

## Section 15

Wednesday, December 7, 2022 8:48 PM

Correction from last time - we can use Chebyshev for bounds.

(Ex) Suppose the mean on a probability exam is 75.

With just this information,

(i) Find an upperbound for probability that someone scored above 90.

(iii) Suppose you gain some information - specifically, the standard deviation was 5. Answer (i)

Chebyshev  $P[X - \mu > \epsilon] \leq \frac{\text{Var}(X)}{\epsilon^2}$ , where  $\epsilon > 0$ ,  $EX = \mu$

$P[X - \mu > \epsilon] \leq \frac{\text{Var}(X)}{\epsilon^2}$

$[X - \mu > \epsilon] \cup [X - \mu < -\epsilon]$

Before stating CLT, a warmup:

(Q) Suppose  $X_1, X_2, \dots, X_n$  are iid w/  $EX_i = \mu$

$\text{Var}(X_i) = \sigma^2$

Consider  $S_n = X_1 + \dots + X_n$ , "sample total"

$\bar{X} = \frac{S_n}{n}$ , "sample mean"

(a) Express  $ES_n$  and  $E\bar{X}$  in terms of above defined parameters.

$ES_n = E\sum_{i=1}^n X_i = n \cdot EX_i = n\mu$

$E\bar{X} = E\frac{S_n}{n} = \mu$

(b) Similarly, compute  $\text{Var}(S_n)$  and  $\text{Var}(\bar{X})$ .

$\text{Var}(\sum_{i=1}^n X_i) = \sum_{i=1}^n \text{Var}(X_i) = n\sigma^2$

$\text{Var}(\frac{\sum_{i=1}^n X_i}{n}) = \frac{1}{n^2} \sum_{i=1}^n \text{Var}(X_i) = \frac{n\sigma^2}{n^2} = \frac{\sigma^2}{n}$

(Def) (Convergence in distribution)

Let  $X_1, X_2, \dots$  be random variables. If there's random variable  $X$  s.t., for all  $x \in \mathbb{R}$ ,  $\lim_{n \rightarrow \infty} P[X_n \leq x] = P[X \leq x]$ , then  $X_n$  converges to  $X$  in distribution, or  $X_n \rightarrow_d X$ .

$F := P[X \leq x]$

(We actually require this for only  $x \in \mathbb{R}$  such that  $P[X \leq x]$  is continuous at  $x$ )

(Rm) So, if we have  $X_n \rightarrow_d X$ , we may assert that  $P[X \leq x] \approx P[X_n \leq x]$ ,

i.e. approximate  $X_n$ 's probability w/  $X$ 's.

### Central Limit Theorem

(Th) Say,  $X_1, X_2, \dots$  are iid w/  $EX_i = \mu < \infty$ ,  $\text{Var}(X_i) = \sigma^2 < \infty$ .

Then, consider  $S_n = \sum X_i$ . We have:

"normalized sample total"  $\frac{S_n - n\mu}{\sigma\sqrt{n}} \rightarrow_d N(0,1)$

$\uparrow$  (Equivalently)

"normalized sample mean"  $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \rightarrow_d N(0,1)$

distribution of standard

$$\text{Sample mean} \leftarrow \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \rightarrow N(0,1)$$

In words,  $P\left[\frac{\bar{S}_n - \mu}{\sigma/\sqrt{n}} < z\right] \approx \phi(z)$ .   
 ↑   
 distribution of standard normal

(A canonical example)

(Ques) Say, 20 of us toss a coin 10 times, and count the number of heads we get. Let  $X_i$  = # of heads that  $i^{\text{th}}$  person gets. Let  $\bar{X} = \frac{1}{20} \sum_{i=1}^{20} X_i$ , "sample mean".

(i) Compute  $E\bar{X}$  and  $\text{Var}(\bar{X})$ .

$$X_i \sim \text{Binomial}(10, 1/2), E X_i = 5, \text{Var}(X_i) = 10 \cdot 1/2 \cdot 1/2 = 10/4 = 2.5$$

$$E\bar{X} = 5$$

$$\text{Var}(\bar{X}) = \text{Var}\left(\frac{1}{20} \sum_{i=1}^{20} X_i\right) = \frac{1}{20^2} \cdot 20 \cdot \text{Var}(X_i) = \frac{1}{20} \cdot \frac{10}{4} = \frac{1}{8} = \text{Var}(\bar{X}) \leftarrow$$

(ii) Approximate, using CLT,  $P[\bar{X} \leq 10]$ .

$$P\left[\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < \frac{10 - \mu}{\sigma/\sqrt{n}}\right] \approx \phi\left(\frac{10 - 5}{\sqrt{1/8}}\right) = \phi\left(\frac{5}{\sqrt{1/8}}\right) \approx 1$$

↓   
  $\sqrt{\text{Var}(X_i)}$

(Playing roulette)

(Ques) • A roulette wheel has 18 red and 18 black slots, and 2 green slots.

- Players can bet \$1 that ball lands in red (or black) slot, and win \$1 if it does.



(a) Let  $X_i$  = winnings on the  $i^{\text{th}}$  play. What is  $\text{prf}$  of  $X_i$ ?

$$X_i \in \{-1, 1\}. P[X_i = 1] = \frac{18}{38}, P[X_i = -1] = \frac{20}{38}.$$

(b) Compute  $E X_i$ ,  $\text{Var}(X_i)$  (You may approximate  $\text{Var}(X_i)$  to 1 significant digit)

$$E X_i = \sum_{x \in \{-1, 1\}} x P[X_i = x] = \frac{18}{38} - \frac{20}{38} = \frac{-2}{38} = \frac{-1}{19}$$

$$\text{Var}(X_i) = \sum_{x \in \{-1, 1\}} (x - E X_i)^2 P[X_i = x] \approx 1 (= 0.9972..)$$

(c) Let  $S_n = X_1 + \dots + X_n$  count winnings up till your  $n^{\text{th}}$  attempt.

If  $n = 19$ , compute  $P[S_n \geq 0]$  i.e. you don't have net loss.

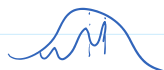
$$P\left[\frac{S_n - E S_n}{(\text{Var}(S_n))^{1/2}} \geq \frac{0 + 1}{\sqrt{19}}\right]$$

$$\text{Now, } E S_n = 19 \times \left(\frac{-1}{19}\right) = -1$$

$$\text{Var}(S_n) = n \text{Var}(X_i) = 19$$

$$\approx P\left[Z \geq \frac{1}{\sqrt{19}}\right] = 1 - \phi\left(\frac{1}{\sqrt{19}}\right) < 0.5$$

> 0.5



(Crazy big numbers)

(Ques) Say,  $U_1, \dots, U_{2000} \sim \text{unif}(0,1)$  iid. Approximate  $P[1980 \leq S_{2000} \leq 2020]$ , where  $S_{2000} = \sum_{i=1}^{2000} U_i$ .

# Jensen's inequality

(Def) (Convexity)

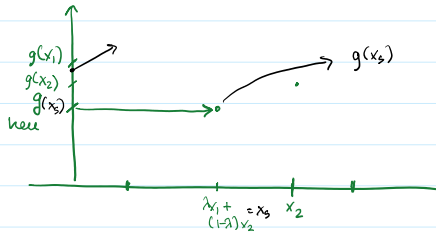
A function  $g: I \rightarrow \mathbb{R}$ , where  $I$  is some interval in  $\mathbb{R}$ , is convex if

for all  $0 \leq \lambda \leq 1$  and  $x_1, x_2 \in I$ ,

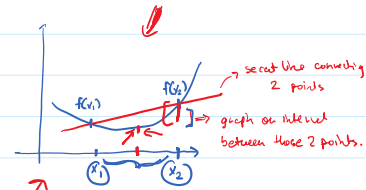
$$g(\lambda x_1 + (1-\lambda)x_2) \leq \lambda g(x_1) + (1-\lambda)g(x_2).$$

Choose  $\lambda > 0$ ,

$x_1, x_2 \in I$ .



"secant line connecting 2 points lies above the graph on the interval between those 2 points"



$$\lambda x_1 + (1-\lambda)x_2$$

(Th) If  $X$  is an r.v. with  $p$  convex,  $p(E[X]) \leq E[p(X)]$ . (Write (pf) from Gabe's book + 1 application, on the support of  $X$ )

(PF) (A1) Say,  $X$  is finite discrete r.v., w/ support  $\{x_1, \dots, x_n\}$

(A2) By induction, we first show:

(Lm) if  $g: I \rightarrow \mathbb{R}$  is convex, then for  $\lambda_1, \dots, \lambda_m$  s.t.  $\sum \lambda_i = 1$ , and  $x_1, \dots, x_m \in I$ ,  $g(\lambda_1 x_1 + \lambda_2 x_2 + \dots + \lambda_m x_m) \leq \lambda_1 g(x_1) + \dots + \lambda_m g(x_m)$

(PF) By induction on  $m$ :

(1)  $m=2$

Then, letting  $\lambda = \lambda_1$ , we see  $\lambda_2 + \lambda_1 = 1$  and the claim  $g(\lambda_1 x_1 + \lambda_2 x_2) \leq \lambda_1 g(x_1) + \lambda_2 g(x_2)$  follows since  $g$  is convex.

(2) Assume true for  $m=k \in \mathbb{N}$

(3) Then, for  $m=k+1$ ...

(a) Say,  $0 \leq \lambda_1, \dots, \lambda_{k+1} \leq 1$  s.t.  $\sum_{i=1}^{k+1} \lambda_i = 1$

(b)  $x_1, \dots, x_{k+1} \in I$ .

(c) To show:  $g(\lambda_1 x_1 + \dots + \lambda_k x_k + \lambda_{k+1} x_{k+1}) \leq \lambda_1 g(x_1) + \dots + \lambda_k g(x_k) + \lambda_{k+1} g(x_{k+1})$

We know  $\lambda_1 + \dots + \lambda_k = 1 - \lambda_{k+1} > 0$ . Then, we rewrite:

$$\lambda_1 x_1 + \dots + \lambda_k x_k + \lambda_{k+1} x_{k+1} = (1 - \lambda_{k+1}) \left[ \frac{\lambda_1 x_1 + \dots + \lambda_k x_k}{1 - \lambda_{k+1}} \right] + \lambda_{k+1} x_{k+1}$$

I claim that  $y \in I$ , (because  $I$  is convex, and  $y$  is convex combo. of points in  $I$ )

Then, since  $g$  is convex, we know  $g((1 - \lambda_{k+1}) y + \lambda_{k+1} x_{k+1}) \leq (1 - \lambda_{k+1}) g(y) + \lambda_{k+1} g(x_{k+1})$

$$= (\lambda_1 + \dots + \lambda_k) g\left(\frac{\lambda_1 x_1 + \dots + \lambda_k x_k}{\lambda_1 + \dots + \lambda_k}\right) + \lambda_{k+1} g(x_{k+1})$$

$$\leq \lambda_1 g(x_1) + \dots + \lambda_k g(x_k) + \lambda_{k+1} g(x_{k+1}), \text{ as desired.}$$

(A3) Now, note that  $E[X] = p_1 x_1 + \dots + p_n x_n$ , where  $p_i = P[X=x_i]$ .

We know  $\sum_{i=1}^n p_i = 1$ ,  $0 \leq p_i \leq 1$ , and so we may apply above lemma:

$$p_1 x_1 + \dots + p_n x_n = E[X] = g(E[X]) \leq E[g(X)] = p_1 g(x_1) + \dots + p_n g(x_n)$$

We know  $\sum_{i=1}^n p_i = 1$ ,  $0 \leq p_i \leq 1$ , and so we may apply above lemma:

$$\mathbb{P}(\mathbb{E}X) = \mathbb{P}(p_1 x_1 + \dots + p_n x_n) \stackrel{\text{lemma}}{=} p_1 \mathbb{P}(x_1) + \dots + p_n \mathbb{P}(x_n) = \mathbb{E}[\mathbb{P}(X)], \text{ as desired.}$$