

# Discrete Choice Analysis

Taro Mieno

AECN 896-003: Applied Econometrics

## Discrete Choice Analysis

- ▶ Focus on understanding choices that are discrete (not continuous)
  - ▶ Whether you own a car or not (binary choice)
  - ▶ Whether you use an iPhone, Android, or other types of cell phones (Multinomial choice)
  - ▶ Which recreation sites you visit this winter (multinomial)
- ▶ Linear models we have seen are often not appropriate

## Binary Response Model

## Binary Response

$y = 0$  (if you do not own a car)

$y = 1$  (if you own at least one car)

## Question we would like to answer

How do independent variables  $x_1, \dots, x_k$  affect the status of  $y$  (the choice of whether to own at least one car or not)?

## Binary Response Model

We try to model the **probability** of  $y = 1$  (own at least one car)

$$Pr(y = 1|x_1, \dots, x_k) = f(x_1, \dots, x_k)$$

as a function of independent variables.

## Linear probability model

$$Pr(y = 1|x_1, \dots, x_k) = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

## Binary Response Model

We try to model the **probability** of  $y = 1$  (own at least one car)

$$Pr(y = 1|x_1, \dots, x_k) = f(x_1, \dots, x_k)$$

as a function of independent variables.

So, how about

$$Pr(y = 1|x_1, \dots, x_k) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$$

such that  $G()$  is a function taking on values strictly between zero and one:  $0 < G(z) < 1$  for all real numbers  $z$ ?

## Notes

Different choices of  $G()$  lead to different models

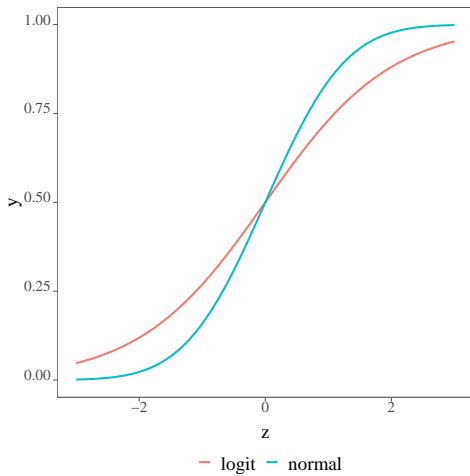
### Logit model

$$G(z) = \exp(z)/[1 + \exp(z)] = \frac{e^z}{1 + e^z}$$

### Probit model

$$G(z) = \Phi(z)$$

where  $\Phi(z)$  is the standard normal cumulative distribution function





## Interpretation

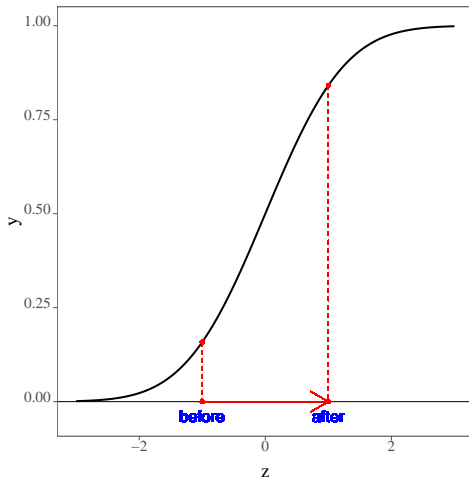
$$Pr(y = 1|x_1, \dots, x_k) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$$

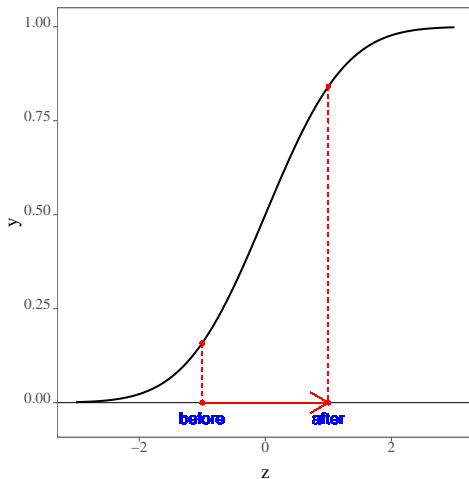
- ▶ What do  $\beta$ s measure?
- ▶ How do we interpret them?

## Before and after

Before:  $x_1 = 0, \dots, x_k = 0 \Rightarrow z = \beta_0$

After:  $x_1 = 1$  and  $x_2 = 0, \dots, x_k = 0 \Rightarrow z = \beta_0 + \beta_1 x_1$





- ▶  $\beta$ s measure how far you move along the x-axis
- ▶  $\beta$ s does not directly measure how independent variables influence the probability of  $y = 1$

Marginal impact of  $x_k$  (continuous) on  $Prob(y = 1)$

$$Pr(y = 1|x_1, \dots, x_k) = G(z)$$

$$z = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

Differentiating both sides with respect to  $x_k$ ,

$$\begin{aligned}\frac{\partial Pr(y = 1|x_1, \dots, x_k)}{\partial x_k} &= G'(z) \times \frac{\partial z}{\partial x_k} \\ &= G'(z) \times \beta_k\end{aligned}$$

## Marginal impact of $x_k$ (continuous) on $Prob(y = 1)$

$$Pr(y = 1|x_1, \dots, x_k) = G(z)$$

$$z = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

Differentiating both sides with respect to  $x_k$ ,

$$\begin{aligned}\frac{\partial Pr(y = 1|x_1, \dots, x_k)}{\partial x_k} &= G'(z) \times \frac{\partial z}{\partial x_k} \\ &= G'(z) \times \beta_k\end{aligned}$$

## Notes

- ▶ The marginal impact of an independent variable depends on the values of all the independent variables:  
 $G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$
- ▶ Since  $G'()$  is always positive, the sign of the marginal impact of an independent variable on  $Prob(y = 1)$  is always the same as the sign of its coefficient

## Estimation of Binary Choice Models

- ▶ Linear models: OLS
- ▶ Binary choice models: Maximum Likelihood Estimation (MLE)

### OLS

Find parameters that makes the sum of residuals squared the smallest

### MLE (very loosely put)

Find parameters ( $\beta$ s) that makes what we observed (collection of binary decisions made by different individuals) most likely (Maximum Likelihood)

## Decisions made by two individuals

- ▶ Individual 1:  $y = 1$  (own at least one car)
- ▶ Individual 2:  $y = 0$  (does not own a car)

## Decisions made by two individuals

- ▶ Individual 1:  $y = 1$  (own at least one car)
- ▶ Individual 2:  $y = 0$  (does not own a car)

## Probability of individual decisions

Individual 1 :  $Prob(y_1 = 1|\mathbf{x}_1) = G(z_1)$

Individual 2 :  $Prob(y_2 = 0|\mathbf{x}_2) = 1 - G(z_2)$



## Decisions made by two individuals

- ▶ Individual 1:  $y = 1$  (own at least one car)
- ▶ Individual 2:  $y = 0$  (does not own a car)

## Probability of individual decisions

Individual 1 :  $Prob(y_1 = 1|\mathbf{x}_1) = G(z_1)$

Individual 2 :  $Prob(y_2 = 0|\mathbf{x}_2) = 1 - G(z_2)$

## Probability of a collection of decisions

The probability that we observe a **collection of choices** made by them (if their decisions are independent)

$$Prob(y_1 = 1|\mathbf{x}_1) \times Prob(y_2 = 0|\mathbf{x}_2) = G(z_1) \times [1 - G(z_2)],$$

which we call **likelihood function**.

## Probability of individual decisions

$$\text{Individual 1 : } \text{Prob}(y_1 = 1 | \mathbf{x}_1) = G(z_1)$$

$$\text{Individual 2 : } \text{Prob}(y_2 = 0 | \mathbf{x}_2) = 1 - G(z_2)$$

## Probability of a collection of decisions

The probability that we observe a **collection of choices** made by them (if their decisions are independent)

$$\text{Prob}(y_1 = 1 | \mathbf{x}_1) \times \text{Prob}(y_2 = 0 | \mathbf{x}_2) = G(z_1) \times [1 - G(z_2)],$$

which we call **likelihood function**.

## MLE

$$\text{Max}_{\beta_1, \dots, \beta_k} G(z_1) \times [1 - G(z_2)]$$

## MLE of Binary Choice Model in General

Maximize the likelihood function:

$$\text{Max}_{\beta_1, \dots, \beta_k} L$$

where  $L = \prod_{i=1}^n \left[ y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right]$  is the likelihood function.

## Log-likelihood function

$$\begin{aligned} LL &= \log \left( \prod_{i=1}^n \left[ y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right] \right) \\ &= \sum_{i=1}^n \log \left( y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right) \end{aligned}$$

## MLE with $LL$

$$\text{argmax}_{\beta_1, \dots, \beta_k} L \equiv \text{argmax}_{\beta_1, \dots, \beta_k} LL$$

## Implementation in R with an example

Participation of females in labor force:

$$Pr(inlf = 1|\mathbf{x}) = G(z)$$

where

$$z = \beta_0 + \beta_1nwifeinc + \beta_2educ + \beta_3exper + \\ + \beta_4exper^2 + \beta_5age + \beta_6kidslt6 + \beta_7kidsge6$$

- ▶ *inlf*: 1 if in labor force in 1975, 0 otherwise
- ▶ *nwifeinc*: earning as a family if she does not work
- ▶ *kidslt6*: # of kids less than 6 years old
- ▶ *kidsge6*: # of kids who are 6-18 year old

## R code: importing the data

```
#--- import the data ---#  
data <- read.dta13('MROZ.dta') %>%  
  mutate(exper2=exper^2)  
  
#--- take a look ---#  
dplyr::select(data,inlf,nwifeinc,kidslt6,kidsge6,educ) %>%  
  head()  
  
inlf  nwifeinc  kidslt6  kidsge6  educ  
1     1 10.910060      1      0    12  
2     1 19.499981      0      2    12  
3     1 12.039910      1      3    12  
4     1  6.799996      0      3    12  
5     1 20.100058      1      2    14  
6     1  9.859054      0      0    12
```

## Estimating binary choice model using *R*

You can use the *glm()* function (no new packages installation necessary) when using cross-sectional data

- ▶ *glm* refers to Generalized Linear Model, which encompass linear models we have been using
- ▶ you specify the *family* option to tell what kind of model you are estimating

## R code: Probit model estimation using glm()

```
probit_lf <- glm(  
  
  #--- formula ---#  
  inlf~nwifeinc+educ+exper+exper2+age+kidslt6+kidsge6,  
  
  #--- data ---#  
  data = data,  
  
  #--- models ---#  
  family=binomial(link='probit')  
)
```

### *family option*

- ▶ *binomial()*: tells *R* that your dependent variable is binary
- ▶ *link = 'probit'*: tells *R* that you want to use the cumulative distribution function of the standard normal distribution as  $G()$  in  $Prob(y = 1|\mathbf{x}) = G(z)$

## R code: Probit model estimation using glm()

```
summary(probit_lf)$coef
```

	Estimate	Std. Error	z value	Pr(> z )
(Intercept)	0.270073573	0.5080781657	0.5315591	5.950314e-01
nwifeinc	-0.012023637	0.0049391713	-2.4343430	1.491885e-02
educ	0.130903969	0.0253987284	5.1539576	2.550456e-07
exper	0.123347168	0.0187586870	6.5754692	4.850000e-11
exper2	-0.001887067	0.0005999272	-3.1454942	1.658065e-03
age	-0.052852442	0.0084623619	-6.2455899	4.222037e-10
kids1t6	-0.868324680	0.1183772702	-7.3352315	2.213386e-13
kidsge6	0.036005611	0.0440302624	0.8177469	4.135017e-01



## R code: Logit model estimation using glm()

```
logit_lf <- glm(  
  #--- formula ---#  
  inlf~nwifeinc+educ+exper+exper2+age+kidslt6+kidsge6,  
  
  #--- data ---#  
  data = data,  
  
  #--- models ---#  
  family=binomial(link='logit')  
)
```

### *family* option

- ▶ *binomial()*: tells *R* that your dependent variable is binary
- ▶ *link = 'logit'*: tells *R* that you want to use the function as  $G()$  in  $Prob(y = 1|\mathbf{x}) = G(z)$

## Logit model estimation results

```
summary(logit_1f)$coef
```

	Estimate	Std. Error	z value	Pr(> z )
(Intercept)	0.425452376	0.860364519	0.4945025	6.209514e-01
nwifeinc	-0.021345174	0.008421380	-2.5346410	1.125626e-02
educ	0.221170370	0.043439281	5.0914832	3.552734e-07
exper	0.205869531	0.032056713	6.4220411	1.344591e-10
exper2	-0.003154104	0.001016107	-3.1041065	1.908546e-03
age	-0.088024375	0.014572890	-6.0402826	1.538446e-09
kidslt6	-1.443354143	0.203582842	-7.0897632	1.343417e-12
kidsge6	0.060112222	0.074789293	0.8037544	4.215388e-01

```
stargazer(probit_lf,logit_lf,type='latex',no.space=TRUE,table.layout='-ldc-t-s-')
```

Table

	<i>Dependent variable:</i>	
	inlf	
nwifeinc	−0.012** (0.005)	−0.021** (0.008)
educ	0.131*** (0.025)	0.221*** (0.043)
exper	0.123*** (0.019)	0.206*** (0.032)
exper2	−0.002*** (0.001)	−0.003*** (0.001)
age	−0.053*** (0.008)	−0.088*** (0.015)
kidslt6	−0.868*** (0.118)	−1.443*** (0.204)
kidsge6	0.036 (0.044)	0.060 (0.075)
Constant	0.270 (0.508)	0.425 (0.860)
Observations	753	753
Log Likelihood	−401.302	−401.765
Akaike Inf. Crit.	818.604	819.530

# Post-estimation operations and diagnostics

## Log-likelihood

$$LL = \sum_{i=1}^n \log \left( y_i \times G(\hat{z}_i) + (1 - y_i) \times (1 - G(\hat{z}_i)) \right)$$

- ▶  $\hat{z}_i = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k$
- ▶  $G(\hat{z}_i)$  is the fitted value of  $Prob(y = 1 | \mathbf{x})$

(The greater (less negative) the LL, the better the fit of the regression)

## McFadden's pseudo- $R^2$

A measure of how much better your model is compared to the model with only the intercept

$$pseudo - R^2 = 1 - LL/LL_0$$

where  $LL_0$  is the log-likelihood when you include only the intercept

## R code: pseudo- $R^2$

```
#--- estimate the model with only the intercept ---#
logit_lf_0 <- glm(inlf~1,data = data,family=binomial(link='logit'))

#--- extract LL using the logLik() function ---#
LL0 <- logLik(logit_lf_0)
LL0
'log Lik.' -514.8732 (df=1)

#--- extract LL using the logLik() function from your preferred model ---#
LL <- logLik(logit_lf)
LL
'log Lik.' -401.7652 (df=8)

#--- pseudo R2 ---#
pR2 <- 1-LL/LL0
pR2
'log Lik.' 0.2196814 (df=8)
```

## R code: alternative (easier) calculation of pseudo- $R^2$

```
#--- or more easily ---#  
1-logit_lf$deviance/logit_lf$null.deviance  
[1] 0.2196814  
  
#--- what are deviances? ---#  
logit_lf$null.deviance # = -2*LL0  
[1] 1029.746  
  
logit_lf$deviance # = -2*LL  
[1] 803.5303
```

## Testing: joint significance

You can do Likelihood Ratio (LR) test:

$$LR = 2(LL_{unrestricted} - LL_{restricted}) \sim \chi^2_{df\_restrictions}$$

where  $df\_restrictions$  is the number of restrictions



## Testing: joint significance

You can do Likelihood Ratio (LR) test:

$$LR = 2(LL_{unrestricted} - LL_{restricted}) \sim \chi^2_{df\_restrictions}$$

where  $df\_restrictions$  is the number of restrictions

## An example

- ▶  $H_0$  : the coefficients on *exper*, *exper2*, and *age* are 0
- ▶  $H_1$  :  $H_0$  is false

## R code: LR test for joint significance

```
#--- unrestricted ---#
logit_ur <- glm(
  inlf~nwifeinc+educ+exper+exper2+age+kidslt6+kidsge6,
  data = data,family=binomial(link='logit')
)

#--- restricted ---#
logit_r <- glm(
  inlf~nwifeinc+educ+kidslt6+kidsge6,
  data = data,family=binomial(link='logit')
)

#--- LR test using lrtest() from the lmtest package ---#
library(lmtest)
lrtest(logit_r,logit_ur)
```

Likelihood ratio test

Model 1: inlf ~ nwifeinc + educ + kidslt6 + kidsge6

Model 2: inlf ~ nwifeinc + educ + exper + exper2 + age + kidslt6 + kidsge6

	#Df	LogLik	Df	Chisq	Pr(>Chisq)
1	5	-464.92			
2	8	-401.77	3	126.32	< 2.2e-16 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

## Prediction

After estimating a binary choice model, you can easily predict the following two

- ▶  $\hat{z} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k$
- ▶  $\widehat{Prob}(y = 1|\mathbf{x}) = G(\hat{z}) = G(\hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k)$

## R code: LR test for joint significance

```
#--- z hat ---#
z <- predict(probit_lf, type='link')
head(z)
      1          2          3          4          5          6
0.5071349 0.6624576 0.5116317 0.7423429 0.1972781 0.8837878

#--- G(z) hat ---#
Gz <- predict(probit_lf, type='response')
head(Gz)
      1          2          3          4          5          6
0.6939699 0.7461610 0.6955456 0.7710602 0.5781950 0.8115946
```

## Marginal effect of an independent variable

- ▶ Coefficient estimates across different models (probit and logit) are not meaningful because the same value of a coefficient estimate means different things
- ▶ They are the estimates of  $\beta$ s, not the direct impact of the independent variables on the  $Prob(y = 1)$

## Calculating the marginal effect of an independent variable

$$\frac{\partial \Pr(y = 1 | x_1, \dots, x_k)}{\partial x_k} = G'(z) \times \beta_k$$

- ▶ the marginal impact depends on the current levels of all the independent variables
- ▶ we typically report one of the two types of marginal impacts
  - ▶ the marginal impact **at the mean** (average person): when all the independent variables take on their respective means
  - ▶ the average of the marginal impacts calculated for each of all the individuals observed

## Marginal impact at the mean

$$\frac{\partial \Pr(y = 1 | \bar{x}_1, \dots, \bar{x}_k)}{\partial x_k} = G'(\beta_0 + \beta_1 \bar{x}_1 + \dots + \beta_k \bar{x}_k) \times \beta_k$$

## Mean marginal impact

$$\sum_{i=1}^n \frac{\partial \Pr(y_i = 1 | \bar{x}_{i,1}, \dots, \bar{x}_{i,k})}{\partial x_k} = \sum_{i=1}^n G'(z_i) \times \beta_k$$

## R code: marginal impact of *educ* at the mean

```
#--- get the coef ---#
probit_lf$coef

(Intercept)      nwifeinc          educ          exper          exper2
0.270073573 -0.012023637  0.130903969  0.123347168 -0.001887067
      age      kidslt6      kidsge6
-0.052852442 -0.868324680  0.036005611

#--- get the mean ---#
means <- summarize(data, mean(nwifeinc), mean(educ), mean(exper), mean(exper2),
  mean(age), mean(kidslt6), mean(kidsge6))

#--- calculate z ---#
z <- probit_lf$coef[1] + sum(probit_lf$coef[-1]*means)

#--- marignal impact ---#
dnorm(z)*probit_lf$coef['educ']

(Intercept)
0.05112843
```

## R code: marginal impact of *educ* at the mean

```
#--- get the coef ---#
probit_lf$coef

(Intercept)      nwifeinc          educ          exper          exper2
0.270073573 -0.012023637  0.130903969  0.123347168 -0.001887067
          age      kidslt6      kidsge6
-0.052852442 -0.868324680  0.036005611

#--- get the mean ---#
means <- summarize(data, mean(nwifeinc), mean(educ), mean(exper), mean(exper2),
  mean(age), mean(kidslt6), mean(kidsge6))

#--- calculate z ---#
z <- probit_lf$coef[1] + sum(probit_lf$coef[-1]*means)

#--- marginal impact ---#
dnorm(z)*probit_lf$coef['educ']

(Intercept)
0.05112843
```

## Interpretation

If your education goes up by 1 year, you are 5% more likely to be in the labor force when you are an average person



## R code: mean marginal impact of *educ*

```
#--- get z for all the individuals ---#  
z <- predict(probit_lf,type='link')  
  
#--- get G'(z) ---#  
Gz_indiv <- dnorm(z)  
  
#--- mean marginal impact of education ---#  
mean(Gz_indiv)*probit_lf$coef['educ']  
educ  
0.03937009
```

## Regression Models for Count Data

## Count as the dependent variable

Count variables take **non-negative discrete** values  $(0, 1, \dots,)$

- ▶ the number of times people get arrested in a year
- ▶ the number of cars owned by a family
- ▶ the number of patents applied for by a firm in a year
- ▶ the number of kids in a family

## Poisson regression

By far the most popular choice to analyze count variables is

### Poisson regression

- ▶ The outcome (count) variable is assumed to be Poisson distributed
- ▶ The mean of the Poisson distribution is assumed to be a function of some variables you believe matter

## Poisson regression

By far the most popular choice to analyze count variables is

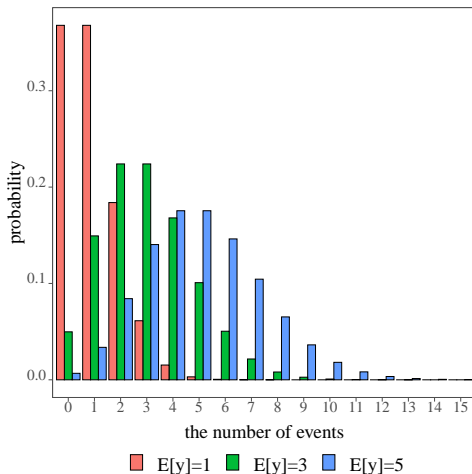
### Poisson regression

- ▶ The outcome (count) variable is assumed to be Poisson distributed
- ▶ The mean of the Poisson distribution is assumed to be a function of some variables you believe matter

## Poisson distribution

Poisson distribution is a discrete probability distribution that describes the probability of the number of events that occur in a fixed interval of time and/or space

$$Prob(y|\lambda) = \frac{\lambda^y e^{-\lambda}}{y!}, \text{ where } \lambda = E[y]$$



## Poisson regression

We try to learn what and how variables affect the expected value (the expected number of events conditional on independent variables)

## Expected number of events conditional on independent variables

$$E[y|\mathbf{x}] = G(\beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k)$$

- ▶ This is exactly the same modeling framework we used
  - ▶ Linear model:  $G(z) = z$
  - ▶ Probit model:  $G(z) = \Phi(z)$

## Expected number of events conditional on independent variables

$$E[y|\mathbf{x}] = G(\beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k)$$

- ▶ This is exactly the same modeling framework we used
  - ▶ Linear model:  $G(z) = z$
  - ▶ Probit model:  $G(z) = \Phi(z)$

## A popular choice of $G()$

$$G(z) = \exp(z)$$

- ▶ ensures that the expected value conditional on  $\mathbf{x}$  is always positive



## The number of events for two individuals

- ▶ Individual 1:  $y = 3$  (own three cars)
- ▶ Individual 2:  $y = 1$  (own one car)

## The number of events for two individuals

- ▶ Individual 1:  $y = 3$  (own three cars)
- ▶ Individual 2:  $y = 1$  (own one car)

## Expected number of events observed

Individual 1 :  $\lambda_1 = \exp(z_1)$

Individual 2 :  $\lambda_2 = \exp(z_2)$

## The number of events for two individuals

- ▶ Individual 1:  $y = 3$  (own three cars)
- ▶ Individual 2:  $y = 1$  (own one car)

## Expected number of events observed

Individual 1 :  $\lambda_1 = \exp(z_1)$

Individual 2 :  $\lambda_2 = \exp(z_2)$

## Probability of observing the number of events we observed

$$\text{Individual 1 : } \text{Prob}(y = 3 | \mathbf{x}_1) = \frac{\lambda_1^3 e^{-\lambda_1}}{3!}$$

$$\text{Individual 2 : } \text{Prob}(y = 1 | \mathbf{x}_2) = \frac{\lambda_2^1 e^{-\lambda_2}}{1!}$$

## Probability of observing the number of events we observed

$$\text{Individual 1 : } Prob(y = 3 | \mathbf{x}_1) = \frac{\lambda_1^3 e^{-\lambda_1}}{3!}$$

$$\text{Individual 2 : } Prob(y = 1 | \mathbf{x}_2) = \frac{\lambda_2^1 e^{-\lambda_2}}{1!}$$

## Probability of observing a series of events by all individuals

The probability that we observe a **collection of choices** made by them (if their events are independent)

$$L = Prob(y_1 = 3 | \mathbf{x}_1) \times Prob(y_2 = 1 | \mathbf{x}_2) = \frac{\lambda_1^3 e^{-\lambda_1}}{3!} \times \frac{\lambda_2^1 e^{-\lambda_2}}{1!}$$

which we call **likelihood function**.

## Probability of observing a series of events by all individuals

The probability that we observe a collection of choices made by them (if their events are independent)

$$L = \text{Prob}(y_1 = 3|\mathbf{x}_1) \times \text{Prob}(y_2 = 1|\mathbf{x}_2) = \frac{\lambda_1^3 e^{-\lambda_1}}{3!} \times \frac{\lambda_2^1 e^{-\lambda_2}}{1!}$$

which we call likelihood function.

## Log-likelihood function

$$LL = \log(L) = \log\left(\frac{\lambda_1^3 e^{-\lambda_1}}{3!}\right) + \log\left(\frac{\lambda_2^1 e^{-\lambda_2}}{1!}\right)$$

(Remember  $\lambda_i = \exp(\beta_0 + \beta_1 x_{i,1} + \dots + \beta_k x_{i,k})$ )

## Probability of observing a series of events by all individuals

The probability that we observe a **collection of choices** made by them (if their events are independent)

$$L = \text{Prob}(y_1 = 3|\mathbf{x}_1) \times \text{Prob}(y_2 = 1|\mathbf{x}_2) = \frac{\lambda_1^3 e^{-\lambda_1}}{3!} \times \frac{\lambda_2^1 e^{-\lambda_2}}{1!}$$

which we call **likelihood function**.

## Log-likelihood function

$$LL = \log(L) = \log\left(\frac{\lambda_1^3 e^{-\lambda_1}}{3!}\right) + \log\left(\frac{\lambda_2^1 e^{-\lambda_2}}{1!}\right)$$

(Remember  $\lambda_i = \exp(\beta_0 + \beta_1 x_{i,1} + \dots + \beta_k x_{i,k})$ )

## MLE

$$\text{Max}_{\beta_1, \dots, \beta_k} LL$$

## Implementation in R with an example

The number of times a man is arrested during 1986:

$$Pr(narr86|\mathbf{x}) = G(z)$$

where

$$z = \beta_0 + \beta_1 pcnv + \beta_2 tottime + \beta_3 qemp86 + \beta_4 inc86 \\ + \beta_5 black + \beta_6 hispan$$

- ▶ *narr86*: # of times arrested in 1986
- ▶ *pcnv*: proportion of prior conviction
- ▶ *tottime*: time in prison since 18
- ▶ *qemp86*: # quarters employed in 1986
- ▶ *inc86*: legal income in 1986 (in \$100)

## R code: importing the data

```
998      0 0.00      3.0 133.1
999      0 1.00      4.0 195.6
1000      0 0.25      4.0 170.7
1001      0 1.00      4.0  54.3
1002      0 1.00      4.0 152.7
1003      0 0.60      0.0   0.0
1004      1 0.00      2.0   2.5
1005      0 0.67      2.0  20.0
1006      0 1.00      4.0 107.3
1007      1 0.33      0.0   0.0
1008      2 0.20      0.0   0.0
1009      0 1.00      4.0  67.1
1010      5 0.25      4.0  61.4
1011      2 0.25      1.0   5.6
1012      0 0.00      1.0   9.6
1013      0 1.00      4.0 175.6
1014      0 0.25      2.0  13.7
1015      0 0.60      0.0   0.0
1016      0 0.00      4.0  28.6
1017      0 0.00      3.0  54.5
1018      0 0.40      2.0   8.2
1019      3 0.56      0.0   0.0
1020      0 0.00      3.0  51.8
1021      0 0.00      0.0   0.0
1022      2 0.56      2.0   4.3
1023      0 0.33      1.0   7.5
1024      2 0.50      3.0  44.5
1025      0 1.00      2.0  27.3
1026      1 0.50      2.0   2.5
```



## R code: Poisson model estimation using glm()

```
pois_lf <- glm(  
  
  #--- formula ---#  
  narr86~pcnv+totttime+qemp86+inc86+black+hispan,  
  
  #--- data ---#  
  data = data,  
  
  #--- models ---#  
  family=poisson(link='log')  
)
```

### *family option*

- ▶ *poisson()*: tells *R* that your dependent variable is Poisson distributed
- ▶ *link = 'log'*: tells *R* that you want to use *exp()* (the inverse of *log()*) as *G()* in  $E(y = 1|\mathbf{x}) = G(z)$

## R code: summary

```
summary(pois_1f)$coef
```

	Estimate	Std. Error	z value	Pr(> z )
(Intercept)	-0.665805562	0.063749782	-10.4440446	1.560185e-25
pcnv	-0.432102017	0.085403627	-5.0595277	4.202962e-07
totttime	-0.001206742	0.005504698	-0.2192204	8.264784e-01
qemp86	-0.009956885	0.028588565	-0.3482821	7.276284e-01
inc86	-0.008395041	0.001039638	-8.0749648	6.749612e-16
black	0.644485369	0.073922145	8.7184344	2.820750e-18
hispan	0.473213044	0.073774349	6.4143303	1.414432e-10

```
stargazer(pois_lf, type='latex', no.space=TRUE, table.layout='-ldc-t-s-')
```

## Table

	<i>Dependent variable:</i>
	narr86
pcnv	−0.432*** (0.085)
tottime	−0.001 (0.006)
qemp86	−0.010 (0.029)
inc86	−0.008*** (0.001)
black	0.644*** (0.074)
hispan	0.473*** (0.074)
Constant	−0.666*** (0.064)
Observations	2,725
Log Likelihood	−2,264.668
Akaike Inf. Crit.	4,543.337

## Calculating the marginal impact of an independent variable

$$\frac{\partial E(y|x_1, \dots, x_k)}{\partial x_k} = \beta_k \times \exp(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$$

- ▶ the marginal impact depends on the current levels of all the independent variables
- ▶ we typically report one of the two types of marginal impacts
  - ▶ the marginal impact **at the mean** (average person): when all the independent variables take on their respective means
  - ▶ the average of the marginal impacts calculated for each of all the individuals observed

## R code: mean marginal impact of income

```
#--- get z for all the individuals ---#  
z <- predict(pois_lf,type='link')  
  
#--- get G'(z) ---#  
Gz_indiv <- exp(z)  
  
#--- mean marignal impact of education ---#  
mean(Gz_indiv)*pois_lf$coef['inc86']  
      inc86  
-0.003394985
```

## R code: mean marginal impact of income

```
#--- get z for all the individuals ---#  
z <- predict(pois_lf, type='link')  
  
#--- get G'(z) ---#  
Gz_indiv <- exp(z)  
  
#--- mean marginal impact of education ---#  
mean(Gz_indiv)*pois_lf$coef['inc86']  
      inc86  
-0.003394985
```

### Interpretation

If your income goes up by \$100, the expected number of getting arrested declines by 0.0034

## Notes

- ▶ Poisson regression model is under the same econometric modeling framework: GLM
- ▶ Codes for testing are exactly the same as those we saw for the binary response models

## Multinomial Choice



## Multinomial Choice

Instead of two options, you may be picking one option out of more than two options

- ▶ which carrier?
  - ▶ Verizon
  - ▶ Sprint
  - ▶ AT&T
  - ▶ T-mobile
- ▶ which transportation means to commute?
  - ▶ own car
  - ▶ Uber
  - ▶ bus
  - ▶ train
  - ▶ bike

## Multinomial logit model

The most popular model to analyze multinomial choice

- ▶ environmental evaluation
- ▶ transportation
- ▶ marketing

# Understanding multinomial logit model through an example

## Choice of trains

1. 10 euros, 30 minutes travel time, one change
2. 20 euros, 20 minutes travel time, one change
3. 22 euros, 22 minutes travel time, no change

## Associated utility

1.  $V_1 = \alpha_1 + \beta 10 + \gamma 30 + \rho 1 + v_1$
2.  $V_2 = \alpha_2 + \beta 20 + \gamma 20 + \rho 1 + v_2$
3.  $V_3 = \alpha_3 + \beta 22 + \gamma 22 + \rho 0 + v_3$

## Choice probability

Logit model assumes that the probability of choosing an alternative is the following:

$$1. P_1 = \frac{e^{V_1}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

$$2. P_2 = \frac{e^{V_2}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

$$3. P_3 = \frac{e^{V_3}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

## Choice probability

Logit model assumes that the probability of choosing an alternative is the following:

$$1. P_1 = \frac{e^{V_1}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

$$2. P_2 = \frac{e^{V_2}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

$$3. P_3 = \frac{e^{V_3}}{e^{V_1} + e^{V_2} + e^{V_3}}$$

## Notes

$$\blacktriangleright 0 < P_j < 1, \forall j = 1, 2, 3$$

$$\blacktriangleright \sum_{j=1}^3 P_j = 1$$

Probability of observing individual  $i$  choosing what  $i$  chose

$$P_i = \prod_{j=1}^3 y_{i,j} \times P_j$$

where  $y_{i,j} = 1$  if  $i$  chose  $j$ , 0 otherwise

An example

$$y_{i,1} = 0, \quad y_{i,2} = 1, \quad y_{i,3} = 0$$

$$P_i = \prod_{j=1}^3 y_{i,j} \times P_j = 0 \times P_1 + 1 \times P_2 + 0 \times P_3$$

Probability of observing a series of choices made by all the subjects

If choices made by the subjects are independent with each other,

$$LL = \prod_{i=1}^n P_i = \prod_{i=1}^n \prod_{j=1}^3 y_{i,j} \times P_j$$

MLE

$$\text{Max}_{\beta, \gamma, \rho} \log(LL)$$

# Interpretation of the coefficients

## Model in general

$$V_{i,j} = \alpha_j + \beta_1 x_{1,i,j} + \cdots + \beta_k x_{k,i,j}$$

$$P_{i,j} = \frac{e^{V_{i,j}}}{\sum_{k=1}^J e^{V_{i,k}}}$$

## Interpretation of the coefficients

$$\frac{\partial P_{i,j}}{\partial x_{k,i,j}} = \beta_k P_{i,j} (1 - P_{i,j})$$

- ▶ A marginal change in  $k$ th variable for alternative  $j$  would change the probability of choosing alternative  $j$  by  $\beta_k P_{i,j} (1 - P_{i,j})$
- ▶ the sign of the impact is the same as the sign of the coefficient



## Implementation in *R*

You can use *mlogit* package to estimate multinomial logit models

- ▶ format your data in a specific manner
- ▶ convert your data using *mlogit.data()*
- ▶ estimate using *mlogit()*

## R code: multinomial logit model

```
#--- library ---#  
library(mlogit)  
  
#--- get the heating data from the mlogit package ---#  
data('TravelMode',package='AER')  
  
#--- take a look at the data ---#  
# first 10 rows  
head(TravelMode,10)
```

	individual	mode	choice	wait	vcost	travel	gcost	income	size
1	1	air	no	69	59	100	70	35	1
2	1	train	no	34	31	372	71	35	1
3	1	bus	no	35	25	417	70	35	1
4	1	car	yes	0	10	180	30	35	1
5	2	air	no	64	58	68	68	30	2
6	2	train	no	44	31	354	84	30	2
7	2	bus	no	53	25	399	85	30	2
8	2	car	yes	0	11	255	50	30	2
9	3	air	no	69	115	125	129	40	1
10	3	train	no	34	98	892	195	40	1

## R code: data preparation

```
#--- convert the data ---#
TM <- mlogit.data(TravelMode,
  shape='long', # what format is the data in?
  choice='choice', # name of the variable that indicates choice made
  chid.var='individual', # name of the variable that indicates who made choices
  alt.var='mode' # the name of the variable that indicates options
)
```

```
#--- take a look at the data ---#
```

```
# first 10 rows
```

```
head(TM,10)
```

	individual	mode	choice	wait	vcost	travel	gcost	income	size
1.air	1	air	FALSE	69	59	100	70	35	1
1.train	1	train	FALSE	34	31	372	71	35	1
1.bus	1	bus	FALSE	35	25	417	70	35	1
1.car	1	car	TRUE	0	10	180	30	35	1
2.air	2	air	FALSE	64	58	68	68	30	2
2.train	2	train	FALSE	44	31	354	84	30	2
2.bus	2	bus	FALSE	53	25	399	85	30	2
2.car	2	car	TRUE	0	11	255	50	30	2
3.air	3	air	FALSE	69	115	125	129	40	1
3.train	3	train	FALSE	34	98	892	195	40	1

# R code: multinomial logit model estimation

```
#--- estimate ---#  
ml_reg <- mlogit(choice~wait+vcost+travel,data=TM)  
  
#--- summary ---#  
summary(ml_reg)
```

Call:

```
mlogit(formula = choice ~ wait + vcost + travel, data = TM, method = "nr")
```

Frequencies of alternatives:

```
      air      train      bus      car  
0.27619 0.30000 0.14286 0.28095
```

nr method

5 iterations, 0h:0m:0s

$g'(-H)^{-1}g = 0.000192$

successive function values within tolerance limits

Coefficients :

	Estimate	Std. Error	z-value	Pr(> z )
train:(intercept)	-0.7866667	0.60260733	-1.3054	0.19174
bus:(intercept)	-1.43363372	0.68071345	-2.1061	0.03520 *
car:(intercept)	-4.73985647	0.86753178	-5.4636	4.665e-08 ***
wait	-0.09688675	0.01034202	-9.3683	< 2.2e-16 ***
vcost	-0.01391160	0.00665133	-2.0916	0.03648 *
travel	-0.00399468	0.00084915	-4.7043	2.547e-06 ***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Log-Likelihood: -192.89

McFadden  $R^2$ : 0.32024

Likelihood ratio test :  $\chi^2 = 181.74$  (p.value = < 2.22e-16)

## R code: understanding the results

```
#--- coefficient ---#
summary(ml_reg)$coef
train:(intercept)    bus:(intercept)    car:(intercept)          wait
      -0.78666672      -1.433633718      -4.739856473      -0.096886747
           vcost           travel
      -0.013911604      -0.003994681
attr(,"names.sup.coef")
character(0)
attr(,"fixed")
train:(intercept)    bus:(intercept)    car:(intercept)          wait
           FALSE           FALSE           FALSE           FALSE
           vcost           travel
           FALSE           FALSE
attr(,"sup")
character(0)
```

### Notes

- ▶ intercept for *air* is dropped (*air* is the base)
  - ▶ train:(intercept) is  $-0.786$  means that train is less likely to be chosen if all the other **included** variables are the same
- ▶ the greater the travel time, the less likely the option is chosen