

# Application of seasonal ARIMA model for forecasting GDP of United Kingdom

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In this paper, an attempt will be made to forecast the value of Gross Domestic Product for the United Kingdom of Great Britain and Northern Ireland for the fourth quarter of 2021. For this purpose, the country's historical GDP data in the form of a time series will be used, covering the range from the first quarter of 1955 to the third quarter of 2021.

The data is contained in the file `gdp_uk.xlsx` and contains 267 observations. For each of them, the values of two variables were given:

- date – year and quarter,
- gdp – GDP value in million GBP, in current prices, not seasonally adjusted.

The data comes from the resources of [Office for National Statistics](#).

The scope of the data was limited by removing the observations to the 1st quarter of 1980 to eliminate the values disturbing the desired series characteristics.

## Initial analysis

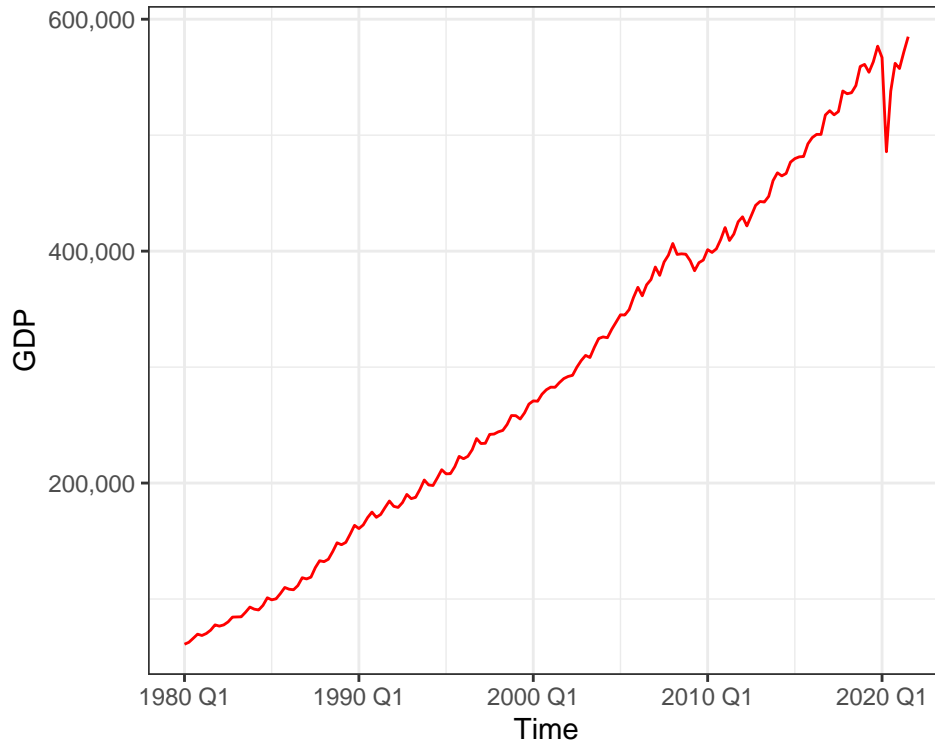
Descriptive statistics for the `gdp` variable:

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
61022	167071	280461	294553	412376	584928
SD	CV	Skewness	Kurtosis		
153515.8	52.11816	0.1999873	1.835088		

The time series average is £ 294,553m with a standard deviation of £ 153,515m. The middle value of the series is £ 280,461m. The smallest value in the series is £ 61,022m and the highest is £ 584,928m. The middle 50% of the series is between the first and third quartiles of £ 167,071m and £ 412,376m, respectively. The volatility coefficient is 52.12, which means the average volatility. A skewness of 0.2 means a slightly right-asymmetric distribution. A positive kurtosis of 1.83 indicates a high concentration of the values around the mean.

The chart shows a relatively steady upward trend with two major collapses around 2008 and 2020.

Chart 1. United Kingdom GDP in 1980-2021



## Determining time series stationarity

### Augmented Dickey Fuller test for integration

$H_0 : \rho = 0$ ,  $y_t$  it is not a stationary process

$H_1 : \rho < 0$ ,  $y_t$  is a stationary process

### For unchanged data

```
[1] "ADF without trend and constant"
      statistic 1pct 5pct 10pct
tau1  3.474337 -2.58 -1.95 -1.62
[1] "ADF with constant"
      statistic 1pct 5pct 10pct
tau2   1.61824 -3.46 -2.88 -2.57
phi1  11.65688  6.52  4.63  3.81
[1] "ADF with constant and trend"
      statistic 1pct 5pct 10pct
tau3  -2.012357 -3.99 -3.43 -3.13
phi2   9.526477  6.22  4.75  4.07
phi3   3.684532  8.43  6.49  5.47
```

Without taking into account the intercept and the trend, the test statistic  $\tau_1 = 3.47$  is greater than the critical value  $-1.95$ , which means that at the significance level of 5% there is no reason to reject the null hypothesis of non-stationarity of the series.

Taking into account the intercept, the test statistic  $\tau_2 = 1.61$  is greater than the critical value  $-2.88$ , which means that at the significance level of 5% there are no grounds to reject the null hypothesis of non-stationarity of the series.

Taking into account the intercept and the trend, the test statistic  $\tau_3 = -2.01$  is greater than the critical value  $-3.43$ , which means that at the significance level of 5% there are no grounds to reject the null hypothesis of series non-stationarity.

### For first seasonal differences

```
[1] "ADF without trend and constant"
      statistic 1pct 5pct 10pct
tau1 -0.9821618 -2.58 -1.95 -1.62
[1] "ADF with constant"
      statistic 1pct 5pct 10pct
tau2 -3.045680 -3.46 -2.88 -2.57
phi1  4.638144  6.52  4.63  3.81
[1] "ADF with constant and trend"
      statistic 1pct 5pct 10pct
tau3 -5.386365 -3.99 -3.43 -3.13
phi2  9.686629  6.22  4.75  4.07
phi3 14.508674  8.43  6.49  5.47
```

Without taking into account the intercept and the trend, the test statistic  $\tau_1 = -0.98$  is greater than the critical value  $-1.95$ , which means that at the significance level of 5% there is no reason to reject the null hypothesis of non-stationarity of the series.

Taking into account the intercept, the test statistic  $\tau_2 = -3.05$  is lower than the critical value  $-2.88$ , which means that at the significance level of 5% we reject the null hypothesis of non-stationarity of the series.

Taking into account the intercept and the trend, the test statistic  $\tau_3 = -5.39$  is lower than the critical value  $-3.43$ , which means that at the significance level of 5% we reject the null hypothesis of non-stationarity of the series.

From the above tests it appears that the series is stationary after the first differentiation,  $X_t \sim I(1)$ .

## HEGY test for seasonal integration

$H_A : \pi_1 = 0$ ,  $y_t$  is non-stationary, non-seasonal

$H_B : \pi_2 = 0$ ,  $y_t$  is a non-stationary process with two-period seasonality

$H_C : \pi_3 = \pi_4 = 0$ ,  $y_t$  is a non-stationary process with four-period seasonality

Call:

```
lm(formula = y4[-1:-4] ~ y1_1 + y2_1 + y3_1 + y3_2 + y4_1 + y4_2 +  
    y4_3 + y4_4)
```

Residuals:

Min	1Q	Median	3Q	Max
-79165	-1849	455	2689	19967

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	6.650e+03	1.602e+03	4.150	5.48e-05 ***
y1_1	2.719e-03	1.084e-03	2.508	0.013175 *
y2_1	-4.263e-01	1.231e-01	-3.463	0.000692 ***
y3_1	3.421e-02	8.478e-02	0.404	0.687117
y3_2	3.805e-02	9.730e-02	0.391	0.696260
y4_1	1.382e-01	9.639e-02	1.434	0.153684
y4_2	4.690e-01	1.805e-01	2.598	0.010287 *
y4_3	-1.457e-01	2.080e-01	-0.700	0.484847
y4_4	-4.995e-01	1.641e-01	-3.043	0.002755 **

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Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7853 on 154 degrees of freedom

Multiple R-squared: 0.5654, Adjusted R-squared: 0.5428

F-statistic: 25.05 on 8 and 154 DF, p-value: < 2.2e-16

Linear hypothesis test

Hypothesis:

y3\_1 = 0

y3\_2 = 0

Model 1: restricted model

Model 2: y4[-1:-4] ~ y1\_1 + y2\_1 + y3\_1 + y3\_2 + y4\_1 + y4\_2 + y4\_3 +  
y4\_4

Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	156	9516718018			
2	154	9496982868	2	19735150	0.16 0.8523

Table 1. HEGY test critical values for quarterly data (Hylleberg et al. 1990).

$n$	$H_A: t\text{-test } \pi_1 = 0$			$H_B: t\text{-test } \pi_2 = 0$			$H_C: F\text{-test } \pi_3 = \pi_4 = 0$		
	1%	5%	10%	1%	5%	10%	1%	5%	10%
48	-3.66	-2.96	-2.62	-2.68	-1.95	-1.60	4.78	3.04	2.32
100	-3.47	-2.88	-2.58	-2.61	-1.95	-1.60	4.77	3.08	2.35
136	-3.51	-2.89	-2.58	-2.60	-1.91	-1.58	4.73	3.00	2.36
200	-3.48	-2.87	-2.57	-2.58	-1.92	-1.59	4.76	3.12	2.37

The t-test statistic for  $y_1(-1)$  variable 2.51 is greater than the critical value -2.89, which means that at the significance level of 5% there are no grounds to reject the hypothesis of non-seasonal non-stationarity.

The t-test statistic for  $y_2(-1)$  variable -3.46 is lower than the critical value -1.91, which means that at the significance level of 5%, we reject the hypothesis of two-period seasonal non-stationarity.

The F-test statistic for the variables  $y_3(-1)$  and  $y_3(-2)$  0.16 is lower than the critical value 3.00, which means that at the significance level of 5% there are no grounds to reject the hypothesis of four-period seasonal non-stationarity.

Table 2. HEGY test hypothesis combinations (Hylleberg et al. 1990).

These hypotheses aren't rejected	These hypotheses are rejected	Stationary variable
$H_A, H_B, H_C$	—	$\Delta_4 y_t$ ( $= y_{4t}$ )
$H_A, H_B$	$H_C$	$\Delta_2 y_t$ ( $= y_{2t}$ )
$H_A, H_C$	$H_B$	$(I-L)(I+L^2)y_t$ ( $= y_{2t}$ )
$H_B, H_C$	$H_A$	$(I+L)(I+L^2)y_t$ ( $= y_{4t}$ )
$H_A$	$H_B, H_C$	$\Delta_1 y_t$
$H_B$	$H_A, H_C$	$(I+L)y_t$
$H_C$	$H_A, H_B$	$(I+L^2)y_t$
—	$H_A, H_B, H_C$	$y_t$

Based on the table above, it can be concluded that the series is characterized by a semi-annual stationarity.

## Seasonal ARIMA model

Based on the tests performed above, it can be determined that the parameters  $d$  and  $D$  in the  $SARIMA(p,d,q)(P,D,Q)$  model will have the value of 1. The optimal values of the remaining parameters are not known, therefore several models with parameter values from 0 to 2 will be estimated.

	p	d	q	P	D	Q	AIC
M1	2	1	2	2	1	2	3385.053
M2	2	1	2	1	1	2	3383.471
M3	2	1	2	1	1	1	3381.501
M4	2	1	1	1	1	1	3383.346
M5	1	1	1	1	1	1	3382.138
M6	1	1	1	0	1	1	3380.691
M7	0	1	1	0	1	1	3383.170

From the above array it can be concluded that, based on the Akaike criterion, the best model is the Model 6.

Forecast method: ARIMA(1,1,1)(0,1,1)[4]

Model Information:

Series: dfts

ARIMA(1,1,1)(0,1,1)[4]

Coefficients:

	ar1	ma1	sma1
	0.5358	-0.8339	-0.8453
s.e.	0.2273	0.1716	0.0638

sigma<sup>2</sup> = 63163093: log likelihood = -1686.35

AIC=3380.69 AICc=3380.95 BIC=3393.04

Error measures:

	ME	RMSE	MAE	MPE	MAPE	MASE
Training set	677.2466	7754.825	3455.684	0.2782024	1.067464	0.2411356

ACF1

Training set -0.03223876

Forecasts:

	Point Forecast	Lo 95	Hi 95
2021 Q4	593756.3	578179.4	609333.1

Standard error: 7947.523

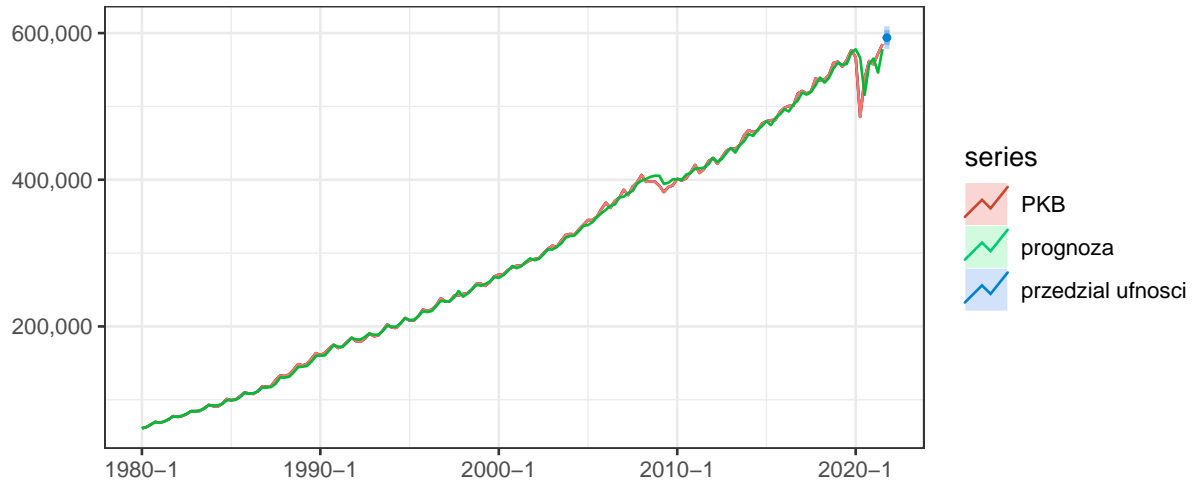
The GDP forecast for the fourth quarter of 2021 for the United Kingdom is £ 593756.3m with an average estimate error of £ 7947.523m, which is 1.3% of the forecast value, which proves that the forecast is acceptable.

The range <578179.4;609333.1> has a 95% probability to cover the unknown United Kingdom GDP for the fourth quarter of 2021.

Based on the arithmetic mean of the forecast error, our forecasts were on average £ 677.247m too low in relation to the real GDP value, which is on average 0.278% of the forecasted variable, which proves that the forecast is not biased.

The value of the GDP forecast deviates from the true value by an average of £ 3455.684m, with forecast errors averaging 1.07% of the true value, which proves that the forecasts are acceptable.

Chart 2. United Kingdom GDP forecast

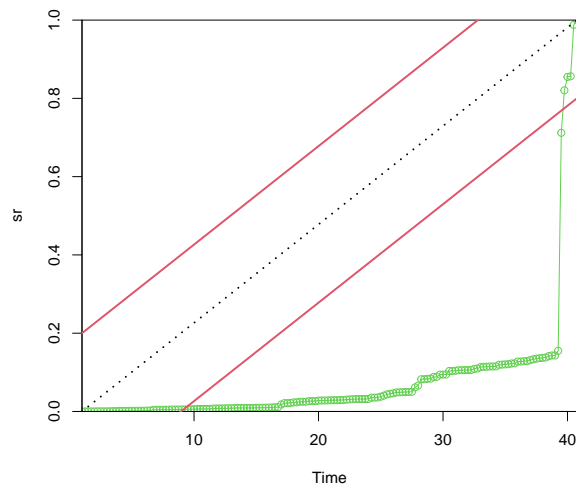
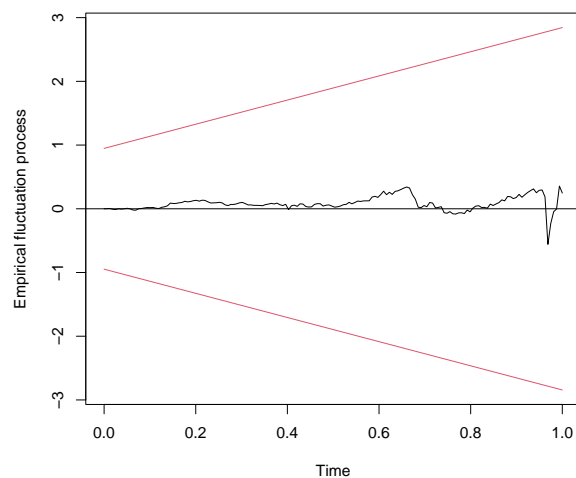


### CUSUM stability test

From Chart 3, it can be concluded that the model is suitable for forecasting because the plot of the cumulative sums has not exceeded the confidence interval.

Based on Chart 4, it can be concluded that the model is characterized by an unstable inference precision.

Chart 3. CUSUM and Chart 4. CUSUM-SQ



## Chow stability test

$$H_0 : \beta_1 = \beta_2$$

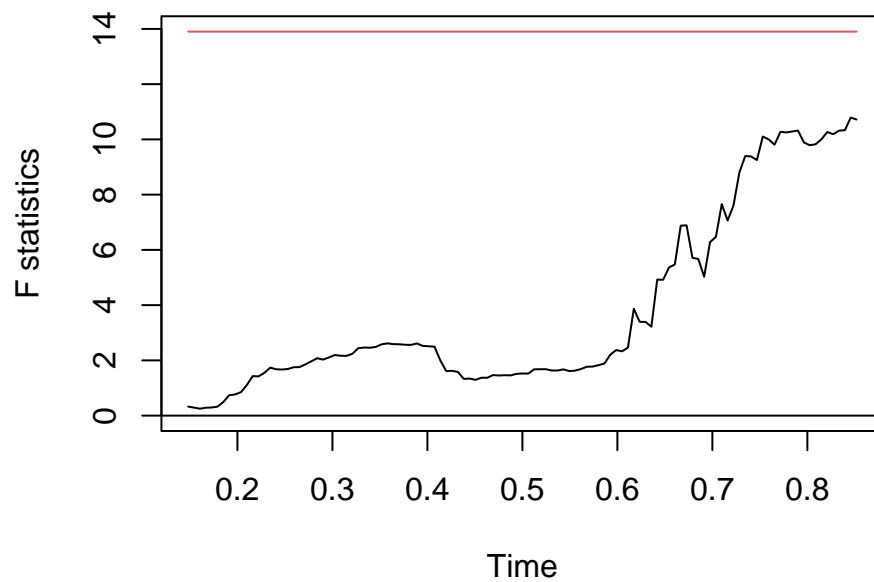
$$H_0 : \beta_1 \neq \beta_2$$

### Chow test

```
data: dls1gdp ~ ma_1 + ma_4  
F = 0.50884, p-value = 0.6768
```

The **p-value** is greater than the significance level of 5% which means there is no structural break. The same conclusion can be drawn from the graph of the QLR test.

Chart 5. QLR test for structural break



## Bibliography

Hylleberg, S. et al. (1990). "Seasonal integration and cointegration". In: *Journal of Econometrics* 44.1–2, pp. 215–238.