

# Assignment

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**Question:** Suppose that  $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n, \mathbf{Y}_1, \mathbf{Y}_2, \dots, \mathbf{Y}_n$  are independent and identically distributed random vectors each having  $N_p(\mu, \Sigma)$  distributions, where  $\Sigma$  is non-singular,  $p > 1$  and  $n > 1$ . If  $\bar{\mathbf{X}} = \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i$  and  $\bar{\mathbf{Y}} = \frac{1}{n} \sum_{i=1}^n \mathbf{Y}_i$ , then which one of the following statements is true?

- (a) There exists  $c > 0$  such that  $c(\bar{\mathbf{X}} - \mu)^T \Sigma^{-1} (\bar{\mathbf{X}} - \mu)$  has  $\chi^2$ -distribution with  $p$  degrees of freedom.
- (b) There exists  $c > 0$  such that  $c(\bar{\mathbf{X}} - \bar{\mathbf{Y}})^T \Sigma^{-1} (\bar{\mathbf{X}} - \bar{\mathbf{Y}})$  has  $\chi^2$ -distribution with  $(p - 1)$  degrees of freedom.
- (c) There exists  $c > 0$  such that  $c \sum_{i=1}^n (\mathbf{X}_i - \bar{\mathbf{X}})^T \Sigma^{-1} (\mathbf{X}_i - \bar{\mathbf{X}})$  has  $\chi^2$ -distribution with  $p$  degrees of freedom.
- (d) There exists  $c > 0$  such that  $c \sum_{i=1}^n (\mathbf{X}_i - \mathbf{Y}_i - \bar{\mathbf{X}} + \bar{\mathbf{Y}})^T \Sigma^{-1} (\mathbf{X}_i - \mathbf{Y}_i - \bar{\mathbf{X}} + \bar{\mathbf{Y}})$  has  $\chi^2$ -distribution with  $p$  degrees of freedom.

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**Solution:**

We are given that,

$$\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n, \mathbf{Y}_1, \mathbf{Y}_2, \dots, \mathbf{Y}_n \sim N_p(\mu, \Sigma) \quad (1)$$

Also,

$$\bar{\mathbf{X}} = \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i \quad (2)$$

$$\bar{\mathbf{Y}} = \frac{1}{n} \sum_{i=1}^n \mathbf{Y}_i \quad (3)$$

The mean of  $\bar{\mathbf{X}}$  is given by:

$$\mu_{\bar{\mathbf{X}}} = E(\bar{\mathbf{X}}) \quad (4)$$

$$= \frac{1}{n} \sum_{i=1}^n E(\mathbf{X}_i) \quad (5)$$

$$= \mu \quad (6)$$

Similarly,

$$\mu_{\bar{\mathbf{Y}}} = \mu \quad (7)$$

The covariance of  $\bar{\mathbf{X}}$  is given by:

$$\Sigma_{\bar{\mathbf{X}}} = E \left[ (\bar{\mathbf{X}} - \mu) (\bar{\mathbf{X}} - \mu)^T \right] \quad (8)$$

$$= E \left[ \left( \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i - \mu \right) \left( \frac{1}{n} \sum_{i=1}^n \mathbf{X}_i - \mu \right)^T \right] \quad (9)$$

$$= \frac{1}{n^2} E \left[ \sum_{i=1}^n (\mathbf{X}_i - \mu_i) \sum_{j=1}^n (\mathbf{X}_j - \mu_j)^T \right] \quad (10)$$

$$= \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n E \left[ (\mathbf{X}_i - \mu) (\mathbf{X}_j - \mu)^T \right] \quad (11)$$

$$= \Sigma \quad (12)$$

Similarly,

$$\Sigma_{\bar{\mathbf{Y}}} = \Sigma \quad (13)$$

To check option (A):  
let us say,

$$\mathbf{A} = (\bar{\mathbf{X}} - \mu)^T \Sigma^{-1} (\bar{\mathbf{X}} - \mu) \quad (14)$$

$$(15)$$

And,

$$\Sigma^{-1} = \mathbf{F}^T \mathbf{F} \quad (16)$$

$$\bar{\mathbf{y}} = \mathbf{F} (\bar{\mathbf{X}} - \mu) \quad (17)$$

$$\implies \mathbf{A} = \bar{\mathbf{y}}^T \bar{\mathbf{y}} \quad (18)$$

$$= \|\bar{\mathbf{y}}\|^2 \quad (19)$$

Equation (19) shows that  $\mathbf{A}$  can have  $\chi^2$ -distribution.

To confirm that we will find the mean and covariance-matrix of  $\bar{\mathbf{y}}$ .

$$\mu_{\bar{\mathbf{y}}} = E(\bar{\mathbf{y}}) \quad (20)$$

$$= E(\mathbf{F} (\bar{\mathbf{X}} - \mu)) \quad (21)$$

$$= \mathbf{F} [E(\bar{\mathbf{X}}) - E(\mu)] \quad (22)$$

$$= \mathbf{F} [\mu - \mu] \quad \text{from (6)} \quad (23)$$

$$= 0 \quad (24)$$

And,

$$\Sigma_{\bar{\mathbf{y}}} = E \left[ (\bar{\mathbf{y}} - \mu_{\bar{\mathbf{y}}}) (\bar{\mathbf{y}} - \mu_{\bar{\mathbf{y}}})^T \right] \quad (25)$$

$$= E \left[ (\mathbf{F} (\bar{\mathbf{X}} - \mu)) (\mathbf{F} (\bar{\mathbf{X}} - \mu))^T \right] \quad (26)$$

$$= E \left[ \mathbf{F} (\bar{\mathbf{X}} - \mu) (\bar{\mathbf{X}} - \mu)^T \mathbf{F}^T \right] \quad (27)$$

$$= \mathbf{F} \left[ E \left[ (\bar{\mathbf{X}} - \mu) (\bar{\mathbf{X}} - \mu)^T \right] \right] \mathbf{F}^T \quad (28)$$

$$= \mathbf{F} \Sigma \mathbf{F}^T \quad (29)$$

$$= \mathbf{I} \quad \text{from (16)} \quad (30)$$

Hence **A** has  $\chi^2$ -distribution.  
So option (A) is correct.