

ECON 6511: Advanced Applied Econometrics

Homework 2
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1. (Based on Wooldridge, Chapter 12, Problem 2) Let $\{e_t : t = -1, 0, 1, \dots\}$ be a sequence of independent, identically distributed random variables with mean zero and variance one. Define a stochastic process by $x_t = e_t - (1/2) e_{t-1} + (1/2) e_{t-2}, t = 1, 2, \dots$

(a) Find $E(x_t)$ and $\text{Var}(x_t)$. Do either of these depend on t ?

$$E(x_t) = E(e_t) - (1/2) E(e_{t-1}) + (1/2) E(e_{t-2})$$

$$E(x_t) = 0 - (1/2) * 0 + (1/2) * 0$$

$$E(x_t) = 0$$

As e_t is a sequence of independent stochastic variables, they are uncorrelated

$$\text{Var}(x_t) = \text{Var}(e_t) + (1/4) \text{Var}(e_{t-1}) + (1/4) \text{Var}(e_{t-2})$$

Also, $\text{Var}(e_t) = 1$ for all t . Therefore,

$$\text{Var}(x_t) = 1 + (1/4) + (1/4) = 3/2$$

(b) Show that $\text{Corr}(x_t, x_{t+1}) = -1/2$ and $\text{Corr}(x_t, x_{t+2}) = 1/3$. Note: For two random variables a and b , $\text{Corr}(a, b) = \text{Cov}(a, b) / \sigma_a \sigma_b$

$$\begin{aligned}
 \text{Cov}(x_t, x_{t+1}) &= E(x_t x_{t+1}) \\
 &= E[(e_t - (1/2)e_{t-1} + (1/2)e_{t-2})(e_{t+1} - (1/2)e_t + (1/2)e_{t-1})] \\
 &= E(e_t * e_{t+1}) - (1/2)E(e_t^2) + (1/2)E(e_t * e_{t-1}) - (1/2)E(e_{t-1} * e_{t+1}) + (1/4)E(e_{t-1} * e_t) - \\
 &\quad (1/4)E(e_{t-1}^2) + (1/2)E(e_{t-2} * e_{t+1}) - (1/4)E(e_{t-2} * e_t) + (1/4)E(e_{t-2} * e_{t-1}) \\
 &= - (1/2)E(e_t^2) - (1/4)E(e_t^2) \\
 &= - (1/2) * 1 - (1/4) * 1 \\
 &= -3/4
 \end{aligned}$$

$$\begin{aligned}
 \text{Corr}(a, b) &= \text{Cov}(a, b) / (\sigma_a * \sigma_b) \\
 \text{Corr}(x_t, x_{t+1}) &= - (3/4) / (3/2) \\
 &= -1/2
 \end{aligned}$$

$$E(e_t^2) = 1/2$$

$$\text{Corr}(x_t, x_{t+2}) = (1/2) / (3/2) = 1/3$$

(c) What is $\text{Corr}(x_t, x_{t+h})$ for $h > 2$?

$$\text{Corr}(x_t, x_{t+h}) \text{ for } h > 2 = 0$$

(d) Is $\{x_t\}$ covariance stationary? Is it weakly dependent?

x_t is weakly dependent because $\text{Correlation}(x_t, x_{t+h}) < 1$ i.e. observations more than two periods apart are uncorrelated

2. (Based on Wooldridge, Chapter 12, Problem 5) For the U.S. economy, let gprice denote the monthly growth in the overall price level and let gwage be the monthly growth in hourly wages. These are both obtained as differences of logarithms: $gprice = \Delta \log(\text{price})$ and $gwage = \Delta \log(\text{wage})$.

(a) Using the monthly data in WAGEPRC.dta, estimate a distributed lag model: $gprice = \alpha + \beta_1 gwage_t + \beta_2 gwage_{t-1} + \beta_3 gwage_{t-2} + \beta_4 gwage_{t-3} + \beta_5 gwage_{t-4} + \beta_6 gwage_{t-5} + \beta_7 gwage_{t-6} + \beta_8 gwage_{t-7} + \beta_9 gwage_{t-8} + \beta_{10} gwage_{t-9} + \beta_{11} gwage_{t-10} + \beta_{12} gwage_{t-11} + \beta_{13} gwage_{t-12} + u_t$ and sketch the estimated lag distribution. At what lag is the effect of gwage on gprice largest? Which lag has the smallest coefficient?

Using stata:

```
. use "/Users/surbhiasati/Desktop/Econometrics/Homework2/WAGEPRC.DTA"

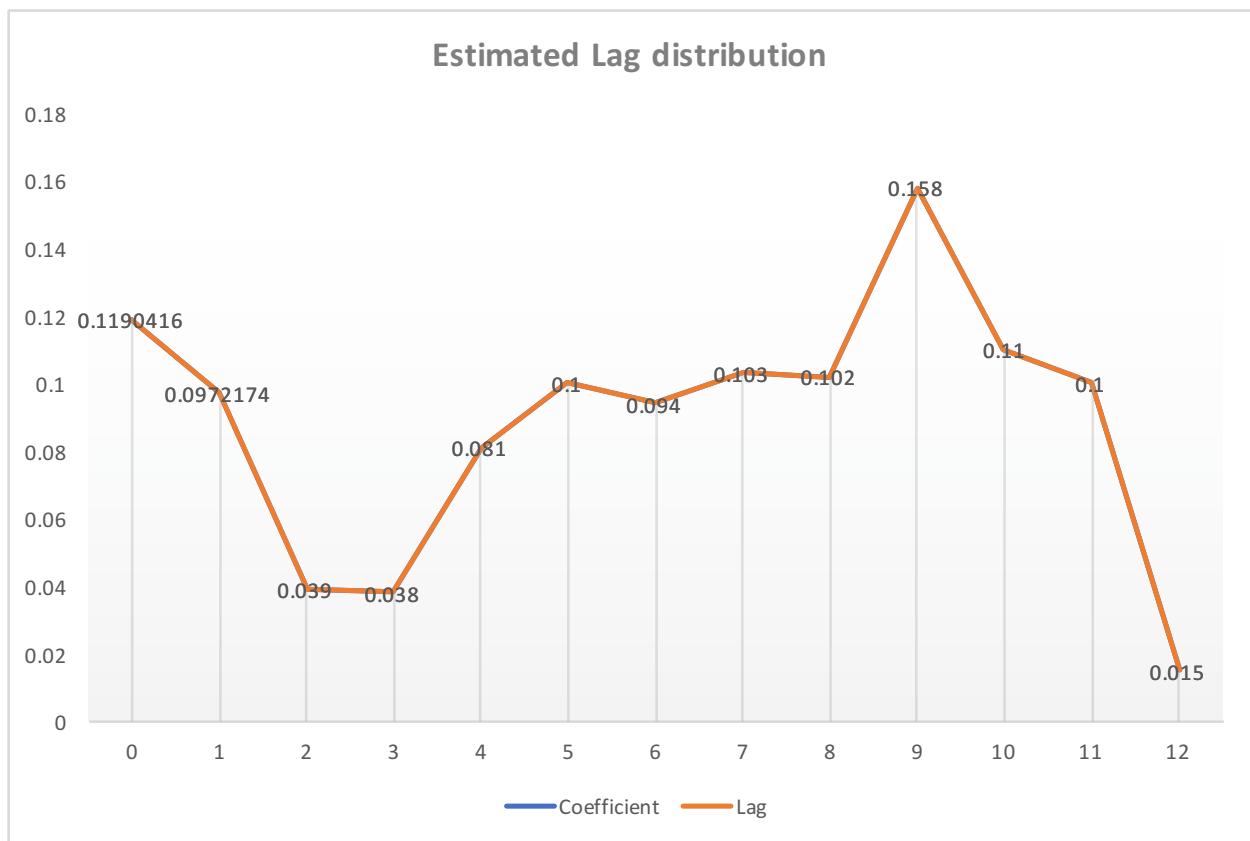
. reg gprice gwage gwage_1 gwage_2 gwage_3 gwage_4 gwage_5 gwage_6 gwage_7 gwage_8 gwage_
> 9 gwage_10 gwage_11 gwage_12
```

Source	SS	df	MS	Number of obs	=	273
Model	.000981458	13	.000075497	F(13, 259)	=	9.25
Residual	.002113658	259	8.1608e-06	Prob > F	=	0.0000
				R-squared	=	0.3171
				Adj R-squared	=	0.2828
Total	.003095116	272	.000011379	Root MSE	=	.00286

gprice	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gwage	.1190416	.0517725	2.30	0.022	.0170929	.2209903
gwage_1	.0972174	.0390409	2.49	0.013	.0203393	.1740954
gwage_2	.0399518	.0390717	1.02	0.307	-.0369869	.1168905
gwage_3	.0382652	.0391513	0.98	0.329	-.0388301	.1153605
gwage_4	.0813362	.0393483	2.07	0.040	.0038528	.1588195
gwage_5	.106852	.0391937	2.73	0.007	.0296731	.1840308
gwage_6	.0949731	.0392186	2.42	0.016	.0177451	.1722011
gwage_7	.1037922	.0393788	2.64	0.009	.0262488	.1813355
gwage_8	.1025629	.0394884	2.60	0.010	.0248037	.180322
gwage_9	.1585079	.0393341	4.03	0.000	.0810526	.2359632
gwage_10	.1104412	.0392229	2.82	0.005	.0332049	.1876776
gwage_11	.1033206	.0394388	2.62	0.009	.0256591	.180982
gwage_12	.0156575	.0518343	0.30	0.763	-.0864128	.1177278
_cons	-.0009296	.0005662	-1.64	0.102	-.0020445	.0001853

Model:

$$\text{gprice} = - .00 + .11 \text{ gwage}_t + .09 \text{ gwage}_{t-1} + .04 \text{ gwage}_{t-2} + .04 \text{ gwage}_{t-3} + .08 \text{ gwage}_{t-4} + .10 \text{ gwage}_{t-5} + .09 \text{ gwage}_{t-6} + .10 \text{ gwage}_{t-7} + .10 \text{ gwage}_{t-8} + .15 \text{ gwage}_{t-9} + .11 \text{ gwage}_{t-10} + .10 \text{ gwage}_{t-11} + .01 \text{ gwage}_{t-12}$$



The largest effect is at the ninth lag 0.158. The smallest effect is at the twelfth lag 0.015.

(b) For which lags are the t statistics less than two?

The t statistics is less than two at the second ($t = 1.02$), third ($t = 0.98$) and twelfth ($t = 0.30$) lag.

(c) What is the estimated long-run propensity? Is it much different than one? Explain what the LRP tells us in this example.

The estimated long-run propensity is the sum of the lag coefficients from zero through twelve. In a distributed lag model, the eventual change in the dependent variable given a permanent, one-unit increase in the independent variable.

$$\text{LRP} = 1.172$$

As, $\text{LRP} > 1$, is a cumulative effect of wage on price. This difference indicates error.

(d) What regression would you run to obtain the standard error of the LRP directly?

Assuming LRP coefficients = $q_1, q_2, q_3, \dots, q_{12}$

$$\text{LRP} = q_0, q_1, q_2, q_3, \dots, q_{12}$$

$$q_0 = \text{LRP} - q_1 - q_2 - q_3 - \dots - q_{12}$$

Plug into the estimated model:

$$\begin{aligned} \text{gprice}_t = & \alpha + (\text{LRP} - q_1 - q_2 - q_3 - \dots - q_{12}) \text{gwage}_t + q_1 \text{gwage}_{t-1} + q_2 \text{gwage}_{t-2} + q_3 \text{gwage}_{t-3} + \\ & q_4 \text{gwage}_{t-4} + q_5 \text{gwage}_{t-5} + q_6 \text{gwage}_{t-6} + q_7 \text{gwage}_{t-7} + q_8 \text{gwage}_{t-8} + q_9 \text{gwage}_{t-9} + \\ & q_{10} \text{gwage}_{t-10} + q_{11} \text{gwage}_{t-11} + q_{12} \text{gwage}_{t-12} + u_t \end{aligned}$$

3. Use the data in PHILLIPS.dta for this exercise.

(a) Estimate an AR(1) model for the unemployment rate. Use this equation to predict the unemployment rate for 2004. Compare this with the actual unemployment rate for 2004. (You can find this information in a recent Economic Report of the President.)

Using stata:

The screenshot shows the Stata/IC 15.1 interface. The main window displays the results of a regression command: `. reg unem unem_1`. The results are presented in two tables: a summary table and a coefficient table.

Source	SS	df	MS	Number of obs	=	55
Model	68.9295284	1	68.9295284	F(1, 53)	=	69.12
Residual	52.8515619	53	.99719928	Prob > F	=	0.0000
Total	121.78109	54	2.25520538	R-squared	=	0.5660
				Adj R-squared	=	0.5578
				Root MSE	=	.9986

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
unem_1	.7423824	.0892927	8.31	0.000	.5632839 .9214809
_cons	1.489685	.5202033	2.86	0.006	.446289 2.53308

Below the tables, the command `. use "/Users/surbhiasati/Desktop/Econometrics/Homework2/phillips.dta", clear` is entered in the Command window.

The right-hand side of the interface shows the Search Help panel and the Properties panel. The Properties panel displays the following information:

- Variables:**
 - Name: year
 - Label: 1948 through
 - Type: int
 - Format: %9.0g
 - Value label: Notes
- Data:**
 - Filename: phillips.dta
 - Label: Notes
 - Variables: 7
 - Observations: 56
 - Size: 1.42K
 - Memory: 64M

The estimated AR model is:

$$\text{Unem}_t = 1.489 + 0.742 \text{Unem}_{t-1}$$

The unemployment rate in 2003 was 6. The predicted unemployment rate for 2004 is

$$1.49 + 0.742(6) = 5.94$$

According to the Economic Report of the President, the unemployment rate was 5.5. Therefore, the equation over predicts the 2004 unemployment rate by 0.44 percentage.

(b) Add a lag of inflation to the AR(1) model from part (a). Is inf_{t-1} statistically significant?

```
. reg unem unem_1 inf_1
```

Source	SS	df	MS	Number of obs	=	55
Model	84.7924146	2	42.3962073	F(2, 52)	=	59.60
Residual	36.9886756	52	.711320685	Prob > F	=	0.0000
				R-squared	=	0.6963
				Adj R-squared	=	0.6846
Total	121.78109	54	2.25520538	Root MSE	=	.8434

unem	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
unem_1	.6494787	.0779388	8.33	0.000	.4930831	.8058743
inf_1	.1830086	.0387537	4.72	0.000	.1052436	.2607735
_cons	1.296416	.4412563	2.94	0.005	.4109709	2.181862

Adding inf_1 to the model we get:

$$\text{Unem}_t = 1.296 + 0.649 \text{Unem}_{t-1} + 0.183 \text{inf}_{t-1}$$

With p value = 0 (less than 0.05) and a t- statistic of 4.72, Lagged inflation is statistically significant.

(c) Use the equation from part (b) to predict the unemployment rate for 2004. Is the result better or worse than in the model from part (a)?

Equation in part (b):

$$\text{Unem}_t = 1.296 + 0.649 \text{Unem}_{t-1} + 0.183 \text{inf}_{t-1}$$

Inflation rate for 2003 = 2.3

Prediction of unemployment in 2003:

$$\text{Unem}_t = 1.296 + 0.649(6) + 0.183(2.3) = 5.6$$

This is very close to the actual rate of 5.5, and it is certainly better than the predication from part (a) that is 5.94

4. (Wooldridge, Chapter 12, C2)

(a) Using the data in WAGEPRC.dta, estimate the distributed lag model from Problem 2 above. Regress \hat{u}_t on \hat{u}_{t-1} to test for AR(1) serial correlation.

Estimated distributed lag model:

$$\text{gprice}_t = -.00 + .11 \text{gwage}_t + .09 \text{gwage}_{t-1} + .04 \text{gwage}_{t-2} + .04 \text{gwage}_{t-3} + .08 \text{gwage}_{t-4} + .10 \text{gwage}_{t-5} + .09 \text{gwage}_{t-6} + .10 \text{gwage}_{t-7} + .10 \text{gwage}_{t-8} + .15 \text{gwage}_{t-9} + .11 \text{gwage}_{t-10} + .10 \text{gwage}_{t-11} + .01 \text{gwage}_{t-12}$$

Regression \hat{u}_t on \hat{u}_{t-1} gives $t = 9.60$, AR(1) correlation is positive

(b) Reestimate the model using iterated Cochrane-Orcutt (CO) estimation. What is your new estimate of the long-run propensity?

The new estimate of the long-run propensity is $= 1.110$

(c) Using iterated CO, find the standard error for the LRP. Determine whether the estimated LRP is statistically different from one at the 5% level.

$$\text{gpricet} = 0 + 0 \text{gwaget} + 1 (\text{gwaget-1} - \text{gwaget}) + 2 (\text{gwaget-2} - \text{gwaget}) + + 12 (\text{gwaget-12} - \text{gwaget}) + u_t$$

Estimating this equation by CO gives $\text{LRP} = 1.110$

The t statistic for testing is $(1.110 - 1)/.191 = 0.57$. This is greater than 5%. Therefore, LRP is not significant.