

Moment Generating Functions (MGFs): $X$  - a r.v.For each  $t \in \mathbb{R}$ ,  $e^{tX}$  is a non-ve r.v. $E(e^{tX}) = m_X(t)$  : moment generating function of  $X$ .

$I = \{t \in \mathbb{R} : m_X(t) < \infty\}.$

$0 \in I.$

Fact:  $I$  is a convex set.Proof:

$t_1, t_2 \in I. \quad m_X(t_1), m_X(t_2) < \infty$

$\Rightarrow \forall 0 < \alpha < 1.$

$E(e^{(\alpha t_1 + (1-\alpha)t_2)X})$

$= E(e^{\alpha t_1 X} \cdot e^{(1-\alpha)t_2 X})$

$\leq (E(e^{\alpha t_1 X}))^\alpha \cdot (E(e^{(1-\alpha)t_2 X}))^{1-\alpha} \quad \begin{array}{l} \text{use Holder's inequality} \\ \text{with } \alpha = \frac{1}{p} \\ \therefore 1-\alpha = \frac{1}{q} \end{array}$

$= (m_X(t_1))^\alpha \cdot (m_X(t_2))^{1-\alpha} < \infty \quad \begin{bmatrix} \alpha \cdot p = 1 \\ (1-\alpha) \cdot q = 1 \end{bmatrix}$

$\therefore t_1, t_2 \in I \Rightarrow \underbrace{\alpha t_1 + (1-\alpha)t_2}_{\text{convex combination}} \in I. \quad \alpha \in (0,1)$

 $\therefore I$  is a convex set.  $\blacksquare$ Corollary:  $\log(m_X(t))$  on  $I$  is a convex function.Note:  $I$  is an interval containing 0. $I = \{0\}$  is possible.

$P(X = \pm n) = c \cdot \frac{1}{n^2}, \quad n \in \mathbb{N}$

Special Case: $n \dots \text{interior point of } I.$

Special Case:

0 is an interior point of I.

Equivalently,  $[-t_0, t_0] \subset I$  for some  $t_0 > 0$ .

Observe: for any  $t \in [-t_0, t_0]$ ,

$$e^{|tx|} \leq e^{t_0 x} + e^{-t_0 x}$$

$$\therefore E(e^{|tx|}) < \infty \quad \forall t \in [-t_0, t_0].$$

Fact: for any  $\theta > 0$ ,  $\exists$  a finite constant  $C (= C_\theta)$

such that,  $x < C \cdot e^{\theta x}$  &  $x \geq 0$ . [Why?  $\because \frac{x}{e^{\theta x}} < c$ ]  
 or,  $|x| < C \cdot e^{\theta |x|}$   
 $\Rightarrow |x|^k < C e^{\theta |x|}$ .  
 $\downarrow$   
 as  $x \rightarrow \infty$ .

$\therefore \forall k \geq 1$ ,  $\exists c$  s.t.

$$|x|^k \leq c \cdot e^{|tx|} \Rightarrow E(|x|^k) < \infty.$$

finite  
in the  
interval  
 $[-t_0, t_0]$ .

$\therefore$  For any  $t \in [-t_0, t_0]$ ,

$$\left| \underbrace{\sum_{k=0}^n \frac{(tx)^k}{k!}} \right| \leq e^{|tx|}.$$

$$\sum_{k=0}^n \frac{(tx)^k}{k!} \longrightarrow e^{tx}.$$

$$Y_n \longrightarrow Y.$$

$$\text{So, by DCT, } E \left( \sum_{k=0}^n \frac{(tx)^k}{k!} \right) \longrightarrow E(e^{tx}) = m_X(t)$$

$$\Rightarrow \sum_{k=0}^n \frac{E(X^k) \cdot t^k}{k!} \longrightarrow m_X(t).$$

$$\Rightarrow m_x(t) = \sum_{n=0}^{\infty} E(X^n) \cdot \frac{t^n}{n!}$$

- \*  $X$  has all moments finite.
- \* For all  $t \in [-t_0, t_0]$ ,  $m_x(t) = \sum_{n=0}^{\infty} E(X^n) \cdot \frac{t^n}{n!}$ .
- \* Thus,  $m_x(t)$  has a power series expansion on an interval around 0 of positive radius of convergence.

$$\Rightarrow E(X^n) = m_x^{(n)}(0)$$

$\nearrow$   $n^{\text{th}}$  (raw) moment  $= n^{\text{th}}$  derivative of mgf evaluated at 0.

Fact:  $h: I \times \mathbb{R} \rightarrow \mathbb{R}$ ,  $h$  - continuous in both variables.  
 $\uparrow$   
open intervals.

Let  $X$  be a real random variable.

Suppose,  
 $\exists$  r.v.  $Y$  with  $E(Y) < \infty$ .

s.t.  $|h(t, x)| \leq Y$   $\forall t \in I$  &  $\forall x$ .

Then,  $g(t) = E(h(t, x))$  is continuous on  $I$

Proof: Exercise (use DCT)

$\forall \{t_n\}_{n \geq 1}$  s.t  $t_n \rightarrow t$

then,  $h(t_n, x) \rightarrow h(t, x)$  [by continuity of  $X$ ]

$\therefore |h(t_n, x)| \leq Y$ .

$\therefore$  By DCT,  $E(h(t_n, x)) \rightarrow E(h(t, x))$ .

$\Rightarrow g(t_n) \rightarrow g(t)$   
 $\therefore g$  is continuous.  $\square$

Fact:

Suppose  $\forall n$ ,

$\frac{\partial}{\partial t}(h(t, x))$  exists at all  $t \in I$ , and

$\exists \dots \exists \dots \exists \dots$  with  $E(Z) < \infty$ ,

$\frac{\partial}{\partial t} (h(t, x))$  exists at all  $t \in I$ , and  
 $\exists$  a r.v.  $Z$  with  $E(Z) < \infty$ ,  
s.t.  $\left| \frac{\partial}{\partial t} (h(t, x)) \right| \leq Z$

Then,  $g(t) = E(h(t, x))$  is differentiable  
&  $g'(t) = E\left(\frac{\partial}{\partial t} h(t, x)\right)$

[ie, in layman's words,  
expectation & differentiation  
can be swapped.]

Proof: (Exercise). Use DCT & MVT.

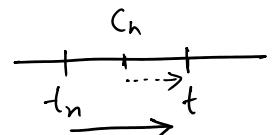
&  $\{t_n\}_{n \geq 1}$  s.t.  $t_n \rightarrow t$ ,

$$\frac{h(t_n, x) - h(t, x)}{t_n - t} = \frac{\partial}{\partial t} h(c_n, x) \quad \begin{array}{l} \text{for some} \\ c_n \in (t_n, t) \\ \text{By MVT} \end{array} \quad \textcircled{1}$$

taking  $E$  on both sides -

$$E\left(\frac{h(t_n, x) - h(t, x)}{t_n - t}\right) = E\left(\frac{\partial}{\partial t} h(c_n, x)\right) \quad \textcircled{2}$$

As  $\frac{\partial}{\partial t} (h(t, x))$  exists &  $t \in I$ ,



$$\lim_{n \rightarrow \infty} \frac{\partial}{\partial t} (h(c_n, x)) = \frac{\partial}{\partial t} (h(t, x))$$

$$\&, \because \left| \frac{\partial}{\partial t} (h(c_n, x)) \right| \leq Z \quad \forall n, \quad \& E(Z) < \infty$$

$$\therefore \text{By DCT, } E\left(\frac{\partial}{\partial t} (h(c_n, x))\right) \rightarrow E\left(\frac{\partial}{\partial t} (h(t, x))\right)$$

using  $\textcircled{2}$

$$E\left(\frac{h(t_n, x) - h(t, x)}{t_n - t}\right) \rightarrow E\left(\frac{\partial}{\partial t} (h(t, x))\right)$$

linearity  
of  
expectation.

$$\frac{E(h(t_n, x)) - E(h(t, x))}{t_n - t} \rightarrow E\left(\frac{\partial}{\partial t} (h(t, x))\right)$$

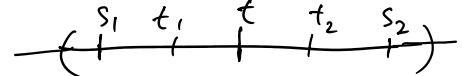
$$\text{as } n \rightarrow \infty, \therefore \frac{\partial}{\partial t} (E(h(t, x))) = E\left(\frac{\partial}{\partial t} (h(t, x))\right)$$



Suppose  $J$  is an open interval  $m_x(t) < \infty \quad \forall t \in J$ .

Then,  $m_x(t)$  is infinitely differentiable on  $J$ , &

$$m_x^{(n)}(t) = E(X^n \cdot e^{tx}) \quad \forall t \in J.$$



Proof: Fix points  $s_1 < t_1 < t_2 < s_2$  in  $J$ .

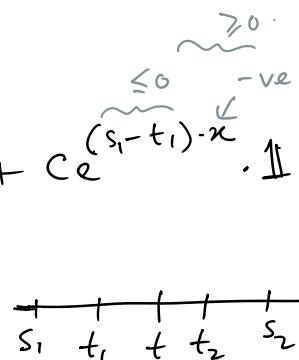
We'll prove:  $m_x(t)$  is differentiable at  $t \in (t_1, t_2)$  &

$$m'_x(t) = E(Xe^{tx}) \quad \forall t \in (t_1, t_2)$$

(the rest of  
the proof  
follows by  
induction.)

We can get a constant  $C$  s.t,

$$|x| \leq Ce^{\frac{(s_2-t_2)x}{\geq 0}} \cdot \mathbb{1}_{x \geq 0} + Ce^{\frac{(s_1-t_1)x}{\leq 0}} \cdot \mathbb{1}_{x < 0}.$$



$\Rightarrow$  for  $t_1 < t < t_2$ :

Substitute  $|x|$  by this.

$$\begin{aligned} \left| \frac{\partial}{\partial x} (e^{tx}) \right| &= |x| \cdot e^{tx} \\ &\leq Ce^{\frac{(s_2-t_2+t)x}{\leq 0}} \cdot \mathbb{1}_{x \geq 0} + Ce^{\frac{(s_1-t_1+t)x}{\geq 0}} \cdot \mathbb{1}_{x < 0} \\ &\leq Ce^{s_2 x} \cdot \mathbb{1}_{x \geq 0} + Ce^{s_1 x} \cdot \mathbb{1}_{x < 0} \rightarrow \end{aligned}$$

this bound is  
free of  
 $t$ .  
Hence,  
can be  
used  
as  $Z$ .

$$\therefore \left| \frac{\partial}{\partial x} (e^{tx}) \right| \leq Ce^{s_2 x} \cdot \mathbb{1}_{x \geq 0} + Ce^{s_1 x} \cdot \mathbb{1}_{x < 0}$$

this is a r.v  
with finite  
expectation.

$\therefore$  By DCT, (2nd part of the "Fact":)

$$E\left(\frac{\partial}{\partial t} (e^{tx})\right) = \frac{\partial}{\partial t} \cdot E(e^{tx}). \text{ exists.}$$

& is equal to  $E(Xe^{tx}) \quad \forall t \in (t_1, t_2)$

for higher order derivatives, just use the same approach and conclude by induction.

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Special Case:

$$X \geq 0.$$

$$\text{This } I \supset (-\infty, 0]$$

$$m_X(t) < \infty \quad \forall t \leq 0.$$

$$\text{Denote } L_X(t) = m_X(-t)$$

$$= E(e^{-tX})$$

$$X > 0, t < 0$$

$$-tX \leq 0$$

$$e^{-tX} \leq 1$$

$$E(e^{-tX}) < \infty \checkmark$$

$$\text{Hence, } I \supset (-\infty, 0]$$

$L_X$  is called the "Laplace Transform" of  $X$ .

We'll prove: (in the problem set- 5)

①  $L_X$  is continuous on  $[0, \infty)$

②  $L_X$  is infinitely differentiable on  $(0, \infty)$

$$\&, L_X^{(n)}(t) = (-1)^n \cdot E(X^n e^{-tX}).$$

Further, the right limit  $L_X^{(n)}(0^+)$  exists & is finite  $\Leftrightarrow E(X^n) < \infty$ .

Digression:

$$\text{Let } Z_{nx} \sim \text{Poi}(nx) \quad n \geq 1 \\ x \geq 0.$$

$$P(Z_{nx} \leq na) \rightarrow ?$$

for  $a > 0$ .

Case-I : ( $x < a$ ):

$$P(Z_{nx} \leq na)$$

$$= P(Z_{nx} - E(Z_{nx}) \leq n(a-x))$$

$$= 1 - P(Z_{nx} - E(Z_{nx}) > n(a-x))$$

$$\leq \frac{V(Z_{nx})}{(n(a-x))^2}$$

$$\therefore X \sim \text{Poi}(\lambda) \Rightarrow E(X) = \lambda, \\ V(X) = \lambda$$

$$E(Z_{nx}) = nx,$$

$$V(Z_{nx}) = nx.$$

By Chebyshev's Inequality

$$\overline{(n(x-a))^2} \quad L \quad \text{Inequality 1}$$

$$\begin{aligned} &\geq 1 - \frac{n^2}{n^2(x-a)^2} \\ &= 1 - \frac{x-a}{n(x-a)^2} \rightarrow 1 \text{ as } n \rightarrow \infty. \end{aligned}$$

Case-II: ( $x > a$ ) .

$$\begin{aligned} P(Z_{nx} \leq na) &= P(-Z_{nx} \geq -na) \\ &= P(-(Z_{nx} - E(Z_{nx})) \geq n(x-a)) \\ &= P(|Z_{nx} - E(Z_{nx})| \geq n(x-a)) \end{aligned}$$

$$(\text{Again, by Chebychev's Inequality}) \leq \frac{x-a}{n(x-a)^2} \rightarrow 0.$$

$$\therefore P(Z_{nx} \leq na) \rightarrow \begin{cases} 1, & x < a \\ 0, & x > a. \end{cases}$$

↓ pmf of  $\text{Poi}(nx)$

$$\sum_{k=0}^{na} e^{-nx} \frac{(nx)^k}{k!} \rightarrow \begin{cases} 1 & \text{if } x < a \\ 0 & \text{if } x > a. \end{cases}$$

$\forall \omega, X(\omega) \in \mathbb{R}.$  hence this holds .

$$\therefore \sum_{k=0}^{na} e^{-nx} \frac{(nx)^k}{k!} \xrightarrow{\text{(pointwise convergence)}} \begin{cases} 1 & \text{on } \{X < a\} \\ 0 & \text{on } \{X > a\} \end{cases}$$

Suppose  $a > 0$  is such that ,

$$P(X=a)=0.$$

$$\therefore P(X > a) + P(X < a) = 1.$$

$$\text{Also, } \left| \sum_{k=0}^{na} e^{-nx} \frac{(nx)^k}{k!} \right| \leq 1 \leftarrow (\text{"Z" for using DCT}\right) \text{ & } E(1) = 1 < \infty.$$

$$\therefore \text{By DCT, } \lim_{n \rightarrow \infty} \sum_{k=0}^{na} e^{-nx} \cdot \frac{(nx)^k}{k!} = \mathbb{1}_{\{x < a\}}$$

(taking  
E on  
both sides)

$$\Rightarrow P(X \leq a) = \lim_{n \rightarrow \infty} \sum_{k=0}^{na} \frac{n^k}{k!} \cdot E(X^k \cdot e^{-nx})$$

$\downarrow$

$$F_X(a) = \lim_{n \rightarrow \infty} \sum_{k=0}^{na} (-1)^k \cdot \frac{n^k}{k!} \cdot L_X^{(k)}(n) \leftarrow \begin{array}{l} \text{mgf of} \\ \text{Laplace} \\ \text{transform} \end{array}$$

as  $P(X=a)=0$

i.e., has no mass.

$\downarrow$   
So, this works at pts of continuity  
of the distribution of  $X$ .

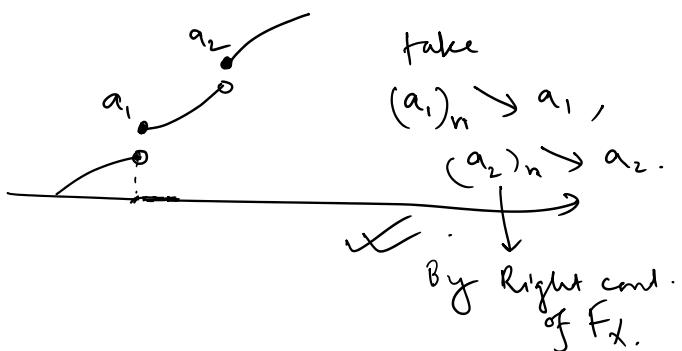
As for the pts of discontinuity,  
we can use right continuity of  
 $F_X$ .

So far we've shown,  $L_X$  determines the CDF of  $X$  at  
only those points where probability Mass = 0.

for points where probability Mass  $> 0$   
at such points there will be discontinuity in  $F_X$ .

&,  $F_X$  is monotonic. Hence, it can have atmost  
countably many discontinuities.

Hence, set of such points must be atmost countable.



Conclusion:

Laplace transform of a non-ve r.v.  
completely determines the distribution of that  
random variable.

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