



Short communication

Quantification of uncertainties associated with space-time estimates of short-term soil CO₂ emissions in a sugar cane areaD.D.B. Teixeira^{a,*}, E.S. Bicalho^{a,b}, C.E.P. Cerri^b, A.R. Panosso^a, G.T. Pereira^a, N. La Scala^a^a Universidade Estadual Paulista, Departamento de Ciências Exatas – FCAV/UNESP, Via de Acesso Prof. Paulo Donato Castellane s/n, 14884-900 Jaboticabal, SP, Brazil^b Universidade de São Paulo, Departamento de Ciência do Solo – ESALQ/USP, Av. Pádua Dias, 11, CP. 9, 13418-900 Piracicaba, SP, Brazil

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ABSTRACT

The characterization of soil CO₂ emissions (FCO₂) is important for the study of the global carbon cycle. This phenomenon presents great variability in space and time, a characteristic that makes attempts at modeling and forecasting FCO₂ challenging. Although spatial estimates have been performed in several studies, the association of these estimates with the uncertainties inherent in the estimation procedures is not considered. This study aimed to evaluate the local, spatial, local-temporal and spatial-temporal uncertainties of short-term FCO₂ after harvest period in a sugar cane area. The FCO₂ was featured in a sampling grid of 60 m × 60 m containing 127 points with minimum separation distances from 0.5 to 10 m between points. The FCO₂ was evaluated 7 times within a total period of 10 days. The variability of FCO₂ was described by descriptive statistics and variogram modeling. To calculate the uncertainties, 300 realizations made by sequential Gaussian simulation were considered. Local uncertainties were evaluated using the probability values exceeding certain critical thresholds, while the spatial uncertainties considering the probability of regions with high probability values together exceed the adopted limits. Using the daily uncertainties, the local-spatial and spatial-temporal uncertainty (Ftemp) was obtained. The daily and mean emissions showed a variability structure that was described by spherical and Gaussian models. The differences between the daily maps were related to variations in the magnitude of FCO₂, covering mean values ranging from $1.28 \pm 0.11 \mu\text{mol m}^{-2} \text{s}^{-1}$ (F197) to $1.82 \pm 0.07 \mu\text{mol m}^{-2} \text{s}^{-1}$ (F195). The Ftemp showed low spatial uncertainty coupled with high local uncertainty estimates. The average emission showed great spatial uncertainty of the simulated values. The evaluation of uncertainties associated with the knowledge of temporal and spatial variability is an important tool for understanding many phenomena over time, such as the quantification of greenhouse gases or the identification of areas with high crop productivity.

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1. Introduction

Human actions related to land management can interfere with the flow of carbon between terrestrial ecosystems and the atmosphere (Morell et al., 2010; Sugihara et al., 2012). It is estimated that in 2005, agriculture was responsible for an emission of 5.1–6.1 Gt of CO₂ eq year⁻¹, corresponding to approximately 10% to 12% of the total anthropogenic emissions of greenhouse gases (IPCC, 2007). Soil CO₂ emissions (FCO₂) combined with erosion processes are the main sources of soil carbon loss (Mchunu and Chaplot, 2012), although erosion may also contribute to the significant increase in soil carbon stocks (Boix-Fayos et al., 2009).

The FCO₂, as well as other soil properties, presents a complex degree of variability in space (Panosso et al., 2009; Herbst et al., 2010; Prolingheuer et al., 2010; Teixeira et al., 2011, 2012; Jurasinski et al., 2012) and time (Herbst et al., 2009; Iqbal et al., 2010; Luan et al., 2012) showing a great variability even in short periods (La Scala et al., 2000; Teixeira et al., 2011). Such variability is a combination of several biological, chemical and physical interactions acting at multiple scales (Parkin, 1993). Tools such as geostatistics have been used for providing the capture, representation and interpretation of the spatial patterns of this variable.

During the process of interpolation of unsampled sites, it is possible to identify some imprecision sources of the estimates, which are evaluated by quantifying the errors and uncertainties of non-sampled locations. Errors are defined as the difference between the calculated value and the real value of a property at the same location in space (Isaaks and Srivastava, 1989). On the other hand, uncertainties correspond to the doubt in estimated value which cannot be compared to the real value, since this is not available. Thus, here it is considered that cross-validation and

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external-validation analysis (see Deutsch and Journel, 1998) are able to provide errors while the cumulative distribution functions (cdf) provide the uncertainty of the studied phenomena.

Several interpolation techniques can be used to generate maps of estimated soil properties. Ordinary kriging (OK), indicator kriging (IK) and sequential Gaussian simulation (SGS) are the main geostatistical interpolations used. Kriging estimators provide the best estimate of the local variables by minimizing the variance estimate, resulting in smoothing of the spatial details. The use of the kriging variance as a local error measure is inappropriate when the data configuration together with the uncertainty in the estimates are considered (Zhao et al., 2009) because only the variogram and the sample configuration are considered in its calculations. As with other analyses, except if a Gaussian model is assumed for errors, kriging provides only an incomplete measure of location accuracy and no consideration of the joint accuracy can be made when several locations are considered together (Deutsch and Journel, 1998). Techniques such as IK only evaluate the uncertainty location and are unable to detect the spatial uncertainties associated with estimates. Such uncertainties are based only on the spatial variation of the known values, without considering the variation in estimates of non-sampled locations (Delbari et al., 2009). Goovaerts (2001) stated that the cdf obtained by IK only provide a measure of local uncertainty and that no quantitative measure of the spatial uncertainty is obtained.

These simulation methods are used in situations where reproducing spatial variability, honoring the inferred heterogeneity in the data, and assessing the impact of uncertainties are required (Goovaerts, 1999). The simulations are stochastic processes of design alternatives that are equally probable joint realizations of a component of a random variable from the model of a random

function (Deutsch and Journel, 1998). Simulations have recently started to be adopted for the evaluation of FCO₂ in agricultural soils (Herbst et al., 2010; Teixeira et al., 2011, 2012), with the aim of improving the quality of the interpolations obtained. In this context, the objective of this paper is to characterize the local, spatial and spatio-temporal uncertainties of FCO₂ estimates at the mesoscopic scale (area of 0.36 ha) over 10 days of evaluation (short period) through many equiprobable realizations of the SGS.

2. Materials and methods

The study was conducted in Guariba in São Paulo State (21°21'S, 48°11'W). According to the Thornthwaite classification, the local climate can be defined as B1rB'4a'-type mesothermal humid, with little water stress, evapotranspiration and less than 48% annual evapotranspiration in the summer. The soil of the area was classified as a high clay Oxisol.

The area has been cultivated with sugar cane (*Saccharum* spp. var. SP86-155) for 8 years under a management system using mechanical harvesting. At the time of the study, the soil was devoid of vegetation and was covered with large amounts of crop residue (12 t ha⁻¹) from mechanical harvesting performed a few days before the start of the experiment. On July 13th, 2010, a regular grid of 60 m × 60 m was installed containing 127 points spaced at minimum distances of 0.5–10 m from the installation of PVC collars, which were used to assess the FCO₂.

The FCO₂ was assessed using three portable systems (LI-COR 8100). The soil chambers are coupled to a system that quantifies the internal concentration of CO₂ through optical absorption spectroscopy in the infrared spectral region. Before starting the experiment, the machines were tested and calibrated. The

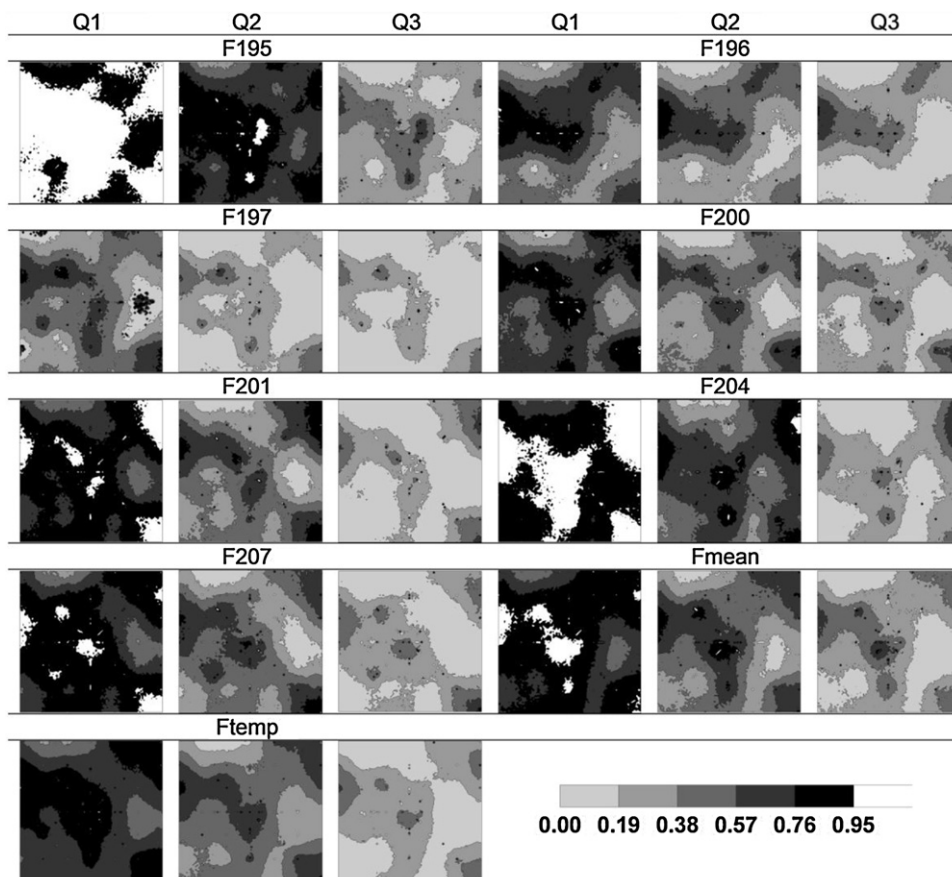


Fig. 1. Probability maps of the values estimated exceeding the values of the quartiles (Q1, Q2, Q3). F... = days of the year that soil CO₂ emission was evaluated.

evaluations were conducted in the morning (8:00–9:30 am) over 7 days on Julian days 195 (F195), 196 (F196), 197 (F197), 200 (F200), 201 (F201), 204 (F204) and 207 (F207) in 2010. After the readings were calculated, the average emission (F_{mean}) considering the values observed in the different days was evaluated.

The short period (10 days) stands out as the period that the FCO_2 is more related to the soil organic carbon losses (Cerri et al., 2011), once the soil is almost bare having few vegetation (Panosso et al., 2009). This post-harvest period also focus the impact of the main agricultural operations due to the harvest and other management practices which contribute largely to the increased rates of soil respiration. As sugar cane is a crop with rapid vegetative growth, the quantification of soil respiration coming mostly from the soil over longer periods is difficult. Moreover, due to the intense root growth, autotrophic respiration quickly integrates the captured emissions from the soil. Thus studies of temporal variations of FCO_2 , obtained directly, are commonly evaluated for short periods after harvest or tillage operations (La Scala et al., 2000; Brito et al., 2010; Panosso et al., 2011).

The FCO_2 variability was described initially by using descriptive statistics and then using the geostatistical modeling of the experimental variogram based on the principles of the intrinsic hypothesis (Deutsch and Journel, 1998) (Eq. (1)).

$$\hat{\gamma}(h) = \frac{1}{2N(h)} \sum_{i=1}^{N(h)} [z(x_i) - z(x_i + h)]^2 \quad (1)$$

where $\hat{\gamma}(h)$ is the experimental semivariance a separation distance h , $z(x_i)$ is the property value of FCO_2 at the i th point, and $N(h)$ is the number of pairs of points separated by distance h . The choice of the variogram model is adjusted based on the residual sum of squared (RSS) and the coefficient of determination (R^2).

After modeling, 300 realizations of SGS were made in accordance with the SGSIM-routine of the Geostatistical Software Library (Deutsch and Journel, 1998). The SGS allows the quantification of local and spatial uncertainties related to simulated values in unsampled locations. Local daily uncertainty was previously evaluated by calculating the probability of the point values exceeding a certain critical limit (Delbari et al., 2009; Zhao et al., 2009). The limits used in this study were 0.79, 1.31 and $1.95 \mu\text{mol m}^{-2} \text{s}^{-1}$, representing the 1st (Q1), 2nd (Q2) and 3rd (Q3) quartiles of the variable F_{mean} , respectively. This procedure quantifies the number of simulated maps in which the variable occasionally exceeds the limit indicated (Eq. (2)).

$$\text{Prob}[z(x_0) \geq z_c] = \frac{n(x_0)}{L} \quad (2)$$

where $z(x_0)$ is the value of the variable z in the location x_0 , z_c is the critical limit adopted, $n(x_0)$ is the number of realizations to which the variable displays a value above z_c and L is the total number of realizations generated. In this study, $L = 300$ for average and daily emissions.

To supplement the location uncertainty, the spatial uncertainty was evaluated based on the calculation of the joint probability (P_j) (Juang et al., 2004). The P_j helps determine the reliability of the areas outlined with a high probability that the simulated values exceed the critical limits ($\text{Prob}[z(x_0) \geq z_c]$). To calculate the P_j , areas that had a probability greater than or equal to 95%, calculated based on Eq. (2), were considered. The procedure consists of counting the number of realizations in which all locations (with probability greater than or equal to 95%) simultaneously appear above the critical value.

$$\text{Prob}[z(x_{01}) \geq z_c, z(x_{02}) \geq z_c, \dots, z(x_{0N}) \geq z_c] = \frac{n(x_{01}, x_{02}, \dots, x_{0N})}{L} \quad (3)$$

where $n(x_{01}, x_{02}, \dots, x_{0N})$ is the number of realizations in which the simulated values at the locations $x_{01}, x_{02}, \dots, x_{0N}$ have joint values above the critical limit adopted (z_c), and L is the number of realizations made in the simulation ($L = 300$). Values closer to 1 provide greater reliability and reduce the spatial uncertainty of the area.

The calculation of the local time uncertainty (F_{temp}) and space-time were performed considering all maps produced daily. Thus, in Eqs. (2) and (3), $L = 2100$ was used for F_{temp} , accounting for all the realizations produced by the SGS on the 07 days evaluated.

3. Results and discussion

The temporal variability of FCO_2 throughout the evaluated day showed small variations, differing significantly ($p < 0.05$) between the emissions related to F195 ($1.82 \pm 0.07 \mu\text{mol m}^{-2} \text{s}^{-1}$) and F197 ($1.28 \pm 0.11 \mu\text{mol m}^{-2} \text{s}^{-1}$) by the Student's t test (Table 1). This result may be related to homogeneous climatic conditions, with few variations in temperature and humidity over the short study period. The FCO_2 daily values were similar to those observed by Brito et al. (2010) in the same soil type and vegetation. Panosso et al. (2011), also in areas of sugar cane, observed average emission values of 2.07 ± 0.06 and $2.74 \pm 0.14 \mu\text{mol m}^{-2} \text{s}^{-1}$ under green and slash-and-burn managements, respectively. The high CV values ($>40\%$) indicate a high variability of the data, demonstrating the need for a more rigorous statistical analysis that enables the evaluation and interpretation of FCO_2 in different compartments of the area. These values are similar to those found by other authors in several soil and cultivation types (Herbst et al., 2009; Brito et al., 2009).

Table 1

Descriptive statistics and the variogram parameters fitted to the data of soil CO_2 emission ($\mu\text{mol m}^{-2} \text{s}^{-1}$).

	Mean	CV ^a	Model	C_0^b	$C_0 + C_1^c$	A^d	$C_0/C_0 + C_1^e$	R^2	RSS ^f
F195	$1.82 \pm 0.07a$	41.70	Sph. ^g	0.62	1.10	34.23	0.56	0.95	1.07E–02
F196	$1.52 \pm 0.12bc$	85.72	Gaus. ^h	0.58	1.19	33.43	0.48	0.94	2.88E–02
F197	$1.28 \pm 0.11c$	93.04	Sph.	0.49	1.07	26.16	0.46	0.69	1.31E–01
F200	$1.53 \pm 0.11bc$	74.29	Sph.	0.53	1.08	29.45	0.49	0.94	3.13E–02
F201	$1.49 \pm 0.07bc$	49.16	Sph.	0.51	1.15	37.34	0.44	0.93	3.52E–02
F204	$1.63 \pm 0.07b$	48.05	Sph.	0.51	1.27	46.32	0.40	0.80	1.74E–01
F207	$1.51 \pm 0.07bc$	53.54	Gaus.	0.52	1.02	26.71	0.51	0.94	1.53E–02
Fmean	1.57 ± 0.07	50.02	Gaus.	0.68	1.19	25.39	0.57	0.90	3.70E–02

$N = 127$; means followed by the same letters on a column do not differ (Student's t -test; $p \leq 0.05$).

^a Coefficient of variation.

^b Nugget effect.

^c Sill.

^d Range (m).

^e Degree of spatial dependence.

^f Residual sum of square.

^g Spherical model.

^h Gaussian model.

The Gaussian and spherical models were adjusted to the FCO₂ experimental variograms (Table 1). The Gaussian model describes phenomena with great continuity at small and large scales, while the spherical model, by presenting a well-defined sill, describes properties with abrupt changes along the ground (Deutsch and Journel, 1998). Brito et al. (2010) fitted spherical and exponential models to the FCO₂ data in different topographic positions in soils under cultivation of sugar cane. Teixeira et al. (2011) used spherical models to describe the diurnal spatial variability of FCO₂ in soils devoid of vegetation.

The use of normalized values for the construction of the variogram used in the process of SGS allows comparison of the effect nugget (C_0) between the daily variogram. The C_0 represents the combination of errors due to variability present on a smaller scale than that evaluated (in this case, less than 0.50 m) and those arising from laboratory determination methods. The Fmedio had the highest C_0 value (0.68) among the models, which was most likely due to its calculations based on average daily values, thus affecting the variability that is inherent in these data. Among the daily values, the higher C_0 was found for the largest daily emission (F195), and the lower C_0 was related to lower emissions (F197). This result indicates a larger error in the estimates of non-sampled locations on the days that FCO₂ has high emission values.

The range values (A) indicates the distance limit of spatial dependence between the samples; in other words, above this limit they are considered independently, showing no spatial relationship to each other. The range values varied from 25.39 (Fmean) to 46.32 m (F204), similar to values found by other studies (Brito et al., 2010; Teixeira et al., 2011). Prolingheuer et al. (2010) studied the spatial and temporal variability of soil respiration in areas under the cultivation of winter wheat, and they identified average range values of 18.0 m. Herbst et al. (2009), in a fine sampling grid of 14 m × 14 m, observed an average range of 2.7 m for FCO₂. These results are lower than those found in this study due to the extremely detailed level adopted by the authors. All models showed a moderate degree of spatial dependence, characterized by the relation $0.25 < C_0/C_0 + C_1 < 0.75$ (Cambardella et al., 1994).

Although the use of different realizations of the SGS allows a visual evaluation of the spatial uncertainties associated with estimates of the property values in question (Grunwald et al., 2007), quantification is possible only through the analysis of P_j (Table 2). Zhao et al. (2009), using the calculation of P_j , evaluated spatial area uncertainty with nutrient deficiency in soil. They observed small spatial uncertainties for organic matter and large uncertainties for the potassium and phosphorus content in the soil. Delbari et al. (2009) evaluated the reliability of areas susceptible to erosion using the same method.

Table 2
Evaluation of the spatial uncertainty of the areas that have values above the limits of 0.79 (Q1), 1.31 (Q2) and 1.95 (Q3) $\mu\text{mol m}^{-2} \text{s}^{-1}$.

	Q1		Q2		Q3	
	NSL ^a	Pj ^b	NSL	Pj	NSL	Pj
F195	5780	0.00	227	0.00	34	100.00
F196	74	57.33	42	100.00	27	100.00
F197	41	100.00	22	100.00	17	100.00
F200	75	73.33	42	100.00	30	100.00
F201	990	0.00	51	100.00	23	100.00
F204	3623	0.00	131	31.33	25	100.00
F207	738	0.00	48	100.00	26	100.00
Fmean	915	0.00	60	67.33	36	88.67
Ftemp	32	79.29	14	100.00	5	100.00

^a Number of simulated locations.

^b Joint probability (%).

Fig. 1 shows probability maps of emission values higher than those observed for the Fmean quartiles. The quartiles of the observed data were used in the absence of reference values for FCO₂ because few studies have promoted their characterization under the same type of soil and cultivation. The maps referring to the Q1 values generally have large areas with a high probability of exceeding this value. However, using Q3 as the limit, the maps do not exhibit continuous areas of high probability, being limited only to reproducing the values measured in the field (greater than Q3, with high probability) because no estimate is generated at these sites.

The difference between the maps of probability is related to daily variations in the magnitude of the emission, where days with higher average emissions (Table 1) have larger areas with a high probability of exceeding the limit used (Fig. 1). Overall, the largest emissions were located in the center and left center of the map, thus providing the greatest probability of having values above the critical limit adopted. The maps of the local temporal uncertainties (Ftemp) do not show continuous areas of high probability, a fact indicative of the great variation and low temporal persistence of FCO₂ even in the short study period (10-day), when no significant changes in the weather conditions, especially temperature and soil moisture, were observed.

Quantifying the spatial uncertainty by P_j shows that as the limit values increase, the area with the probability of showing values above these decreases, thus decreasing the number of simulated locations (NSL) (Table 2). With the exception of Fmean, in any variable values were simulated with a high probability of being larger than Q3, so only the punctual observations were considered in the calculations, resulting in P_j values equal to 100%. F195 and F204 had the highest spatial uncertainties, both with a 0.00% probability of exceeding Q1 and 0.00 and 31.33% probability of exceeding Q2, respectively. This greater uncertainty was due to a larger area of values above the simulated Q1 and Q2 limits (NSL value). The Fmean shows great spatial uncertainty even with small areas considered, while the Ftemp has low spatial uncertainty related to the high local uncertainty of the values simulated. Teixeira et al. (2011), evaluating the daily changes in FCO₂, observed great persistence in the pattern of spatial variability in different periods, indicating that areas with higher emissions during the morning had the highest emissions in the afternoon.

4. Conclusion

FCO₂ showed low spatial-temporal uncertainty coupled with high local-temporal uncertainty estimates. These procedures of uncertainty evaluation allow the identification and characterization of areas susceptible to high emission values. These areas are persistent or not on time or, for example, are influenced by the management of agricultural areas, relief or other soil properties. Thus some measures, such as sample design using denser sampling in regions with the greatest temporal uncertainty, can be adopted to decrease the uncertainty and increase the accuracy in the estimates. However, further studies are needed to establish reference values for the FCO₂ and other important greenhouse gases in agricultural soils. The combination of local and spatial uncertainty in different periods for the temporal assessment of these phenomena, a technique proposed in this paper, represents an important tool not yet discussed in the literature. These temporal data can be used in various areas, such as the identification of areas with high yields of crops, areas susceptible to intense rainfall and temperature, and agricultural pest management. This information will also be useful for the determination of specific types of management in each tested environment.

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