

MTL 106 (Introduction to Probability Theory and Stochastic Processes)
Assignment 2 Report

Name: Arpit Saxena

Entry Number: 2018MT10742

1. Basic Probability
2. Random Variable/Function of a Random Variable
3. Stochastic Processes
4. Stochastic Processes
5. DTMC

(a) The transition matrix P of a Discrete Time Markov Chain is given as:

$$P = \begin{pmatrix} 0.6 & 0.2 & 0.1 & 0 & 0 & 0.1 & 0 & 0 \\ 0.1 & 0.1 & 0.2 & 0.2 & 0 & 0 & 0.4 & 0 \\ 0.2 & 0.3 & 0 & 0 & 0 & 0.1 & 0.2 & 0.2 \\ 0.5 & 0.2 & 0.1 & 0.1 & 0 & 0.1 & 0 & 0 \\ 0.1 & 0.1 & 0.1 & 0.1 & 0.1 & 0.1 & 0.1 & 0.3 \\ 0.3 & 0 & 0.3 & 0.2 & 0.1 & 0 & 0 & 0.1 \\ 0 & 0 & 0 & 0 & 0.5 & 0.4 & 0 & 0.1 \\ 0 & 0.1 & 0.1 & 0.1 & 0 & 0.2 & 0.2 & 0.3 \end{pmatrix}$$

Without calculating the stationary probabilities, find out whether stationary distribution exists for this system or not.

- (b) Consider a DTMC of n states. We know that the transition matrix for the DTMC is diagonalisable and absolute value of all of its eigenvalues is less than 1. Will the limiting distribution of this system exist?

Use the result to find the existence of limiting distribution of a DTMC with the following transition matrix:

$$\begin{pmatrix} 0 & \frac{2}{3} & \frac{1}{3} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix}$$

Answer

- (a) Let π be the stationary probabilities for the system. i.e. $\pi = (\pi_0, \pi_1, \dots, \pi_{n-1})$

Also, we have $\pi P = \pi$ since the distribution is taken as stationary

$$\begin{aligned} \pi P &= \pi \\ \implies \pi(P - I) &= 0 \end{aligned}$$

Thus we have a homogenous system of linear equations. Since we are after a non-trivial solution (as row sum of $\pi = 1$), we have the condition

$$\det(P - I) = 0$$

Note that we have an overconstrained system of linear equations, given by $\pi P = \pi$, $\sum_{i=0}^{n-1} \pi_i = 1$ along with the condition that $\pi_i \geq 0 \forall i = 0, 1, \dots, n-1$

Thus, the condition that $\det(P - I) = 0$ is necessary for stationary distribution to exist but not sufficient.

For the given transition matrix:

$$\det(P - I) = \begin{vmatrix} -0.4 & 0.2 & 0.1 & 0 & 0 & 0.1 & 0 & 0 \\ 0.1 & -0.9 & 0.2 & 0.2 & 0 & 0 & 0.4 & 0 \\ 0.2 & 0.3 & -1 & 0 & 0 & 0.1 & 0.2 & 0.2 \\ 0.5 & 0.2 & 0.1 & -0.9 & 0 & 0.1 & 0 & 0 \\ 0.1 & 0.1 & 0.1 & 0.1 & -0.9 & 0.1 & 0.1 & 0.3 \\ 0.3 & 0 & 0.3 & 0.2 & 0.1 & -1 & 0 & 0.1 \\ 0 & 0 & 0 & 0 & 0.5 & 0.4 & -1 & 0.1 \\ 0 & 0.1 & 0.1 & 0.1 & 0 & 0.2 & 0.2 & -0.7 \end{vmatrix} \neq 0$$

This tells us that the given DTMC will not have a stationary distribution.

- (b) Let P denote the transition probability matrix for the DTMC. Then we know that P is a $n \times n$ matrix. Given that it is diagonalisable, \exists invertible $A_{n \times n}$ such that $P = A\Lambda A^{-1}$ where

$$\begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & \dots & 0 & \lambda_n \end{pmatrix}$$

where $\lambda_1, \dots, \lambda_n$ are the eigenvalues of the matrix P . We are given that $|\lambda_i| < 1, i = 1, 2, \dots, n$

$$\begin{aligned} P &= A\Lambda A^{-1} \\ \implies P^2 &= A\Lambda A^{-1} A\Lambda A^{-1} \\ &= A\Lambda^2 A^{-1} \end{aligned}$$

Continuing like this, we get

$$P^m = A\Lambda^m A^{-1}$$

Since Λ is a diagonal matrix, Λ^m can be easily calculated and is given by:

$$\Lambda^m = \begin{pmatrix} \lambda_1^m & 0 & \dots & 0 \\ 0 & \lambda_2^m & \dots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & \dots & 0 & \lambda_n^m \end{pmatrix}$$

$$\Rightarrow \lim_{m \rightarrow \infty} \Lambda^m = \begin{pmatrix} \lim_{m \rightarrow \infty} \lambda_1^m & 0 & \dots & 0 \\ 0 & \lim_{m \rightarrow \infty} \lambda_2^m & \dots & 0 \\ \vdots & & \ddots & \vdots \\ 0 & \dots & 0 & \lim_{m \rightarrow \infty} \lambda_n^m \end{pmatrix}$$

$|\lambda_i| \leq 1 \Rightarrow \lim_{m \rightarrow \infty} \lambda_i^m = 0$ thus we have

$$\lim_{m \rightarrow \infty} \Lambda^m = 0$$

$$\Rightarrow \lim_{m \rightarrow \infty} P^m = \lim_{m \rightarrow \infty} A \Lambda^m A^{-1} = 0$$

Thus, the m -step transition probability matrix $P^{(m)} = P^m$ goes to 0 as $m \rightarrow \infty$ which implies that the limiting distribution does not exist.

Now, given the transition matrix

$$\begin{pmatrix} 0 & \frac{2}{3} & \frac{1}{3} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix}$$

The eigenvalues of the given matrix are $\pm \frac{1}{2}, 1$. Since it has 3 distinct eigenvalues, it must be diagonalisable. Two diagonal entries of the eigenvalue matrix will go to zero, and the other would be 1. Thus $\lim_{m \rightarrow \infty} P^m$ exists and thus the limiting distribution for the given DTMC exists.

6. DTMC

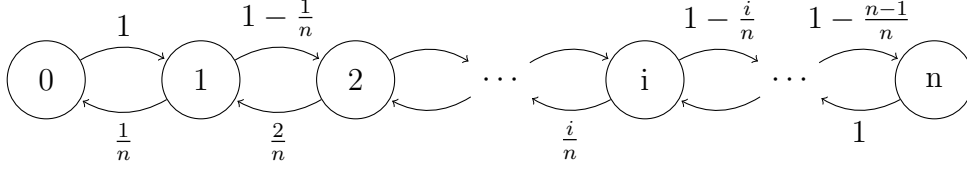
7. CTMC

Consider a modified version of Conway's game of life. In this simplified version, we model using a birth death process where the maximum number of entities allowed in the system is n . When the number of entities in the system is i , the birth rate is given as $1 - \frac{i}{n}$ and the death rate is given as $\frac{i}{n}$. These model the qualities that birth rate decreases with increasing population and death rate increases due to overpopulation.

Give the diagrammatic representation of the process. Find the steady state distribution of the system, and the expected number of entities in steady state.

Answer

The diagrammatic representation is given below:



Let X_t denote the number of entities in the system at time t . Let π denote the stationary probabilities of the system, i.e., $\pi = (\pi_0, \pi_1, \dots, \pi_n)$ where $\pi_i = \lim_{t \rightarrow \infty} P(X_t = i)$

Using the general formula for a birth death process, we have

$$\begin{aligned}
 \pi_i &= \frac{\lambda_0 \lambda_1 \dots \lambda_{i-1}}{\mu_1 \mu_2 \dots \mu_i} \pi_0 \\
 &= \frac{1(1 - \frac{1}{n}) \dots (1 - \frac{i-1}{n})}{\frac{1}{n} \frac{2}{n} \dots \frac{i}{n}} \pi_0 \\
 &= \frac{n(n-1) \dots (n-i+1)}{i!} \pi_0 \\
 &= \frac{n!}{i!(n-i)!} \pi_0 \\
 \implies \pi_i &= \binom{n}{i} \pi_0, \quad i = 1, 2, \dots, n
 \end{aligned}$$

Now we use the normalising condition:

$$\begin{aligned}
 \sum_{i=0}^n \pi_i &= 1 \\
 \implies \pi_0 + \sum_{i=1}^n \pi_i &= 1 \\
 \implies \pi_0 + \sum_{i=1}^n \binom{n}{i} \pi_0 &= 1 \\
 \implies \sum_{i=0}^n \binom{n}{i} \pi_0 &= 1 \\
 \implies 2^n \pi_0 &= 1 \\
 \implies \pi_0 &= 2^{-n} \\
 \implies \pi_i &= \binom{n}{i} 2^{-n}, \quad i = 0, 1, \dots, n
 \end{aligned}$$

Let X be the random variable denoting the number of entities in the system in steady state. Then $P(X = i) = \pi_i$, $i = 0, 1, \dots, n$

$$\begin{aligned}
E(X) &= \sum_{i=0}^n i \pi_i \\
&= \sum_{i=0}^n i \binom{n}{i} 2^{-n} \\
&= 2^{-n} \sum_{i=0}^n i \binom{n}{i}
\end{aligned}$$

To calculate $\sum_{i=0}^n i \binom{n}{i}$ we use the binomial formula as:

$$(1+x)^n = \sum_{i=0}^n \binom{n}{i} x^i$$

Differentiating wrt x on both sides, we get

$$n(1+x)^{n-1} = \sum_{i=1}^n i \binom{n}{i} x^{i-1}$$

Now putting $x = 1$ implies

$$\begin{aligned}
n2^{n-1} &= \sum_{i=1}^n i \binom{n}{i} \\
n2^{n-1} &= \sum_{i=0}^n i \binom{n}{i}
\end{aligned}$$

Substituting in the above equation, we get

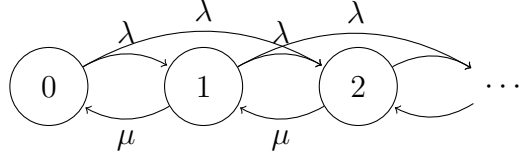
$$E(X) = 2^{-n} \cdot n2^{n-1} = \frac{n}{2}$$

8. CTMC

Consider a Markov process where states are $0, 1, 2, \dots$ and any state i can move to state $i+1$ and $i+2$ both with rate λ . Also, a state i can move to $i-1$ with rate μ . Give a diagrammatic representation of the process. Find the expected state in which the process would be in steady state

Answer

The diagrammatic representation of the process can be shown as:



At steady state, assume the probability vector to be π , then using the Kolmogorov forward equations and setting rate of change of P to be 0, we get

$$\begin{aligned} 0 &= -2\lambda\pi_0 + \mu\pi_1 \\ 0 &= \lambda\pi_0 + \mu\pi_2 - (2\lambda + \mu)\pi_1 \\ 0 &= \lambda\pi_{i-2} + \lambda\pi_{i-1} + \mu\pi_{i+1} - (2\lambda + \mu)\pi_i \end{aligned}$$

Let us define $P(x) = \sum_{n=0}^{\infty} \pi_n x^n$. Then, we have

$$\begin{aligned} P(x) &= \pi_0 + \pi_1 x + \pi_2 x^2 + \cdots + \pi_{i-2} x^{i-2} + \cdots \\ \lambda x^3 P(x) &= \lambda\pi_0 x^3 + \lambda\pi_1 x^4 + \lambda\pi_2 x^5 + \cdots + \lambda\pi_{i-2} x^{i+1} + \cdots \\ \lambda x^2 P(x) &= \lambda\pi_0 x^2 + \lambda\pi_1 x^3 + \lambda\pi_2 x^4 + \cdots + \lambda\pi_{i-1} x^{i+1} + \cdots \\ -(2\lambda + \mu)xP(x) &= -(2\lambda + \mu)\pi_0 x - (2\lambda + \mu)\pi_1 x^2 - \cdots - (2\lambda + \mu)\pi_i x^{i+1} + \cdots \\ \mu P(x) &= \mu\pi_0 + \mu\pi_1 x + \mu\pi_2 x^2 + \cdots + \mu\pi_{i+1} x^{i+1} + \cdots \end{aligned}$$

Adding the last 4 equations and using the equations obtained earlier we get

$$\begin{aligned} \{\lambda x^3 + \lambda x^2 - (2\lambda + \mu)x + \mu\}P(x) &= \mu\pi_0 - \mu\pi_0 x \\ \implies P(x) &= \frac{\mu\pi_0(1-x)}{\lambda x^3 + \lambda x^2 - (2\lambda + \mu)x + \mu} \\ \implies P(x) &= \frac{\mu\pi_0(1-x)}{(x-1)(\lambda x^2 + 2\lambda x - \mu)} \\ \implies P(x) &= \frac{-\mu\pi_0}{\lambda x^2 + 2\lambda x - \mu} \end{aligned}$$

Using the definition of $P(x)$, we get $P(1) = \sum_{n=0}^{\infty} \pi_n = 1$ and let X be the distribution in steady state, then $P'(1) = \sum_{n=0}^{\infty} n\pi_n = E(X)$

Using these conditions, we have

$$\begin{aligned} P(1) &= 1 \\ \implies \frac{-\mu\pi_0}{3\lambda - \mu} &= 1 \\ \implies \pi_0 &= 1 - \frac{3\lambda}{\mu} \\ P'(x) &= \frac{\mu\pi_0}{(\lambda x^2 + 2\lambda x - \mu)^2} (2\lambda x + 2\lambda) \\ \implies P'(1) &= \frac{\mu\pi_0}{(3\lambda - \mu)^2} \cdot 4\lambda \end{aligned}$$

Substituting $\pi_0 = \frac{\mu-3\lambda}{\mu}$, we get

$$\begin{aligned} P'(1) &= \frac{\mu - 3\lambda}{(3\lambda - \mu)^2} \cdot 4\lambda \\ P'(1) &= \frac{4\lambda}{\mu - 3\lambda} \\ \implies E(X) &= \frac{4\lambda}{\mu - 3\lambda} \end{aligned}$$

Obviously, this only exists when $\mu - 3\lambda > 0 \implies \frac{\lambda}{\mu} < \frac{1}{3}$

9. Queueing Models

A company has one 16-core machine, two 8-core machines and two 4-core machines. They want to use them as servers. The inter arrival time of queries is exponentially distributed with mean 0.1ms. They estimate that the time taken by a core per query would be exponentially distributed with mean time 3, 2, 4 milliseconds for the 16-core, 8-core and 4-core machines respectively. They want to set up a simple static load balancer in front of these machines, which will schedule the queries on a core of a machine with a probability to assign the query to each core.

Determine the probabilities by which the load balancer should schedule queries on each type of core to minimise the maximum expected waiting time for a query. Note that one query would occupy the core on which it's running for the entire time it's running.

Answer

Basically, what we want to do here is to divide the incoming queries into 3 different types of cores (they are different since they have different service time distributions) We can tabulate the information as follows:

Number of machines	Number of cores	Total number of cores	Mean service time(ms)
1	16	16	3
2	8	16	2
2	4	8	4

Number the types of cores given in the above table as 1, 2, 3 and let p_1, p_2, p_3 denote the probabilities by which a query will be sent to a core of that type by the load balancer.

Given that the incoming queries form a Poisson process with parameter $\frac{1}{0.1 \text{ ms}} = 10 \text{ ms}^{-1}$ and the load balancer is decomposing this Poisson process into separate streams. Then, the queries going to cores of type 1, 2, 3 will form a Poisson process with parameters $10p_1, 10p_2, 10p_3$ respectively and $16p_1 + 16p_2 + 8p_3 = 1$ since each query will be routed to one of the given cores.

We now model each core as a M/M/1 queue. Generically, let's take the arrival process parameter as λ and the service time parameter as μ .

Let W be the waiting time for a query, then $P(W < 0) = 0$ and $P(W = 0) = \rho$ where $\rho = \frac{\lambda}{\mu}$.

For $W > 0$, there has to be atleast one person in the system. Assuming there are n people in the system, we have $W = \tilde{S}_1 + S_2 + \dots + S_n$, where \tilde{S}_1 is the time left for query that is already running, and $S_2 \dots S_n$ are the running times for the remaining queued queries.

Since service times are exponentially distributed, which has memoryless property, \tilde{S}_1 is also exponentially distributed with parameter μ

Therefore, $W/N = n \sim \text{Gamma}(n, \mu)$ since it is a sum of n independent exponentially distributed random variables with parameter μ .

For $t > 0$,

$$\begin{aligned} P(W \leq t) &= \sum_{n=1}^{\infty} P(W \leq t/N = n) P(N = n) \\ &= 1 - \rho + \sum_{n=1}^{\infty} \int_0^t \frac{\mu^n x^{n-1} e^{-\mu x}}{(n-1)!} dx (1 - \rho) \rho^n \end{aligned}$$

Now, taking summation inside the integral, we get

$$\begin{aligned} P(W \leq t) &= 1 - \rho + \int_0^t \sum_{n=1}^{\infty} \frac{(\mu x \rho)^{n-1}}{(n-1)!} e^{-\mu x} dx (1 - \rho) \mu \rho \\ &= 1 - \rho + \int_0^t \mu \rho e^{\mu x \rho} e^{-\mu x} dx (1 - \rho) \\ &= 1 - \rho + \mu \rho (1 - \rho) \frac{e^{-\mu(1-\rho)t} - 1}{\mu(\rho - 1)} \\ &= 1 - \rho + \rho(1 - e^{-(\mu-\lambda)t}) \\ &= 1 - \rho e^{-(\mu-\lambda)t} \end{aligned}$$

Therefore, the CDF of W in the steady state is given by

$$P(W \leq t) = \begin{cases} 0 & t < 0 \\ 1 - \rho & t = 0 \\ 1 - \rho e^{-(\mu-\lambda)t} & 0 < t < \infty \end{cases}$$

where $\rho = \frac{\lambda}{\mu}$ and the steady state solution is only possible when $\rho < 1 \implies \lambda < \mu$

Then the pdf of W is given by $f_W(t) = \rho(\mu - \lambda)e^{-(\mu-\lambda)t}$ when $0 < t < \infty$ and 0 otherwise.

$$\begin{aligned} E(W) &= \int_0^{\infty} t \rho(\mu - \lambda) e^{-(\mu-\lambda)t} dt \\ &= \frac{\rho}{\mu - \lambda} \int_0^{\infty} e^{-(\mu-\lambda)t} (\mu - \lambda) t d[(\mu - \lambda)t] \end{aligned}$$

Since $\mu - \lambda > 0$, $(\mu - \lambda)t \rightarrow \infty$ as $t \rightarrow \infty$

$$\begin{aligned}
E(W) &= \frac{\rho}{\mu - \lambda} \int_0^\infty t e^{-t} dt \\
&= \frac{\rho}{\mu - \lambda} \\
&= \frac{\lambda}{\mu^2 - \lambda\mu} \quad \left(\text{Since } \rho = \frac{\lambda}{\mu} \right)
\end{aligned}$$

Now substituting the values of μ as $\frac{1}{3}, \frac{1}{2}, \frac{1}{4}$ for cores 1, 2, 3 respectively and also using the query process parameters as calculated above, we get the expected waiting times for cores 1, 2, 3 respectively as:

$$\frac{90p_1}{1 - 30p_1}, \frac{40p_2}{1 - 20p_2}, \frac{160p_3}{1 - 40p_3}$$

We also need to have $\rho < 1$ for all the cores, i.e. $\frac{10p_i}{\mu_i} < 1$ for all the cores, which gives $p_1 < \frac{1}{30}, p_2 < \frac{1}{20}, p_3 < \frac{1}{40}$

So our problem reduces to the following optimisation problem:

$$\text{Minimise } \max \left\{ \frac{90p_1}{1 - 30p_1}, \frac{40p_2}{1 - 20p_2}, \frac{160p_3}{1 - 40p_3} \right\}$$

$$\text{In the domain } 16p_1 + 16p_2 + 8p_3 = 1, p_1 < \frac{1}{30}, p_2 < \frac{1}{20}, p_3 < \frac{1}{40}$$

Solving the equations with the aid of computational tools available, we find that the minimum expected waiting time is approximately 4.77 ms, when $p_1 \approx 0.020, p_2 \approx 0.035, p_3 \approx 0.013$

So, we have the probabilities by which the load balancer should send the queries to cores of type 1 as $16p_1 \approx 0.327$, cores of type 2 as $16p_2 \approx 0.564$ and cores of type 3 as $8p_3 \approx 0.109$ for minimisation of the maximum expected waiting time for each query.

10. Queueing Models

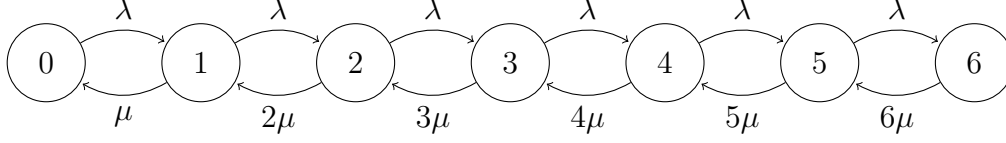
A restaurant on a very busy main road has a parking capacity of 6 cars. People whose car does not get a parking space can't go to eat in the restaurant due to the busyness of the road and the unavailability of any other parking nearby. People in a car are binomially distributed as $B(5, 0.5)$ (0 included since some people just park their car and not eat). Arrival of cars follows a Poisson process with average inter arrival time 10 minutes. Time spent in the restaurant by one cars passengers is exponentially distributed with average as 40 minutes.

Draw the state diagram of the underlying process, and derive the steady state equations. What is the expected number of people in the restaurant in steady state?

Answer

By the description, we can observe the cars follow a M/M/6/6 queue. Let λ be the parameter of the arrival process and μ be the parameter of the service time exponential distribution. Then, by the given information, $\lambda = \frac{1}{10} \text{ min}^{-1}$ and $\mu = \frac{1}{40} \text{ min}^{-1}$

The state diagram of the underlying birth death process is then given by:



Here the numbers in each state indicate the number of cars parked in the parking lot. Let $\{X_t | t \in \mathbb{R}\}$ be the underlying stochastic process where X_t denotes the number of cars parked in the parking lot at time t .

Define $\pi_n = \lim_{t \rightarrow \infty} \text{Prob}(X_t = n)$. Then using Kolmogorov, forward equations and letting $t \rightarrow \infty$, we get:

$$\begin{aligned}
 0 &= -\lambda\pi_0 + \mu\pi_1 \\
 0 &= \lambda\pi_{i-1} + (i+1)\mu\pi_{i+1} - (\lambda + i\mu)\pi_i, 1 \leq i \leq 5 \\
 0 &= \lambda\pi_5 - 5\mu\pi_6
 \end{aligned}$$

From the first equation, we get $\pi_1 = \frac{\lambda}{\mu}\pi_0$. For $i = 1$:

$$\begin{aligned}
 0 &= \lambda\pi_0 + 2\mu\pi_2 - (\lambda + \mu)\pi_1 \\
 \implies 0 &= \lambda\pi_0 + 2\mu\pi_2 - (\lambda + \mu)\frac{\lambda}{\mu}\pi_0 \\
 \implies 2\mu\pi_2 &= \frac{\lambda^2}{\mu}\pi_0 \\
 \implies \pi_2 &= \frac{\lambda^2}{2\mu^2}\pi_0
 \end{aligned}$$

By induction, we can show that

$$\pi_i = \frac{\lambda^i}{i! \mu^i} \pi_0, 1 \leq i \leq 5$$

And then using the last equation, we get:

$$\begin{aligned}
0 &= \lambda\pi_5 - 5\mu\pi_6 \\
\implies 0 &= \lambda \frac{\lambda^5}{5!\mu^5} \pi_0 - 6\mu\pi_6 \\
\implies 6\mu\pi_6 &= \frac{\lambda^6}{5!\mu^5} \pi_0 \\
\implies \pi_6 &= \frac{\lambda^6}{6!\mu^6} \pi_0
\end{aligned}$$

Define $\rho = \frac{\lambda}{\mu}$ Using the fact that the total probability must be 1, which is also called the normalising condition, we get:

$$\begin{aligned}
\sum_{n=0}^6 \frac{\rho^n}{n!} \pi_0 &= 1 \\
\implies \pi_0 &= \frac{1}{\sum_{n=0}^6 \frac{\rho^n}{n!}}
\end{aligned}$$

Let X be the random variable denoting the number of cars in the parking lot at steady state. Then,

$$\begin{aligned}
E(X) &= \sum_{n=0}^6 n \cdot \frac{\rho^n}{n!} \pi_0 \\
&= \sum_{n=0}^6 n \cdot \frac{\frac{\rho^n}{n!}}{\sum_{n=0}^6 \frac{\rho^n}{n!}} \\
&= \rho \cdot \frac{\sum_{n=1}^6 \frac{\rho^{n-1}}{(n-1)!}}{\sum_{n=0}^6 \frac{\rho^n}{n!}} \\
&= \rho \cdot \frac{\sum_{n=0}^5 \frac{\rho^n}{n!}}{\sum_{n=0}^6 \frac{\rho^n}{n!}} \\
&= \rho \left(1 - \frac{\frac{\rho^6}{6!}}{\sum_{n=0}^6 \frac{\rho^n}{n!}} \right)
\end{aligned}$$

Using $\lambda = \frac{1}{10}, \mu = \frac{1}{40}$ we get $\rho = 4$. Then we can solve the solution using a calculator and get $E(X) \approx 3.531$

Now, we have the expected number of cars in steady state, but we need to calculate the expected number of customers in the restaurant in steady state.

Let N be the random variable denoting the number of customers in the restaurant in steady state. Let N_i be the number of passengers in car i , then we have $N = \sum_{i=0}^X N_i$, where X is

the random variable denoting the number of cars in steady state. We know that N_i 's are iid random variables which are distributed as $B(5, 0.5)$ from which we get $E(N_i) = 5 \times 0.5 = 2.5$. Using Wald's equation since N_i 's are iid and also independent from X , we get

$$\begin{aligned} E(N) &= E(N_1) \times E(X) \\ &\approx 2.5 \times 3.531 \\ &= 8.828 \end{aligned}$$

Therefore, the expected number of customers in the restaurant in steady state is approximately 9.