

ECON 6511: Econometrics
Midterm Exam
Points: 60, Time: 210 minutes

Instructions

- Use 3 or more decimal places unless otherwise stated
- No notes or cellphones

Section A: Multiple-choice (15 points)

1. The model: $y_t = \beta_0 + \beta_1 c_t + u_t$, $t = 1, 2, \dots, n$, is an example of a(n):
 - (a) autoregressive conditional heteroskedasticity model.
 - (b) static model.
 - (c) finite distributed lag model.
 - (d) infinite distributed lag model.
2. Which of the following is an assumption on which time series regression is based?
 - (a) A time series process follows a model that is nonlinear in parameters.
 - (b) In a time series process, no independent variable is a perfect linear combination of the others.
 - (c) In a time series process, at least one independent variable is a constant.
 - (d) For each time period, the expected value of the error u_t , given the explanatory variables for all time periods, is positive.
3. If an explanatory variable is strictly exogenous it implies that:
 - (a) changes in the lag of the variable does not affect future values of the dependent variable.
 - (b) the variable is correlated with the error term in all future time periods.
 - (c) the variable cannot react to what has happened to the dependent variable in the past.
 - (d) the conditional mean of the error term given the variable is zero.
4. A study which observes whether a particular occurrence (e.g. a new policy) influences some outcome is referred to as a(n):
 - (a) event study.
 - (b) exponential study.

- (c) laboratory study.
 - (d) comparative study.
5. A covariance stationary time series is weakly dependent if:
- (a) the correlation between the independent variable at time t and the dependent variable at time $t + h$ goes to ∞ as $h \rightarrow 0$.
 - (b) the correlation between the independent variable at time t and the dependent variable at time $t + h$ goes to 0 as $h \rightarrow \infty$.
 - (c) the correlation between the independent variable at time t and the independent variable at time $t + h$ goes to ∞ as $h \rightarrow 0$.
 - (d) the correlation between the independent variable at time t and the independent variable at time $t + h$ goes to 0 as $h \rightarrow \infty$.
6. The model $y_t = e_t + \beta_1 e_{t-1} + \beta_2 e_{t-2}$, $t = 1, 2, \dots$, where e_t is an i.i.d. sequence with zero mean and variance σ_e^2 represents a(n):
- (a) static model.
 - (b) moving average process of order one.
 - (c) moving average process of order two.
 - (d) autoregressive process of order two.
7. The model $y_t = y_{t-1} + e_t$, $t = 1, 2, \dots$ represents a:
- (a) AR(2) process.
 - (b) MA(1) process.
 - (c) random walk process.
 - (d) random walk with a drift process.
8. If a process is said to be integrated of order one, or I(1), then:
- (a) it is covariance stationary
 - (b) averages of such processes already satisfy the standard limit theorems
 - (c) the first difference of the process is weakly dependent
 - (d) it does not have a unit root
9. In the presence of serial correlation:
- (a) estimated standard errors remain valid.

- (b) estimated test statistics remain valid.
 - (c) estimated OLS values are not BLUE.
 - (d) estimated variance does not differ from the case of no serial correlation.
10. Which of the following identifies an advantage of first differencing a time-series?
- (a) First differencing eliminates most of the serial correlation.
 - (b) First differencing eliminates most of the heteroskedasticity.
 - (c) First differencing eliminates most of the multicollinearity.
 - (d) First differencing eliminates the possibility of spurious regression.
11. A regression model exhibits unobserved heterogeneity if
- (a) there are unobserved factors affecting the dependent variable that change over time
 - (b) there are unobserved factors affecting the dependent variable that do not change over time
 - (c) there are unobserved factors which are correlated with the observed independent variables
 - (d) there are no unobserved factors affecting the dependent variable
12. Ordinary least squares estimation is subject to (heterogeneity) bias if
- (a) the regression model exhibits heteroskedasticity
 - (b) the unobserved effect is correlated with the observed explanatory variables
 - (c) the regression model includes a lagged dependent variable
 - (d) the explanatory variables do not change over time
13. What should be the degrees of freedom (df) for fixed effects estimation if the data set includes N cross sectional units over T time periods and the regression model has k independent variables?
- (a) $N - kT$
 - (b) $NT - k$
 - (c) $NT - N - k$
 - (d) $N - T - k$
14. The general approach to obtaining fully robust standard errors and test statistics in the context of panel data is known as

- (a) confounding
 - (b) differencing
 - (c) clustering
 - (d) attenuating
15. Which of the following assumptions is needed for the usual standard errors to be valid when differencing with more than two time periods?
- (a) The regression model exhibits heteroskedasticity.
 - (b) The differenced idiosyncratic error or Δu_{it} is uncorrelated over time.
 - (c) The unobserved factors affecting the dependent variable are time-constant.
 - (d) The regression model includes a lagged independent variable.

Section B: Written Answer (45 points)

1. (20 points) Consider the following model that relates wages to productivity:

$$\log(hr\text{wage}_t) = \beta_0 + \beta_1 \log(outphr_t) + u_t,$$

where the variable *hrwage* is average hourly wage in the U.S. economy, and the variable *outphr* is output per hour. The data are annual data between 1947 and 1987 with $n = 41$. The output from this regression and many other specifications are included below.

- (a) Interpret (with one perfect sentence) what the estimate of β_1 in the above model means.
- (b) **Formally** test the hypothesis that $\beta_1 = 1$ against a two-sided alternative, i.e. test whether productivity growth is fully passed onto workers in the form of higher wages.
- (c) Included in the attached output are regressions of both $\log(hr\text{wage})$ and $\log(output)$ regressed on a simple linear time trend, t . What do these regressions suggest about the likely bias in the estimate of β_1 above? Explain.
- (d) Consider a second specification that includes a time trend:

$$\log(hr\text{wage}_t) = \beta_0 + \beta_1 \log(outphr_t) + \beta_2 t + u_t$$

Interpret the coefficients on $\log(outphr_t)$ and t . Are they both statistically significant?

- (e) Explain how you would linearly de-trend the variable $\log(hr\text{wage})$, i.e. remove its time trend?
- (f) After removing time-trends, suppose the first order autocorrelations of $\log(hr\text{wage})$ and $\log(outphr)$ are $\hat{\rho} = 0.967$ and $\hat{\rho} = 0.945$ respectively. What do these findings imply? What is an appropriate remedy?
- (g) Consider a model in first differences:

$$\Delta \log(hr\text{wage}_t) = \beta_0 + \beta_1 \Delta \log(outphr_t) + \Delta u_t$$

where $ghr\text{wage} = \Delta \log(hr\text{wage}_t)$ and $goutphr = \Delta \log(outphr_t)$ in the Stata output. Interpret the coefficient on $\Delta \log(outphr_t)$. **Formally** test the hypothesis that $\beta_1 = 1$ against a two-sided alternative.

- (h) Consider a final model:

$$ghr\text{wage}_t = \beta_0 + \beta_1 goutphr_t + \beta_2 goutphr_{t-1} + u_t$$

that includes one lag of output growth. Is the lagged value statistically significant?

- (i) Test $\beta_1 + \beta_2 = 1$ against a two-sided alternative.

$$ghrwage_t = \beta_0 + \theta goutphr_t + \beta_2(goutphr_{t-1} - goutphr_t) + u_t$$

- (j) Does $goutphr_{t-2}$ need to be in the model? Explain.

. reg lhrwage loutphr

Source	SS	df	MS	Number of obs =	41
Model	.956010689	1	.956010689	F(1, 39) =	311.95
Residual	.119521723	39	.00306466	Prob > F =	0.0000
				R-squared =	0.8889
				Adj R-squared =	0.8860
Total	1.07553241	40	.02688831	Root MSE =	.05536

lhrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
loutphr	.6891012	.039016	17.66	0.000	.6101839	.7680185
_cons	-1.534398	.1716633	-8.94	0.000	-1.88162	-1.187176

. reg loutphr t

Source	SS	df	MS	Number of obs =	41
Model	1.91976939	1	1.91976939	F(1, 39) =	800.97
Residual	.09347595	39	.002396819	Prob > F =	0.0000
Total	2.01324534	40	.050331133	R-squared =	0.9536
				Adj R-squared =	0.9524
				Root MSE =	.04896

loutphr	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	.0182881	.0006462	28.30	0.000	.0169811	.0195951
_cons	4.010183	.0155758	257.46	0.000	3.978678	4.041688

. reg lhrwage t

Source	SS	df	MS	Number of obs =	41
Model	.79327836	1	.79327836	F(1, 39) =	109.61
Residual	.282254052	39	.007237283	Prob > F =	0.0000
Total	1.07553241	40	.02688831	R-squared =	0.7376
				Adj R-squared =	0.7308
				Root MSE =	.08507

lhrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	.0117559	.0011229	10.47	0.000	.0094847	.0140272
_cons	1.246799	.0270657	46.07	0.000	1.192053	1.301544

. reg lhrwage loutphr t

Source	SS	df	MS	Number of obs =	41
Model	1.04458064	2	.522290318	F(2, 38) =	641.22
Residual	.030951776	38	.00081452	Prob > F =	0.0000
Total	1.07553241	40	.02688831	R-squared =	0.9712
				Adj R-squared =	0.9697
				Root MSE =	.02854

lhrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
loutphr	1.639639	.0933471	17.56	0.000	1.450668	1.828611
t	-.01823	.0017482	-10.43	0.000	-.021769	-.0146909
_cons	-5.328454	.3744492	-14.23	0.000	-6.086487	-4.570421

. reg ghrwage goutphr

Source	SS	df	MS	Number of obs =	40
Model	.006255013	1	.006255013	F(1, 38) =	21.77
Residual	.01091799	38	.000287316	Prob > F =	0.0000
Total	.017173003	39	.000440333	R-squared =	0.3642
				Adj R-squared =	0.3475
				Root MSE =	.01695

ghrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
goutphr	.809316	.1734537	4.67	0.000	.4581773	1.160455
_cons	-.0036621	.00422	-0.87	0.391	-.0122051	.0048808

. reg ghrwage goutphr goutph_1

Source	SS	df	MS	Number of obs = 39		
Model	.008456778	2	.004228389	F(2, 36) = 17.51		
Residual	.008692555	36	.00024146	Prob > F = 0.0000		
Total	.017149333	38	.000451298	R-squared = 0.4931		
				Adj R-squared = 0.4650		
				Root MSE = .01554		

ghrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
goutphr	.7283641	.1672223	4.36	0.000	.3892217	1.067507
goutph_1	.4576351	.1656126	2.76	0.009	.1217571	.793513
_cons	-.010425	.0045439	-2.29	0.028	-.0196404	-.0012096

. gen new_var = goutph_1 - goutphr
(2 missing values generated)

. reg ghrwage goutphr new_var

Source	SS	df	MS	Number of obs = 39		
Model	.008456778	2	.004228389	F(2, 36) = 17.51		
Residual	.008692555	36	.00024146	Prob > F = 0.0000		
Total	.017149333	38	.000451298	R-squared = 0.4931		
				Adj R-squared = 0.4650		
				Root MSE = .01554		

ghrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
goutphr	1.185999	.2031416	5.84	0.000	.7740089	1.59799
new_var	.4576351	.1656126	2.76	0.009	.1217571	.793513
_cons	-.010425	.0045439	-2.29	0.028	-.0196404	-.0012096

. reg ghrwage goutphr goutph_1 goutph_2

Source	SS	df	MS	Number of obs =	38
Model	.008048615	3	.002682872	F(3, 34) =	12.04
Residual	.007575626	34	.000222813	Prob > F =	0.0000
Total	.015624241	37	.000422277	R-squared =	0.5151
				Adj R-squared =	0.4724
				Root MSE =	.01493

ghrwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
goutphr	.7464273	.1615037	4.62	0.000	.4182123	1.074642
goutph_1	.3740461	.1665312	2.25	0.031	.0356139	.7124783
goutph_2	.0653486	.1597067	0.41	0.685	-.2592144	.3899116
_cons	-.0112838	.0048409	-2.33	0.026	-.0211217	-.0014458

2. **(15 points)** Consider the following model that relates traffic deaths per 100 million miles to two policies that attempt to reduce drunk driving. The two laws are *open container laws* – which make it illegal for travel with an open alcoholic drink and *administrative per se laws* – which enable courts to suspend driver licenses immediately after an arrest. Consider a cross-sectional model relating the death rate ($dthrte$) in a state to dummy variables indicating whether the state has any of these laws (*open* and *admin*):

$$dthrte_i = \beta_0 + \beta_1 open_i + \beta_2 admin_i + u_i.$$

This equation is estimated separately for 1985 where $open_{85}$ and $admin_{85}$ are the dummy variables for 1985. Results are provided below.

- (a) How do you interpret the coefficients on $open_{85}$ and $admin_{85}$? Do the signs of the coefficients make sense? Explain.
- (b) Are the estimates likely to be unbiased? Explain.
- (c) Consider an alternative unobserved effects model using panel data with $T = 2$ periods:

$$dthrte_{it} = \beta_0 + \beta_1 d2_t + \beta_2 open_{it} + \beta_3 admin_{it} + a_i + u_{it}.$$

where $d2_t$ is a dummy variable equal to 1 in the second period. List two factors that might be included in a_i .

- (d) Explain how we can obtain unbiased or at least consistent estimates of β_1 and β_2 assuming we have data from states before and after these laws were passed.
- (e) Write down the Stata code that you would use to generate the change in $dthrte$ between two periods assuming the data was organized so that you had one observation per state and period.
- (f) Results from the following model:

$$\Delta dthrte_i = \beta_0 + \beta_1 \Delta open_i + \beta_2 \Delta admin_i + \Delta u_i.$$

using two years of data (1985 and 1990) are included below. In 1985, 19 states had open container laws, while 22 states had such laws in 1990. In 1985, 21 states had per se laws; the number had grown to 29 by 1990. Interpret the estimated coefficient on $copen$, which is equivalent to $\Delta open_i$ in the above model. Interpret the estimate of β_0 .

- (g) Does the sign on $cadmin$ make sense? Is it statistically significant?
- (h) What do we need to assume for the estimates in part (f) to be “correct?”

. reg dthrte85 open85 admn85

Source	SS	df	MS	Number of obs = 51		
Model	1.17595166	2	.58797583	F(2, 48) = 1.66		
Residual	17.0240489	48	.354667685	Prob > F = 0.2013		
Total	18.2000005	50	.36400001	R-squared = 0.0646		
				Adj R-squared = 0.0256		
				Root MSE = .59554		

dthrte85	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
open85	-.2131043	.1725004	-1.24	0.223	-.5599395	.1337309
adm85	.229711	.1694617	1.36	0.182	-.1110146	.5704366
_cons	2.684805	.1257889	21.34	0.000	2.431889	2.93772

. reg cdthrt copen cadmn

Source	SS	df	MS	Number of obs = 51		
Model	.762579785	2	.381289893	F(2, 48) = 3.23		
Residual	5.66369475	48	.117993641	Prob > F = 0.0482		
Total	6.42627453	50	.128525491	R-squared = 0.1187		
				Adj R-squared = 0.0819		
				Root MSE = .3435		

cdthrt	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
copen	-.4196787	.2055948	-2.04	0.047	-.8330547	-.0063028
cadmn	-.1506024	.1168223	-1.29	0.204	-.3854894	.0842846
_cons	-.4967872	.0524256	-9.48	0.000	-.6021959	-.3913784

3. **(15 points)** Consider the following model of rental prices in one of 64 college towns for the years 1980 and 1990. The idea is to see whether a stronger presence of students affects rental rates. The unobserved effects model is:

$$\log(\text{rent}_{it}) = \beta_0 + \delta_0 y90_t + \beta_1 \log(\text{pop}_{it}) + \beta_2 \log(\text{avginc}_{it}) + \beta_3 \text{pctstu}_{it} + a_i + u_{it},$$

where *pop* is city population, *avginc* is average income, and *pctstu* is student population as a percentage (0 to 100) of city population (during the school year). Results by pooled OLS, random effects, and fixed effects are given below.

- (a) Using pooled OLS, what is the interpretation of the 1990 dummy variable? What does the estimate of β_3 imply?
- (b) Are the standard errors for the pooled OLS estimates likely to be valid? Explain.
- (c) What does the estimate of β_3 using fixed effects imply? Why is it so different to the pooled OLS estimate?
- (d) List one factor that might be contained in a_i and explain how it might be correlated with *pctstu*.
- (e) What does the estimate of β_3 under random effects imply about the likely magnitude of $\hat{\theta}$ (the quasi demeaning parameter)?
- (f) Which model do you prefer and why?
- (g) Suppose you estimated the model using OLS but included dummy variables for each city as opposed to time demeaning or quasi demeaning the data. What would you expect to get for your estimate of β_3 ?

	Pooled OLS	Fixed Effects	Random Effects
$\log(pop)$	0.041 (0.023)	0.072 (0.088)	0.056 (0.029)
$\log(avginc)$	0.571** (0.053)	0.310** (0.066)	0.447** (0.052)
$pctstu$	0.005** (0.001)	0.011** (0.004)	0.005** (0.001)
$y90$	0.262** (0.035)	0.386** (0.037)	0.323** (0.029)
Constant	-0.569 (0.535)	1.409 (1.167)	0.445 (0.566)
Observations	128	128	128
R^2	0.861	0.977	

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$

Formulae

- For simple regression model: $\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$
- Estimated slope parameter when regression equation passes through origin: $\frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n x_i^2}$
- $SSR = \sum_{i=1}^n \hat{u}_i^2$
- $SSE = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$
- $SST_j = \sum_{i=1}^n (x_{ij} - \bar{x}_j)^2$
- $R^2 = \frac{SSE}{SST} = 1 - \frac{SSR}{SST}$
- For simple regression model: $se(\hat{\beta}_1) = \frac{\hat{\sigma}}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2}}$
- $\widehat{Var}(\hat{\beta}_j) = \frac{\hat{\sigma}^2}{SST_j(1-R_j^2)}$
- t statistic: $t_{\hat{\beta}_j} \equiv \frac{\hat{\beta}_j - \alpha_j}{se(\hat{\beta}_j)} \sim t_{n-k-1} = t_{df}$
- 95% confidence interval: $P\left(\hat{\beta}_j - c_{0.05} \cdot se(\hat{\beta}_j) \leq \beta_j \leq \hat{\beta}_j + c_{0.05} \cdot se(\hat{\beta}_j)\right) = 0.95$
- F statistic = $\frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/(n-k-1)}$
- R^2 form of F statistic = $\frac{(R_{ur}^2 - R_r^2)/q}{(1-R_{ur}^2)/(n-k-1)}$
- $\bar{R}^2 = 1 - \frac{SSR/(n-k-1)}{SST/(n-1)}$
- Chow test statistic: $F = \frac{[SSR_P - (SSR_1 + SSR_2)]/(k+1)}{(SSR_1 + SSR_2)/[n-2(k+1)]}$

Cumulative Areas under the Standard Normal Distribution

<i>z</i>	0	1	2	3	4	5	6	7	8	9
-3.0	0.0013	0.0013	0.0013	0.0012	0.0012	0.0011	0.0011	0.0011	0.0010	0.0010
-2.9	0.0019	0.0018	0.0018	0.0017	0.0016	0.0016	0.0015	0.0015	0.0014	0.0014
-2.8	0.0026	0.0025	0.0024	0.0023	0.0023	0.0022	0.0021	0.0021	0.0020	0.0019
-2.7	0.0035	0.0034	0.0033	0.0032	0.0031	0.0030	0.0029	0.0028	0.0027	0.0026
-2.6	0.0047	0.0045	0.0044	0.0043	0.0041	0.0040	0.0039	0.0038	0.0037	0.0036
-2.5	0.0062	0.0060	0.0059	0.0057	0.0055	0.0054	0.0052	0.0051	0.0049	0.0048
-2.4	0.0082	0.0080	0.0078	0.0075	0.0073	0.0071	0.0069	0.0068	0.0066	0.0064
-2.3	0.0107	0.0104	0.0102	0.0099	0.0096	0.0094	0.0091	0.0089	0.0087	0.0084
-2.2	0.0139	0.0136	0.0132	0.0129	0.0125	0.0122	0.0119	0.0116	0.0113	0.0110
-2.1	0.0179	0.0174	0.0170	0.0166	0.0162	0.0158	0.0154	0.0150	0.0146	0.0143
-2.0	0.0228	0.0222	0.0217	0.0212	0.0207	0.0202	0.0197	0.0192	0.0188	0.0183
-1.9	0.0287	0.0281	0.0274	0.0268	0.0262	0.0256	0.0250	0.0244	0.0239	0.0233
-1.8	0.0359	0.0351	0.0344	0.0336	0.0329	0.0322	0.0314	0.0307	0.0301	0.0294
-1.7	0.0446	0.0436	0.0427	0.0418	0.0409	0.0401	0.0392	0.0384	0.0375	0.0367
-1.6	0.0548	0.0537	0.0526	0.0516	0.0505	0.0495	0.0485	0.0475	0.0465	0.0455
-1.5	0.0668	0.0655	0.0643	0.0630	0.0618	0.0606	0.0594	0.0582	0.0571	0.0559
-1.4	0.0808	0.0793	0.0778	0.0764	0.0749	0.0735	0.0721	0.0708	0.0694	0.0681
-1.3	0.0968	0.0951	0.0934	0.0918	0.0901	0.0885	0.0869	0.0853	0.0838	0.0823
-1.2	0.1151	0.1131	0.1112	0.1093	0.1075	0.1056	0.1038	0.1020	0.1003	0.0985
-1.1	0.1357	0.1335	0.1314	0.1292	0.1271	0.1251	0.1230	0.1210	0.1190	0.1170
-1.0	0.1587	0.1562	0.1539	0.1515	0.1492	0.1469	0.1446	0.1423	0.1401	0.1379
-0.9	0.1841	0.1814	0.1788	0.1762	0.1736	0.1711	0.1685	0.1660	0.1635	0.1611
-0.8	0.2119	0.2090	0.2061	0.2033	0.2005	0.1977	0.1949	0.1922	0.1894	0.1867
-0.7	0.2420	0.2389	0.2358	0.2327	0.2296	0.2266	0.2236	0.2206	0.2177	0.2148

(continued)

<i>z</i>	0	1	2	3	4	5	6	7	8	9
−0.6	0.2743	0.2709	0.2676	0.2643	0.2611	0.2578	0.2546	0.2514	0.2483	0.2451
−0.5	0.3085	0.3050	0.3015	0.2981	0.2946	0.2912	0.2877	0.2843	0.2810	0.2776
−0.4	0.3446	0.3409	0.3372	0.3336	0.3300	0.3264	0.3228	0.3192	0.3156	0.3121
−0.3	0.3821	0.3783	0.3745	0.3707	0.3669	0.3632	0.3594	0.3557	0.3520	0.3483
−0.2	0.4207	0.4168	0.4129	0.4090	0.4052	0.4013	0.3974	0.3936	0.3897	0.3859
−0.1	0.4602	0.4562	0.4522	0.4483	0.4443	0.4404	0.4364	0.4325	0.4286	0.4247
−0.0	0.5000	0.4960	0.4920	0.4880	0.4840	0.4801	0.4761	0.4721	0.4681	0.4641
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7704	0.7734	0.7764	0.7794	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8907	0.8925	0.8944	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
1.7	0.9554	0.9564	0.9573	0.9582	0.9591	0.9599	0.9608	0.9616	0.9625	0.9633
1.8	0.9641	0.9649	0.9656	0.9664	0.9671	0.9678	0.9686	0.9693	0.9699	0.9706
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
2.1	0.9821	0.9826	0.9830	0.9834	0.9838	0.9842	0.9846	0.9850	0.9854	0.9857
2.2	0.9861	0.9864	0.9868	0.9871	0.9875	0.9878	0.9881	0.9884	0.9887	0.9890
2.3	0.9893	0.9896	0.9898	0.9901	0.9904	0.9906	0.9909	0.9911	0.9913	0.9916
2.4	0.9918	0.9920	0.9922	0.9925	0.9927	0.9929	0.9931	0.9932	0.9934	0.9936
2.5	0.9938	0.9940	0.9941	0.9943	0.9945	0.9946	0.9948	0.9949	0.9951	0.9952
2.6	0.9953	0.9955	0.9956	0.9957	0.9959	0.9960	0.9961	0.9962	0.9963	0.9964
2.7	0.9965	0.9966	0.9967	0.9968	0.9969	0.9970	0.9971	0.9972	0.9973	0.9974
2.8	0.9974	0.9975	0.9976	0.9977	0.9977	0.9978	0.9979	0.9979	0.9980	0.9981
2.9	0.9981	0.9982	0.9982	0.9983	0.9984	0.9984	0.9985	0.9985	0.9986	0.9986
3.0	0.9987	0.9987	0.9987	0.9988	0.9988	0.9989	0.9989	0.9989	0.9990	0.9990

Examples: If $Z \sim \text{Normal}(0,1)$, then $P(Z \leq -1.32) = .0934$ and $P(Z \leq 1.84) = .9671$.

Source: This table was generated using the Stata® function `normprob`.

Critical Values of the *t* Distribution

Significance Level						
1-Tailed: 2-Tailed:		.10 .20	.05 .10	.025 .05	.01 .02	.005 .01
D e g r e e s o f F r e e d o m	1	3.078	6.314	12.706	31.821	63.657
	2	1.886	2.920	4.303	6.965	9.925
	3	1.638	2.353	3.182	4.541	5.841
	4	1.533	2.132	2.776	3.747	4.604
	5	1.476	2.015	2.571	3.365	4.032
	6	1.440	1.943	2.447	3.143	3.707
	7	1.415	1.895	2.365	2.998	3.499
	8	1.397	1.860	2.306	2.896	3.355
	9	1.383	1.833	2.262	2.821	3.250
	10	1.372	1.812	2.228	2.764	3.169
	11	1.363	1.796	2.201	2.718	3.106
	12	1.356	1.782	2.179	2.681	3.055
	13	1.350	1.771	2.160	2.650	3.012
	14	1.345	1.761	2.145	2.624	2.977
	15	1.341	1.753	2.131	2.602	2.947
	16	1.337	1.746	2.120	2.583	2.921
	17	1.333	1.740	2.110	2.567	2.898
	18	1.330	1.734	2.101	2.552	2.878
	19	1.328	1.729	2.093	2.539	2.861
	20	1.325	1.725	2.086	2.528	2.845
	21	1.323	1.721	2.080	2.518	2.831
	22	1.321	1.717	2.074	2.508	2.819
	23	1.319	1.714	2.069	2.500	2.807
	24	1.318	1.711	2.064	2.492	2.797
	25	1.316	1.708	2.060	2.485	2.787
	26	1.315	1.706	2.056	2.479	2.779
	27	1.314	1.703	2.052	2.473	2.771
	28	1.313	1.701	2.048	2.467	2.763
	29	1.311	1.699	2.045	2.462	2.756
	30	1.310	1.697	2.042	2.457	2.750
	40	1.303	1.684	2.021	2.423	2.704
	60	1.296	1.671	2.000	2.390	2.660
	90	1.291	1.662	1.987	2.368	2.632
	120	1.289	1.658	1.980	2.358	2.617
	∞	1.282	1.645	1.960	2.326	2.576

Examples: The 1% critical value for a one-tailed test with 25 *df* is 2.485. The 5% critical value for a two-tailed test with large (> 120) *df* is 1.96.

Source: This table was generated using the Stata® function `invttail`.

F Values for $\alpha = 0.05$

d_2	d_1								
	1	2	3	4	5	6	7	8	9
1	161.4	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5
2	18.51	19.00	19.16	19.25	19.3	19.33	19.35	19.37	19.38
3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81
4	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00
5	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77
6	5.99	5.14	4.76	4.53	4.39	4.28	4.21	4.15	4.10
7	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68
8	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39
9	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18
10	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02
11	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90
12	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80
13	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71
14	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65
15	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59
16	4.49	3.63	3.24	3.01	2.85	2.74	2.66	2.59	2.54
17	4.45	3.59	3.20	2.96	2.81	2.70	2.61	2.55	2.49
18	4.41	3.55	3.16	2.93	2.77	2.66	2.58	2.51	2.46
19	4.38	3.52	3.13	2.90	2.74	2.63	2.54	2.48	2.42
20	4.35	3.49	3.10	2.87	2.71	2.60	2.51	2.45	2.39
21	4.32	3.47	3.07	2.84	2.68	2.57	2.49	2.42	2.37
22	4.30	3.44	3.05	2.82	2.66	2.55	2.46	2.40	2.34
23	4.28	3.42	3.03	2.80	2.64	2.53	2.44	2.37	2.32
24	4.26	3.40	3.01	2.78	2.62	2.51	2.42	2.36	2.30
25	4.24	3.39	2.99	2.76	2.60	2.49	2.40	2.34	2.28
26	4.23	3.37	2.98	2.74	2.59	2.47	2.39	2.32	2.27
27	4.21	3.35	2.96	2.73	2.57	2.46	2.37	2.31	2.25
28	4.20	3.34	2.95	2.71	2.56	2.45	2.36	2.29	2.24
29	4.18	3.33	2.93	2.70	2.55	2.43	2.35	2.28	2.22
30	4.17	3.32	2.92	2.69	2.53	2.42	2.33	2.27	2.21
40	4.08	3.23	2.84	2.61	2.45	2.34	2.25	2.18	2.12
60	4.00	3.15	2.76	2.53	2.37	2.25	2.17	2.10	2.04
120	3.92	3.07	2.68	2.45	2.29	2.17	2.09	2.02	1.96
inf	3.84	3.00	2.60	2.37	2.21	2.10	2.01	1.94	1.88