

CS6210 - Homework/Assignment-4

Arnab Das(u1014840)

November 7, 2016

1: Chapter-6, question-4

(a) The techniques discussed in this chapter are for polynomial data fitting and not exponential data fitting, hence cannot be applied directly to $u(t)$.

(b) Given approximation of the form

$$u(t) = \gamma_1 \exp(\gamma_2 t)$$

and provided data points as $(t_1, z_1), (t_2, z_2), \dots, (t_m, z_m)$, where $z_i > 0, i = 1, 2, \dots, m$ and $m > 0$. Considering instead the following approximation:

$$v(t) = \ln u(t) = (\ln \gamma_1) + \gamma_2 t$$

such that the data points become $(t_1, b_1), (t_2, b_2), \dots, (t_m, b_m)$ where $b_i = \ln z_i$ and

$$v(t) = x_1 + x_2 t$$

such that $x_1 = \ln \gamma_1$ and $x_2 = \gamma_2$ and $v(t) = \ln u(t)$

Then for solving $A^T A x = A^T b$, we define the following matrices:

$$A = \begin{bmatrix} 1 & t_1 \\ 1 & t_2 \\ \dots & \dots \\ 1 & t_m \end{bmatrix} \text{ leading to } B = A^T A = \begin{bmatrix} \sum_{i=1}^m 1 = m & \sum_{i=1}^m t_i \\ \sum_{i=1}^m t_i & \sum_{i=1}^m t_i^2 \end{bmatrix} \text{ and } b = \begin{bmatrix} t_1 \\ t_2 \\ \dots \\ t_m \end{bmatrix} \text{ leading to } A^T b = \begin{bmatrix} \sum_{i=1}^m b_i \\ \sum_{i=1}^m t_i b_i \end{bmatrix} \text{ Thus solving}$$

$$B = A^T A x = A^T b$$

we get the following two equations:

$$x_1 + x_2 = 1 \tag{1}$$

and

$$3x_1 + 5x_2 = 4.9 \tag{2}$$

Solving (i) and (ii) , we get: $x_1 = \ln \gamma_1 = 0.5$ and $x_2 = \gamma_2 = 0.95$

Thus, we have ,

$$\begin{aligned} v(t) &= \ln u(t) = \ln(0.5) + 0.95t \\ \Rightarrow u(t) &= 0.5 \times \exp(0.95t) \end{aligned}$$

(Answer).

2: Chapter-6, question-5

(a) For tall and skinny matrices, A , of the system of equations , $Ax = b$, the number of rows is much larger than the number of columns. Let the number of rows be m and the number of columns n , then in such cases generally $m \gg n$. These systems are overdetermined and b is generally not in the range space of A . Thus applying LU does not gives a unique solution to $Ax=b$. Instead, a better way to approach such problems is to minimize the residual $\|b - Ax\|$, such that within the tolerance of the residual we have the best solution for x . When transforming to the normal equation, $A^T A x = A^T b$, here one can use LU decomposition since we have $n \times n$ matrix, however cholesky decomposition is a better choice here since the matrix $A^T A$ is symmetric positive definite. In case of the QR decomposition, which has an upper triangular part in R , however the Q matrix allows us to extract the upper-triangular system of equation whose solution leads to a solution of the least square problem. Here, the orthonormal behaviour of Q is used to transform into

equivalent set of equations such that the **norms are not affected**. Finally, SVD is used more in cases where A is rank deficient or nearly rank deficient, in which cases LU cannot be used. Thus, in all the cases LU directly doesn't fit the scenario for application except that LU requires to be slightly modified so that for every column it zeroes out, the vector of its remaining elements is orthonormal to the previous columns. With this introduction of orthonormality, it can be used in QR decomposition.

(b) In the normal equation: the way condition number came to be $K(B) = K^2(A)$. This was because of the following derivation:

In normal equation we were solving:

$$A^T A x = A^T b$$

and

$$(A^T A)^{-1} A^T = V S^{-2} V^T V \sum^T U^T = V (S^{-1} 0) U^T$$

where $A = U \sum V^T$, V and U are orthogonal matrices and \sum is a diagonal matrix with singular values of A along the diagonal, and hence $A^T A = V \sum U^T U \sum V^T = V \sum^2 V^T$. Thus $\|(A^T A)^{-1} A^T\| = \|(S^{-1} 0)\|$ and $\|((A^T A)^{-1} A^T)^{-1}\| = \|V(S 0) U^T\|$. Thus the condition number becomes:

$$K(B = A^T A) = \frac{\lambda_1}{\lambda_n} = \frac{\sigma_1^2}{\sigma_n^2} = K^2(A)$$

Now, for QR, we write $x = (A^T A)^{-1} A^T b$, and we have represented $A=QR$ by a QR decomposition where Q is orthonormal and of same dimension as A. The following translation results in the final form:

$$x = (A^T A)^{-1} A^T b = (R^T Q^T Q R)^{-1} R^T Q^T b = (R^T R)^{-1} R^T Q^T b = R^{-1} Q^T b$$

Hence, we get the following relation,

$$x = R^{-1} Q^T b$$

Notice, that multiplication by orthogonal matrices do not affect the norms, thus,

$$\|R^{-1} Q^T\| = \|R^{-1}\| = \|A^{-1}\|$$

and also,

$$\|(Q^T)^{-1} R\| = \|R\| = \|A\|$$

Thus the condition number in this case comes to be $\|A\| \|A^{-1}\| = K(A)$.

The main saving comes due to usage of the orthonormal decomposition for the transformations, since the transformations using q only transform them in space without affecting the norm values, while for the previous normal equation, the singular values were getting multiplied during creation $A^T A$, introducing change in norm values during the transformations.

3: Chapter-8, question-8

(a) Given a rank deficient matrix, we analyse here its effect on the Gram-Schmidt process. Consider the span of three vectors v_1, v_2 and v_3 , such that v_3 is linearly dependent on v_1 and v_2 . We can write v_3 as a linear combination of v_1 and v_2 , as:

$$\mathbf{v}_3 = \mathbf{a}\mathbf{v}_1 + \mathbf{b}\mathbf{v}_2$$

Let \mathbf{u}_i denote the orthonormal unit vectors we generate along the Gram-Schmidt process. Then, for the first vector, \mathbf{v}_1 , we have

$$\mathbf{u}_1 = \frac{\mathbf{v}_1}{|\mathbf{v}_1|}$$

And the orthogonal transformation of \mathbf{v}_2 that is orthogonal to \mathbf{v}_1 is

$$\mathbf{y}_2 = \mathbf{v}_2 - (\mathbf{v}_2 \cdot \mathbf{u}_1) \mathbf{u}_1$$

Then, the next orthonormal vector, \mathbf{u}_2 will be along \mathbf{y}_2 , such that $\mathbf{u}_2 = \frac{\mathbf{y}_2}{|\mathbf{y}_2|}$.

Now the orthogonal transformation of \mathbf{v}_3 , say \mathbf{y}_3 , that is orthogonal to the span($\mathbf{u}_1, \mathbf{u}_2$), will be,

$$\begin{aligned} \mathbf{y}_3 &= \mathbf{v}_3 - \left((\mathbf{v}_3 \cdot \mathbf{u}_1) \mathbf{u}_1 + (\mathbf{v}_3 \cdot \mathbf{u}_2) \mathbf{u}_2 \right) \\ \mathbf{y}_3 &= \mathbf{v}_3 - \left(((\mathbf{a}\mathbf{v}_1 + \mathbf{b}\mathbf{v}_2) \cdot \mathbf{u}_1) \mathbf{u}_1 + ((\mathbf{a}\mathbf{v}_1 + \mathbf{b}\mathbf{v}_2) \cdot \mathbf{u}_2) \mathbf{u}_2 \right) \\ \mathbf{y}_3 &= \mathbf{v}_3 - \left(\mathbf{a}\mathbf{v}_1 + (\mathbf{b}\mathbf{v}_2 \cdot \mathbf{u}_1) \mathbf{u}_1 + (\mathbf{a}\mathbf{v}_1 \cdot \mathbf{u}_2) \mathbf{u}_2 + (\mathbf{b}\mathbf{v}_2 \cdot \mathbf{u}_2) \mathbf{u}_2 \right) \end{aligned}$$

Here:

$$(\mathbf{b}\mathbf{v}_2 \cdot \mathbf{u}_1) \mathbf{u}_1 = \mathbf{b}(\mathbf{v}_2 - \mathbf{y}_2)$$

$$(\mathbf{a}\mathbf{v}_1 \cdot \mathbf{u}_2) \mathbf{u}_2 = \mathbf{0}$$

$$(\mathbf{b}\mathbf{v}_2 \cdot \mathbf{u}_2) \mathbf{u}_2 = \mathbf{b}\mathbf{y}_2$$

Replacing them and the value of \mathbf{v}_3 in the above equation we get:

$$\mathbf{y}_3 = \mathbf{a}\mathbf{v}_1 + \mathbf{b}\mathbf{v}_2 - (\mathbf{a}\mathbf{v}_1 + \mathbf{b}\mathbf{v}_2 - \mathbf{b}\mathbf{y}_2 + \mathbf{0} + \mathbf{b}\mathbf{y}_2) = \mathbf{0}$$

Thus \mathbf{y}_3 , comes out to be zero, when we encounter the linearly dependent vector. (Answer).

(b) The classical pseudoCode of Gram-Schmidt is as below:

Classical-GS

```

for k=1:n
    w = a_k
    for j=1:k-1
        r_jk = q_j^T w
    end
    for j=1:k-1
        w = w - r_jk q_j
    end
    r_kk = ||w||
    q_k = w / r_kk
end

```

Modified-GS

```

for k=1:n
    w = a_k
    for j=1:k-1
        r_jk = q_j^T w
        w = w - r_jk q_j
    end
    r_kk = ||w||
    q_k = w / r_kk
end

```

Suppose at the k'th iteration the orthonormal q's already calculated are $Q_{k-1} = [q_1, q_2, \dots, q_{k-1}]$. In the **classical** case, we first calculate the projections, so suppose the calculated values of the projections in the

k'th iteration are

$$[r_{ik}, r_{2k}, \dots, r_{k-1,k}]$$

where, $r_{jk} = q_j^T w$, where w in the k'th iteration is initialised to a_k , the vector whose orthogonal transformation is done in the k'th step. Then, these projections along the corresponding orthonormal directions are subtracted from w , in the second inner loop, resulting in:

$$w = w - r_{1k}q_1 - r_{2k}q_2 - \dots - r_{k-1,k}q_{k-1} \quad (3)$$

In the modified version, instead of precomputing all the r_{jk} 's at a time, we compute r_{jk} and subtract from w , thus always *orthogonalizing* against the currently computed version. Suppose, we are in the k'th iteration, and $j=1$ results in the following computation

$$w_1 = w - r_{1k}q_1 = w - (q_1^T w)q_1$$

Then for $j=2$ in the k'th iteration:

$$w_2 = w_1 - r_{2k}q_2 = w - (q_1^T w)q_1 - q_2^T (w - (q_1^T w)q_1)q_2$$

$$w_2 = w_1 - r_{2k}q_2 = w - (q_1^T w)q_1 - (q_2^T w)q_2 - 0$$

$$w_2 = w_1 - r_{2k}q_2 = w - r_{1k}q_1 - r_{2k}q_2 - 0$$

And similarly for increasing j 's (j limits to k) it holds true since q_i and q_j are orthonormal for $i \neq j$. Thus, in exact arithmetic the modified and classical version are numerically the same. (Proved).

(c)

4: Chapter-7, Question-9

For the iterative scheme, we have $x_{k+1} = x_k + \alpha_k p_k$, where, p_k is the search direction and α_k is the step size. This includes the basic stationary methods as well of the form, $x_{k+1} = x_k + M^{-1}r_k$.

(a) Now, consider the given iterative scheme:

$$x_{k+1} = x_k + \alpha(b - Ax_k)$$

Since, $r_k = b - Ax_k$ is the residual at the k'th step, and in gradient descent, the search direction is in the reverse direction of the residual, Hence, $p_k = r_k$. Hence, for a fixed α we get:

$$M^{-1} = \alpha I$$

$$M = (\alpha I)^{-1}$$

(answer).

The iteration matrix(T) is defined as

$$T = I - M^{-1}A = I - \alpha IA = I - \alpha A$$

(answer).

(b) (i) Given that A is symmetric positive definite and its eigen values follows the inequality:

$$\lambda_1 > \lambda_2 > \dots > \lambda_n > 0 \quad (4)$$

Scaling equation(4) by α , we get the following

$$\begin{aligned}\alpha\lambda_1 &> \alpha\lambda_2 > \cdots > \alpha\lambda_n > 0 \\ -\alpha\lambda_1 &< -\alpha\lambda_2 < \cdots < -\alpha\lambda_n < 0 \\ 1 - \alpha\lambda_1 &< 1 - \alpha\lambda_2 < \cdots < 1 - \alpha\lambda_n < 1\end{aligned}\tag{5}$$

Since, $T = I - \alpha A$, then for an eigen value λ_i of A , their will be a corresponding eigen value of T as $(1 - \alpha\lambda_i)$. Thus, equation-5 gives the eigen values of T and the order of their relative magnitudes.

The theorem for Statioanry method converegence says, that if the spectral radius of the iteration matrix is less than 1, then the system converges. The spectral radius of a matrix is defined as :

$\rho(B) = \max|\lambda_i| : \lambda_i$ are the eigen values of B

Equation(5) shows that the spectral radius of T will be $|1 - \alpha\lambda_n|$ and for convergence it should be less than 1. Hence,

$$\begin{aligned}-1 &< 1 - \alpha\lambda_n < 1 \\ 0 &< \alpha < \frac{2}{\lambda_n}\end{aligned}\tag{6}$$

(condition on α for convergence).

(ii) The condition number of a symmetric positive definite matrix, A , is given by:

$$\begin{aligned}K(A) &= \frac{\lambda_{max}}{\lambda_{min}} = \frac{\lambda_1}{\lambda_n} \\ \lambda_n &= \frac{\lambda_1}{K(A)}\end{aligned}$$

and substituting this in equation(6)

$$0 < \alpha < \frac{2K(A)}{\lambda_1}\tag{7}$$

(Condition on α in terms of the condition number)

(iii) For the statioanry method, at the i 'th ieration and for iteration matrix T , the following error relation holds:

$$\|e_i\| \leq \|T\|^i e_0$$

Also, for the steepest descent at the i 'th iteration we have the following relation,

$$\|e_i\| \leq \left(\frac{k-1}{k+1}\right)^i e_0$$

where k is the condition number. Since steepest descent is the one that converges at the best speed for the class of gradient descent, we can equate these two relations to get:

$$\begin{aligned}\|T\| &= \frac{k-1}{k+1} \\ (1 - \alpha\lambda_n) &= \frac{k-1}{k+1}\end{aligned}$$

$$\alpha\lambda_n = \frac{2}{k+1} = \frac{2\lambda_n}{\lambda_1 + \lambda_n}$$

$$\alpha = \frac{2}{\lambda_1 + \lambda_n}$$

(Best value for α for maximizing speed) .

(c) "If A is strictly diagonally dominant and $\alpha = 1$, then the iterative scheme converges to the solution for any initial guess."

Since, $\alpha = 1$, our M is basically the identity matrix. Note that, for any matrix equation, $Ax=b$, we can scale each row of A and corresponding element in B by the corresponding diagonal element of that row such that the system of equation remains unaltered. By performing such an operation we can ensure that all the diagonal elements are 1, but the diagonal dominance of the original matrix still remains. Since, our M is the identity matrix, and the diagonal matrix of A is also the identity matrix, this means in our problem we chose the splitting of M to be the diagonal matrix, which means we are performing the Jacobi iteration. Thus it just remains to prove that Jacobi iterations converge if the matrix is diagonally dominant.

Proof:

By definition of diagonal dominance, we have:

$$|a_{ii}| > \sum_{j \neq i} |a_{ij}|$$

$$\sum_{j \neq i} \frac{|a_{ij}|}{|a_{ii}|} < 1$$

For Convergence, we are required to show that $\|M^{-1}N\| < 1$. Here we pick the infinity norm:

$$\|G\|_{\infty} = \|M^{-1}N\| = \|D^{-1}(L+U)\|_{\infty} = \max_{1 \leq i \leq m} \sum_{j \neq i} \frac{|a_{ij}|}{|a_{ii}|} < 1$$

Hence Jacobi converges, and by consequence of that, if a Matrix is strictly diagonally dominant and $\alpha = 1$, then the iterative scheme converges to the solution for any initial guess. (Proved).

5: Chapter-7, question-12

Given a linear system, $Ax=b$, where A is symmetric. Suppose M-N is a splitting of A, where M is symmetric positive definite and N is symmetric.

Since, M is symmetric Positive Definite, hence we can write the condition number of A as,

$$K(A) = \|M\| \|M^{-1}\| \frac{\lambda_{\max}(M)}{\lambda_{\min}(M)}$$

or, from here we get:

$$\|M^{-1}\| = \frac{\lambda_{\max}(M)}{\lambda_{\min}(M) \times \|M\|} \quad (8)$$

Now, for an iterative scheme with a splitting of $A=M-N$, we have:

$$x_{k+1} = M^{-1}Nx_k + M^{-1}b$$

This converges if the norm of $M^{-1}N$ is less than 1.

$$\|M^{-1}N\| \leq \|M^{-1}\| \|N\| = \frac{\lambda_{\max}(M)}{\lambda_{\min}(M) \times \|M\|} \|N\|$$

Now the norm of a symmetric positive definite matrix is the magnitude of its highest eigenvector. Hence, $\|M\| = \lambda_{max}$, and hence they cancel each other. Since, it is given that $\lambda_{min}(M) > \rho(N)$, and for a symmetric matrix N , $\rho(N) \leq \|N\|$. Thus, we get the following:

$$\|M^{-1}N\| \leq \frac{\|N\|}{\lambda_{min}(M)} \leq \frac{\rho(N)}{\lambda_{min}(M)} < 1$$

Hence, the system converges if $\rho(N) < \lambda_{min}(M)$ (Proved).

6: GMRES —

7: Chapter-8, question-7

(a) Given a column stochastic matrix, P , of size $n \times n$ whose all entries are non-negative and each column sum to 1.

$$A(\alpha) = \alpha P + (1 - \alpha)E$$

For this matrix, A , the entry at its i 'th row and j 'th column will be

$$a_{ij} = \alpha p_{ij} + \frac{1 - \alpha}{n}$$

Then the sum of the elements of its j 'th column will be:

$$\begin{aligned} \sum_{i=1}^n a_{ij} &= \sum_{i=1}^n \left(\alpha p_{ij} + \frac{1 - \alpha}{n} \right) \\ \sum_{i=1}^n a_{ij} &= \alpha \sum_{i=1}^n p_{ij} + \frac{1 - \alpha}{n} \sum_{i=1}^n 1 \end{aligned}$$

Since P is also stochastic, hence $\sum_{i=1}^n p_{ij} = 1$, thus we get:

$$\sum_{i=1}^n a_{ij} = \alpha + \frac{1 - \alpha}{n} \times n = 1$$

Thus, the sum of elements of a column in A is also 1. Hence, A is also stochastic. (Proved).

(b) A column stochastic matrix, or a Markov matrix, ' A ', has an eigen value of 1 and all others are less than 1. So, $\lambda = 1$, is the dominant eigen value. If A is column stochastic, then A^T is row-stochastic, thus the sum of the elements of each row of A^T is 1. So:

$$A^T \begin{bmatrix} 1 \\ 1 \\ \dots \\ 1 \end{bmatrix} = \begin{bmatrix} 1 \\ 1 \\ \dots \\ 1 \end{bmatrix}$$

Hence, A^T has an eigen value of 1. Since the determinants of $\det(A - \lambda_n I)$ and $\det(A^T - \lambda_n I)$ are the same, hence A and A^T has the same eigen value. Thus, A has an eigen value of 1.

Now, suppose A has an eigen vector, say v , whose eigen-value λ is greater than 1, that is, $|\lambda| > 1$. This implies that $A^n v = \lambda^n v$ has exponentially growing length for large $n \rightarrow \infty$. This further implies that there is a large coefficient $[A^n]_{i,j}$ which is larger than 1, since A is non-negative. Since, matrix multiplication of two stochastic matrices results in stochastic matrix, hence as A is stochastic, so A^n is also stochastic which indicates that all its entries has to be ≤ 1 , providing a contradiction. So, all eigen values other than 1, has to be less than 1. Hence,

the dominant eigen value is 1(Proved).

Let the dominant eigen-vector be v , and its corresponding eigen value is 1.

Hence,

$$A(\alpha)v = v = \alpha Pv + (1 - \alpha)Ev$$

$$(I - \alpha P)v = (1 - \alpha)Ev$$

Now, $Ev = \begin{bmatrix} 1/n & 1/n & \dots & 1/n \\ 1/n & 1/n & \dots & 1/n \\ \dots & \dots & \dots & \dots \\ 1/n & 1/n & \dots & 1/n \end{bmatrix} v = \frac{1}{n} \begin{bmatrix} 1 \\ 1 \\ \dots \\ 1 \end{bmatrix}$, as the eigen-vectors are also stochastic vectors.

Substituting them in the original equation, we get:

$$(I - \alpha P)v = (1 - \alpha) \frac{1}{n} \begin{bmatrix} 1 \\ 1 \\ \dots \\ 1 \end{bmatrix}$$

$$v = \frac{1 - \alpha}{n} (I - \alpha P)^{-1} \begin{bmatrix} 1 \\ 1 \\ \dots \\ 1 \end{bmatrix} \quad \text{This is the dominant eigen-vector of } A(\alpha).$$

(c) Suppose the second dominant eigen value of A is λ_2 and its corresponding eigen-vector is x . Then:

$$\alpha Px + (1 - \alpha)Ex = \lambda_2 x \quad (9)$$

We will come back to equation(9) after proving the following lemma.

lemma-1: If x_i is the eigen vector of A with eigen value λ_i and y_j is the eigen vector of A^T with eigen value λ_j , then if $\lambda_i \neq \lambda_j$, then $x_i^T y_j = 0$

proof: We can write the following :

$$y_j^T A = \lambda_j y_j^T$$

$$Ax_i = \lambda_i x_i$$

Transposing and then Multiplying the first by y_j and the second by x_i^T , we get

$$x_i^T A^T y_j = \lambda_i x_i^T y_j$$

$$x_i^T A^T y_j = \lambda_j x_i^T y_j$$

Subtracting the above leads to the fact that is $\lambda_i \neq \lambda_j$, then:

$$x_i^T y_j = 0$$

(Proved)

Coming back to our original equation(9), notice that E 's columns are vectors $e^T = \frac{1}{n}[1, 1, \dots, 1]$.

This all 1 vector is an eigen-vector of A^T and corresponds to eigen-value of 1. Thus, applying the lemma we just proved, we then get $Ex_2 = 0$. Now the equation(9) becomes,

$$\alpha Px_2 = \lambda_2 x_2$$

$$Px_2 = \frac{\lambda_2}{\alpha} x_2$$

Since, P is also stochastic and the x_2 is also an eigen vector of P and 1 is the dominant eigen vector, hence:

$$\frac{\lambda_2}{\alpha} \leq 1$$

$$\lambda_2 \leq \alpha$$

Thus, the second largest wigenvalue of $A(\alpha)$ is upper bounded by α .