

# 第三次作业

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### 2.12 题

(1)

解: 该分布是区间  $(\theta - 1, \theta + 1)$  上的均匀分布,  $\theta$  是位置参数族, 从而  $\theta$  的无信息先验为  $\pi(\theta) \equiv 1$

(2)

解: 该分布的密度函数为

$$\begin{aligned} p(x|\beta) &= \frac{\beta}{\pi(x^2 + \beta^2)} \\ &= \frac{1}{\pi\beta[(\frac{x}{\beta})^2 + 1]} \\ &= \beta^{-1}\varphi(\frac{x}{\beta}) \end{aligned}$$

$\beta$  是刻度参数, 从而无信息先验为  $\pi(\beta) = \frac{1}{\beta}$

(3)

**解:** 考虑  $(\mu, \sigma^2)$  整体, 既不是位置参数, 也不是刻度参数, 如果考虑两个参数独立的情况, 此时无信息先验为  $\pi(\mu, \sigma^2) = \frac{1}{\sigma^2}$

(4)

**解:** 帕累托分布  $P(x_0, \alpha)$  的密度函数为

$$\pi(x|x_0) = \begin{cases} 0 & x < x_0 \\ \frac{\alpha x_0^\alpha}{x^{\alpha+1}} & x \geq x_0 \end{cases}$$

也即

$$\pi(x|x_0) = \begin{cases} 0 & \frac{x}{x_0} < 1 \\ \frac{\alpha}{x_0(\frac{x}{x_0})^{\alpha+1}} & \frac{x}{x_0} \geq 1 \end{cases}$$

可以看出  $x_0$  为刻度参数, 从而无信息先验为  $\pi(x_0) = \frac{1}{x_0}$

## 2.13 题

(1)

**解:** 泊松分布  $P(\lambda)$  的密度函数  $f(x, \lambda) = \frac{\lambda^x}{x!} e^{-\lambda}$  使用 Jeffreys 方法写出对数似然函数

$$\begin{aligned} l(\lambda|x) &= \ln[f(x|\lambda)] \\ &= -\lambda + x \ln \lambda - \ln(x!) \end{aligned}$$

求二阶导  $\frac{\partial^2 l}{\partial \lambda^2} = -\frac{x}{\lambda^2}$

求费希尔信息量

$$I(\lambda) = E_{X|\lambda} \left( \frac{X}{\lambda^2} \right) = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda}$$

从而 Jeffrey 先验为  $\pi(\lambda) = \sqrt{\frac{1}{\lambda}}$  这是一个广义先验分布

## (2)

**解:** 负二项分布  $Nb(r, \theta)$  的密度函数  $f(x, \theta) = \binom{x-1}{r-1} \theta^r (1-\theta)^{x-r}$  则对数似然函数为

$$\begin{aligned} l(\theta|x) &= \ln[f(x|\theta)] \\ &= \ln\binom{x-1}{r-1} + r\ln\theta + (x-r)\ln(1-\theta) \end{aligned}$$

求二阶导  $\frac{\partial^2 l}{\partial \theta^2} = -\frac{r}{\theta^2} - \frac{x-r}{(1-\theta)^2}$

求费希尔信息量

$$\begin{aligned} I(\theta) &= E_{X|\theta} \left( \frac{r}{\theta^2} + \frac{x-r}{(1-\theta)^2} \right) \\ &= \frac{r}{\theta^2} + \frac{\frac{r}{\theta} - r}{(1-\theta)^2} \\ &= \frac{r}{\theta^2(1-\theta)} \end{aligned}$$

$\theta$  的先验分布  $\pi(\theta) \propto \frac{1}{\theta(1-\theta)^{\frac{1}{2}}}$  ( $0 < \theta < 1$ )

## (3)

**解:** 指数分布  $Exp(\frac{1}{\lambda})$  的密度函数  $f(x, \lambda) = \frac{1}{\lambda} e^{-\frac{x}{\lambda}}$  从而对数似然函数为

$$\begin{aligned} l(\lambda|x) &= \ln[f(x|\lambda)] \\ &= -\ln\lambda - \frac{x}{\lambda} \end{aligned}$$

求二阶导  $\frac{\partial^2 l}{\partial \lambda^2} = \frac{1}{\lambda^2} - 2\frac{x}{\lambda^3}$

求费希尔信息量

$$\begin{aligned} I(\lambda) &= E_{X|\lambda} \left( -\frac{1}{\lambda^2} + 2\frac{x}{\lambda^3} \right) \\ &= -\frac{1}{\lambda^2} + 2\frac{\lambda}{\lambda^3} \\ &= \frac{1}{\lambda^2} \end{aligned}$$

从而  $\lambda$  的先验分布为  $\pi(\lambda) = \frac{1}{\lambda}, \lambda > 0$  这是一个广义先验

(4)

解: 伽马分布  $\Gamma(\alpha, \lambda)$  密度函数为  $f(x|\lambda) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$  对数似然为

$$\begin{aligned} l(\lambda|x) &= \ln[f(x|\lambda)] \\ &= \alpha \ln \lambda - \lambda x - \ln(\Gamma(\alpha)) + (\alpha - 1) \ln x \end{aligned}$$

求二阶导为  $\frac{\partial^2 l}{\partial \lambda^2} = -\frac{\alpha}{\lambda^2}$  从而

$$\begin{aligned} I(\lambda) &= E_{X|\lambda} \left( \frac{\alpha}{\lambda^2} \right) \\ &= \frac{\alpha}{\lambda^2} \end{aligned}$$

从而  $\lambda$  的先验分布为  $\pi(\lambda) = \frac{1}{\lambda}, \lambda > 0$  这是一个广义先验

(5)

解: 多项分布  $M(n, p)$  密度函数为  $f(x|p) = \frac{n!}{x_1! \dots x_k!} p_1^{x_1} \dots p_k^{x_k}$  其中  $\sum_{i=1}^k x_i = n, \sum_{i=1}^k p_i = 1$  对数似然函数为

$$\begin{aligned} l(p|x) &= \ln[f(x|p)] \\ &= \sum_{i=1}^{k-1} x_i \ln p_i + x_k \ln \left( 1 - \sum_{i=1}^{k-1} p_i \right) + \ln \left( \frac{n!}{x_1! \dots x_k!} \right) \end{aligned}$$

对每个  $i, j$  求偏导导结果为

$$\frac{\partial^2 l}{\partial p_i \partial p_j} = \begin{cases} -\frac{x_k}{p_k^2} & i \neq j \\ -\frac{x_i}{p_i^2} - \frac{x_k}{p_k^2} & i = j \end{cases}$$

从而有

$$I_{ij}(p) = \begin{cases} \frac{n}{p_i} + \frac{n}{p_k}, i = j \\ \frac{n}{p_k}, i \neq j \end{cases}$$

其中  $i, j$  在  $1, \dots, k-1$  中取值算得  $\det(I) = n^{k-1} \prod_{i=1}^k \frac{1}{p_i}$  从而  $p$  的先验分布密度为  $\pi(p) \propto \sqrt{\prod_{i=1}^k \frac{1}{p_i}}$  为迪利克雷分布  $D(\frac{1}{2}, \dots, \frac{1}{2})$  这里有  $k$  项。

## 2.20 题

证明: 考虑先验分布取  $h(\theta)$  的后验分布

$$\begin{aligned}\pi(\theta|x) &\propto h(\theta)f(x|\theta) \\ &\propto Ae^{k_1 a(\theta)+k_2 c(\theta)}e^{a(\theta)b(x)+c(\theta)+d(x)} \\ &\propto e^{(k_1+b(x))a(\theta)+(k_2+1)c(\theta)+d(x)}\end{aligned}$$

其中  $k_1 + b(x), d(x)$  为  $x$  的函数  $a(\theta), (k_2 + 1)c(\theta)$  是  $\theta$  的函数, 其后验仍在原来的分布族中, 从而  $h(\theta)$  为  $\theta$  的共轭先验分布。

## 2.24 题

证明: 考虑分成两段, 在  $a < \theta < z$  上, 令  $g_1(\theta) = 1$ , 有  $E^\pi(g_1(\theta)) = \frac{1}{2}$  从而

$$\begin{aligned}\tilde{\pi}(\theta) &= \frac{\pi_0(\theta)e^{\lambda_1 g_1(\theta)}}{\int_{\Theta} \pi_0(\theta)e^{\lambda_1 g_1(\theta)} d\theta} \\ &= \frac{\frac{1}{\theta}e^{\lambda_1}}{\int_{\Theta} \frac{1}{\theta}e^{\lambda_1} d\theta} \\ &= a_1 \frac{\frac{1}{\theta}}{\int_a^z \frac{1}{\theta} d\theta} \\ &= a_1 \frac{1}{\theta} \ln\left(\frac{z}{a}\right)^{-1}\end{aligned}$$

其中  $a_1$  需要满足

$$\int_a^z a_1 \frac{1}{\theta} \ln\left(\frac{z}{a}\right)^{-1} d\theta = \frac{1}{2}$$

从而  $a_1 = \frac{1}{2}$  从而  $a < \theta < z$  时, 最大熵先验为  $\pi(\theta) = \frac{1}{\theta}(2\ln\frac{z}{a})^{-1}$  同理考虑  $z < \theta < b$  在其上令  $g_2(\theta) = 1$ , 有  $E^\pi(g_2(\theta)) = \frac{1}{2}$  相同的步骤得到  $z < \theta < b$  时, 最大熵先验为  $\pi(\theta) = \frac{1}{\theta}(2\ln\frac{b}{z})^{-1}$  综上

$$\pi(\theta) = \begin{cases} \frac{1}{\theta}(2\ln\frac{z}{a})^{-1} & 0 < a < \theta < z \\ \frac{1}{\theta}(2\ln\frac{b}{z})^{-1} & z < \theta < b \end{cases}$$

## 2.25 题

证明: 由  $\varphi_1 = \frac{\mu_2}{\mu_1}, \varphi_2 = \mu_1\mu_2$  解出  $\mu_1 = \sqrt{\frac{\varphi_2}{\varphi_1}}, \mu_2 = \sqrt{\varphi_1\varphi_2}$  变换的雅各比矩阵为

$$\frac{\partial(\varphi_1, \varphi_2)}{\partial(\mu_1, \mu_2)} = \begin{pmatrix} -\frac{1}{2}\varphi_1^{-\frac{3}{2}}\varphi_2^{\frac{1}{2}} & \frac{1}{2}\varphi_1^{-\frac{1}{2}}\varphi_2^{-\frac{1}{2}} \\ \frac{1}{2}\varphi_1^{-\frac{1}{2}}\varphi_2^{\frac{1}{2}} & \frac{1}{2}\varphi_1^{\frac{1}{2}}\varphi_2^{-\frac{1}{2}} \end{pmatrix}$$

从而对数似然

$$\begin{aligned} l(\varphi_1, \varphi_2|x) &= \ln(f(\varphi_1, \varphi_2|x)|\det(\frac{\partial(\varphi_1, \varphi_2)}{\partial(\mu_1, \mu_2)})) \\ &= -\ln\mu_1\mu_2 - \frac{x_1}{\mu_1} - \frac{x_2}{\mu_2} + \ln(\frac{1}{2}\varphi_1^{-1}) \\ &= -\ln\varphi_1 - \ln\varphi_2 - \frac{x_1}{\sqrt{\frac{\varphi_2}{\varphi_1}}} - \frac{x_2}{\sqrt{\varphi_1\varphi_2}} - \ln 2 \end{aligned}$$

$\varphi_2$  为多余参数, 求偏导得到

$$\frac{\partial^2 l}{\partial \varphi_1^2} = \varphi_1^{-2} + \frac{1}{4}x_1\varphi_1^{-\frac{3}{2}}\varphi_2^{-\frac{1}{2}} - \frac{3}{4}x_2\varphi_1^{-\frac{5}{2}}\varphi_2^{-\frac{1}{2}}$$

$$\frac{\partial^2 l}{\partial \varphi_2^2} = \varphi_2^{-2} - \frac{3}{4}(x_1\varphi_1^{\frac{1}{2}} + x_2\varphi_1^{-\frac{1}{2}})\varphi_2^{-\frac{5}{2}}$$

$$\frac{\partial^2 l}{\partial \varphi_2 \partial \varphi_1} = \frac{1}{4}(x_1\varphi_1^{-\frac{1}{2}} - x_2\varphi_1^{-\frac{3}{2}})\varphi_2^{-\frac{5}{2}}$$

所以

$$E^{X|\varphi}(-\frac{\partial^2 l}{\partial \varphi_2^2}) = \frac{1}{2}\varphi_2^{-2}$$

$$E^{X|\varphi}(-\frac{\partial^2 l}{\partial \varphi_1^2}) = -\frac{1}{2}\varphi_1^{-2}$$

$$E^{X|\varphi}(-\frac{\partial^2 l}{\partial \varphi_2 \partial \varphi_1}) = 0$$

可得到费希尔信息阵是对角阵

$$I(\varphi_1, \varphi_2) = \begin{pmatrix} -\frac{1}{2}\varphi_1^{-2} & 0 \\ 0 & \frac{1}{2}\varphi_2^{-2} \end{pmatrix}$$

从而可以得到条件 reference 先验为  $\pi(\varphi_2|\varphi_1) \propto \sqrt{\varphi_2^{-2}} = \varphi_2^{-1}$  这是一个广义先验

选择参数空间  $\Phi = R_+ \times R_+$  上的单调增子集  $\Omega_i = L_i \times S_i$  则  $\Omega_{i,\varphi_1} = L_I = [l_{i1}, l_{i2}]$

$$K_i(\varphi_1) = \frac{1}{\int_{\Omega_{i,\varphi_1}} \pi(\varphi_2|\varphi_1) d\varphi_2} = \frac{1}{\ln l_{i2} - \ln l_{i1}}$$

$$\pi_i(\varphi_2|\varphi_1) = K_i(\varphi_1) \pi(\varphi_2|\varphi_1) I_{[l_{i1}, l_{i2}]} = \frac{1}{(\ln l_{i2} - \ln l_{i1}) \varphi_2}$$

求边缘 reference 先验

$$\begin{aligned} \pi_i(\varphi_1) &= \exp\left(\frac{1}{2} \int_{\Omega_{i,\varphi_1}} \pi_i(\varphi_2|\varphi_1) \ln \frac{|I(\varphi_1, \varphi_2)|}{|I_{22}(\varphi_1, \varphi_2)|} d\varphi_2\right) \\ &= \exp\left(\frac{1}{2} \int_{\Omega_{i,\varphi_1}} \frac{1}{(\ln l_{i2} - \ln l_{i1}) \varphi_2} \ln \frac{1}{2} \varphi_1^{-2} d\varphi_2\right) \\ &= A \varphi_1^{-1} \end{aligned}$$

其中 A 为常数, 下面取  $\varphi_{10} = 1$ , 求极限可得

$$\begin{aligned} \pi(\varphi_1, \varphi_2) &= \lim_{i \rightarrow \infty} \frac{K_i(\varphi_1) \pi_i(\varphi_1)}{K_i(\varphi_{10}) \pi_i(\varphi_{10})} \pi(\varphi_2, \varphi_1) \\ &= \lim_{i \rightarrow \infty} \frac{\frac{1}{\ln l_{i2} - \ln l_{i1}} A \varphi_1^{-1}}{\frac{1}{\ln l_{i2} - \ln l_{i1}} A} \varphi_2^{-1} \\ &= (\varphi_1 \varphi_2)^{-1} \end{aligned}$$

从而我们得到参数  $(\varphi_1, \varphi_2)$  的 reference 先验为  $\pi(\varphi_1, \varphi_2) = (\varphi_1 \varphi_2)^{-1}$