### Recitation 10

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### Linear Regression

- ▶ Deep topic, there are entire courses on Linear Regression and friends
- ▶ Very nice to analyze mathematically. Guaranteed solutions via convexity.
- ▶ Combine with *convex* functions for regularization
- ightharpoonup Lasso,  $L_1$  penalty
  - ▶ Good for variable selection
- ightharpoonup Ridge,  $L_2$  penalty
  - ▶ The go-to baseline in most cases
- ▶ In practice, don't use linear regression w/out regularization.
  - ► See Intro to Data Science, Machine Learning

# Questions: Linear Regression Warm Up

When solving the least squares problem, the optimization problem is  $\min_{\beta} ||X\beta - y||_2^2$ ,  $X \in \mathbb{R}^{n \times d}$ ,  $y \in \mathbb{R}^n$ ,  $\beta \in \mathbb{R}^d$ .  $n \ge d$ 

- 1. Explain what  $X, \beta, y$  represent
- 2. Geometrically, what are we trying to do?
- 3. How can we obtain the normal equation  $X^T X \beta = X^T y$  from this geometric intuition?
  - Hint  $Im(A)^{\perp} = Ker(A^T)$
- 4. Under what conditions is  $X^TX$  invertible? If  $X^TX$  is not invertible, do the normal equations still have a solution?

### Solutions 1: Linear Regression Warm Up

When solving the least squares problem, the optimization problem is  $\min_{\beta} ||X\beta - y||_2^2$ ,  $X \in \mathbb{R}^{n \times d}$ ,  $y \in \mathbb{R}^n$ ,  $\beta \in \mathbb{R}^d$ .  $n \geq d$ 

#### Solution

- 1. Explain what  $X, \beta, y$  represent.
  - X contains the independent variable observations on in each row, and the features in each column.
  - y contains the corresponding dependent variable observations. β contains the coefficients that transform the independent
  - variable into the dependent variable.
- 2. Geometrically, what are we trying to do? Thinking of this from the framework of linear transformations, we are trying to find a point  $\hat{\beta} \in \mathbb{R}^d$ , s.t  $X\beta \in Im(X) \subset \mathbb{R}^n$  is closest to y.

### Solutions 2: Linear Regression Warm Up

When solving the least squares problem, the optimization problem is  $\min_{\beta} ||X\beta - y||_2^2$ ,  $X \in \mathbb{R}^{n \times d}$ ,  $y \in \mathbb{R}^n$ ,  $\beta \in \mathbb{R}^d$ .  $n \geq d$ 

### Solution

3. How can we obtain the normal equation  $X^T X \beta = X^T y$  from this geometric intuition?

The point closest to y in  $\operatorname{Im}(X)$  is the projection of y onto  $\operatorname{Im}(X)$ . Let  $P_{\operatorname{Im}(X)}y=X\hat{\beta}$ 

$$X\hat{\beta} - y \perp \operatorname{Im}(X)$$

$$X\hat{\beta} - y \in \operatorname{Ker}(X^T) \quad \text{since } Im(A)^{\perp} = Ker(A^T)$$

$$X^T(X\hat{\beta} - y) = 0$$

$$X^T X \hat{\beta} = X^T y$$
  
Index what conditions is  $X^T$ 

4. Under what conditions is  $X^TX$  invertible? If  $X^TX$  is not invertible, do the normal equations still have a solution?  $X^TX$  is invertible if rank(X) = d. There is always a solution since  $Im(X^TX) = Im(X^T)$  (use SVD, check Rec 8).

# Questions: Linear Regression vs PCA

Let  $(\vec{x_1}, y_1), ..., (\vec{x_n}, y_n) \in \mathbb{R}^{d+1}$  be a centered dataset. Each  $\vec{x_i} \in \mathbb{R}^d$ .

Let  $\beta \in \mathbb{R}^d$ . Let  $X \in \mathbb{R}^{n \times d}$ , n > d have full rank).

Let y be the vector containing  $y_1, ..., y_n$ .

The OLS solution is given by  $\hat{\beta} = (X^T X)^{-1} X^T y$ .

We can use this to generate predictions  $\hat{y} = X(X^TX)^{-1}X^T\vec{y}$ .

- 1. Recall that  $X(X^TX)^{-1}X^T$  is an orthogonal projection, which subspace is this an orthogonal projection onto?
- 2. Let n=1; consider the subspace generated by the first principal component (from PCA) and the line generated by linear regression solution for  $\vec{y} = \vec{x}\beta$ . Are these the same line? If not, what is the difference?

# Solutions 1: Linear Regression vs PCA

Let  $(\vec{x_1}, y_1), ..., (\vec{x_n}, y_n) \in \mathbb{R}^{d+1}$  be a centered dataset.

Each  $\vec{x_i} \in \mathbb{R}^d$ .

Let  $\beta \in \mathbb{R}^d$ . Let  $X \in \mathbb{R}^{n \times d}$ , n > d have full rank).

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We can use this to generate predictions  $\hat{y} = X(X^TX)^{-1}X^T\vec{y}$ .

1. Recall that  $X(X^TX)^{-1}X^T$  is an orthogonal projection, which subspace is this an orthogonal projection onto?

### Solution

Using SVD, let X have SVD  $X = U\Sigma V^T$ , then

$$X(X^TX)^{-1}X^T = U_dU_d^T$$
 (Per Recitation 8)

So  $X(X^TX)^{-1}X^T$  is an orthogonal projection onto the columns of  $U_d$ , which span Im(X).

Note:  $X \in \mathbb{R}^{n \times d}$ , and Im(X) is a d dimensional subspace in  $\mathbb{R}^n$ .

Question for you! How interpretable is this? (Answer... not very)

## Solutions 2: Linear Regression vs PCA

Let  $(\vec{x_1}, y_1), ..., (\vec{x_n}, y_n) \in \mathbb{R}^{d+1}$  be a centered dataset. Each  $\vec{x_i} \in \mathbb{R}^d$ .

Let  $\beta \in \mathbb{R}^d$ . Let  $X \in \mathbb{R}^{n \times d}$ , n > d have full rank).

Let y be the vector containing  $y_1, ..., y_n$ .

The OLS solution is given by  $\hat{\beta} = (X^T X)^{-1} X^T y$ .

We can use this to generate predictions  $\hat{y} = X(X^TX)^{-1}X^T\vec{y}$ .

2. Let n=1; consider the subspace generated by the first principal component (from PCA) and the line generated by linear regression solution for  $\vec{y} = \vec{x}\beta$ . Are these the same line? If not, what is the difference?

#### Solution

(Check notebook on github)

They are not the same line. PCA is an orthogonal projection that minimizes the  $L_2$  orthogonal distance to the line, while linear regression minimizes the  $L_2$  distance parallel to the y-axis to the line.

## Questions: Ridge Regression

Let  $X \in \mathbb{R}^{n \times d}$ , n > d, and not have full rank. (X is a data matrix) Recall that the OLS solution is  $\hat{x} = (X^T X)^{-1} X^T y$ .

- 1. Since X is not full rank, what does this say about the features?
- 2. What is the issue with the OLS solution?
- 3. The ridge regression solution is given by  $(X^TX + \lambda Id_d)^{-1}X^Ty$ . How does this fix the issue?
- 4. Suppose that X has SVD  $X = U\Sigma V^T$ , and X has singular values  $\sigma_1, ..., \sigma_d$ . What are the eigenvalues of  $X^TX + \lambda Id_d$ ?
- 5. How does increasing  $\lambda$  affect the condition number of  $(X^TX + \lambda Id_d)$ ?

## Solutions: Ridge Regression and Multicollinearity

Let  $X \in \mathbb{R}^{n \times d}$ , n > d, and not have full rank. (X is a data matrix) Recall that the OLS solution is  $\hat{x} = (X^T X)^{-1} X^T y$ .

#### Solution

- 1. Since X is not full rank, what does this say about the features?

  Columns of X are not linearly independent, so some of the features can be perfectly explained by other features.
- 2. What is the issue with the OLS solution? Since X does not have full rank,  $X^TX$  doesn't have full rank and is not invertible. So the OLS solution is not well-defined.
- 3. The ridge regression solution is given by  $(X^TX + \lambda Id_d)^{-1}X^Ty$ . How does this fix the issue? Adding  $\lambda Id_d$  to  $X^TX$  shifts its eigenvalues up, which makes  $(X^TX + \lambda Id_d)$  invertible.

## Solutions: Ridge Regression and Multicollinearity

Let  $X \in \mathbb{R}^{n \times d}$ , n > d, and not have full rank. (X is a data matrix) Recall that the OLS solution is  $\hat{x} = (X^T X)^{-1} X^T y$ .

### Solution

4. Suppose that X has SVD  $X = U\Sigma V^T$ , and X has singular values  $\sigma_1, ..., \sigma_d$ . What are the eigenvalues of  $X^TX + \lambda Id_d$ ?

Note that  $X^TX = V\Sigma^T\Sigma V^T$ 

Eigvals of  $X^T X$ :  $\sigma_1^2, ..., \sigma_d^2$ , (Note: X isn't full rank, so  $\sigma_d = 0$ ) Eigvals of  $X^T X + \lambda Id_d$ :  $\sigma_1^2 + \lambda, ..., \sigma_d^2 + \lambda$ .

5. How does increasing  $\lambda$  affect the condition number of  $(X^TX + \lambda Id_d)$  vs  $X^TX$ ?

Condition number of  $X^T X = \frac{\sigma_1^2}{\sigma_d^2} = \infty$ 

Condition number of  $(X^TX + \lambda Id_d) = \frac{\sigma_1^2 + \lambda}{\sigma_d^2 + \lambda}$ 

Furthermore, for  $\lambda_1 > \lambda_2$ , we get the relationship  $\frac{\sigma_1^2 + \lambda_1}{\sigma_d^2 + \lambda_1} < \frac{\sigma_1^2 + \lambda_2}{\sigma_d^2 + \lambda_2}$