

# Discrete Choice

AECN 396/896-002

## Discrete Choice Analysis

- Focus on understanding choices that are discrete (not continuous)
  - Whether you own a car or not (binary choice)
  - Whether you use an iPhone, Android, or other types of cell phones (Multinomial choice)
  - Which recreation sites you visit this winter (multinomial)
- Linear models we have seen are often not appropriate

# Binary Response Model

## Binary response

$y = 0$  (if you do not own a car)

$y = 1$  (if you own at least one car)

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$y = 0$  (if you do not own a car)

$y = 1$  (if you own at least one car)

## Question

How do variables  $x_1, \dots, x_k$  affect the status of  $y$  (the choice of whether to own at least one car or not)?

## Binary response

We try to model the **probability** of  $y = 1$  (own at least one car)

$$Pr(y = 1|x_1, \dots, x_k) = f(x_1, \dots, x_k)$$

as a function of independent variables.

### Linear Probability Model

$$Pr(y = 1|x_1, \dots, x_k) = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

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$$Pr(y = 1|x_1, \dots, x_k) = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

### Drawback

There is no guarantee that the predicted probability is bounded within  $[0, 1]$ .

### How about this?

$$Pr(y = 1|x_1, \dots, x_k) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$$

where  $0 < G(z) < 1$  for all real numbers  $z$



Different choices of  $G()$  lead to different models.

### Logit model

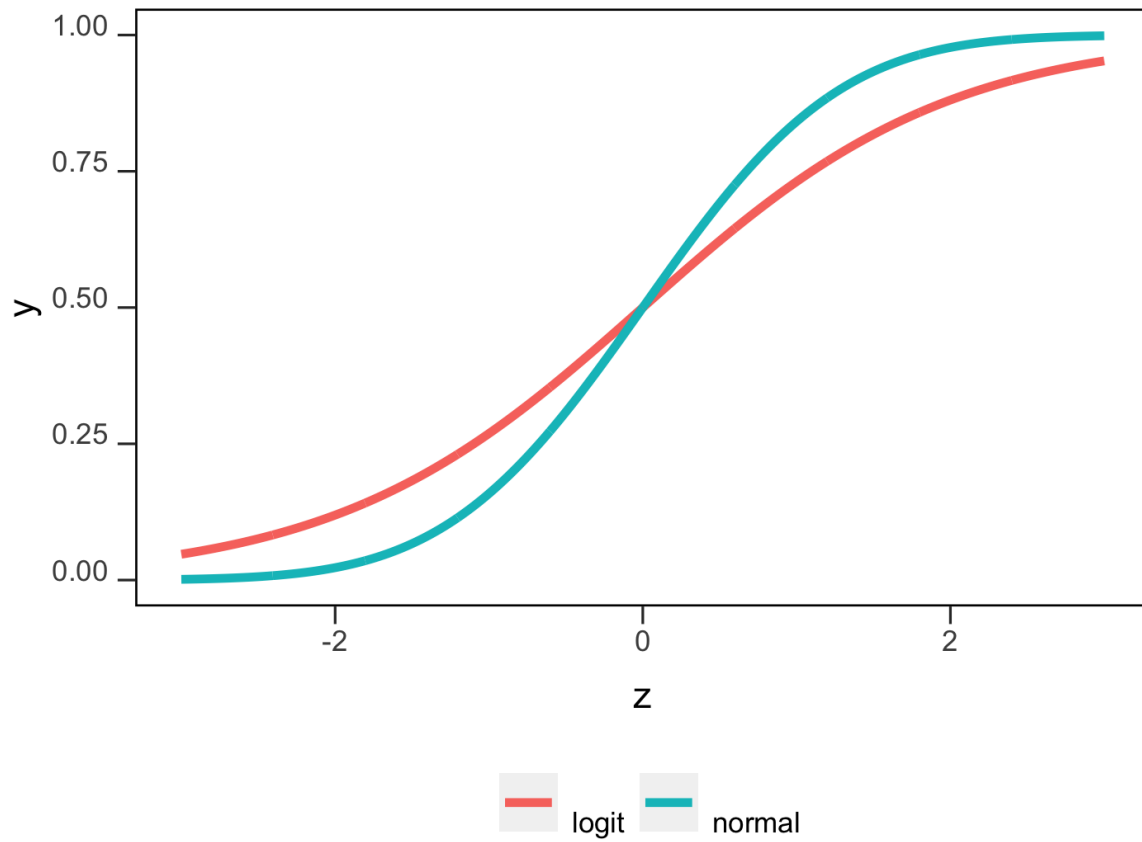
$$G(z) = \exp(z) / [1 + \exp(z)] = \frac{e^z}{1+e^z}$$

where  $z = \beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k$

### Probit model

$$G(z) = \Phi(z)$$

where  $\Phi(z)$  is the standard normal cumulative distribution function

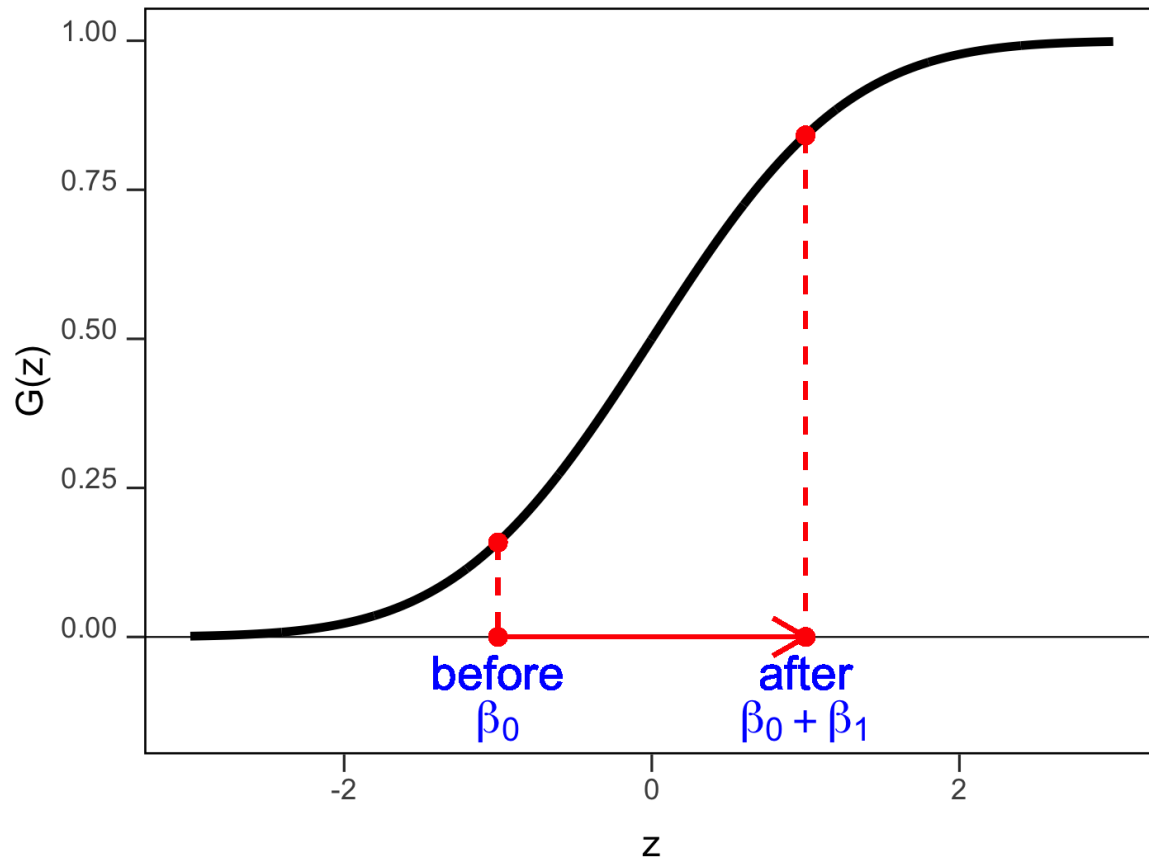


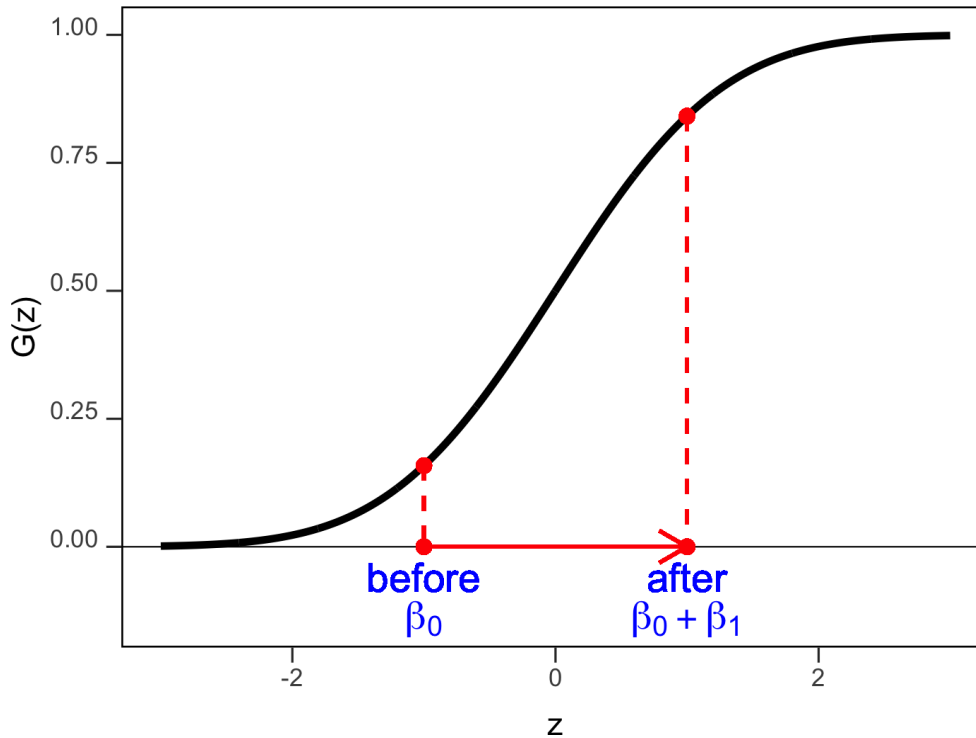
$$Pr(y = 1|x_1, \dots, x_k) = G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$$

- What do  $\beta$ s measure?
- How do we interpret them?

Before:  $x_1 = 0, \dots, x_k = 0 \Rightarrow z = \beta_0$

After:  $x_1 = 1$  and  $x_2 = 0, \dots, x_k = 0 \Rightarrow z = \beta_0 + \beta_1$





- $\beta$ s measure how far you move along the x-axis
- $\beta$ s does not directly measure how independent variables influence the probability of  $y = 1$

To understand the marginal impact of  $x_k$  on  $Prob(y = 1)$  (how a change in  $x_k$  affects the likelihood of owning a car), you need to do a bit of math.

### Model

$$Pr(y = 1|x_1, \dots, x_k) = G(z)$$

$$z = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k$$

### marginal impact

Differentiating both sides with respect to  $x_k$ ,

$$\begin{aligned} \frac{\partial Pr(y=1|x_1, \dots, x_k)}{\partial x_k} &= G'(z) \times \frac{\partial z}{\partial x_k} \\ &= G'(z) \times \beta_k \end{aligned}$$

### marginal impact

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$$\begin{aligned}\frac{\partial \Pr(y=1|x_1, \dots, x_k)}{\partial x_k} &= G'(z) \times \frac{\partial z}{\partial x_k} \\ &= G'(z) \times \beta_k\end{aligned}$$

### Notes

- The marginal impact of an independent variable depends on the values of all the independent variables:  $G(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$
- Since  $G'()$  is always positive, the sign of the marginal impact of an independent variable on  $\Pr(y = 1)$  is always the same as the sign of its coefficient

## Estimation of Binary Choice Models

- Linear models: OLS
- Binary choice models: [Maximum Likelihood Estimation \(MLE\)](#)



## OLS

Find parameters that makes the sum of residuals squared the smallest

## MLE (very loosely put)

Find parameters ( $\beta$ s) that makes what we observed (collection of binary decisions made by different individuals) most likely ([Maximum Likelihood](#))

### Observed decisions made by two individuals

- Individual 1:  $y = 1$  (own at least one car)
- Individual 2:  $y = 0$  (does not own a car)

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### Probability of individual decisions

- Individual 1 :  $Prob(y_1 = 1|\mathbf{x}_1) = G(z_1)$
- Individual 2 :  $Prob(y_2 = 0|\mathbf{x}_2) = 1 - G(z_2)$

where

- $\mathbf{x}_i$  is a collection of independent variables for individual  $i$  ( $x_{1,i}, \dots, x_{k,i}$ ).
- $z_i = \beta_0 + \beta_1 x_{1,i} + \dots + \beta_k x_{k,i}$

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- $z_i = \beta_0 + \beta_1 x_{1,i} + \dots + \beta_k x_{k,i}$

### Probability of a collection of decisions

The probability that we observe a **collection of choices** made by them (if their decisions are independent)

$$Prob(y_1 = 1|\mathbf{x}_1) \times Prob(y_2 = 0|\mathbf{x}_2) = G(z_1) \times [1 - G(z_2)]$$

which we call **likelihood function**.

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which we call **likelihood function**.

### MLE

$$Max_{\beta_1, \dots, \beta_k} G(z_1) \times [1 - G(z_2)]$$

## MLE of Binary Choice Model in General

Maximize the likelihood function:

$$\text{Max}_{\beta_1, \dots, \beta_k} L$$

where  $L = \prod_{i=1}^n \left[ y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right]$  is the likelihood function.

## Log-likelihood function

$$\begin{aligned} LL &= \log \left( \prod_{i=1}^n \left[ y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right] \right) \\ &= \sum_{i=1}^n \log \left( y_i \times G(z_i) + (1 - y_i) \times (1 - G(z_i)) \right) \end{aligned}$$

## MLE with (LL)

$$\text{argmax}_{\beta_1, \dots, \beta_k} L \equiv \text{argmax}_{\beta_1, \dots, \beta_k} LL$$

## Implementation in R with an example

Participation of females in labor force:

$$Pr(inlf = 1|\mathbf{x}) = G(z)$$

where

$$z = \beta_0 + \beta_1 nwifeinc + \beta_2 educ + \beta_3 exper \\ + \beta_4 exper^2 + \beta_5 age + \beta_6 kidslt6 + \beta_7 kidsge6$$

- *inlf*: 1 if in labor force in 1975, 0 otherwise
- *nwifeinc*: earning as a family if she does not work
- *kidslt6*: # of kids less than 6 years old
- *kidsge6*: # of kids who are 6-18 year old



```
### import the data ###
data <- read.dta13("MROZ.dta") %>%
  mutate(exper2 = exper^2)

### take a look ###
dplyr::select(data, inlf, nwifeinc, kidslt6, kidsge6, educ) %>%
  head()
```

```
##   inlf  nwifeinc kidslt6 kidsge6 educ
## 1    1 10.910060      1      0   12
## 2    1 19.499981      0      2   12
## 3    1 12.039910      1      3   12
## 4    1  6.799996      0      3   12
## 5    1 20.100058      1      2   14
## 6    1  9.859054      0      0   12
```

##	inlf	nwifeinc	kidslt6	kidsge6	educ
## 1	1	10.910060	1	0	12
## 2	1	19.499981	0	2	12
## 3	1	12.039910	1	3	12
## 4	1	6.799996	0	3	12
## 5	1	20.100058	1	2	14
## 6	1	9.859054	0	0	12

For individual 1 (row 1 of the data),

$$z_1 = \beta_0 + \beta_1 10.91 + \beta_2 12 + \beta_3 14 + \beta_4 196 + \beta_5 32 + \beta_6 1 + \beta_7 0$$

The probability that individual 1 would make the decision he/she made is:

$G(z_1)$  (a function of  $\beta$ s)

##		inlf	nwifeinc	kidslt6	kidsge6	educ
## 748	0	5.330	0	2	12	
## 749	0	28.200	0	2	13	
## 750	0	10.000	2	3	12	
## 751	0	9.952	0	0	12	
## 752	0	24.984	0	0	12	
## 753	0	28.363	0	3	9	

For individual 753 (row 753 of the data),

$$z_{753} = \beta_0 + \beta_1 28.36 + \beta_2 9 + \beta_3 12 + \beta_4 144 + \beta_5 39 + \beta_6 0 + \beta_7 3$$

The probability that individual 753 would make the decision he/she made is:

$$1 - G(z_{753}) \text{ (a function of } \beta\text{s)}$$

Multiply all the probabilities of observed choices given  $\beta$ s,

$$L = G(z_1) \times G(z_2) \times \dots [1 - G(z_753)]$$

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$$LL = \log\left(G(z_1) \times G(z_2) \times \dots [1 - G(z_753)]\right)$$

Multiply all the probabilities of observed choices given  $\beta$ s,

$$L = G(z_1) \times G(z_2) \times \dots [1 - G(z_753)]$$

$$LL = \log\left(G(z_1) \times G(z_2) \times \dots [1 - G(z_753)]\right)$$

Solve the following problems to estimate  $\beta$ s:

$$\text{Max}_{\beta_1, \dots, \beta_7} \quad LL$$

## Estimating binary choice model using (R)

You can use the `glm()` function (no new packages installation necessary) when using cross-sectional data

- `glm` refers to Generalized Linear Model, which encompass linear models we have been using
- you specify the `family` option to tell what kind of model you are estimating

## Probit model estimation

```
probit_lf <- glm(  
  #--- formula ---#  
  inlf ~ nwifeinc + educ + exper + exper2 + age + kidslt6 + kidsge6,  
  #--- data ---#  
  data = data,  
  #--- models ---#  
  family = binomial(link = "probit")  
)
```

## family option

- `binomial()`: tells R that your dependent variable is binary
- `link = "probit"`: tells R that you want to use the cumulative distribution function of the standard normal distribution as  $G()$  in  $Prob(y = 1|\mathbf{x}) = G(z)$



```
msummary(
  probit_lf,
  stars = TRUE,
  gof_omit = "IC",
  output = "flextable"
) %>%
  fontsize(
    size = 18,
    part = "all"
  ) %>%
  autofit()
```

	Model 1
(Intercept)	0.270 (0.508)
nwifeinc	-0.012 (0.005)
educ	0.131** (0.025)
exper	0.123* (0.019)
exper2	-0.002 (0.001)
age	-0.053 (0.008)
kidslt6	-0.868 (0.118)

+  $p < 0.1$ ,  $p < 0.05$ ,  **$p < 0.01$** ,  $p < 0.001$

## Logit model estimation

```
logit_lf <- glm(  
  #--- formula ---#  
  inlf ~ nwifeinc + educ + exper + exper2 + age + kidslt6 + kidsge6,  
  #--- data ---#  
  data = data,  
  #--- models ---#  
  family = binomial(link = "logit")  
)
```

## family option

- `binomial()`: tells R that your dependent variable is binary
- `link = "logit"`: tells R that you want to use  $G(z) = \frac{e^z}{1+e^z}$  in  $Prob(y = 1|\mathbf{x}) = G(z)$

```

msummary(
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  stars = TRUE,
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+  $p < 0.1$ ,  $p < 0.05$ ,  **$p < 0.01$** ,  $p < 0.001$

## Post-estimation operations and diagnostics

### Log-likelihood (fitted)

$$LL = \sum_{i=1}^n \log\left(y_i \times G(\hat{z}_i) + (1 - y_i) \times (1 - G(\hat{z}_i))\right)$$

- $\hat{z}_i = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k$
- $G(\hat{z}_i)$  is the fitted value of  $Prob(y = 1|\mathbf{x})$

### Example

- $G(\hat{z}_i) = 0.9$ : predicted that individual  $i$  is very likely to own a car
- $y_i = 0$ : in reality, individual  $i$  does not own a car

$\Rightarrow$

$$\log\left(0 \times 0.9 + (1 - 0) \times (1 - 0.9)\right) = \log(0.1) = -2.3$$

### Log-likelihood (fitted)

$$LL = \sum_{i=1}^n \log\left(y_i \times G(\hat{z}_i) + (1 - y_i) \times (1 - G(\hat{z}_i))\right)$$

- $\hat{z}_i = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k$
- $G(\hat{z}_i)$  is the fitted value of  $Prob(y = 1|\mathbf{x})$

### Example

- $G(\hat{z}_i) = 0.9$ : predicted that individual  $i$  is very likely to own a car
- $y_i = 1$ : in reality, individual  $i$  indeed owns a car

$\Rightarrow$

$$\log\left(1 \times 0.9 + (1 - 1) \times (1 - 0.9)\right) = \log(0.9) = -0.11$$

### Log-likelihood (fitted)

So, the better your prediction (model fit) is the (The greater (less negative) the LL is.

### McFadden's pseudo- $R^2$

A measure of how much better your model is compared to the model with only the intercept.

$$pseudo - R^2 = 1 - LL/LL_0$$

where  $LL_0$  is the log-likelihood when you include only the intercept.



## R code

```
logit_lf_0 <- glm(  
  inlf ~ 1,  
  data = data,  
  family = binomial(link = "logit")  
)  
  
### extract LL using the logLik() function ###  
(LL0 <- logLik(logit_lf_0))
```

```
## 'log Lik.' -514.8732 (df=1)
```

```
### extract LL using the logLik() function from your preferred model ###  
(LL <- logLik(logit_lf))
```

```
## 'log Lik.' -401.7652 (df=8)
```

```
### pseudo R2 ###  
1 - LL / LL0
```

```
## 'log Lik.' 0.2196814 (df=8)
```

## Alternatively

```
#--- or more easily ---#  
1 - logit_lf$deviance / logit_lf$null.deviance
```

```
## [1] 0.2196814
```

```
#--- what are deviances? ---#  
logit_lf$null.deviance # =  $-2 \times LL_0$ 
```

```
## [1] 1029.746
```

```
logit_lf$deviance # =  $-2 \times LL$ 
```

```
## [1] 803.5303
```

- `null.deviance` =  $-2 \times LL_0$
- `deviance` =  $-2 \times LL$

## Testing joint significance

You can do Likelihood Ratio (LR) test:

$$LR = 2(LL_{unrestricted} - LL_{restricted}) \sim \chi^2_{df\_restrictions}$$

where  $df\_restrictions$  is the number of restrictions.

### Note

LR test is very similar conceptually to F-test.

### Example

- $H_0$  : the coefficients on *exper*, *exper2*, and *age* are 0
- $H_1$  :  $H_0$  is false

```

#--- unrestricted ---#
logit_ur <- glm(
  inlf ~ nwifeinc + educ + exper + exper2 + age + kidslt6 + kidsge6,
  data = data, family = binomial(link = "logit")
)

#--- restricted ---#
logit_r <- glm(
  inlf ~ nwifeinc + educ + kidslt6 + kidsge6,
  data = data, family = binomial(link = "logit")
)

#--- LR test using lrtest() from the lmtest package ---#
library(lmtest)
lrtest(logit_r, logit_ur)

```

```

## Likelihood ratio test
##
## Model 1: inlf ~ nwifeinc + educ + kidslt6 + kidsge6
## Model 2: inlf ~ nwifeinc + educ + exper + exper2 + age + kidslt6 + kidsge6
##   #Df  LogLik Df  Chisq Pr(>Chisq)
## 1    5 -464.92
## 2    8 -401.77  3 126.32  < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

## Prediction

After estimating a binary choice model, you can easily predict the following two

- $\hat{z} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k$
- $\widehat{Prob}(y = 1 | \mathbf{x}) = G(\hat{z}) = G(\hat{\beta}_0 + \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_k x_k)$

## R code

```
#--- z hat ---#  
z <- predict(probit_lf, type = "link")  
head(z)
```

```
##           1           2           3           4           5           6  
## 0.5071349 0.6624576 0.5116317 0.7423429 0.1972781 0.8837878
```

```
#--- G(z) hat ---#  
Gz <- predict(probit_lf, type = "response")  
head(Gz)
```

```
##           1           2           3           4           5           6  
## 0.6939699 0.7461610 0.6955456 0.7710602 0.5781950 0.8115946
```

## Marginal effect of an independent variable

- Coefficient estimates across different models (probit and logit) are not meaningful because the same value of a coefficient estimate means different things
- They are the estimates of  $\beta$ s, not the direct impact of the independent variables on the  $Prob(y = 1)$

## Marginal effect of an independent variable

$$\frac{\partial \Pr(y=1|x_1, \dots, x_k)}{\partial x_k} = G'(z) \times \beta_k$$

- the marginal impact depends on the current levels of all the independent variables
- we typically report one of the two types of marginal impacts
  - (becoming obsolete) the marginal impact **at the mean** (average person): when all the independent variables take on their respective means
  - the average of the marginal impacts calculated for each of all the individuals observed



### Marginal impact at the mean

$$\frac{\partial Pr(y=1|\bar{x}_1, \dots, \bar{x}_k)}{\partial x_k} = G'(\beta_0 + \beta_1 \bar{x}_1 + \dots + \beta_k \bar{x}_k) \times \beta_k$$

### Mean marginal impact (MME)

$$\sum_{i=1}^n \frac{\partial Pr(y_i=1|x_{i,1}, \dots, x_{i,k})}{\partial x_k} = \sum_{i=1}^n G'(z_i) \times \beta_k$$

## R codes to get MME

```
### get z for all the individuals ###
z <- predict(probit_lf, type = "link")

### get G'(z) ###
Gz_indiv <- dnorm(z)

### mean marginal impact of education ###
mean(Gz_indiv) * probit_lf$coef["educ"]
```

```
##      educ
## 0.03937009
```

Fortunately, the `margins` package provides you with a more convenient way of calculating MMEs.

```
library(margins)

#--- calculate MME based on the probit estimation ---#
mme_lf <- margins(probit_lf, type = "response")

#--- get the summary ---#
summary(mme_lf)
```

##	factor	AME	SE	z	p	lower	upper
##	age	-0.0159	0.0024	-6.7392	0.0000	-0.0205	-0.0113
##	educ	0.0394	0.0073	5.4186	0.0000	0.0251	0.0536
##	exper	0.0371	0.0052	7.1779	0.0000	0.0270	0.0472
##	exper2	-0.0006	0.0002	-3.2050	0.0014	-0.0009	-0.0002
##	kidsge6	0.0108	0.0132	0.8189	0.4129	-0.0151	0.0367
##	kidslt6	-0.2612	0.0319	-8.1860	0.0000	-0.3237	-0.1986
##	nwifeinc	-0.0036	0.0015	-2.4604	0.0139	-0.0065	-0.0007

## Multinomial Choice Model

---

## Multinomial Choice

Instead of two options, you are picking one option out of more than two options

- which carrier?
  - Verizon
  - Sprint
  - AT\&T
  - T-mobile
- which transportation means to commute?
  - drive
  - Uber
  - bus
  - train
  - bike

## Multinomial logit model

The most popular model to analyze multinomial choice

- environmental evaluation
- transportation
- marketing

# Understanding multinomial logit model

## Choice of train route options

- 10 euros, 30 minutes travel time, one change
- 20 euros, 20 minutes travel time, one change
- 22 euros, 22 minutes travel time, no change

# Understanding multinomial logit model

## Choice of train route options

- 10 euros, 30 minutes travel time, one change
- 20 euros, 20 minutes travel time, one change
- 22 euros, 22 minutes travel time, no change

## Associated utility

- $V_1 = \alpha_1 + \beta_{10} + \gamma_{30} + \rho_1 + v_1$
- $V_2 = \alpha_2 + \beta_{20} + \gamma_{20} + \rho_1 + v_2$
- $V_3 = \alpha_3 + \beta_{22} + \gamma_{22} + \rho_0 + v_3$



### Choice probability

Logit model assumes that the probability of choosing an alternative is the following:

- $P_1 = \frac{e^{V_1}}{e^{V_1} + e^{V_2} + e^{V_3}}$
- $P_2 = \frac{e^{V_2}}{e^{V_1} + e^{V_2} + e^{V_3}}$
- $P_3 = \frac{e^{V_3}}{e^{V_1} + e^{V_2} + e^{V_3}}$

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- $P_1 = \frac{e^{V_1}}{e^{V_1} + e^{V_2} + e^{V_3}}$
- $P_2 = \frac{e^{V_2}}{e^{V_1} + e^{V_2} + e^{V_3}}$
- $P_3 = \frac{e^{V_3}}{e^{V_1} + e^{V_2} + e^{V_3}}$

## Notes

- $0 < P_j < 1, \forall j = 1, 2, 3$
- $\sum_{j=1}^3 P_j = 1$

### Modeled probability of choices

Modeled probability of observing individual  $i$  choosing the option  $i$  chose

$$P_i = \prod_{j=1}^3 y_{i,j} \times P_j$$

where  $y_{i,j} = 1$  if  $i$  chose  $j$ , 0 otherwise.

### Example

$$y_{i,1} = 0, \quad y_{i,2} = 1, \quad y_{i,3} = 0$$

$$P_i = \prod_{j=1}^3 y_{i,j} \times P_j = 0 \times P_1 + 1 \times P_2 + 0 \times P_3$$

The probability of observing a series of choices made by all the subjects is

$$LL = \prod_{i=1}^n P_i = \prod_{i=1}^n \prod_{j=1}^3 y_{i,j} \times P_j$$

if choices made by the subjects are independent with each other.

The probability of observing a series of choices made by all the subjects is

$$LL = \prod_{i=1}^n P_i = \prod_{i=1}^n \prod_{j=1}^3 y_{i,j} \times P_j$$

if choices made by the subjects are independent with each other.

**MLE**

$$\text{Max}_{\beta, \gamma, \rho} \log(LL)$$

## Interpretation of the coefficients

### Model in general

$$V_{i,j} = \alpha_j + \beta_1 x_{1,i,j} + \cdots + \beta_k x_{k,i,j}$$

$$P_{i,j} = \frac{e^{V_{i,j}}}{\sum_{k=1}^J e^{V_{i,k}}}$$

# Interpretation of the coefficients

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## Interpretation of the coefficients

$$\frac{\partial P_{i,j}}{\partial x_{k,i,j}} = \beta_k P_{i,j}(1 - P_{i,j})$$

- A marginal change in  $k$ th variable for alternative  $j$  would change the probability of choosing alternative  $j$  by  $\beta_k P_{i,j}(1 - P_{i,j})$
- the sign of the impact is the same as the sign of the coefficient

## Implementation in (R)

You can use *mlogit* package to estimate multinomial logit models:

- format your data in a specific manner
- convert your data using *mlogit.data()*
- estimate the model using *mlogit()*



```

#--- library ---#
library(mlogit)

#--- get the travel mode data from the mlogit package ---#
data("TravelMode", package = "AER")

#--- take a look at the data ---#
# first 10 rows
head(TravelMode, 10)

```

##	individual	mode	choice	wait	vcost	travel	gcost	income	size
## 1	1	air	no	69	59	100	70	35	1
## 2	1	train	no	34	31	372	71	35	1
## 3	1	bus	no	35	25	417	70	35	1
## 4	1	car	yes	0	10	180	30	35	1
## 5	2	air	no	64	58	68	68	30	2
## 6	2	train	no	44	31	354	84	30	2
## 7	2	bus	no	53	25	399	85	30	2
## 8	2	car	yes	0	11	255	50	30	2
## 9	3	air	no	69	115	125	129	40	1
## 10	3	train	no	34	98	892	195	40	1

## R code: data preparation

```
#--- convert the data ---#
TM <- mlogit.data(TravelMode,
  shape = "long", # what format is the data in?
  choice = "choice", # name of the variable that indicates choice made
  chid.var = "individual", # name of the variable that indicates who made choices
  alt.var = "mode" # the name of the variable that indicates options
)
```

```
### take a look at the data ###  
# first 10 rows  
head(TM, 10)
```

```
## ~~~~~  
## first 10 observations out of 840  
## ~~~~~  
## individual mode choice wait vcost travel gcost income size idx  
## 1 1 air FALSE 69 59 100 70 35 1 1:air  
## 2 1 train FALSE 34 31 372 71 35 1 1:rain  
## 3 1 bus FALSE 35 25 417 70 35 1 1:bus  
## 4 1 car TRUE 0 10 180 30 35 1 1:car  
## 5 2 air FALSE 64 58 68 68 30 2 2:air  
## 6 2 train FALSE 44 31 354 84 30 2 2:rain  
## 7 2 bus FALSE 53 25 399 85 30 2 2:bus  
## 8 2 car TRUE 0 11 255 50 30 2 2:car  
## 9 3 air FALSE 69 115 125 129 40 1 3:air  
## 10 3 train FALSE 34 98 892 195 40 1 3:rain
```

```
#--- estimate ---#  
ml_reg <- mlogit(choice ~ wait + vcost + travel, data = TM)
```

```
summary(ml_reg)
```

```
##
## Call:
## mlogit(formula = choice ~ wait + vcost + travel, data = TM, method = "nr")
##
## Frequencies of alternatives:choice
##      air      train      bus      car
## 0.27619 0.30000 0.14286 0.28095
##
## nr method
## 5 iterations, 0h:0m:0s
## g'(-H)^-1g = 0.000192
## successive function values within tolerance limits
##
## Coefficients :
##              Estimate Std. Error z-value Pr(>|z|)
## (Intercept):train -0.7866667  0.60260733 -1.3054  0.19174
## (Intercept):bus   -1.43363372  0.68071345 -2.1061  0.03520 *
## (Intercept):car   -4.73985647  0.86753178 -5.4636 4.665e-08 ***
## wait              -0.09688675  0.01034202 -9.3683 < 2.2e-16 ***
```

## Understanding the results

```
summary(ml_reg)$coef
```

```
## (Intercept):train (Intercept):bus (Intercept):car wait \
## -0.786666672 -1.433633718 -4.739856473 -0.096886747 -0.01391
## attr(,"names.sup.coef")
## character(0)
## attr("fixed")
## (Intercept):train (Intercept):bus (Intercept):car wait \
## FALSE FALSE FALSE FALSE F
## attr(,"sup")
## character(0)
```

- intercept for *air* is dropped (*air* is the base)
  - train:(intercept) is  $-0.786$  means that train is less likely to be chosen if all the other **included** variables are the same
- the greater the travel time, the less likely the option is chosen