\Lambda_j - \frac{1}{2} \left(B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c. \right)^{-1} (\eta[\omega]^* B_i[\omega] + c.c.) \right) \\
& - \frac{1}{4} \left(\eta[\omega]^* B_j[\omega] + c.c. \right) \left(B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c. \right)^{-1} \\
& \left(\eta[\omega]^* B_j[\omega] + c.c. \right) \\
& \left. \right)
\end{aligned}$$

Therefore,

$$\begin{aligned}
(S1) &= \int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\frac{1}{T} \int d\omega W[\omega] \times (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j)^* \times (A[\omega] - \eta[\omega] + B_i[\omega] d\Lambda_i + C_{ij}[\omega] d\Lambda_i d\Lambda_j)} = \\
&= \sqrt{\frac{1}{\text{Det}_{ij} \left[\frac{1}{T} \int d\omega W[\omega] \left(B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c. \right) \right]}} \times \\
&\times e^{-\frac{1}{T} \int d\omega W[\omega] \times (A[\omega] - \eta[\omega])^* \times (A[\omega] - \eta[\omega])} \\
&\times e^{-\frac{1}{T} \frac{1}{4} \int d\omega W[\omega] \times (\eta[\omega]^* B_j[\omega] + c.c.) \left(B_i[\omega]^* B_j[\omega] + (A[\omega] - \eta[\omega])^* C_{ij}[\omega] + c.c. \right)^{-1} (\eta[\omega]^* B_j[\omega] + c.c.)}
\end{aligned}$$

The last two factors are equivalent to gaussian distribution of $\eta[\omega]$

$\text{const} \times e^{-\int d\omega \frac{\eta[\omega]^* \eta[\omega]}{p[\omega]}}$, where $p[\omega] \propto T$ and, therefore, it vanishes in the $T \rightarrow 0$ limit.

Therefore, we can assume that the normalization

factor $e^{-\frac{G\theta_T \{z[\omega] + \eta[\omega]\}}{T}}$ depends weakly on $\eta[\omega]$ in the $T \rightarrow 0$ limit.

$$6. \int d\eta d\eta^* e^{-\frac{\frac{T_1}{T_2} \eta^* \eta - (\eta - a)^* (\eta - a)}{T_1}}$$

the integral is done using identity

$$\begin{aligned}
& \int d\eta d\eta^* e^{-\frac{\frac{T_1}{T_2} \eta^* \eta - (\eta - a)^* (\eta - a)}{T_1}} \\
&= \int dx dy e^{-\frac{\frac{T_1}{T_2} (x^2 + y^2) + (x - i y - a_1 + i a_2) (x + i y - a_1 - i a_2) + c.c.}{T_1}} \\
&= \int dx dy e^{-\frac{\frac{T_1}{T_2} (x^2 + y^2) + 2 ((a_1 - x)^2 + (a_2 - y)^2)}{T_1}}
\end{aligned}$$

$$\begin{aligned}
&= \int dx dy e^{-\frac{\frac{2a_1^2 T_1}{T_1+2T_2} + \left(2 + \frac{T_1}{T_2}\right) \left(x - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2 + (x-y, a_1 \rightarrow a_2)}{T_1}} \\
&= \frac{\pi T_1}{\left(2 + \frac{T_1}{T_2}\right)} e^{-\frac{2(a_1^2 + a_2^2)}{T_1+2T_2}}
\end{aligned}$$

$$\begin{aligned}
&(x - iy - a_1 + ia_2)(x + iy - a_1 - ia_2) + \\
&(x + iy - a_1 - ia_2)(x - iy - a_1 + ia_2) \\
&2((a_1 - x)^2 + (a_2 - y)^2) \\
&\frac{T_1}{T_2}(x^2 + y^2) + 2((a_1 - x)^2 + (a_2 - y)^2)
\end{aligned}$$

$$\left(2 + \frac{T_1}{T_2}\right) \left(\frac{2a_1^2 T_1 T_2}{(T_1 + 2T_2)^2} + \left(x - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2\right)$$

$$\left(2 + \frac{T_1}{T_2}\right) \left(\frac{2a_1^2 T_1 T_2}{(T_1 + 2T_2)^2} + \left(x - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2\right) = \left(\frac{T_1}{T_2} x^2 + 2(a_1 - x)^2\right)$$

0

$$\left(2 + \frac{T_1}{T_2}\right) \left(\frac{2a_1^2 T_1 T_2}{(T_1 + 2T_2)^2}\right) = \frac{2a_1^2 T_1}{T_1 + 2T_2}$$

0

$$\frac{2a_1^2 T_1}{T_1 + 2T_2} + \left(2 + \frac{T_1}{T_2}\right) \left(x - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2$$

$$(16) \int dx dy e^{-\frac{\frac{T_1}{T_2}(x^2 + y^2) + 2((a_1 - x)^2 + (a_2 - y)^2)}{T_1}} =$$

$$\int dx dy e^{-\frac{\frac{2a_1^2 T_1}{T_1+2T_2} + \left(2 + \frac{T_1}{T_2}\right) \left(x - \frac{2a_1}{\left(2 + \frac{T_1}{T_2}\right)}\right)^2 + (x-y, a_1 \rightarrow a_2)}{T_1}} = \frac{\pi T_1}{\left(2 + \frac{T_1}{T_2}\right)} e^{-\frac{2(a_1^2 + a_2^2)}{T_1+2T_2}}$$

7. × 2 x2 matrix diagonalization and ellipse

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$$\begin{pmatrix} \cos[\phi] \\ \sin[\phi] \end{pmatrix}^T \cdot \begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} \cdot \begin{pmatrix} \cos[\phi] \\ \sin[\phi] \end{pmatrix} // \text{Tr} // \text{FullSimplify}$$


$$a + b \cos[2(\theta - \phi)]$$


$$\begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} // \text{Eigensystem} // \text{FullSimplify}$$


$$\{\{a - b, a + b\}, \{-\tan[\theta], 1\}, \{\cot[\theta], 1\}\}$$


$$\{\{a - b, a + b\}, \{-\sin[\theta], \cos[\theta]\}, \{\cos[\theta], \sin[\theta]\}\}$$


$$\left( \begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \sqrt{a-b} & 0 \\ 0 & \sqrt{a+b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\phi] \\ \sin[\phi] \end{pmatrix} \right)^T \cdot \begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} \cdot$$


$$\begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \sqrt{a-b} & 0 \\ 0 & \sqrt{a+b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\phi] \\ \sin[\phi] \end{pmatrix} // \text{Tr} // \text{FullSimplify}$$


$$(a - b)(a + b)$$


$$\left( \begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \frac{\sqrt{a-b}}{\sqrt{(a-b)(a+b)}} & 0 \\ 0 & \frac{\sqrt{a+b}}{\sqrt{(a-b)(a+b)}} \end{pmatrix} \right)^T \cdot \begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} \cdot$$


$$\begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \frac{\sqrt{a-b}}{\sqrt{(a-b)(a+b)}} & 0 \\ 0 & \frac{\sqrt{a+b}}{\sqrt{(a-b)(a+b)}} \end{pmatrix} // \text{FullSimplify} // \text{MatrixForm}$$


$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$


$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$


$$\left( \left( \begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \frac{1}{\sqrt{a+b}} & 0 \\ 0 & \frac{1}{\sqrt{a-b}} \end{pmatrix} \right)^T // \text{Inverse} \right) -$$


$$\left( \begin{pmatrix} \sqrt{a+b} & 0 \\ 0 & \sqrt{a-b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\theta] & \sin[\theta] \\ -\sin[\theta] & \cos[\theta] \end{pmatrix} \right)^T // \text{FullSimplify} // \text{MatrixForm}$$


$$\begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix}$$


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$$\begin{aligned}
& \left(\begin{pmatrix} \sqrt{a+b} & 0 \\ 0 & \sqrt{a-b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\theta] & \sin[\theta] \\ -\sin[\theta] & \cos[\theta] \end{pmatrix} \right)^T \cdot \\
& \left(\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \cdot \left(\begin{pmatrix} \sqrt{a+b} & 0 \\ 0 & \sqrt{a-b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\theta] & \sin[\theta] \\ -\sin[\theta] & \cos[\theta] \end{pmatrix} \right) \right) - \\
& \begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} // \text{FullSimplify} // \text{MatrixForm} \\
& \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix} \\
& \begin{pmatrix} a + b \cos[2\theta] & b \sin[2\theta] \\ b \sin[2\theta] & a - b \cos[2\theta] \end{pmatrix} = \left(\begin{pmatrix} \sqrt{a+b} & 0 \\ 0 & \sqrt{a-b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\theta] & \sin[\theta] \\ -\sin[\theta] & \cos[\theta] \end{pmatrix} \right)^T \cdot \\
& \left(\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \cdot \left(\begin{pmatrix} \sqrt{a+b} & 0 \\ 0 & \sqrt{a-b} \end{pmatrix} \cdot \begin{pmatrix} \cos[\theta] & \sin[\theta] \\ -\sin[\theta] & \cos[\theta] \end{pmatrix} \right) \right)
\end{aligned}$$

errors

$$\begin{aligned}
\frac{\delta G}{G} &= \left(\frac{1}{a} \frac{\delta \lambda}{\lambda} \right)^2 \\
\begin{pmatrix} \frac{\delta \lambda 1}{\lambda 1} \\ \frac{\delta \lambda 2}{\lambda 2} \end{pmatrix} &= \\
\frac{1}{30} \begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \frac{1}{\sqrt{a+b}} & 0 \\ 0 & \frac{1}{\sqrt{a-b}} \end{pmatrix} \cdot \begin{pmatrix} 1 \\ 1 \end{pmatrix} &= \frac{1}{30} \begin{pmatrix} \cos[\theta] & -\sin[\theta] \\ \sin[\theta] & \cos[\theta] \end{pmatrix} \cdot \begin{pmatrix} \frac{1}{\sqrt{a+b}} \\ \frac{1}{\sqrt{a-b}} \end{pmatrix}
\end{aligned}$$

check

$$\begin{aligned}
& \int d\eta d\eta^* e^{-\frac{\eta^* \eta}{T_2} - \frac{1}{2} \xi^* \eta} \\
&= \int d\eta d\eta^* e^{-\frac{\left(\eta - \frac{T_2}{2} \frac{1}{2} \xi\right)^* \left(\eta - \frac{T_2}{2} \frac{1}{2} \xi\right)}{T_2} + \frac{T_2}{4} \xi^* \xi} \\
&= \pi T_2 e^{\frac{T_2}{4} \xi^* \xi}
\end{aligned}$$

8. digression on $\pi[\omega]$

Noise with power spectrum $\pi[\omega] = \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle$ has a probability distribution function

$$(2) \quad \mathbb{P} \{ \eta[\omega] \} = \text{const} \times e^{-\int d\omega \frac{\eta[\omega]^* \eta[\omega]}{\pi[\omega]}}$$

whats missing is the total time . We assume that
the total time is Ω -- this is analogous to the crystal size

$$\eta[\omega_n] = \int dt \eta[t] e^{-i \omega_n t}$$

$$\eta[t] = \int \frac{d\omega}{2\pi} \eta[\omega] e^{i \omega t} = \frac{\Delta\omega}{2\pi} \sum \eta[\omega_n] e^{i \omega_n t}$$

the size of the "unit cell" in time is $\Delta\omega = \frac{1}{2\pi\Omega}$ -- this provides regularization

$\langle \langle \eta[t_2] \times \eta[t_1] \rangle \rangle$ is intensive (not proportional to "volume" Ω)

$\langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle$ is extensive (yes proportional to "volume" Ω)

$$\langle \text{power spectrum} \rangle = G[\omega_n] = \frac{1}{\Omega} \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle = (2\pi\Delta\omega) \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle$$

$\langle \text{power spectrum} \rangle$

is intensive (not proportional to "volume" Ω) -- it is "per unit volume"

$$\begin{aligned} \langle \langle \eta[t_2] \times \eta[t_1] \rangle \rangle &= \left(\frac{\Delta\omega}{2\pi} \right)^2 \sum \langle \langle \eta[\omega_n]^* \eta[\omega_m] \rangle \rangle e^{i(\omega_n + \omega_m) \frac{(t_1-t_2)}{2} + (\omega_n - \omega_m) \frac{(t_1+t_2)}{2}} \\ &= \left(\frac{\Delta\omega}{2\pi} \right)^2 \sum \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle \delta_{nm} e^{i(\omega_n + \omega_m) \frac{(t_1-t_2)}{2} + (\omega_n - \omega_m) \frac{(t_1+t_2)}{2}} \\ &= \left(\frac{\Delta\omega}{2\pi} \right)^2 \sum \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle e^{i\omega_n (t_1-t_2)} \\ &= \left(\frac{\Delta\omega}{2\pi} \right) \int \frac{d\omega}{2\pi} \langle \langle \eta[\omega]^* \eta[\omega] \rangle \rangle e^{i\omega_n (t_1-t_2)} \end{aligned}$$

9. relation between power spectrum and $\pi[\omega]$

the probability distribution $\mathbb{P} \{ \eta[\omega_n] \} = \text{const} \times e^{-\int d\omega \frac{\eta[\omega]^* \eta[\omega]}{\pi[\omega]}}$ in a discrete form

$$\int d\omega \frac{\eta[\omega]^* \eta[\omega]}{\pi[\omega]} = \sum \Delta\omega \frac{\eta[\omega_n]^* \eta[\omega_n]}{\pi[\omega_n]}$$

$$\mathbb{P} \{ \eta[\omega_n] \} = \text{const} \times e^{-\sum \Delta\omega \frac{\eta[\omega_n]^* \eta[\omega_n]}{\pi[\omega_n]}}$$

$$e^{-X \{ \xi[\omega_n] \}} = \int \mathcal{D}\eta[\omega_n] \times \mathcal{D}\eta[\omega_n]^* e^{-\sum \Delta\omega \frac{\eta[\omega_n]^* \eta[\omega_n]}{\pi[\omega_n]} - \sum \Delta\omega (\xi[\omega_n]^* \eta[\omega_n] + \xi[\omega_n] \eta[\omega_n]^*)}$$

$$= \prod \int d\eta[\omega_n] \times d\eta[\omega_n]^* e^{-\Delta\omega \frac{\eta[\omega_n]^* \eta[\omega_n]}{\pi[\omega_n]} - \Delta\omega (\xi[\omega_n]^* \eta[\omega_n] + \xi[\omega_n] \eta[\omega_n]^*)}$$

$$\begin{aligned}
& \int d\eta[\omega_n] \times d\eta[\omega_n]^* e^{-\Delta\omega \frac{\eta[\omega_n]^* \eta[\omega_n]}{\pi[\omega_n]} - \Delta\omega (\xi[\omega_n]^* \eta[\omega_n] + \xi[\omega_n] \eta[\omega_n]^*)} \\
&= \int d\eta[\omega_n] \times d\eta[\omega_n]^* e^{-\frac{\Delta\omega}{\pi[\omega_n]} (\eta[\omega_n]^* \eta[\omega_n] - \pi[\omega_n] \times (\xi[\omega_n]^* \eta[\omega_n] + \xi[\omega_n] \eta[\omega_n]^*))} \\
&= \int d\eta[\omega_n] \times d\eta[\omega_n]^* e^{-\frac{\Delta\omega}{\pi[\omega_n]} ((\eta[\omega_n] - \pi[\omega_n] \times \xi[\omega_n])^* (\eta[\omega_n] - \pi[\omega_n] \times \xi[\omega_n]) - \pi[\omega_n]^2 (\xi[\omega_n])^* (\xi[\omega_n]))} \\
&= e^{\Delta\omega \pi[\omega_n] (\xi[\omega_n])^* (\xi[\omega_n])} \int d\eta[\omega_n] \times d\eta[\omega_n]^* e^{-\frac{\Delta\omega}{\pi[\omega_n]} ((\eta[\omega_n] - \pi[\omega_n] \times \xi[\omega_n])^* (\eta[\omega_n] - \pi[\omega_n] \times \xi[\omega_n]))} \\
&= e^{\Delta\omega \pi[\omega_n] (\xi[\omega_n])^* (\xi[\omega_n])} \frac{\pi[\omega_n]}{\Delta\omega}
\end{aligned}$$

$$\begin{aligned}
e^{-X\{\xi[\omega_n]\}} &= \left(\prod \frac{\pi[\omega_n]}{\Delta\omega} \right) e^{\sum \Delta\omega \pi[\omega_n] \times \xi[\omega_n]^* \times \xi[\omega_n]} \\
&= \left(\prod \frac{\pi[\omega_n]}{\Delta\omega} \right) e^{\int d\omega \pi[\omega] \times \xi[\omega]^* \times \xi[\omega]} \\
\langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle &= \langle \eta[\omega_n]^* \eta[\omega_n] \rangle - \langle \eta[\omega_n]^* \rangle \langle \eta[\omega_n] \rangle \\
&= \left(e^{X\{\xi[\omega_n]\}} \frac{1}{\Delta\omega^2} \frac{d}{d\xi[\omega_n]} \frac{d}{d\xi^*[\omega_n]} e^{-X\{\xi[\omega_n]\}} - \left(e^{X\{\xi[\omega_n]\}} \frac{1}{\Delta\omega} \frac{d}{d\xi[\omega_n]} e^{-X\{\xi[\omega_n]\}} \right)^2 \right) \Big|_{\{\xi^*[\omega_n]=0\}} \\
&= \frac{1}{\Delta\omega^2} \frac{d}{d\xi[\omega_n]} \frac{d}{d\xi^*[\omega_n]} X\{\xi[\omega_n]\} = \frac{1}{\Delta\omega} \pi[\omega_n] \\
\langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle &= \frac{1}{\Delta\omega} \pi[\omega_n]
\end{aligned}$$

$$\begin{aligned}
S[\omega_n] &= \frac{1}{\Omega} \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle \\
\frac{1}{\Delta\omega} \pi[\omega_n] &= \langle \langle \eta[\omega_n]^* \eta[\omega_n] \rangle \rangle = S[\omega_n] \Omega \\
\frac{1}{\Omega \Delta\omega} \pi[\omega_n] &= S[\omega_n] \\
(2\pi) \pi[\omega_n] &= S[\omega_n] \\
\pi[\omega_n] &= \frac{1}{2\pi} S[\omega_n]
\end{aligned}$$

$$\left(e^{-a x^2 - b x - c} D[D[e^{a x^2 + b x + c}, x], x] - (e^{-a x^2 - b x - c} D[e^{a x^2 + b x + c}, x])^2 \right) /. x \rightarrow 0 // FullSimplify$$

2 a

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We now assume that noise is weak and we can Taylor

expand the fitting function $Z_{\{\lambda_i\}}[\omega]$ in deviations $\delta\lambda_i$

$$Z_{\{\Lambda_i + \lambda_i\}}[\omega]$$

$$\text{const} \int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\int d\omega \frac{(Z[\omega] - z[\omega])^* \cdot (Z[\omega] - z[\omega])}{T_2}} =$$

$$\text{const} \int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\int d\omega \frac{\left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)^* \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)}{T}} =$$

$$\begin{aligned} & \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)^* \\ & \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right) \\ & (Z_{\{\Lambda_i\}}[\omega] - z[\omega])^* (Z_{\{\Lambda_i\}}[\omega] - z[\omega]) \\ & + \left(\frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} \right)^* (Z_{\{\Lambda_i\}}[\omega] - z[\omega]) d\Lambda_i + \text{h.c.} (\Lambda_i + d\Lambda_i) \xi_i \\ & + \left(\left(\frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} \right)^* \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_j} + \left(\frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} \right)^* (Z_{\{\Lambda_i\}}[\omega] - z[\omega]) \right) d\Lambda_i d\Lambda_j \end{aligned}$$

$$\begin{aligned} & (B_i + \text{h.c.} \xi_i) d\Lambda_i + A_{ij} d\Lambda_i d\Lambda_j = d\Lambda_i A d\Lambda_j + (B_i + \text{h.c.} \xi_i) d\Lambda_i \\ & = (d\Lambda_j + (B_i + \text{h.c.} \xi_i) A^{-1}) A (d\Lambda_j + A^{-1} (B_i + \text{h.c.} \xi_i)) - (B_i + \text{h.c.} \xi_i) A^{-1} (B_i + \text{h.c.} \xi_i) \end{aligned}$$

$$\int d\lambda_1 d\lambda_2 \dots d\lambda_N e^{-\int d\omega \frac{\left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)^* \left(Z_{\{\Lambda_i\}}[\omega] - z[\omega] + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i} d\Lambda_i + \frac{dZ_{\{\Lambda_i\}}[\omega]}{d\Lambda_i d\Lambda_j} d\Lambda_i d\Lambda_j \right)}{T}} =$$

$$e^{\frac{1}{T} \int d\omega (B_i + \text{h.c.} \xi_i) A^{-1} (B_i + \text{h.c.} \xi_i)} =$$