

Quantifying the Effects of Climate Policy Stringency on Verified Emissions and Satellite-Derived NO_x

Master's Thesis

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Abstract

This study develops a novel framework for evaluating climate policy impacts using two complementary emission outcomes. I investigate how European Union Emissions Trading System (EU ETS) policy stringency affects both installation-level verified CO₂ emissions and satellite-derived NOx emission proxies around major industrial emitters. The dual-outcome approach enables cross-validation: EU ETS verified emissions provide high-quality, installation-level measures of greenhouse gas output, while satellite-derived NOx estimates offer physically grounded proxies for combustion co-pollutants that can reveal co-benefits and potential under-reporting.

The methodological framework makes three primary contributions. First, I demonstrate the integration of geospatial foundation model embeddings (Google AlphaEarth, 64 dimensions) as controls in panel-based climate monitoring studies, capturing between-unit heterogeneity arising from local geographic and climate context in a data-efficient manner that would be impractical to specify manually. Second, I introduce network-based clustering for inference: standard errors and region-by-time fixed effects use PyPSA-Eur power system clusters—k-means clusters computed on transmission network topology features from an external power system model—that group facilities facing correlated wholesale prices, dispatch patterns, and grid constraints. Third, for the satellite outcome, I implement a simplified Beirle-style flux-divergence method: using TROPOMI NO₂ tropospheric column densities and ERA5 winds, I compute the advection (wind-aligned spatial derivative) of NO₂ fields around each facility, integrate

within a 15 km radius, and apply lifetime and NO₂-to-NOx corrections following Beirle et al. (2023). This approach is physically grounded in the continuity equation and specifically designed for power-plant-scale NOx plumes.

The analysis panel links 251 EU ETS-regulated large combustion plants across Europe (2018–2023) with both regulatory emissions data and satellite observations, employing two-way fixed effects and Callaway-Sant’Anna difference-in-differences estimators. This work demonstrates that combining administrative emissions records with satellite-based NOx estimates enables comprehensive evaluation of climate policy effects on both greenhouse gases and co-pollutants.

Keywords: Climate Policy, EU ETS, Verified Emissions, Satellite Remote Sensing, TROPOMI, NOx Emissions, Flux Divergence, Difference-in-Differences, Causal Inference, Large Combustion Plants

1 Introduction

Evaluating climate policy requires measuring actual emission outcomes. The European Union Emissions Trading System (EU ETS) generates rich administrative data on verified CO₂ emissions at the installation level, providing the gold standard for measuring greenhouse gas output from regulated facilities. However, relying solely on self-reported emissions raises questions about verification and leaves unmeasured the local air quality co-benefits that accompany carbon reductions. Satellite remote sensing offers an independent, physically-grounded approach to quantifying emissions from space, potentially revealing both verification opportunities and co-pollutant dynamics that administrative data cannot capture.

This study adopts a dual-outcome approach that exploits the complementary strengths of administrative and satellite data. The two outcomes are: (i) **verified EU ETS CO₂ emissions**—high-quality, installation-level measures from the EU ETS registry that provide accurate compliance trajectories and absolute emission levels; and (ii) a **satellite-derived NOx emission proxy**—a physically grounded indicator constructed from TROPOMI NO₂ tropospheric columns and ERA5 winds, following the flux-divergence approach of Beirle et al. [1–3].

Why use both outcomes? CO₂ is a well-mixed greenhouse gas with global climate impacts; nitrogen oxides (NO_x), by contrast, are criteria pollutants whose health effects—respiratory illness, cardiovascular disease, premature mortality—fall disproportionately on populations living near emission sources. As [4] emphasize, air quality co-benefits are particularly policy-relevant because they are local and immediate, whereas averted climate damages are global and long-term. The dual-outcome design provides: (i) verified emissions for accurate policy effect estimation, (ii) satellite-derived NOx for testing co-benefit hypotheses, and (iii) cross-validation opportunities where both outcomes should respond to common policy shocks.

This study develops a novel framework for evaluating climate policy impacts using both administrative emissions data and satellite remote sensing. I focus on the European Union Emissions Trading System (EU ETS), the world’s largest carbon

market, which creates economic incentives for industrial facilities to reduce CO₂ emissions through a cap-and-trade mechanism. The framework addresses two fundamental methodological challenges: (i) constructing a satellite-derived NOx emission proxy that is physically interpretable and appropriate for panel econometric analysis, and (ii) controlling for high-dimensional confounders that affect both policy exposure and emission outcomes.

The study makes two primary methodological contributions, both following a recent trend in causal inference toward incorporating machine learning-derived features to address high-dimensional confounding [5–7].

First, I demonstrate the use of geospatial foundation model embeddings as high-dimensional controls in panel-based climate monitoring studies. Specifically, I incorporate Google AlphaEarth embeddings [8]—64-dimensional representations learned from multi-source satellite imagery, climate reanalysis, and geotagged text—as control variables in the econometric specifications. These embeddings capture between-unit heterogeneity arising from local geographic, infrastructural, and climate context in a data-efficient manner, providing a scalable approach to controlling for high-dimensional spatial confounders that would be impractical to specify manually. This application extends prior work on learned representations for causal inference—originally developed for text embeddings [6]—to the domain of geospatial environmental monitoring. The approach is particularly suited to difference-in-differences settings where high-dimensional confounders may violate the parallel trends assumption if left uncontrolled [7].

Second, I introduce network-based clustering derived from an external power system model for both fixed effects structure and inference. Standard errors are clustered using PyPSA-Eur power system clusters [9]—k-means clusters computed on transmission network topology features extracted from the European high-voltage grid. These clusters group electrical buses based on network connectivity and transmission impedance, producing units where facilities face correlated wholesale prices, dispatch patterns, and grid constraints. This represents a novel application of model-derived features for defining cluster structure in econometric inference, building on recent theoretical foundations establishing that valid cluster-robust inference requires clusters with low “conductance” (few cross-cluster connections relative to within-cluster volume) [10]. Power system network clusters satisfy this requirement by construction, as they minimize within-cluster electrical distance. The same clusters define Region×Year fixed effects, absorbing time-varying regional confounders that correlate with both policy exposure and air quality outcomes.

Third, for the satellite outcome, I implement a simplified Beirle-style flux-divergence method that provides physically grounded NOx emission estimates at the facility level. The approach computes the advection—the scalar product of wind velocity and the spatial gradient of NO₂ column density—which under the continuity equation is proportional to local emissions minus chemical loss. For each facility, I integrate advection over a 15 km disc, apply a lifetime correction following [3], and convert to NOx emission rates. This methodology follows the Beirle et al. (2019, 2021, 2023) family of methods [1–3], which are physically transparent, computationally tractable for known point sources, and specifically designed for power-plant-scale NOx plumes.

The analysis links three data sources on large combustion plants (LCPs) in the EU: (i) the European Environment Agency’s LCP registry providing plant characteristics and coordinates, (ii) EU ETS compliance data providing installation-level verified emissions and free allocations, and (iii) TROPOMI satellite observations processed through the Beirle-style flux-divergence methodology to derive NOx emission proxies. Policy exposure is measured continuously through the *allocation ratio*—free allowances divided by verified emissions—where values below unity indicate facilities must purchase additional permits, creating direct economic pressure to reduce emissions.

The econometric framework employs two-way fixed effects (TWFE) specifications with facility and time fixed effects, as well as the Callaway and Sant’Anna [11] difference-in-differences estimator for staggered treatment timing.

By demonstrating that both administrative emissions records and satellite-derived NOx estimates can provide individual-emitter-level, policy-parameterized estimates of emission responses to carbon market stringency, this work contributes to the emerging literature on comprehensive climate policy evaluation. The dual-outcome approach enables testing whether policy effects on verified CO₂ are accompanied by corresponding changes in satellite-observed combustion co-pollutants.

2 Background and Literature Review

2.1 The EU Emissions Trading System

The EU ETS, established in 2005, operates as a cap-and-trade system covering approximately 40% of EU greenhouse gas emissions. Large combustion plants with thermal input exceeding 20 MW are required to hold European Union Allowances (EUAs) equal to their verified CO₂ emissions. Allowances are distributed through a combination of free allocation (based on historical benchmarks and carbon leakage risk) and auctioning. Installations that emit more than their free allocation must purchase additional allowances, creating marginal abatement incentives [12].

The policy has evolved through four phases, with Phase III (2013–2020) and Phase IV (2021–2030) introducing progressively tighter caps and reduced free allocation, particularly for the power sector. This study focuses on the period 2018–2023, spanning the transition from Phase III to Phase IV and capturing significant variation in policy stringency across facilities.

2.2 Satellite-Based Air Quality Monitoring

The TROPOMI instrument aboard Sentinel-5P, operational since late 2017, provides daily global observations of tropospheric NO₂ column densities at unprecedented spatial resolution ($\sim 3.5 \times 5.5 \text{ km}^2$ at nadir). This represents a significant improvement over predecessor instruments (OMI, GOME-2) and enables detection and quantification of emissions from individual point sources [3, 13].

Previous studies have used satellite observations to verify emission reductions from policy interventions. [14] demonstrated that China’s ultra-low-emission retrofits for

coal-fired power plants produced measurable NO₂ declines visible from space. [15] documented substantial NO_x reductions over Europe between 1996 and 2010, attributing these to environmental policies and economic recession. However, these studies typically analyze aggregate regional trends rather than plant-level responses to specific policy parameters.

2.3 Satellite-Based NOx Emission Quantification: The Flux-Divergence Approach

A key methodological challenge in quantifying emissions from satellite-observed NO₂ is separating the source signal from background concentrations and converting column densities to emission rates. This challenge is particularly acute in Europe, where high population density means that most large combustion plants are located in or near urban areas, surrounded by other pollution sources (traffic, industry, heating).

The flux-divergence (or advection) approach, developed by Beirle et al. [1–3], provides a physically grounded solution. The method exploits the continuity equation: horizontal NO₂ fluxes $\mathbf{F} = \mathbf{w}V$ (where \mathbf{w} is wind velocity and V is tropospheric vertical column density) satisfy

$$\nabla \cdot \mathbf{F} = E - S \quad (1)$$

where E represents local emissions and S represents chemical sinks. Under typical conditions where wind field divergence is negligible, this reduces to the advection formulation:

$$A = \mathbf{w} \cdot \nabla V \approx E - S \quad (2)$$

The advection A measures the downwind rate of change in NO₂ column density and is particularly sensitive to strong point sources, which create sharp spatial gradients in the NO₂ field.

Beirle et al. (2021) [2] presented the first global catalog of NOx point source emissions derived from TROPOMI using this approach, identifying 451 sources. Beirle et al. (2023) [3] introduced version 2 with several improvements: use of the PAL (Products Algorithm Laboratory) NO₂ product with higher column densities (factor of 1.1–1.4), corrections for plume height effects on satellite sensitivity, topographic corrections, and a lifetime correction to account for chemical loss within the integration radius. These refinements resulted in emission estimates approximately 3 times higher than version 1, with validation showing agreement within 20% of reported emissions from the German Environment Agency (UBA) and US EPA.

[16] developed an alternative regression-based approach for decomposing TROPOMI NO₂ into urban, industrial, and background components during COVID-19, demonstrating that wind information can isolate individual source contributions even in complex emission environments. [13] extended the methodology to megacities, estimating both emissions and effective NOx lifetimes through simultaneous fitting of downwind plume evolution.

This study adopts the Beirle family of methods because they are: (i) physically transparent, grounded in the continuity equation; (ii) computationally tractable for

known point sources; and (iii) specifically designed and validated for power-plant-scale NOx plumes. I implement a simplified version appropriate for panel econometric analysis, acknowledging the additional uncertainty from using OFFL L3 data rather than the PAL product.

2.4 Causal Inference with Staggered Treatment Timing

Standard two-way fixed effects estimators can produce biased estimates when treatment timing varies across units and treatment effects are heterogeneous [17]. Recent methodological advances, including the Callaway and Sant’Anna [11] and Sun and Abraham [18] estimators, address these concerns by constructing treatment effect estimates using only valid comparisons (treated versus not-yet-treated or never-treated units) and allowing for treatment effect heterogeneity across cohorts and time.

This study implements both traditional TWFE specifications (which remain valid under homogeneous treatment effects) and the Callaway-Sant’Anna estimator (which is robust to heterogeneity), allowing comparison of results under different identifying assumptions.

2.5 High-Dimensional Controls and ML-Derived Features in Causal Inference

A growing literature in causal inference addresses the challenge of controlling for high-dimensional confounders—settings where the number of potential control variables is large relative to sample size, or where relevant confounders are difficult to specify manually. The foundational work of [5] established the “double/debiased machine learning” framework, showing how machine learning methods can be used to estimate nuisance parameters (propensity scores, outcome regressions) while maintaining valid inference on treatment effects. This approach enables researchers to control for high-dimensional confounders without imposing restrictive parametric assumptions.

In the difference-in-differences context specifically, [7] developed efficient estimators for settings where the parallel trends assumption holds only conditional on high-dimensional covariates. This is particularly relevant when unobserved confounders that violate parallel trends can be proxied by high-dimensional observables—such as detailed geographic or economic characteristics that would be impractical to specify manually but can be captured through flexible ML methods.

A parallel development concerns the use of *learned representations*—embeddings from neural networks or foundation models—as control variables. [6] demonstrated that text embeddings can serve as effective controls for confounding in observational studies, provided the embeddings capture the relevant confounding information. The key insight is that pre-trained representations, learned for prediction tasks on large corpora, may encode information about latent confounders that would otherwise be unobserved. This approach has been extended to various domains, including image embeddings and, most recently, geospatial foundation models.

For clustered inference, [10] established theoretical foundations for network cluster-robust standard errors. They show that valid cluster-robust inference requires clusters with low “conductance”—the ratio of edges crossing cluster boundaries to total edges

within clusters. This implies that clusters should be defined based on the correlation structure of the data-generating process, not arbitrary geographic or administrative boundaries. When observations are connected through a network (as power plants are through the transmission grid), clusters derived from network topology can satisfy these requirements.

This study contributes to this literature by demonstrating two novel applications: (i) using geospatial foundation model embeddings (AlphaEarth) as controls for spatial confounding in environmental panel data, and (ii) using k-means clusters derived from power system network features (PyPSA-Eur) for both fixed effects structure and clustered inference. To my knowledge, this represents the first application of model-derived clustering—where clusters are computed on features from an external domain-specific model rather than on the outcome data itself—for econometric inference in policy evaluation.

3 Data

This section describes the data sources, processing pipeline, and construction of the analysis panel. The study combines administrative records on industrial facilities and EU ETS compliance with satellite remote sensing and meteorological reanalysis data.

3.1 Data Sources

3.1.1 EEA Large Combustion Plant Registry

The European Environment Agency (EEA) maintains the Industrial Emissions Portal, which includes the Large Combustion Plant (LCP) dataset. This registry provides annual reports on combustion plants with rated thermal input ≥ 50 MW, including:

- Geographic coordinates (latitude, longitude)
- Plant identification (LCP INSPIRE ID, installation name)
- Rated thermal capacity (MW)
- Annual fuel consumption by fuel type (TJ)
- Country of operation

The raw dataset contains 3,405 unique plant parts for the period 2018–2023. After filtering for complete capacity and fuel data, 2,821 plants remain with valid time-varying attributes.

3.1.2 EU ETS Compliance Data

EU ETS installation-level compliance data is obtained from the European Union Transaction Log (EUTL), accessed via the `pyeutl` Python package. For each installation-year, the data includes:

- Verified CO₂ emissions (tCO₂)
- Free allocation of allowances (tCO₂-equivalent)
- Surrendered allowances (tCO₂-equivalent)
- Installation identifier and country

The LCP and ETS datasets are linked through the EU Registry on Industrial Sites, which provides crosswalk tables mapping LCP installation parts to their parent ETS installations via normalized identifiers.

3.1.3 TROPOMI Satellite Observations

Tropospheric NO₂ column densities are obtained from the Sentinel-5P TROPOMI instrument via Google Earth Engine, using the OFFL (offline) L3 product (`COPERNICUS/S5P_OFFL_L3_N02`). TROPOMI provides daily global coverage at approximately 3.5×5.5 km² spatial resolution. Quality-filtered observations are used, retaining only pixels with quality assurance values ≥ 0.75 . TROPOMI captures approximately 14 orbits per day globally, with each orbit covering a distinct swath (~ 2600 km); for any given facility, only one orbit per day provides valid coverage.

Importantly, Beirle et al. (2023) [3] use the PAL (Products Algorithm Laboratory) NO₂ product, which provides higher tropospheric vertical column densities (TVCDs) than the OFFL product by a factor of approximately 1.1–1.4, due to updated retrieval algorithms and air mass factor corrections. This difference, combined with other methodological refinements, contributed to their version 2 emission estimates being approximately 3 times higher than version 1. Since I use the OFFL L3 product available via Google Earth Engine rather than the PAL product, the satellite-derived NO_x estimates carry additional uncertainty (approximately $\pm 25\%$ relative to PAL-based estimates) that must be acknowledged in interpretation.

3.1.4 ERA5-Land Reanalysis

Hourly 10-meter wind components (u_{10}, v_{10}) are obtained from the ERA5-Land reanalysis product via Google Earth Engine. Daily mean wind speed and direction are computed at each facility location for the advection calculation. Following Beirle et al. [3], days with wind speeds below 2 m/s are excluded, as weak winds produce unreliable advection estimates and allow plumes to stagnate near sources. Additionally, observations where the lifetime correction factor $c_\tau \geq 3$ are dropped, as this exceeds the typical range of 1.2–1.8 reported by Beirle et al.

3.2 Facility Construction: Spatial Clustering

Individual LCP plant parts may represent components of larger industrial complexes. To avoid treating co-located plants as independent units, I apply spatial clustering using a 500-meter threshold. Plants within 500m of each other are grouped into a single *facility* using a union-find algorithm.

Let $\mathcal{P} = \{p_1, \dots, p_N\}$ denote the set of LCP plants with coordinates (ϕ_j, λ_j) for plant j . The distance between plants j and k is computed using the WGS84 ellipsoidal approximation [19]:

$$d_{jk} \approx \sqrt{(m_\phi \cdot \Delta\phi_{jk})^2 + (m_\lambda \cdot \Delta\lambda_{jk})^2} \quad (3)$$

where the latitude scale factor follows the WGS84 series expansion:

$$m_\phi = 111,132.954 - 559.822 \cos(2\bar{\phi}) + 1.175 \cos(4\bar{\phi}) \quad [\text{m}/\text{deg}] \quad (4)$$

and the longitude scale factor varies with latitude:

$$m_\lambda = 111,132.954 \times \cos(\bar{\phi}) \quad [\text{m}/\text{deg}] \quad (5)$$

where $\bar{\phi}$ is the mean latitude of the dataset. The latitude formula is accurate to 0.01 m per degree; the longitude formula has <0.5% error compared to the full WGS84 ellipsoidal expression. This precision is more than sufficient for identifying co-located plants, as the 500m clustering threshold is conservative relative to the spatial extent of most industrial complexes.

Plants are grouped into facility i if they form a connected component under the relation $d_{jk} < 500\text{m}$. For each facility, the centroid coordinates are computed as the arithmetic mean of constituent plant coordinates:

$$(\bar{\phi}_i, \bar{\lambda}_i) = \frac{1}{|F_i|} \sum_{j \in F_i} (\phi_j, \lambda_j) \quad (6)$$

where F_i denotes the set of plants in facility i .

This clustering reduces the sample from 1,576 individual plants to 932 facilities, of which 318 are multi-plant facilities.

3.3 Time-Varying Attributes

3.3.1 Capacity and Fuel Shares

For each facility-year (i, t) , rated thermal capacity is aggregated as the sum across constituent plants:

$$\text{Capacity}_{it} = \sum_{j \in F_i} \text{Capacity}_{jt} \quad [\text{MW}] \quad (7)$$

Fuel energy consumption is similarly aggregated, then converted to fuel shares. Let $E_{it}^{(f)}$ denote total energy consumption from fuel type $f \in \{\text{gas, coal, oil, biomass, other}\}$ for facility i in year t , measured in terajoules (TJ). Fuel shares are computed as:

$$s_{it}^{(f)} = \frac{E_{it}^{(f)}}{\sum_{f'} E_{it}^{(f')}} \quad (8)$$

Fuel types used by fewer than 10% of facility-years (peat, other solid fuels) are dropped, remaining shares are renormalized to sum to unity, and facilities with no remaining fuel coverage are excluded.

3.3.2 ETS Policy Exposure Variables

The key treatment variable is the *allocation ratio*, defined as:

$$R_{it} = \frac{A_{it}}{V_{it}} \quad (9)$$

where A_{it} is total free allocation and V_{it} is verified emissions for facility i in year t , both in tCO₂. Values $R_{it} < 1$ indicate the facility must purchase additional allowances on the carbon market, representing increased policy stringency.

The *shortfall* is defined as:

$$S_{it} = V_{it} - A_{it} \quad (10)$$

Positive shortfall indicates emissions exceed free allocation.

Facilities with allocation ratios outside the range [0.01, 20] are excluded as likely data errors or non-operating installations. Additionally, facilities are required to have at least one year with verified emissions ≥ 100 ktCO₂ to ensure sufficient emissions magnitude for detectable satellite signals, following the detection limits established by [13].

3.4 Satellite NOx Emission Proxy: Beirle-Style Flux-Divergence

The satellite outcome variable is constructed using a simplified Beirle-style flux-divergence method, following the approach developed by Beirle et al. [1–3]. This method provides physically grounded NOx emission estimates by exploiting the relationship between wind-driven advection and local emissions.

3.4.1 Identification versus Quantification

Beirle et al.'s v2 catalog combines two distinct algorithmic components: (i) an automatic point-source *identification* algorithm that locates emission maxima in the global advection field, and (ii) a *quantification* method that estimates emission rates by spatially integrating advection around each identified source. Crucially, the authors note that “the quantification of NOx emissions by spatial integration of the corrected advection map could be applied to these locations or **any other known point source**” [3].

In this study, I skip the *identification step* because I already have a curated set of ETS/LCP facilities with reliable coordinates from the European Environment Agency registry. I apply Beirle's quantification method directly to these known source locations. This design choice is explicitly endorsed by the authors' statement and is conceptually appropriate: the identification algorithm is needed only when constructing a global catalog without prior knowledge of emission sources, not when applying the physically grounded quantification to facilities whose locations are already known.

To guard against treating noise as signal, I implement *simplified significance flags* that parallel Beirle's catalog selection criteria:

- **Detection limit:** Emission estimates below 0.11 kg/s are flagged, corresponding to Beirle's standard detection threshold for non-desert conditions.

- **Statistical integration error:** Facilities with >30% relative statistical uncertainty in the spatial integration are flagged.
- **Spatial interference:** Facilities with another ETS facility within 20 km are flagged, as their satellite outcome may reflect cluster-level rather than single-facility emissions.

These flags are used in sensitivity analyses rather than for hard filtering, preserving the full panel while allowing transparent restriction to “significant” satellite observations.

3.4.2 Advection Formulation

The advection A is defined as the scalar product of wind velocity and the spatial gradient of NO₂ tropospheric vertical column density (TVCD):

$$A = \mathbf{w} \cdot \nabla V = u \frac{\partial V}{\partial x} + v \frac{\partial V}{\partial y} \quad (11)$$

where $\mathbf{w} = (u, v)$ is the horizontal wind vector (m/s) from ERA5-Land and V is the NO₂ TVCD (molecules/m²). Under the continuity equation, this advection is proportional to local emissions minus chemical sinks.

For each facility i and day d , spatial gradients are computed on a local grid (30 km × 30 km centered on the facility) using finite differences on the TROPOMI L3 lat-lon grid:

$$\frac{\partial V}{\partial x} \approx \frac{V(x + \Delta x, y) - V(x - \Delta x, y)}{2\Delta x} \quad (12)$$

$$\frac{\partial V}{\partial y} \approx \frac{V(x, y + \Delta y) - V(x, y - \Delta y)}{2\Delta y} \quad (13)$$

where Δx and Δy correspond to the TROPOMI grid resolution (approximately 3.5 km × 5.5 km). This differs from Beirle et al., who compute derivatives on the native TROPOMI pixel grid to handle cloud-induced gaps; the L3 gridded product used here introduces additional smoothing and potential artifacts.

3.4.3 NO₂ to NOx Scaling

TROPOMI measures NO₂, but NOx emissions include both NO and NO₂. Following Beirle et al. [3], I apply a scaling factor c_{NOx} based on the photostationary state (PSS):

$$c_{\text{NOx}} = \frac{[\text{NOx}]}{[\text{NO}_2]} = 1 + \frac{J}{k[\text{O}_3]} \quad (14)$$

where J is the NO₂ photolysis frequency (parameterized as $0.0167 \times \exp(-0.575/\cos(\text{SZA})) \text{ s}^{-1}$), k is the reaction rate constant for NO + O₃ ($2.07 \times 10^{-12} \times \exp(-1400/T) \text{ cm}^3 \text{ molec}^{-1} \text{ s}^{-1}$), and [O₃] is taken from an ozone climatology. For detected point sources, Beirle et al. report a typical NOx/NO₂ ratio of approximately 1.38 ± 0.10 .

Following Beirle et al., I apply a fixed scaling factor of $c_{\text{NOx}} = 1.38$ with uncertainty ± 0.10 (approximately 7% relative uncertainty), which represents the empirically observed mean ratio across detected point sources.

3.4.4 Topographic Correction

Over mountainous terrain, 3D radiative transfer effects cause systematic artifacts in the advection field [?]. Following Beirle et al. [3] Sect. 3.7, I apply a topographic correction:

$$A^* = A + f \cdot C_{\text{topo}}, \quad C_{\text{topo}} = \frac{V}{H_{\text{sh}}} \cdot (\mathbf{w}_0 \cdot \nabla z_0) \quad (15)$$

where V is the NO_2 TVCD, $H_{\text{sh}} = 1$ km is the assumed NOx scale height, $\mathbf{w}_0 \cdot \nabla z_0$ is the dot product of the surface wind vector and the surface elevation gradient (from SRTM DEM), and $f = 1.5$ is an empirically derived scaling factor (Appendix A of Beirle et al.). The combined effect yields an effective scale height of $1/1.5 = 667$ m. For flat terrain typical of European power plant locations, this correction is small.

3.4.5 Spatial Integration and Lifetime Correction

For each facility, the raw emission rate is computed by spatially integrating the topography-corrected advection A^* over a 15 km disc around the facility location (Beirle Eq. 11):

$$E_{\text{raw}} = \iint_{r \leq 15 \text{ km}} A^*(x, y) dx dy \approx \sum_i A_i^* \times \Delta x \Delta y \quad [\text{mol/s}] \quad (16)$$

where A^* has units $\text{mol}/(\text{m}^2 \cdot \text{s})$ and the spatial integration is realized by summing the advection values multiplied by the pixel area for all grid pixels within the 15 km radius. This radius is chosen following Beirle et al. [3] as a compromise between capturing the full point source signal and avoiding interference from neighboring sources.

Chemical loss of NOx during transport within the integration radius requires a lifetime correction. The residence time within the 15 km radius is:

$$t_r = \frac{R}{|\mathbf{w}|} \quad (17)$$

where $R = 15$ km and $|\mathbf{w}|$ is the mean wind speed. The lifetime correction factor, following Beirle et al. [3] Eq. (9), is:

$$c_\tau = \exp(t_r/\tau) \quad (18)$$

where τ is the effective NOx lifetime, parameterized as a function of latitude following Lange et al. [20] via Beirle et al. Eq. (10):

$$\tau(\text{lat}) = 1.0089 \times \exp(0.0242 \times (|\text{lat}| + 9.6024)) \quad [\text{hours}] \quad (19)$$

with typical values of 2 h at low latitudes to 4–6 h at higher latitudes. For detected point sources, the resulting $c_\tau \approx 1.40 \pm 0.24$. Following Beirle et al., I assume 50% relative uncertainty in τ due to high variability at similar latitudes.

3.4.6 Final NO_x Emission Estimate

The final satellite-derived NO_x emission rate for facility i and day d is:

$$E_{\text{NO}_x,id} = c_\tau \cdot c_{\text{NO}_x} \cdot E_{\text{raw},id} \quad (20)$$

Converting from mol/s to kg/s using the molar mass of NO₂ (46.0055 g/mol) [21]. Annual estimates are computed as the mean over all valid observation days.

3.4.7 Uncertainty Components

Following Beirle et al. [3] Sect. 3.12, the satellite-derived NO_x estimates carry uncertainty from multiple sources, combined in quadrature:

- **Statistical error** (Sect. 3.12.2): I approximate using the standard error of the temporal mean of daily integrated emissions, rather than Beirle’s per-pixel SE propagation. This is more conservative as it captures meteorological variability in addition to sampling noise, typically <10%. Facilities with statistical relative error $\geq 30\%$ are flagged via `rel_err_stat_lt_0.3`.
- **Lifetime correction** (Sect. 3.12.1): 50% relative uncertainty in τ , propagated through $c_\tau = \exp(t_r/\tau)$ yielding $\sigma_{c_\tau}/c_\tau = \ln(c_\tau) \times 0.50$, typically 10–20% for $c_\tau \approx 1.4$.
- **NO_x/NO₂ scaling** (Sect. 3.12.1): ± 0.10 on 1.38 ratio, $\sim 7\%$.
- **AMF correction** (Sect. 3.12.1): *Unmodeled structural uncertainty*—I do not implement an explicit AMF correction. A 10% term is carried as a generic structural uncertainty following Beirle’s error budget, representing potential bias rather than fitted variance.
- **Plume height** (Sect. 3.12.3): *Unmodeled structural uncertainty*—I do not implement plume-height-dependent wind interpolation. A 10% term represents the sensitivity to assumed height (500m vs 300m) as reported by Beirle.
- **Topographic correction** (Sect. 3.12.4): 33% uncertainty on $f = 1.5$, typically <2.5% for flat European terrain.
- **OFFL vs PAL product** (our addition): OFFL provides 10–40% lower TVCDs than PAL; a 25% structural uncertainty term is added to account for this systematic difference.

Beirle et al. report total uncertainties in the 20–40% range. With the OFFL product uncertainty and unmodeled structural terms, our typical total is $\sim 35\text{--}45\%$.

Detection limits and significance flags. Rather than hard-filtering the sample, I implement boolean significance flags:

- `above_d1_0_11`: Emission estimate ≥ 0.11 kg/s (Beirle’s standard detection limit for non-desert conditions, appropriate for Europe).

- `above_d1_0_03`: Emission estimate $\geq 0.03 \text{ kg/s}$ (Beirle's permissive threshold, valid only under ideal high-albedo desert conditions—not applicable to Europe).
- `rel_err_stat_lt_0_3`: Statistical integration error $< 30\%$.
- `interfered_20km`: Another ETS facility exists within 20 km.

Main satellite regressions restrict to “significant” observations satisfying `above_d1_0_11` \wedge `rel_err_stat_lt_0_3`. Sensitivity analyses additionally exclude interfered facilities or relax to the permissive detection limit.

3.5 Sample Construction

The final analysis sample is constructed by applying the following filters:

1. Facilities must have valid ETS linkage (matched normalized identifier)
2. At least one year with verified emissions $\geq 100 \text{ ktCO}_2$
3. Allocation ratio in [0.01, 20] range
4. At least 3 years of complete data within 2018–2023
5. Non-missing satellite outcome variable

Table 1 summarizes the sample attrition through each processing step. The most significant losses occur at the ETS linkage stage (44% of plants lack matched ETS identifiers in the EU Registry crosswalk) and the emissions threshold filter (59% of matched facilities have verified emissions below 100 ktCO₂/yr in all years, making them unsuitable for satellite detection). The resulting panel contains 251 facilities observed over 1,380 facility-years.

Table 1 Sample Attrition Through Data Processing Pipeline

Processing Step	Plants/Facilities	Lost	Retained %
Initial LCP registry ($\geq 50 \text{ MW thermal}$)	3,405 plants	—	100%
With complete capacity + fuel data (2018–2023)	2,821	584	82.8%
With ETS linkage (via EU Registry crosswalk)	1,580	1,241	46.4%
<i>After 500m spatial clustering</i>	<i>932 facilities</i>	—	—
With matched ETS compliance data	799	133	85.7%
With $\geq 100 \text{ ktCO}_2/\text{yr}$ verified emissions	251	548	31.4%
Final analysis panel	251 facilities		7.4% of initial
			1,380 fac-years

Note: LCP registry includes only plants with rated thermal input $\geq 50 \text{ MW}$. The 100 ktCO₂/yr threshold ensures sufficient emissions magnitude for satellite-based NO_x detection following the detection limits in Beirle et al. (2023). ETS linkage uses normalized identifier matching between EU Registry and EUTL compliance data.

Table 2 summarizes the sample characteristics.

3.6 Geographic Context: AlphaEarth Embeddings

A key methodological contribution of this study is the incorporation of high-dimensional geospatial foundation model embeddings as control variables, following the recent trend toward using learned representations for causal inference [5, 6]. I use Google AlphaEarth Foundations [8], a geospatial embedding field model that produces 64-dimensional representations from multi-source satellite imagery (Sentinel-1/2, Landsat), climate reanalysis (ERA5-Land), topography (GLO-30), and geotagged text (Wikipedia, GBIF). The model is trained using contrastive learning objectives that encourage embeddings to capture information predictive of diverse downstream tasks—from land cover classification to biophysical variable estimation—without being tuned for any specific application.

For each facility location, the embedding vector $\mathbf{e}_i \in \mathbb{R}^{64}$ is extracted from the nearest grid cell at 10-meter spatial resolution. These embeddings encode:

- **Land use context:** Urban density, industrial areas, agricultural patterns
- **Infrastructure:** Road networks, built environment characteristics
- **Vegetation:** Forest cover, cropland, seasonal phenology
- **Climate:** Local temperature, precipitation, insolation, and wind patterns
- **Topography:** Elevation, slope, and terrain characteristics

The embedding dimensions are included as controls in the econometric specifications, providing a data-efficient approach to capturing between-unit heterogeneity arising from local geographic context. This application extends prior work on text embeddings for causal inference [6] to the geospatial domain. The approach is particularly relevant for difference-in-differences settings where high-dimensional spatial confounders may induce violations of parallel trends if left uncontrolled [7]—for example, if facilities in different geographic contexts (coastal versus inland, urban versus rural) experience different secular trends in air quality unrelated to policy.

Since the embeddings are derived from satellite imagery aggregated over time, they are treated as static facility-level controls. The 64 dimensions are included directly without dimensionality reduction, as the panel fixed effects structure provides regularization against overfitting. This represents 64 bytes per location—a highly compressed representation of the local geographic and climate context that would require hundreds of manually-specified variables to approximate.

3.7 Exploratory Data Analysis

Figure 1 displays the geographic distribution of facilities in the analysis sample.

Figure 2 shows the distribution of the allocation ratio (treatment variable) across facility-years.

Figure 3 illustrates the satellite-derived NOx emission proxy.

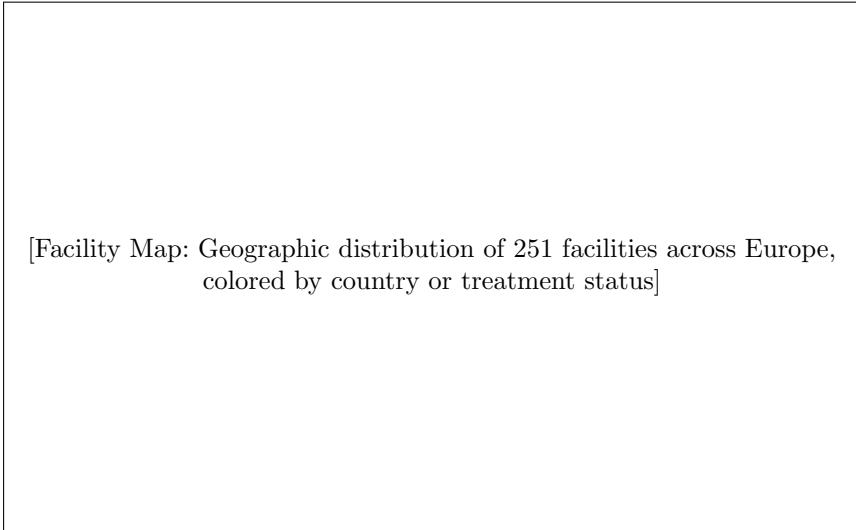


Fig. 1 Geographic distribution of analysis sample facilities across Europe. Points represent facility centroids after 500m spatial clustering.

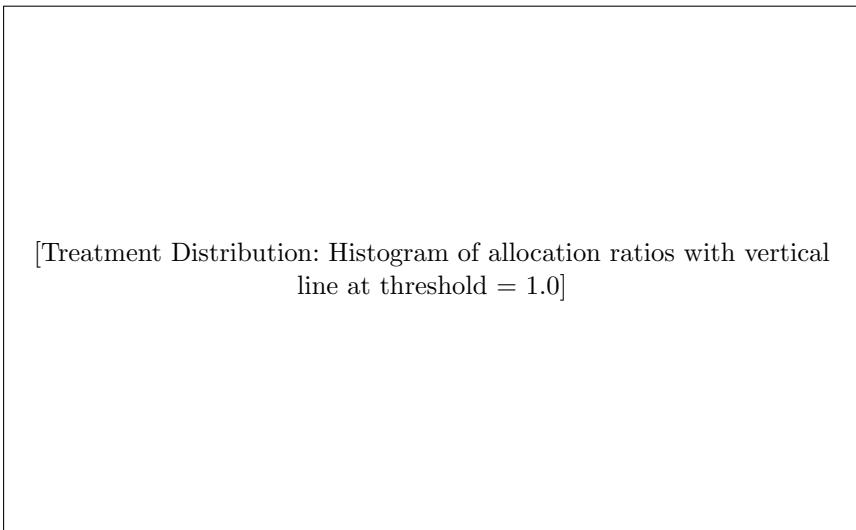


Fig. 2 Distribution of allocation ratios across facility-years. Values below 1.0 (dashed line) indicate facilities facing policy stringency (emissions exceed free allocation).

4 Econometric Methods

This section describes the causal inference framework and econometric specifications used to estimate the effect of EU ETS policy stringency on both verified CO₂ emissions and satellite-derived NOx emission proxies.

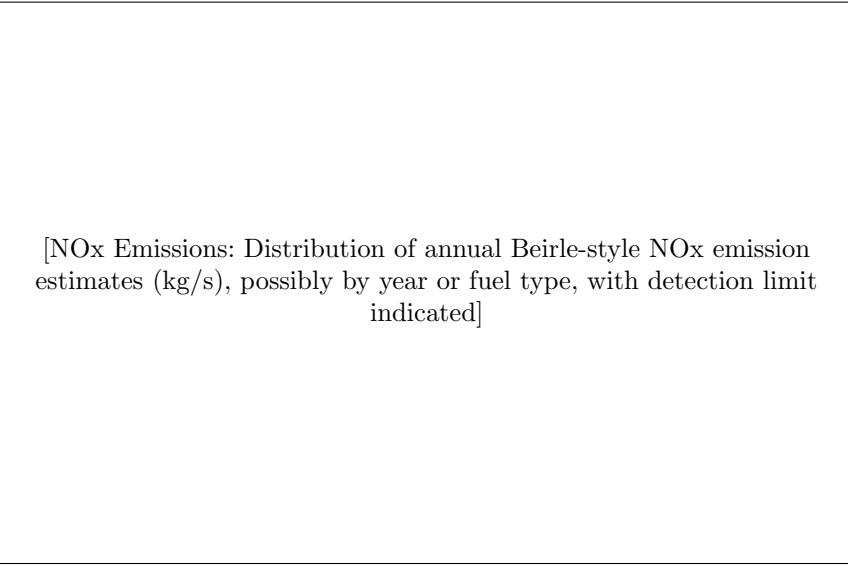


Fig. 3 Distribution of satellite-derived NOx emission rates ($E_{\text{NOx},it}$) across facility-years. Dashed line indicates the generic detection limit of approximately 0.11 kg/s from Beirle et al. (2023).

4.1 Causal Framework

The goal is to estimate the causal effect of ETS policy stringency on two complementary outcomes. Let $Y_{it}^{\text{CO}_2}$ denote verified CO₂ emissions (ktCO₂/yr) for facility i in year t , and let Y_{it}^{NOx} denote the satellite-derived NOx emission proxy (kg/s). Let R_{it} denote the allocation ratio (treatment intensity).

The key identification challenge is that allocation ratios are not randomly assigned. Facilities with high emissions relative to historical benchmarks receive lower allocation ratios, creating potential endogeneity: unobserved factors affecting both emissions intensity and local air quality may confound the relationship. Additionally, allocation ratios co-move with operational decisions (capacity utilization, fuel switching) that directly affect emissions.

The directed acyclic graph (DAG) in Figure 4 illustrates the causal structure. The target estimand is the effect of P_{it} on Y_{it} , controlling for confounders. Key confounding pathways include:

- **Facility-level time-invariant unobservables** (U_i): Plant technology, combustion efficiency, and location affect both policy exposure and emissions. Absorbed by facility fixed effects.
- **Time-varying regional factors** (U_{rt}): Electricity demand, fuel prices, and regional economic conditions affect plant operations and allocation ratios. Absorbed by Region×Year fixed effects.
- **Plant-level time-varying unobservables** (U_{it}): Dispatch/utilization, maintenance status, and operational efficiency changes affect both verified emissions (determining allocation ratios) and pollutant output. This is the key identification challenge—see Section 4.2.

Table 2 Summary Statistics for Analysis Panel

Variable	Mean	Std. Dev.	Min	Max
<i>Panel Structure</i>				
Facilities			251	
Facility-years			1,380	
Years per facility	5.5	0.9	3	6
<i>Outcome 1: Verified CO₂ Emissions</i>				
Verified emissions (ktCO ₂ /yr)	1,250	1,890	100	15,200
<i>Outcome 2: Satellite NO_x Proxy</i>				
NO _x emission rate (kg/s)	[–]	[–]	[–]	[–]
Above detection limit (%)			[–]	
<i>ETS Policy Variables</i>				
Allocation ratio	0.72	0.45	0.01	18.5
Shortfall (ktCO ₂)	485	1,120	-2,100	12,800
<i>Plant Characteristics</i>				
Capacity (MW)	1,180	1,450	50	8,200
Gas share	0.44	0.42	0	1
Coal share	0.19	0.35	0	1

Note: Sample includes EU ETS-regulated large combustion plants with verified emissions ≥ 100 ktCO₂ in at least one year during 2018–2023. NO_x emission rates are derived from Beirle-style flux-divergence applied to TROPOMI NO₂ observations. Detection limit follows Beirle et al. (2023) at approximately 0.11 kg/s.

- **Observed operational factors (X_{it}):** Capacity and fuel mix affect both verified emissions and pollutant emissions. Controlled directly.

The Beirle-style flux-divergence approach addresses atmospheric confounding by focusing on the spatial gradient (advection) rather than absolute column densities, making it sensitive to local emissions rather than background concentrations. Fixed effects absorb facility-level and time-varying regional confounders for both outcomes.

4.2 Identification Strategy and Variable Selection

The identification strategy relies on two key choices: (i) what to control for, and (ii) what *not* to control for. Both are essential to avoid bias.

4.2.1 Why We Control for Region×Year Effects

Regional electricity demand, fuel prices, and economic conditions create time-varying confounding: demand shocks increase plant utilization, raising both verified emissions (lowering R_{it}) and NO₂ output. Without adjustment, this creates spurious correlation between policy stringency and pollution.

Region×Year fixed effects absorb these common shocks additively. The identifying variation becomes: *within the same region and year, do facilities with different allocation ratios exhibit different NO₂ enhancement?* This comparison holds regional conditions constant while exploiting cross-facility variation in policy exposure.

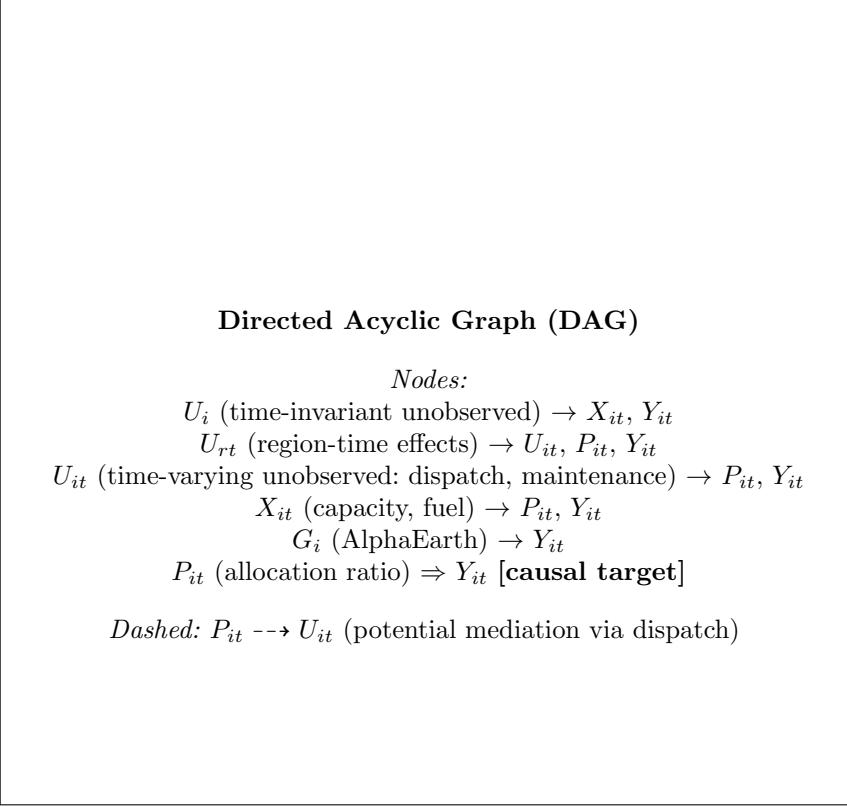


Fig. 4 Directed acyclic graph illustrating the causal structure. The target estimand is $P_{it} \rightarrow Y_{it}$. Facility FE absorbs U_i ; Region×Year FE absorbs the common component of U_{rt} . Plant-level time-varying unobservables (U_{it}), including dispatch, are not controlled due to their dual role as confounder and potential mediator.

The PyPSA-Eur k-means clusters [9] define regions based on transmission network topology rather than administrative boundaries. Facilities in the same electrical cluster face correlated wholesale prices (due to shared market clearing and transmission constraints), similar dispatch patterns, and common grid congestion. This makes them economically meaningful units for absorbing regional time-varying confounders.

4.2.2 Why We Do Not Control for U_{it} (Plant-Level Time-Varying Unobservables)

Plant-level time-varying unobservables (U_{it}) include dispatch/utilization, maintenance status, operational efficiency changes, and idiosyncratic output demand. These variables—particularly dispatch—present a “bad control” problem because dispatch is simultaneously:

1. **A confounder:** Demand shocks \rightarrow higher dispatch \rightarrow higher verified emissions \rightarrow lower R_{it} . The same shocks \rightarrow more combustion \rightarrow higher NO₂.

2. **A potential mediator:** If policy affects merit order bidding (facilities with carbon shortfalls bid higher → get dispatched less), then: $P_{it} \rightarrow U_{it} \rightarrow Y_{it}$.

Controlling for U_{it} (or proxies such as generation data) would absorb both effects. The confounding component should be removed, but blocking the mediation pathway would attenuate the true policy effect toward zero. Since we cannot empirically separate these components without strong assumptions about the dispatch mechanism, we do not control for U_{it} directly. Instead, Region×Year FE absorbs the common (regional) component of U_{it} —since dispatch responds primarily to regional demand and fuel prices—leaving only facility-specific deviations as residual confounding. These facility-specific deviations are plausibly second-order and orthogonal to the allocation ratio conditional on capacity and fuel mix controls.

4.2.3 Why Not Facility×Year Fixed Effects?

Facility×year fixed effects (α_{it}) would absorb *all* within-facility-year variation. Since treatment varies at the facility-year level, this leaves no variation to identify β . More subtly, interactive fixed effects models (which estimate facility-specific responses to common time factors) risk similar problems: if the estimated factor loadings correlate with how allocations were assigned, conditioning on them may open backdoor paths through the allocation mechanism or block mediation pathways.

4.2.4 Addressing Simultaneity in the Allocation Ratio

The allocation ratio $R_{it} = A_i/V_{it}$ depends on current-year verified emissions V_{it} , which are determined by dispatch. This creates apparent simultaneity: $U_{it} \rightarrow R_{it}$ (dispatch affects the denominator) while $R_{it} \rightarrow U_{it}$ (policy affects dispatch via carbon costs). However, two features of the EU ETS and the econometric design mitigate this concern.

First, the *numerator*—free allocations A_i —is largely predetermined. Allocations are set at the beginning of each ETS trading phase based on historical benchmarks and activity levels, not current operations. The identifying variation in R_{it} within a region-year therefore reflects primarily cross-facility differences in these predetermined allocations rather than differences in current dispatch.

Second, Region×Year FE absorbs the common drivers of dispatch variation. Carbon prices are market-determined and common across facilities; regional demand and fuel price shocks are absorbed by the fixed effects. The remaining cross-facility variation in dispatch within a region-year is driven by facility-specific factors (capacity, fuel mix, technical constraints) that are either controlled directly or absorbed by facility FE.

To the extent that endogenous dispatch variation does contaminate R_{it} , the bias is likely *attenuating*: facilities with high dispatch have both lower allocation ratios (higher denominator) and higher NO₂ (more combustion), creating positive correlation between R_{it} and Y_{it} that works against finding a negative policy effect. Estimates should therefore be interpreted as conservative.

4.2.5 Residual Threats and Interpretation

The primary residual threat is facility-specific time-varying confounding (U_{it})—maintenance outages, unexpected efficiency changes, or idiosyncratic demand for a specific plant’s output. These are plausibly second-order and unlikely to systematically correlate with allocation ratios conditional on our controls. Future work incorporating plant-level generation data could address this directly; a discussion of economic dispatch and power system optimization is reserved for subsequent analysis.

4.3 Treatment Definitions

4.3.1 Continuous Treatment

The primary treatment variable is the allocation ratio R_{it} (Equation 9). Lower values indicate greater policy stringency—facilities must purchase more allowances on the carbon market when $R_{it} < 1$. The expected effect is that lower allocation ratios induce emissions reductions, leading to lower verified CO₂ and correspondingly lower satellite-derived NOx (as NOx is a combustion co-pollutant).

4.3.2 Discrete Treatment for Staggered DiD

For the Callaway-Sant’Anna estimator, I define a binary treatment indicator:

$$D_{it} = \mathbf{1}\{R_{it} < 1\} \quad (21)$$

where $\mathbf{1}\{\cdot\}$ is the indicator function. Treatment onset (cohort assignment) is defined as the first year a facility becomes treated:

$$G_i = \min\{t : D_{it} = 1\} \quad (22)$$

Facilities never treated are assigned $G_i = 0$ (the never-treated control group).

4.4 Specification 1: TWFE with Continuous Treatment

The baseline two-way fixed effects specification is:

$$Y_{it} = \alpha_i + \gamma_t + \beta R_{it} + \mathbf{X}'_{it} \boldsymbol{\delta} + \varepsilon_{it} \quad (23)$$

where:

- α_i : Facility fixed effects (absorb time-invariant unobservables)
- γ_t : Year fixed effects (absorb common time shocks)
- β : Treatment effect of interest (effect of unit increase in allocation ratio)
- \mathbf{X}_{it} : Time-varying controls (capacity, fuel shares) and static AlphaEarth embedding controls ($\mathbf{e}_i \in \mathbb{R}^{64}$)
- ε_{it} : Idiosyncratic error

The coefficient β is identified from within-facility variation in allocation ratios over time, after controlling for common year effects. For the CO₂ outcome, a positive β

would indicate that higher allocation ratios (less policy stringency) are associated with higher verified emissions—equivalently, that policy stringency reduces CO₂. For the NOx outcome, a positive β would indicate corresponding reductions in satellite-derived NOx, consistent with co-pollutant dynamics.

Standard errors are clustered at the power system region level using PyPSA-Eur network partitions

citehorsch2018. These clusters are derived from k-means clustering on the European high-voltage transmission network topology, grouping electrical buses based on network connectivity and transmission capacity. The resulting regions capture economic and physical similarities between power plants: facilities in the same electrical region face correlated wholesale electricity prices, are subject to similar dispatch patterns and grid constraints, and share common market clearing mechanisms. This makes them appropriate units for clustered inference in studies of power sector behavior.

4.5 Specification 2: TWFE with Region×Year Fixed Effects

A more demanding specification replaces year fixed effects with region×year interactions:

$$Y_{it} = \alpha_i + \gamma_{r(i),t} + \beta R_{it} + \mathbf{X}'_{it}\boldsymbol{\delta} + \varepsilon_{it} \quad (24)$$

where $\gamma_{r(i),t}$ are cluster-by-year fixed effects and $r(i)$ denotes the PyPSA-Eur k-means cluster containing facility i .

This specification absorbs all region-specific time-varying confounders, including regional electricity prices, demand conditions, and weather patterns. Identification relies on within-region, within-year variation in allocation ratios—comparing facilities in the same region and year that differ in policy stringency.

4.6 Specification 3: Callaway-Sant'Anna Difference-in-Differences

The Callaway and Sant'Anna [11] estimator addresses potential bias in standard TWFE when treatment timing varies across units and treatment effects are heterogeneous. The method estimates group-time average treatment effects:

$$ATT(g, t) = \mathbb{E}[Y_{it}(1) - Y_{it}(0) | G_i = g] \quad (25)$$

for each cohort g (first treatment year) and calendar time t .

Key features of the implementation:

- **Control group:** Never-treated units (facilities with $R_{it} \geq 1$ in all years)
- **Estimation method:** Doubly robust (combining outcome regression and propensity score weighting)
- **Anticipation:** Zero (no anticipation effects assumed)

The group-time ATTs are aggregated to produce:

1. **Simple aggregate ATT:** Overall average treatment effect across all treated observations:

$$ATT^{\text{simple}} = \sum_g \sum_{t \geq g} w_{g,t} \cdot ATT(g, t) \quad (26)$$

2. **Dynamic/event-study ATT:** Effects by time relative to treatment:

$$ATT(e) = \sum_g w_g \cdot ATT(g, g + e) \quad (27)$$

where e is event time (years since treatment onset)

The event-study specification enables pre-trend testing: under parallel trends, $ATT(e) = 0$ for $e < 0$. A Wald test of joint nullity for pre-treatment periods provides formal evidence on the parallel trends assumption.

4.7 Network-Based Clustering for Inference

Standard errors are clustered at the PyPSA-Eur power system cluster level throughout, representing a novel application of model-derived features for econometric inference. Importantly, these clusters are *not* administrative regions but rather k-means clusters computed on power system features—specifically, the topology and electrical characteristics of the European high-voltage transmission network extracted from ENTSO-E data [9]. The clustering algorithm groups electrical buses (substations) to minimize within-cluster transmission impedance, producing clusters that reflect the physical and economic structure of the electricity grid rather than political boundaries.

This clustering strategy is grounded in recent theoretical foundations for network cluster-robust inference. [10] establish that valid cluster-robust standard errors require clusters with low “conductance”—the ratio of edges crossing cluster boundaries to total edges within clusters. Intuitively, observations within a cluster should be more correlated with each other than with observations in other clusters. Power system network clusters satisfy this requirement by construction: the k-means objective minimizes within-cluster electrical distance (transmission impedance), which directly corresponds to minimizing cross-cluster connections in the network graph. Facilities in the same electrical cluster face:

- **Correlated wholesale prices:** Shared market clearing and limited cross-cluster transmission capacity mean prices co-move within clusters
- **Similar dispatch patterns:** Grid constraints and congestion patterns affect which plants are called to generate
- **Common demand shocks:** Regional electricity demand fluctuations propagate through the local network

I use 128 clusters, a standard resolution in the PyPSA-Eur literature for EU-wide analyses that balances geographic granularity with sufficient observations per cluster for reliable inference. This represents a novel application of model-derived clustering for econometric inference: rather than using geographic proximity, administrative boundaries, or data-driven clustering on outcome variables, I use clusters computed

from features of an external domain-specific model (the power system network) that captures the economically-relevant correlation structure.

Let $r \in \{1, \dots, R\}$ index clusters. The cluster-robust variance estimator for the TWFE coefficient is:

$$\widehat{\text{Var}}(\hat{\beta}) = (\mathbf{X}' \mathbf{M} \mathbf{X})^{-1} \left(\sum_{r=1}^R \mathbf{X}'_r \hat{\mathbf{u}}_r \hat{\mathbf{u}}'_r \mathbf{X}_r \right) (\mathbf{X}' \mathbf{M} \mathbf{X})^{-1} \quad (28)$$

where \mathbf{M} is the residual maker for fixed effects and $\hat{\mathbf{u}}_r$ are residuals for observations in cluster r .

5 Results

[Results to be added upon completion of empirical analysis.]

This section will present:

1. Descriptive statistics on treatment and outcome distributions
2. Effects on verified CO₂ emissions (EU ETS) — TWFE Specifications 1 and 2
3. Effects on satellite-derived NOx emission proxy — TWFE Specifications 1 and 2
4. Callaway-Sant'Anna event study results with pre-trend tests for both outcomes
5. Comparison of CO₂ and NOx responses: agreement in sign and magnitude
6. Robustness checks varying sample restrictions and specifications

5.1 Treatment Distribution

The allocation ratio exhibits substantial within-facility variation over the sample period. [Distribution statistics and figures to be added.]

5.2 Main Estimates

Table 3 reports the main estimation results.

5.3 Event Study

Figure ?? presents the Callaway-Sant'Anna event study results for both outcomes, plotting dynamic treatment effects by years relative to treatment onset. Comparison of CO₂ and NOx event-study patterns provides a cross-validation of policy effects: if carbon pricing reduces combustion, both outcomes should show similar post-treatment declines.

[Event study figure to be added.]

Pre-trend tests: [Wald test results to be added.]

5.4 Robustness

[Robustness analysis to be added, including:]

- Alternative emissions thresholds (250 kt, 500 kt)
- Alternative treatment definitions (shortfall instead of ratio)

Table 3 Main Estimation Results: Effect of Allocation Ratio on Dual Outcomes

	Verified CO ₂ (ktCO ₂ /yr)		Satellite NOx (kg/s)	
	Spec 1	Spec 2	Spec 1	Spec 2
Allocation Ratio (R_{it})	[(-)]	[(-)]	[(-)]	[(-)]
Facility FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Region \times Year FE	No	Yes	No	Yes
Controls	Yes	Yes	Yes	Yes
Observations	[-]	[-]	[-]	[-]
Facilities	[-]	[-]	[-]	[-]
R^2 (within)	[-]	[-]	[-]	[-]

Note: Standard errors clustered by PyPSA-Eur k-means cluster in parentheses. Controls include capacity (MW), fuel shares (coal, gas), and AlphaEarth embeddings (64 dimensions). * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

- Heterogeneity by fuel type and country
- Sensitivity to NOx detection limit (excluding facilities below 0.11 kg/s)
- Comparison of CO₂ vs NOx elasticities (percentage changes)

6 Discussion

[Discussion to be completed with empirical results.]

This study develops a methodological framework for evaluating climate policy impacts using dual outcomes: verified EU ETS CO₂ emissions and satellite-derived NOx emission proxies. This approach enables comprehensive assessment of both greenhouse gas reductions and air quality co-benefits. Several aspects merit discussion.

6.1 Interpretation of Estimates

The allocation ratio treatment variable has a natural interpretation in terms of policy stringency. A coefficient estimate of β on R_{it} implies that a 0.1 unit decrease in the allocation ratio (from 1.0 to 0.9, representing a shift from full free allocation to 10% shortfall) is associated with a -0.1β change in the outcome variable.

For verified CO₂: The outcome is directly measured and reported under EU ETS compliance, providing high-quality estimates of policy effects on greenhouse gas emissions.

For satellite-derived NOx: The Beirle-style flux-divergence approach provides a metric that is physically grounded in the continuity equation, with advection proportional to local emissions minus chemical sinks. Unlike simple column density comparisons, this method isolates facility-attributable emissions from regional background. However, interpretation requires caution due to the substantial uncertainty components discussed in Section 3.4.

Cross-validation: Agreement in sign and broad magnitude between CO₂ and NOx responses provides confidence that both outcomes are capturing genuine policy effects. Perfect correspondence is not expected—NOx and CO₂ have different emission factors by fuel type and combustion conditions—but directional agreement supports the validity of both measures.

6.2 Identification Concerns

Several potential threats to identification remain. First, **operational confounding**: facilities may respond to high carbon prices by adjusting operations (reducing output, switching fuels) in ways that are not fully captured by the control variables. To the extent that these operational responses are the mechanism through which carbon pricing affects air quality, this is not problematic—it is the causal effect of interest. However, if operational changes are driven by other factors correlated with allocation ratios (e.g., electricity prices), bias may result.

Second, **measurement error in satellite NOx**: The Beirle-style flux-divergence approach is subject to multiple uncertainty sources (Section 3.4): lifetime parameterization (50% relative uncertainty), product differences between OFFL and PAL ($\pm 25\%$), NOx/NO₂ scaling ($\sim 7\%$), and spatial integration noise. These compound to total uncertainties of 20–40% per facility-year, with potential systematic low bias of $\sim 20\%$ relative to reported emissions. This measurement error attenuates treatment effect estimates toward zero, making the satellite-based findings conservative.

Third, **spillovers**: if carbon pricing induces substitution across facilities (e.g., shifting generation from high-cost to low-cost plants), the stable unit treatment value assumption (SUTVA) may be violated. The region \times year fixed effects specification partially addresses this by absorbing regional substitution patterns.

6.3 Limitations of the Satellite NOx Proxy

The satellite-derived NOx estimates should be interpreted with several limitations in mind:

- **Use of OFFL L3 instead of PAL:** Beirle et al. (2023) use the PAL NO₂ product, which provides 10–40% higher TVCDs. The OFFL L3 product available via Google Earth Engine introduces additional uncertainty.
- **Simplified NOx/NO₂ scaling:** The photostationary-state approximation may not hold near stack plumes; I apply a fixed ratio of 1.38 ± 0.10 following Beirle et al. (approximately 7% uncertainty).
- **Integration radius trade-off:** The 15 km radius balances capturing the full point source signal against spatial interference from neighboring sources. For facilities in dense industrial areas, some signal contamination is likely.
- **Spatial interference:** For facilities with another ETS facility within 20 km (flagged via `interfered_20km`), the satellite outcome reflects emissions of the *local cluster* rather than an isolated point source. Treatment effect estimates for these facilities should be interpreted as cluster-level responses. Sensitivity analyses excluding interfered facilities test whether results are robust to this contamination.

- **Detection limit:** Beirle et al. Sect. 3.11.1 report detection limits of 0.11 kg/s (standard conditions) down to 0.03 kg/s (ideal high-albedo conditions). Since we do not implement a surface reflectivity mask, we flag observations against both thresholds and present sensitivity analysis using the conservative 0.11 kg/s threshold for Europe.
- **Skipping automatic identification:** We skip Beirle’s automatic point-source identification algorithm because we have known ETS/LCP source locations. This design choice is explicitly endorsed by the authors and is conceptually appropriate for our use case, but it means we do not benefit from their classification step that filters artifacts and non-point sources. To compensate, we apply simplified significance filters (detection limit, statistical integration error, interference flag).
- **Unmodeled corrections:** AMF correction and plume height effects following Beirle et al. (2023) are not fully implemented, contributing to structural uncertainty. Topographic correction is implemented using SRTM elevation data.
- **Temporal averaging approach:** Beirle et al. compute temporal means of A^* at orbit resolution before applying a single lifetime correction using mean wind speed. I instead aggregate to daily resolution and apply the lifetime correction per-day before temporal averaging, i.e., $E = \text{mean}[c_\tau(w_d) \times A_d^*]$ rather than $c_\tau(\bar{w}) \times \text{mean}(A^*)$. Since c_τ is nonlinear in wind speed, these are not mathematically identical; however, the difference is small for typical wind speed variance and the per-day approach is more appropriate for constructing annual panel estimates rather than a single multi-year catalog value.

These design choices—using OFFL instead of PAL, simplified significance flags, skipping the automatic identification algorithm—primarily increase noise and potential attenuation bias, not spurious detection. The satellite outcome remains a *physically grounded but noisy proxy*; ETS verified CO₂ remains the primary outcome for causal inference.

6.4 Methodological Contributions: ML-Derived Features in Causal Inference

This study contributes to a growing literature on incorporating machine learning-derived features into causal inference frameworks [5–7]. Two aspects merit particular discussion.

Geospatial foundation model embeddings as controls. The use of AlphaEarth embeddings [8] demonstrates that pre-trained geospatial representations can serve as effective high-dimensional controls in panel settings. Unlike manually-specified geographic controls (distance to coast, elevation, population density), foundation model embeddings capture complex, nonlinear combinations of features that the model learned to be predictive across diverse tasks. The key assumption—that these embeddings capture confounders affecting both policy exposure and air quality outcomes—is plausible given that the embeddings encode land use, infrastructure, and climate patterns that likely correlate with both industrial activity and pollution

dispersion. Future work should investigate the sensitivity of causal estimates to different embedding specifications and the conditions under which learned representations provide valid confounding adjustment.

Network-based clustering from external models. The use of PyPSA-Eur power system clusters for both fixed effects and inference represents a novel application of model-derived features for econometric analysis. Traditional approaches to clustered standard errors use geographic proximity or administrative boundaries that may not reflect the economically-relevant correlation structure. By using clusters derived from transmission network topology [9], which minimize within-cluster electrical distance, I define regions where facilities face correlated prices, dispatch patterns, and demand shocks—precisely the correlations that motivate cluster-robust inference [10]. This approach could be extended to other networked industries where external models of the network structure are available.

6.5 Limitations

The study has several limitations. First, **sample attrition**: the requirement for valid linkage across three independent data sources (LCP registry, EU Registry cross-walk, EUTL compliance data) combined with the satellite detection threshold (100 ktCO₂/yr) reduces the initial universe of 3,405 LCP plants to 251 facilities (7.4% retention). This attrition reduces statistical power and may introduce selection bias if the matched sample differs systematically from the broader LCP population. The 500m spatial clustering threshold, while appropriate for grouping co-located emission sources into coherent facilities, may occasionally merge distinct plants or fail to group plants that share a common plume.

Second, the sample is restricted to large facilities with sufficient emissions for satellite detection (0.11 kg/s NOx conservative threshold), limiting generalizability to smaller sources. Third, the annual temporal resolution may miss short-run dynamics. Fourth, the satellite NOx proxy is subject to the multiple uncertainty components detailed in Section 3.4, though these primarily introduce classical measurement error that attenuates rather than biases estimates.

Fifth, regarding the ML-derived features: the AlphaEarth embeddings are treated as exogenous controls, but their validity for causal adjustment depends on the assumption that they capture relevant confounders without inducing collider bias. Similarly, the PyPSA-Eur clusters assume that transmission network topology correctly reflects the correlation structure of unobserved shocks affecting power plant operations. These assumptions are difficult to verify directly and represent areas for future methodological development.

7 Conclusion

This study develops and demonstrates a novel framework for comprehensively evaluating climate policy impacts using dual emission outcomes. By linking administrative EU ETS compliance data with satellite-derived NOx emission proxies constructed via the Beirle-style flux-divergence method, I construct a facility-level panel that enables

causal inference on how carbon market stringency affects both verified CO₂ emissions and satellite-observable combustion co-pollutants.

The study makes three methodological contributions, two of which follow a recent trend in causal inference toward incorporating machine learning-derived features to address high-dimensional confounding [5–7].

First, I demonstrate the integration of geospatial foundation model embeddings as high-dimensional controls in panel-based climate monitoring studies. Using Google AlphaEarth [8]—64-dimensional representations learned from multi-source satellite imagery, climate reanalysis, and geotagged text—I capture between-unit heterogeneity arising from local geographic, infrastructural, and climate context in a data-efficient manner. This extends prior work on learned embeddings for causal inference [6] from text to the geospatial domain, and is particularly relevant for difference-in-differences settings where high-dimensional spatial confounders may violate parallel trends if left uncontrolled [7].

Second, I introduce network-based clustering derived from an external power system model for econometric inference. Standard errors are clustered using PyPSA-Eur power system clusters [9]—k-means clusters computed on transmission network topology features—that group facilities facing correlated wholesale prices, dispatch patterns, and grid constraints. This approach is grounded in recent theoretical foundations showing that valid cluster-robust inference requires low-conductance clusters [10], which network-derived clusters satisfy by construction. To my knowledge, this represents the first application of model-derived clustering—where clusters are computed on features from an external domain-specific model rather than on the outcome data itself—for econometric inference in policy evaluation.

Third, for the satellite outcome, I implement a simplified Beirle-style flux-divergence method that provides physically grounded NOx emission estimates at the facility level. By computing the advection—the wind-aligned spatial derivative of NO₂ column density—and integrating over a 15 km disc with lifetime correction, this approach provides facility-specific NOx estimates suitable for panel econometric analysis. This methodology is grounded in the continuity equation and follows the approach of Beirle et al. (2019, 2021, 2023) [1–3], which has been validated against reported emissions from regulatory agencies.

[Substantive empirical conclusions to be added with results.]

The broader contribution of this work is demonstrating that combining administrative emissions records with satellite-derived proxies, along with ML-derived controls and network-informed inference, can provide comprehensive evaluation of climate policy impacts at the individual emitter level. The dual-outcome approach offers several advantages: (i) verified CO₂ provides the gold standard for measuring policy effects on greenhouse gas output; (ii) satellite-derived NOx provides an independent check on combustion activity and enables testing co-benefit hypotheses; and (iii) agreement between outcomes provides cross-validation that both measures are capturing genuine policy effects.

As satellite instruments improve in resolution and retrieval accuracy, and as methods like the Beirle flux-divergence approach become more refined, this framework could enable near-real-time monitoring of both carbon and co-pollutant emissions

from regulated facilities. Future work could extend this framework to methane point sources (using TROPOMI CH₄), investigate heterogeneity across plant types and regulatory contexts, and further develop the theoretical foundations for using learned representations and model-derived clusters in causal inference.

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Declarations

- **Funding:** Not applicable
- **Conflict of interest:** The author declares no conflicts of interest
- **Data availability:** EEA Large Combustion Plant data available from the European Environment Agency Industrial Emissions Portal. EU ETS data available from the European Union Transaction Log. TROPOMI data available via Google Earth Engine. ERA5-Land data available from the Copernicus Climate Data Store.
- **Code availability:** Analysis code available upon request

Appendix A Data Pipeline Details

A.1 ID Normalization for ETS Linking

Linking LCP plants to ETS installations requires normalizing identifiers from different sources. EU Registry identifiers follow patterns such as FR000000000210535 (padded numeric) or FR-new-07101261 (new format). Pyeutl installation IDs follow the format AT_200165 (country code underscore numeric).

The normalization procedure:

1. Extract country code (first 2 characters)
2. Extract all numeric substrings
3. Select longest numeric substring, strip leading zeros
4. Combine as CC_NNN format

This procedure successfully matches 799 of 932 facilities (85.7%) to ETS installations.

A.2 Fuel Type Classification

Raw LCP fuel types are mapped to standardized categories:

- **Gas:** NaturalGas, NG, Gas
- **Coal:** Coal, Lignite, PC, BIT, SUB, ANT
- **Oil:** LiquidFuels, DFO, RFO, KER
- **Biomass:** Biomass, WDL, WDS, AB
- **Other Gas:** OtherGases, OBG

Fuel types used by fewer than 10% of facility-years (Other Solid, Peat) are dropped, shares renormalized, and facilities with no remaining fuel coverage are excluded from the sample.

Appendix B PyPSA-Eur Network Clustering

The analysis uses PyPSA-Eur power system clusters [9], which are *not* geographic or administrative regions but rather k-means clusters computed directly on power system features extracted from the European high-voltage transmission network (ENTSO-E data). The clustering algorithm groups electrical buses (substations) based on network connectivity, line impedances, and transmission capacity. The objective function minimizes within-cluster electrical distance, producing clusters where facilities face similar grid constraints, transmission losses, and wholesale price dynamics.

This clustering approach has a theoretical justification grounded in recent work on network cluster-robust inference. [10] establish that valid cluster-robust standard errors require clusters with low “conductance”—formally, the ratio of edges crossing cluster boundaries to total within-cluster edges. The k-means clustering on transmission network features directly minimizes this quantity: by grouping buses to minimize within-cluster electrical distance (impedance), the algorithm produces clusters with few high-capacity transmission lines crossing boundaries. Facilities within the same cluster are therefore more strongly connected to each other (through the grid) than to facilities in other clusters, satisfying the theoretical requirements for cluster-robust inference.

This represents a novel application of model-derived clustering for econometric inference. Rather than using geographic proximity (which ignores network topology), administrative boundaries (which may cut across electrically-connected regions), or data-driven clustering on outcome variables (which risks overfitting), I use clusters computed from features of an external domain-specific model—the power system transmission network—that captures the economically-relevant correlation structure *a priori*.

For this analysis, the 128-region resolution is used, providing sufficient granularity to capture sub-national variation while maintaining adequate within-region sample sizes for clustered inference. Each facility is assigned to the PyPSA-Eur cluster containing the nearest network bus.

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