

## Topic: Expectation of random variables

In today's problem class we will be computing expectations of various discrete and continuous random variables.

1. Compute the mean of the following random variables.

- (a)  $X \sim \text{Poi}(\lambda)$ ,
- (b)  $X \sim \text{Exp}(\lambda)$ ,

**Solution:** Recall that we have

$$E(X) = \begin{cases} \sum_n nP(X = n), & \text{if } X \text{ is discrete,} \\ \int x f_X(x) dx, & \text{if } X \text{ is continuous.} \end{cases}$$

(a) In the case when  $X \sim \text{Poi}(\lambda)$ , we have

$$\begin{aligned} E(X) &= \sum_{n=0}^{\infty} nP(X = n) = \sum_{n=0}^{\infty} n \frac{\lambda^n}{n!} e^{-\lambda} = e^{-\lambda} \sum_{n=1}^{\infty} \frac{\lambda^n}{(n-1)!} \\ &= e^{-\lambda} \lambda \sum_{n=0}^{\infty} \frac{\lambda^n}{n!} = e^{-\lambda} e^{\lambda} \lambda = \lambda. \end{aligned}$$

(b) In the case when  $X \sim \text{Exp}(\lambda)$ , we have

$$E(X) = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \frac{1}{\lambda} \int_0^{\infty} z e^{-z} dz = \Gamma(2) \frac{1}{\lambda} = \frac{1}{\lambda}.$$

2. Let  $X$  be a continuous random variable with the following p.d.f.:

$$f_X(x) = \begin{cases} \theta \lambda e^{-\lambda x}, & x \geq 0; \\ (1 - \theta) \lambda e^{\lambda x}, & x < 0. \end{cases}$$

where  $\lambda$  and  $\theta$  are constants such that  $\lambda > 0$  and  $0 \leq \theta \leq 1$ .

- (a) Show that  $f_X(x)$  is a valid p.d.f..
- (b) Find  $E(X)$ . *Hint: You might find it useful to look up the definition of the Gamma function.*
- (c) Find  $\text{Var}(X)$

**Solution:**

- (a)  $f_X(x) \geq 0$  for all  $x \in \mathbb{R}$ , since  $\lambda > 0, \theta \in (0, 1)$  and the exponential function is always positive, and

$$\begin{aligned} \int_{-\infty}^{\infty} f_X(x) dx &= \int_{-\infty}^0 (1 - \theta) \lambda e^{\lambda x} dx + \int_0^{\infty} \theta \lambda e^{-\lambda x} dx \\ &= (1 - \theta) e^{\lambda x} \Big|_{-\infty}^0 - \theta e^{-\lambda x} \Big|_0^{\infty} = 1 - \theta + \theta = 1. \end{aligned}$$

- (b) For the expectation, we have

$$E(X) = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^0 (1 - \theta) \lambda x e^{\lambda x} dx + \int_0^{\infty} \theta \lambda x e^{-\lambda x} dx.$$

Recall that the Gamma function is defined as

$$\Gamma(t) = \int_0^{\infty} x^{t-1} e^{-x} dx, \quad t > 0.$$

In the case, when  $t \in \mathbb{N}$ , we have that  $\Gamma(t) = (t-1)!$ .

We have, using the transformation  $z = -\lambda x$ ,  $dz = -\lambda dx$ :

$$\begin{aligned} \int_{-\infty}^0 (1-\theta)\lambda x e^{\lambda x} dx &= (1-\theta) \int_{\infty}^0 (-z) e^{-z} (-1) dz \frac{1}{\lambda} = (1-\theta)(-1) \int_0^{\infty} z e^{-z} dz \frac{1}{\lambda} \\ &= -\frac{(1-\theta)}{\lambda} \Gamma(2) = \frac{(\theta-1)}{\lambda}. \end{aligned}$$

Similarly, using the transformation  $z = \lambda x$ ,  $dz = \lambda dx$ :

$$\begin{aligned} \int_0^{\infty} \theta \lambda x e^{-\lambda x} dx &= \int_0^{\infty} \theta z e^{-z} dz \frac{1}{\lambda} \\ &= \frac{\theta}{\lambda} \Gamma(2) = \frac{\theta}{\lambda}. \end{aligned}$$

Hence,

$$E(X) = \frac{(\theta-1)}{\lambda} + \frac{\theta}{\lambda} = \frac{(2\theta-1)}{\lambda}.$$

(c) We compute the second moment first:

$$E(X^2) = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_{-\infty}^0 (1-\theta)\lambda x^2 e^{\lambda x} dx + \int_0^{\infty} \theta \lambda x^2 e^{-\lambda x} dx.$$

For the first integral, using the same transformation as above  $z = -\lambda x$ ,  $dz = -\lambda dx$ , we get:

$$\begin{aligned} \int_{-\infty}^0 (1-\theta) \frac{1}{\lambda} \lambda^2 x^2 e^{\lambda x} dx &= (1-\theta) \frac{1}{\lambda} \int_{\infty}^0 z^2 e^{-z} (-1) dz \frac{1}{\lambda} = \frac{(1-\theta)}{\lambda^2} \int_0^{\infty} z^2 e^{-z} dz \\ &= \frac{(1-\theta)}{\lambda^2} \Gamma(3) = \frac{2(1-\theta)}{\lambda^2}. \end{aligned}$$

Similarly, for the second integral, using the transformation  $z = \lambda x$ ,  $dz = \lambda dx$ :

$$\begin{aligned} \int_0^{\infty} \theta \lambda x^2 e^{-\lambda x} dx &= \frac{\theta}{\lambda^2} \int_0^{\infty} z^2 e^{-z} dz \\ &= \frac{\theta}{\lambda^2} \Gamma(3) = \frac{2\theta}{\lambda^2}. \end{aligned}$$

So altogether we have

$$E(X^2) = \frac{2}{\lambda^2}$$

and hence

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{2}{\lambda^2} - \frac{(2\theta-1)^2}{\lambda^2} = \frac{1-4\theta^2+4\theta}{\lambda^2}.$$