

# Linear models

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- ① We consider linear model

$$y = \beta_0 + \beta_1 x_1 + \cdots + \beta_j x_j + u$$

- ② We use t-test in order to test hypotheses about a particular  $\beta_k$
- ③ Remark:  $\beta_k$  are unknown features of the population and we will never know them with certainty. Nevertheless, we can hypothesize about the value of  $\beta_k$  and then use statistical inference to test our hypothesis.

# Testing against one-sided alternatives

- We consider null hypothesis

$$H_0 : \beta_k = 0$$

- Intuition: since  $\beta_k$  measures the partial effect of  $x_k$  on  $y$ ,  $H_0$  means that once  $x_1, x_2, \dots, x_{k-1}, x_{k+1}, \dots, x_j$  have been accounted for,  $x_k$  has no effect on the expected value of  $y$
- We test  $H_0$  against  $H_1 : \beta_k > 0$ .

- The statistic we use to test  $H_0$  is called the  $t$  statistic or the  $t$  ratio of  $\hat{\beta}_k$  and is defined as

$$t_{\hat{\beta}_k} := \frac{\hat{\beta}_k}{se(\hat{\beta}_k)}.$$

- It is reasonable to use  $t_{\hat{\beta}_k}$  to detect  $\beta_j \neq 0$  since
  - $se(\hat{\beta}_k)$  is always positive
  - $t_{\hat{\beta}_k}$  has the same sign as  $\hat{\beta}_k$
  - for a given value of  $se(\hat{\beta}_k)$  a larger value of  $\hat{\beta}_k$  leads to larger values of  $t_{\hat{\beta}_k}$ .

Few remarks:

- Since we are testing  $H_0 : \beta_k = 0$  it is only natural to look at our unbiased estimator of  $\beta_k$ .
- In practice the point estimate  $\hat{\beta}_k$  will be never exactly zero
- A sample value of  $\hat{\beta}_k$  very far from zero provides evidence against  $H_0$
- $t_{\hat{\beta}_k}$  measures how many estimated standard deviations  $\hat{\beta}_k$  is away from zero
- Values of  $t_{\hat{\beta}_k}$  sufficiently far from zero will result in rejection of  $H_0$ .
- Determining a rule for rejecting  $H_0$  at a given significance level, that is a probability of rejecting  $H_0$  when it is true, requires knowing the sample distribution of  $t_{\hat{\beta}_k}$  which is  $t_{n-k-1}$ , where  $k+1$  is a number of unknown parameters.

# Choice of rejection rule

- Firstly, decide on a significance level or the probability of rejecting  $H_0$  when it is in fact true
- For example: suppose we have decided on a 5% significance level. It means that we are willing to mistakenly reject  $H_0$  when it is true 5% of time
- We are looking at sufficiently large positive value of  $t_{\hat{\beta}_k}$  in order to reject  $H_0$ .
- The definition sufficiently large with a 5% significance level is the 95th percentile in a  $t$  distribution with  $n - k - 1$  degrees of freedom, denote this by  $c$ .
- The rejection rule is that  $H_0$  is rejected in favor of  $H_1$  at the 5% significance level if

$$t_{\hat{\beta}_k} > c.$$

- By our choice of the critical value  $c$ , rejection of  $H_0$  will occur for 5% of all random samples when  $H_0$  is true.

## Two-Sided alternatives

- In applications, it is common to test the null hypothesis  $H_0 : \beta_k = 0$  against a two-sided alternative that is

$$H_1 : \beta_k \neq 0.$$

- When the alternative is two-sided, we are interested in the absolute value of the  $t$  statistic. The rejection rule for  $H_0$  is

$$|t_{\hat{\beta}_k}| > c.$$

- In order to find  $c$ , we again specify a significance level, let say 5%. For a two-tailed test,  $c$  is chosen to make an area in each tail of the  $t$  distribution equal to 2.5%.
- $p$  is the 97.5%th percentile in the  $t$  distribution with  $n - k - 1$  degrees of freedom.
- Check, if  $n - k - 1 = 25$ , the critical value for a two-sided test is  $c = 2.060$ .

- If  $H_0$  is rejected in favor of  $H_1$  at the 5% level, we say that  $x_k$  is statistically significant or statistically different from zero, at the 5% level.
- If  $H_0$  is not rejected, we say that  $x_k$  is statistically insignificant at the 5% level.



## Testing other hypotheses about $\beta_j$

- $H_0 : \beta_j = a_j$ .
- the appropriate  $t$  statistic is

$$t = (\hat{\beta}_j - a_j)/se(\hat{\beta}_j).$$

- As before,  $t$  measures how many estimated standard deviations  $\hat{\beta}_j$  is from the hypothesized value of  $\beta_j$ .
- The general statistic  $t$  is usefully written as

$$t = \frac{\text{estimate} - \text{hypothesized value}}{\text{standard error}}.$$

# Computing p-values for t tests

- Rather than testing a different significance levels, it is more informative to answer the following question: Given the observed value of the  $t$  statistic, what is the smallest level at which the null hypothesis would be rejected?
- this level is known as  $p$ -value for the test.
- The  $p$  value for testing the null hypothesis  $H_0 : \beta_j = 0$  against two-sided alternative is given by

$$\mathbb{P}(|T| > |t|),$$

where for clarity we let  $T$  denote a  $t$  distributed random variable with  $n - j - 1$  degrees of freedom and  $t$  is numerical value of the test statistic.

- the  $p$ -value is the probability of observing a  $t$  statistic as extreme as we did if the null hypothesis is true. That means that small  $p$ -values are evidence against null, large  $p$ -values provide little evidence against  $H_0$ .

- Explain  $wage|educ, exper, tenure$
- load WAGE1.raw

$$y = wage1(:, 1)$$
$$[n, k] = size(wage1)$$
$$X = [ones(n, 1), wage1(:, [2, 3, 4])]$$
$$[n, k] = size(X)$$

- $\beta = (X' \times X)^{-1} \times X' \times y$
- $u = y - X \times \beta$
- $sig2 = u' \times u / (n - 4)$  because we have 3 variables and intercept
- $std = sqrt(diag(sig2 \times inv(X' \times X)))$

•  $\beta =$

-2.8727

0.5990

0.0223

0.1693

•  $std =$

0.7290

0.0513

0.0121

0.0216

- Download modul *jplv6*, decompress it in your working directory
- Test  $H_0 : \beta_{exper} = 0$  (one-sided and two-sided test)
- Calculate the test statistic

$$t = \frac{\beta}{std}.$$

# T-statistic

- Calculate:  $t = \beta./std$
- $t =$

−3.9408

11.6795

1.8528

7.8204

- Formulate  $H_0$  : 'One year of studies brings additional 60 centimes of hourly wage'
- Test  $H_0$
- $H_1$  ? Positive or...? Remember of 95% percentile.



# Critical values and $p$ -values

- Command *tdis\_inv*
- Calculate  $p$ -value for two tests
- Validation of  $H_0$ ?
- Command *tdis\_prb*.

- Trace the histogram of  $u$ .
- What are the properties of distribution? It is normal distribution?

- Perform a regression for the logarithm of the salary and calculate the parameter beta. Draw a new histogram of residuals.
- Detect 10 observations which outlie the most ( 5 the smallest, and 5 the largest), discard them, and recalculate beta.
- How to interpret wage changes for an additional year of education? Test  $H_0 : \beta_{educ} = 0.1$ .
- Reject or accept?
- Calculate the  $p$  value

•  $\beta =$

0.2844

0.0920

0.0041

0.0221

•  $std =$

0.1042

0.0073

0.0017

0.0031

•  $t =$

2.7292

12.5552

2.3914

7.1331

# Testing hypotheses about a single linear combination of the parameters

- Testing hypotheses concerning two parameters  $H_0 : \beta_i = \beta_k$
- For the most part, the alternative is one-sided  $H_1 : \beta_i < \beta_k$
- $t$ -statistic is of the following form

$$t = \frac{\hat{\beta}_k - \hat{\beta}_i}{se(\hat{\beta}_k - \hat{\beta}_i)}$$

- Once we have the  $t$  statistic, testing proceeds as before. We choose significance level for the test, based on df obtain the critical value. Because of the form of  $H_1$ , the rejection rule is of the form  $t < -c$ . Or, we compute  $t$  statistic, and then the  $p$ -value.

# Testing multiple linear restrictions: the F test

- We wish to test multiple hypotheses about underlying parameters  $\beta_1, \dots, \beta_k$ . We want to test whether a set of independent variables has no partial effect on a dependent variable.
- Unrestricted model with  $k$  independent variables

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$$

Number of parameters in unrestricted model is  $k + 1$ .

- The null hypothesis is stated as

$$H_0 : \beta_{k-q+1} = 0, \dots, \beta_k = 0,$$

which puts  $q$  restrictions on the model.

- $H_1 : H_0$  is not true.
- Restricted model

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_{k-q} x_{k-q} + u.$$



$$F := \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n - k - 1)},$$

where  $SSR_r$  is the sum of squared residuals from the restricted model and  $SSR_{ur}$  is the sum of squared residuals from the unrestricted model.



$$q = \text{numerator degrees of freedom} = df_r - df_{ur}$$



$$n - k - 1 = \text{denominator degrees of freedom} = df_{ur}$$

- One can show that under  $H_0$ ,  $F$  is distributed as a  $F$  random variable with  $(q, n - k - 1)$  degrees of freedom
- We will reject  $H_0$  in favor of  $H_1$  when  $F$  is 'sufficiently' large (how large it depends on chosen significance level). The critical value depends on  $q$  and  $n - k - 1$



- Once  $c$  has been obtained, we reject  $H_0$  in favor of  $H_1$  at the chosen significance level if

$$F > c.$$

- If  $H_0$  is rejected, then we say that  $x_{k-q+1}, \dots, x_k$  are jointly statistically significant at the appropriate significance level.
- Remark: The test alone does not allow us to say which of the variables has a partial effect on  $y$ ; they may all affect  $y$  or maybe just only one affects  $y$ . If  $H_0$  is not rejected, then the variables are jointly insignificant

# The R-squared Form of the F Statistic

- In most applications it is convenient to use a form of the F statistic that can be computed using the  $R$ -squareds from the restricted and unrestricted models. The reason is that  $R$ -squared is always between zero and one
- $R$ -squared form of the F statistic

$$F := \frac{(R_{ur}^2 - R_r^2)/q}{(1 - R_{ur}^2)/(n - k - 1)}$$

# Computing $p$ -values for $F$ Tests

- For reporting the outcomes of  $F$  tests,  $p$ -values are especially useful. Since the  $F$  distribution depends on the numerator and denominator  $df$ , it is difficult to get a feel for how strong or weak the evidence is against the null hypothesis simply by looking at the value of the  $F$  statistic and one or two critical values. In the  $F$  testing context, the  $p$ -value is defined as

$$p - value = \mathbb{P}(\mathcal{F} > F),$$

where  $\mathcal{F}$  is an  $F$  random variable with  $(q, n - k - 1)$  degrees of freedom, and  $F$  is actual value of the test statistic.

- The same interpretation as it did for  $t$  statistics: it is the probability of observing a value of the  $F$  at least as large as we did, given that the null hypothesis is true. A small  $p$ -value is evidence against  $H_0$ .

# Joint hypothesis testing with matlab

- Effect (on a wage) of education = effect of professional experience
- Test  $H_0 : \beta_{educ} = \beta_{exper}$
- Create a variable  $capitaltot = educ + exper$
- $Log(wage)|educ, capitaltot, tenure$
- Test the nullity of the coefficient associated with  $capitaltot$

# Construction of the test variable

- Calculate  $\text{educ} + \text{exper}$


$$\text{test} = X(:, 2) + X(:, 3);$$
$$X = [X(:, [1, 2, 4]), \text{test}];$$

•  $\beta =$

0.2844

0.0879

0.0221

0.0041

•  $std =$

0.1042

0.0070

0.0031

0.0017

•  $t =$

2.7292

12.5880

7.1331

2.3914

- Test  $H_0 : \beta_{educ} = 0, \beta_{exper} = 0$
- 2 restrictions
- Estimate the unrestricted model

$$\log(wage)|educ, exper, tenure,$$

calculate SSR0

- Estimate the restricted model

$$\log(wage)|tenure,$$

calculate SSR1

- Calculate F



- Sums of squares of errors  $SSR0 = u'u$
- $SSR0 = 101.4556$

# Constrained model

- Remove variables educ and exper
- $X = X(:, [1, 4]);$
- $SSR1 = 132.6105$

- Compare sum of squares of deviations of 2 models (constrained and unconstrained)



$$F = ((SSR1 - SSR0)/SSR0)(n - k)/2 = 80.1478 > F_{2,n-4}$$

- We reject  $H_0$ .
- $fdis\_prb(F, 2, n - k)$

- We test  $\beta(educ) = 0$
- Using the restricted model  $SSR2 = 132.09$

- Compare sum of squares of deviations of two models (restricted and unrestricted)
- 

$$F = ((SSR2 - SSR0)/SSR0)(n - k)/1 \Rightarrow F_{1,522}$$

- We reject  $H_0$
- Remark  $F_{1,522} = t_{522}^2$

## Binary observations

Qualitative factors often come in the form of binary information: a person is female or male; a person does or does not own a personal computer; a firm offers a certain kind of employee pension plan or it does not; a state administers capital punishment or it does not. In all of these examples, the relevant information can be captured by defining a binary variable or a zero-one variable. In econometrics, binary variables are most commonly called dummy variables, although this name is not especially descriptive.

## Example

Consider the following simple model of hourly wage determination:

$$wage = \beta_0 + \delta_0 female + \beta_1 educ + u.$$

We use  $\delta_0$  as the parameter on female in order to highlight the interpretation of the parameters multiplying dummy variables; later, we will use whatever notation is most convenient. In above model, only two observed factors affect wage: gender and education. Since  $female = 1$  when the person is female, and  $female = 0$  when the person is male, the parameter  $\delta_0$  has the following interpretation:  $\delta_0$  is the difference in hourly wage between females and males, given the same amount of education (and the same error term  $u$ ). Thus, the coefficient  $\delta_0$  determines whether there is discrimination against women: if  $\delta_0 < 0$ , then, for the same level of other factors, women earn less than men on average.

- load wage1.raw
- Carry out regression:

$$\text{Log}(\text{wage}) | \text{female}, \text{married}, \text{educ}, \text{exper}, \text{tenure}$$

- Test  $\beta_{\text{female}} = 0$



- $y = \text{wage1}(:, 1);$
- $[n, k] = \text{size}(\text{wage1});$
- $X = [\text{ones}(n, 1), \text{wage1}(:, [2, 3, 4, 6, 7])];$
- $[n, k] = \text{size}(X)$
- $y = \log(y);$
- $\text{beta} = \text{inv}(X' * X) * X' * y$
- $u = y - X * \text{beta};$
- $\text{sig2} = u' * u / (n - k)$
- $\text{std} = \text{sqrt}(\text{diag}(\text{sig2} * \text{inv}(X' * X)))$
- $t = \text{beta} ./ \text{std}$

# Results

- $\beta =$

0.4901

0.0839

0.0031

0.0169

-0.2855

0.1257

- std =

0.1011

0.0070

0.0017

0.0030

0.0373

- Sometimes it is natural for the partial effect, elasticity, or semi-elasticity of the dependent variable with respect to an explanatory variable to depend on the magnitude of yet another explanatory variable.
- Consider

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_j x_i x_j + \cdots + \beta_i x_i + \cdots + \beta_k x_k$$

- the partial effect of  $x_i$  on  $y$  is

$$\frac{\Delta(y)}{\Delta(x_i)} = \beta_i + \beta_j x_j$$

- If  $\beta_j > 0$ , then, there is an interaction effect between  $x_i$  and  $x_j$ .
- We want to test if  $\beta_j = 0$ .

- Create :

$$marrmale = (1 - female) * married$$

$$marrfem = female * married$$

$$-singfem = female * (1 - married)$$

- Carry out regression:

$$\log(wage) | marrmale, marrfem, singfem, educ, exper, tenure$$

- Test the effect of being a women

```
educ = X(:, 2);  
exper = X(:, 3);  
tenure = X(:, 4);  
female = X(:, 5);  
married = X(:, 6);  
marrmale = (1 - female) . * married;  
marrfem = female . * married;  
singfem = female . * (1 - married);
```

```
 $X = [\text{ones}(n, 1) | \text{educ}, \text{exper}, \text{tenure}, \text{marrmale}, \text{marrfem}, \text{singfem}];$   
 $[n, k] = \text{size}(X)$ 
```

Then we do the regression with the new matrix.

•  $\beta =$

0.3878

0.0835

0.0032

0.0157

0.2921

-0.1202

-0.0967

•  $std =$

0.1022

0.0069

0.0017

0.0029

0.0553

0.0579

0.0574



•  $t =$

3.7938

12.1870

1.9136

5.3747

5.2785

−2.0756

−1.6853

$$SSR0 = u' * u$$

$$SSR0 = 85.4648$$

$$X = [\text{ones}(n,1)|educ, exper, tenure, marrmale];$$

$$SSR1 = u' * u$$

$$SSR1 = 98.46$$

$$F = ((SSR1 - SSR0)/SSR0) * (n - k)/2$$

$$F = 39.57$$

$$fdist_{prob}(F, n - k, 2)$$

- Interaction between female and educ
- Calculate  $fmeduc = female * educ$
- Perform the regression:

$$\log(wage) | female, educ, fmeduc, exper, tenure$$

- Test the significance of effect of being a woman

```
femeduc = female. * educ;  
X = [ones(n,1)|educ, exper, tenure, female, femeduc];
```

•  $\beta =$

0.4647

0.0903

0.0046

0.0174

-0.2104

-0.0072

•  $std =$

0.1228

0.0087

0.0016

0.0030

0.1738

0.0135

•  $t =$

3.7851

10.3685

2.8530

5.8545

-1.2104

-0.5347

$$SSR0 = 90.0950$$

$$SSR1 = 101.4556$$

$$F = 32.7848$$



# Logarithmic transformation

- Use WAGE1.RAW
- Perform regression

$$\text{Log}(\text{wage}) | \text{educ}, \text{exper}, \text{tenure}, \text{educ} * \text{fem}, \text{exper} * \text{fem}, \text{tenure} * \text{fem}, \text{fem}$$

- Test the coefficients associated with  $\text{women} = 0$

# Logarithmic transformation

- Use WAGE1.RAW
- $\text{Log}(\text{wage}) | \text{educ}, \text{exper}, \text{tenure}$
- Test the stability of coefficients men/women
- unrestricted model:  $SSR0 = 88.6$
- restricted model:  $SSR1 = 101.3$

$$F = (101.3 - 88.6)/88.6/(4/(526 - 8)) = 18.8$$

$$F_{dis_{prb}}(18.8, 4, 516) < .01$$

## 2 separate regressions

- Create sample consisting with men and women separately (see: TP Matlab 1)
- Do two regressions and add sum of squares of deviations : sum of squares of deviations of unrestricted model.

$$\text{Log}(\text{wage}) | \text{educ}, \text{exper}, \text{tenure}$$

for men, SSR01

$$\text{Log}(\text{wage}) | \text{educ}, \text{exper}, \text{tenure},$$

for women, SSR02

- $SSR0 = SSR01 + SSR02 = 88.6$