

Stochastic

Ashan Jayamal

2024-08-18



# Contents

<b>1</b>	<b>Premilaries</b>	<b>5</b>
1.1	Brownian Motion . . . . .	5
1.2	Definition of Brownian Motion . . . . .	5
1.3	Simple Properties of Brownian Motion . . . . .	6
1.4	Wiener Integral . . . . .	6
<b>2</b>	<b>Introduction</b>	<b>7</b>
2.1	Events and Probability . . . . .	7
2.2	Random Variables . . . . .	12
2.3	Conditional Probability and Independence . . . . .	20
<b>3</b>	<b>Random Walk to Brownier Motion.</b>	<b>25</b>
<b>4</b>	<b>Conditional Expectation</b>	<b>29</b>
4.1	Conditioning on an Event . . . . .	29
4.2	Conditioning on a Discrete Random Variable . . . . .	32
4.3	Conditioning on an Arbitrary Random Variable . . . . .	41
<b>5</b>	<b>Les</b>	<b>43</b>
5.1	Integrals . . . . .	43
5.2	Random Walks . . . . .	43
<b>6</b>	<b>Book</b>	<b>45</b>



# Chapter 1

## Premilaries

**Definition 1.1.** Consider a set  $X$ . An  $\sigma$ -algebra  $\mathcal{F}$  of subsets of  $X$  is a collection  $\mathcal{F}$  of subsets of  $X$  satisfying the following conditions:

- $\emptyset \in \mathcal{F}$
- If  $B \in \mathcal{F}$ , then its complement  $B^c$  is also in  $\mathcal{F}$
- If  $B_1, B_2, \dots$  is a countable collection of sets in  $\mathcal{F}$ , then their union  $\bigcup_{n=1}^{\infty} B_n$  is also in  $\mathcal{F}$ .

### 1.1 Brownian Motion

Let  $(\Omega, \mathcal{F}, P)$  be a probability space. A stochastic process is a measurable function  $X(t, \omega)$  defined on the product space  $[0, \infty) \times \Omega$ . In particular:

- For each  $t$ ,  $X(t, \cdot)$  is a random variable.
- For each  $\omega$ ,  $X(\cdot, \omega)$  is a measurable function (called a sample path).

For convenience, the random variable  $X(t, \cdot)$  will be written as  $X(t)$  or  $X_t$ . Thus, a stochastic process  $X(t, \omega)$  can also be expressed as  $X(t)(\omega)$  or simply as  $X(t)$  or  $X_t$ .

### 1.2 Definition of Brownian Motion

**Definition 1.2.** A stochastic process  $B(t, \omega)$  is called a Brownian motion if it satisfies the following conditions:

1.  $P(\{\omega : B(0, \omega) = 0\}) = 1$ .

2. For any  $0 \leq s < t$ , the random variable  $B(t) - B(s)$  is normally distributed with mean 0 and variance  $t - s$ , i.e., for any  $a < b$ ,

$$P(a \leq B(t) - B(s) \leq b) = \frac{1}{\sqrt{2\pi(t-s)}} \int_a^b e^{-x^2/2(t-s)} dx.$$

3.  $B(t, \omega)$  has independent increments, i.e., for any  $0 \leq t_1 < t_2 < \dots < t_n$ , the random variables

$$B(t_1), B(t_2) - B(t_1), \dots, B(t_n) - B(t_{n-1})$$

are independent.

4. Almost all sample paths of  $B(t, \omega)$  are continuous functions, i.e.,

$$P(\{\omega : B(\cdot, \omega) \text{ is continuous}\}) = 1$$

### 1.3 Simple Properties of Brownian Motion

Let  $B(t)$  be a fixed Brownian motion. We give below some simple properties that follow directly from the definition of Brownian motion.

**Proposition 1.1.** *For any  $t > 0$ ,  $B(t)$  is normally distributed with mean 0 and variance  $t$ . For any  $s, t \geq 0$ , we have  $\mathbb{E}[B(s)B(t)] = \min\{s, t\}$ .*

*Remark.* Regarding Definition @ref(def:Brownian\_Motion), it can be proved that condition (2) and  $\mathbb{E}[B(s)B(t)] = \min\{s, t\}$  imply condition (3).

*Proof.* By condition (1), we have  $B(t) = B(t) - B(0)$  and so the first assertion follows from condition (2). With out loss of generlity, assume that  $s < t$ .

$$\mathbb{E}[B(s)B(t)] = \mathbb{E}[B(s)(B(t) - B(s)) + B(s)^2] = 0 + s = s$$

which is equal to  $\min\{s, t\}$ . □

**Proposition 1.2** (Translation Invariance). *For a fixed  $t_0 \geq 0$ , the stochastic process  $B(t) = B(t + t_0) - B(t_0)$  is also a Brownian motion.*

**Proposition 1.3** (Scaling invariance). *For any real number  $\lambda > 0$ , the stochastic process  $B(t) = \frac{B(\lambda t)}{\sqrt{\lambda}}$  is also a Brownian motion.*

### 1.4 Wiener Integral

# Chapter 2

## Introduction

### 2.1 Events and Probability

**Definition 2.1.** Let  $\Omega$  be a non-empty set. A  **$\sigma$ -field**  $\mathcal{F}$  on  $\Omega$  is a family of subsets of  $\Omega$  such that:

- The empty set  $\emptyset$  belongs to  $\mathcal{F}$ ;
- If  $A$  belongs to  $\mathcal{F}$ , then so does the complement  $\Omega \setminus A$ ;
- If  $A_1, A_2, \dots$  is a sequence of sets in  $\mathcal{F}$ , then their union  $A_1 \cup A_2 \cup \dots$  also belongs to  $\mathcal{F}$ .

**Example 2.1.** The family of Borel sets  $\mathcal{F} = B(\mathbb{R})$  is a  $\sigma$ -field on  $\mathbb{R}$ . We recall that  $B(\mathbb{R})$  is the smallest  $\sigma$ -field containing all intervals in  $\mathbb{R}$ .

**Definition 2.2.** Let  $\mathcal{F}$  be a  $\sigma$ -field on  $\Omega$ . A probability measure  $P$  is a function  $P : \mathcal{F} \rightarrow [0, 1]$  such that

1.  $P(\Omega) = 1$ ;
2. if  $A_1, A_2, \dots$  are pairwise disjoint sets (that is,  $A_i \cap A_j = \emptyset$  for  $i \neq j$ ) belonging to  $\mathcal{F}$ , then

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i);$$

- The triple  $(\Omega, \mathcal{F}, P)$  is called a probability space. \
- The sets belonging to  $\mathcal{F}$  are called events. \
- An event  $A$  is said to occur almost surely (a.s.) whenever  $P(A) = 1$ .

**Example 2.2.** Let consider,

- $\Omega = [0, 1]$  with the
- $\sigma$ -field  $= \mathcal{F} = \mathcal{B}([0, 1])$  of Borel sets  $B \subseteq [0, 1]$ , and
- Lebesgue measure  $P = \text{Leb}$  on  $[0, 1]$ .

Then  $(\Omega, \mathcal{F}, P)$  is a probability space.

Recall that Leb is the unique measure defined on Borel sets such that

$$\text{Leb}[a, b] = b - a$$

for any interval  $[a, b]$ . (In fact, Leb can be extended to a larger  $\sigma$ -field, but we shall need Borel sets only.)

**Exercise 2.1.** Show that if  $A_1, A_2, \dots$  is an expanding sequence of events, that is

$$A_1 \subseteq A_2 \subseteq A_3 \subseteq \dots$$

then

$$P\left(\bigcup_{n=1}^{\infty} A_n\right) = \lim_{n \rightarrow \infty} P(A_n).$$

Similarly, if  $A_1, A_2, \dots$  is a contracting sequence of events, that is,

$$A_1 \supseteq A_2 \supseteq A_3 \supseteq \dots$$

then

$$P\left(\bigcap_{n=1}^{\infty} A_n\right) = \lim_{n \rightarrow \infty} P(A_n).$$

**Hint:** Write  $A_1 \cup A_2 \cup \dots$  as the union of a sequence of disjoint events: start with  $A_1$ , then add a disjoint set to obtain  $A_1 \cup A_2$ , then add a disjoint set again to obtain  $A_1 \cup A_2 \cup A_3$ , and so on. Now that you have a sequence of disjoint sets, you can use the definition of a probability measure. To deal with the product  $A_1 \cap A_2 \cap \dots$ , write it as a union of some events with the aid of De Morgan's law.

Suppose that  $A_1, A_2, \dots$  is an expanding sequence of events, i.e:

$$A_1 \subseteq A_2 \subseteq \dots$$

Now observe

$$A_1 \cup A_2 \cup A_3 \cup \dots = A_1 \cup (A_2 \setminus A_1) \cup (A_3 \setminus A_2) \cup \dots \quad (*)$$

Note that  $A_1, (A_2 \setminus A_1), (A_3 \setminus A_2), \dots$  are pairwise disjoint. (Because this expanding sequence)

$$\begin{aligned} \text{Let } S_n &:= P(A_1) + P(A_2 \setminus A_1) + \dots + P(A_n \setminus A_{n-1}) \\ &= \Pr(A_1 \cup (A_2 \setminus A_1) \cup (A_3 \setminus A_2) \cup \dots) \\ &= \Pr(A_1 \cup A_2 \cup \dots \cup A_n) \\ &= \Pr(A_n) \end{aligned}$$

$$\lim_{n \rightarrow \infty} S_n = \lim_{n \rightarrow \infty} \Pr(A_n) \quad \text{(By *)}$$

$$\Pr(A_1) + \Pr(A_2 \setminus A_1) + \dots = \lim_{n \rightarrow \infty} \Pr(A_n) \quad (1)$$

$$\begin{aligned} \text{By } (*) \\ \Pr(A_1 \cup A_2 \cup \dots) &= \Pr(A_1 \cup (A_2 \setminus A_1) \cup (A_3 \setminus A_2) \cup \dots) \\ &= \Pr(\lim_{n \rightarrow \infty} \Pr(A_n)) \quad (\text{By (1)}) \end{aligned}$$

Suppose that  $A_1 \supseteq A_2 \supseteq A_3 \supseteq \dots$

$$\Pr(A_1 \cap A_2 \cap \dots \cap A_n)$$

By De-Morgan's Law

$$\Pr(\complement(A_1 \cap A_2 \cap \dots \cap A_n)) = \Pr(\complement A_1) \cup \Pr(\complement A_2) \cup \dots \cup \Pr(\complement A_n)$$

$$\Pr(\complement(A_1 \cap A_2 \cap \dots \cap A_n)) = \Pr(\complement A_1) \cup \Pr(\complement A_2) \cup \dots \cup \Pr(\complement A_n)$$

$$1 - \Pr(A_1 \cap A_2 \cap \dots \cap A_n) = \Pr(\complement A_n) \quad (A_1 \supseteq A_2 \supseteq \dots)$$

$$1 - \Pr(A_1 \cap A_2 \cap \dots \cap A_n) = 1 - \Pr(A_n)$$

$$\Pr(A_1 \cap A_2 \cap \dots \cap A_n) = \Pr(A_n)$$

$$\lim_{n \rightarrow \infty} \Pr(A_1 \cap A_2 \cap \dots \cap A_n) = \Pr(A_n)$$

Thus,

$$\Pr(A_1 \cap A_2 \cap \dots) = \Pr(A_n)$$

**Lemma 2.1** (Borel-Cantelli). Let  $A_1, A_2, \dots$  be a sequence of events such that  $\Pr(A_1) + \Pr(A_2) + \dots < \infty$  and let  $B_n = A_n \cup A_{n+1} \cup \dots$ . Then  $\Pr(B_1 \cap B_2 \cap \dots) = 0$ .

**Exercise 2.2.** Prove the Borel-Cantelli lemma above.

**Hint:**  $B_1, B_2, \dots$  is a contracting sequence of events.

Let  $A_1, A_2, \dots$  be a sequence of events such that  $P(A_1) + P(A_2) + \dots < \infty$  and let  $B_n := A_1 \cup A_{n+1} \cup \dots$ . Observe that  $B_1 \supseteq B_2 \supseteq B_3 \dots$

Using the previous result from exercise

$$\begin{aligned} \Pr(B_1 \cap B_2 \cap \dots) &= \lim_{n \rightarrow \infty} \Pr(B_n) \\ &= \lim_{n \rightarrow \infty} \Pr(A_1 \cup A_{n+1} \cup \dots) \\ &\leq \lim_{n \rightarrow \infty} [P(A_1) + P(A_{n+1}) + \dots] \quad (\text{subadditive property}) \end{aligned}$$

Given that  $\sum_{i=1}^{\infty} P(A_i)$  is convergent.  $\checkmark$

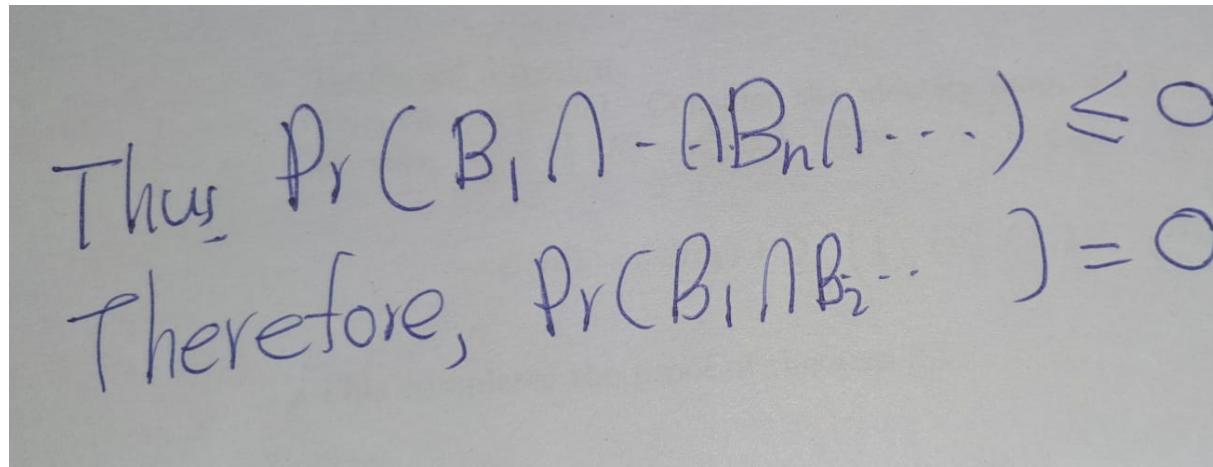
~~This~~  $\sum_{i=1}^{\infty} P(A_{n+i}) + \sum_{j=n+1}^{\infty} P(A_j) < \infty$

Most right hand side part in equation  $\checkmark$   
is often called tail of series.

If  $\sum_{n=1}^{\infty} a_n$  is cgt then tail of cgt

We know that tail of cgt series tend to zero.

Thus right hand side of equation  $\checkmark$   
is 0



## 2.2 Random Variables

**Definition 2.3.** If  $\mathcal{F}$  is a  $\sigma$ -field on  $\Omega$ , then a function  $X : \Omega \rightarrow \mathbb{R}$  is said to be  **$\mathcal{F}$ -measurable** if

$$\{\omega \in \Omega : X(\omega) \in B\} = X^{-1}(B)$$

for every Borel set  $B \in \mathcal{B}(\mathbb{R})$ .

If  $(\Omega, \mathcal{F}, P)$  is a probability space, then such a function  $X$  is called a **random variable**.

**Definition 2.4.** The  $\sigma$ -field  $\sigma(X)$  **generated by a random variable**  $X : \Omega \rightarrow \mathbb{R}$  consists of all sets of the form  $\{\omega \in \Omega : X(\omega) \in B\}$ , where  $B$  is a Borel set in  $\mathbb{R}$ .

**Definition 2.5.** The  $\sigma$ -field  $\sigma(\{X_i : i \in I\})$  generated by a family  $\{X_i : i \in I\}$  of random variables is defined to be the smallest  $\sigma$ -field containing all events of the form  $\{X_i \in B\}$ , where  $B$  is a Borel set in  $\mathbb{R}$  and  $i \in I$ .

**Exercise 2.3.** We call  $f : \mathbb{R} \rightarrow \mathbb{R}$  a **Borel function** if the inverse image  $f^{-1}(B)$  of any Borel set  $B$  in  $\mathbb{R}$  is a Borel set. Show that if  $f$  is a Borel function and  $X$  is a random variable, then the composition  $f(X)$  is  $\sigma(X)$ -measurable.

*Hint:* Consider the event  $\{f(X) \in B\}$ , where  $B$  is an arbitrary Borel set. Can this event be written as  $\{X \in A\}$  for some Borel set  $A$ ?

Page } .....  
888

Let  $B$  an Borel set

Let  $f: \mathbb{R} \rightarrow \mathbb{R}$  bean Borel function

Then  $f^{-1}(B)$  is a Borel set.

~~So, Let  $\{\xi \in f^{-1}(B)\}$~~

So,

$$\{\xi \in f^{-1}(B)\} = \{f(\xi) \in B\}$$

$$= \{f \circ \xi \in B\}$$

Thus,  $f(\xi)$  is  $\sigma(\xi)$  measurable.

**Lemma 2.2** (Doob-Dynkin). *Let  $X$  be a random variable. Then each  $\sigma(X)$ -measurable random variable  $\eta$  can be written as*

$$\eta = f(X)$$

for some Borel function  $f: \mathbb{R} \rightarrow \mathbb{R}$ .

*Proof.* Omiited  $\square$

**Definition 2.6.** Every random variable  $X: \Omega \rightarrow \mathbb{R}$  gives rise to a probability measure

$$P_X(B) = P\{X \in B\}$$

on  $\mathbb{R}$  defined on the  $\sigma$ -field of Borel sets  $B \in \mathcal{B}(\mathbb{R})$ . We call  $P_X$  the distribution of  $X$ . The function  $F_X : \mathbb{R} \rightarrow [0, 1]$  defined by

$$F_X(x) = P\{X \leq x\}$$

is called the cumulative distribution function (CDF) of  $X$ .

**Exercise 2.4.** Show that the distribution function  $F$  is non-decreasing, right-continuous, and

$$\lim_{x \rightarrow -\infty} F_\xi(x) = 0, \quad \lim_{x \rightarrow +\infty} F_\Xi(x) = 1.$$

**Hint:** For example, to verify right-continuity show that  $F_\xi(x_n) \rightarrow F_\xi(x)$  for any decreasing sequence  $x_n$  such that  $x_n \rightarrow x$ . You may find the results of Exercises useful.

### • Non-decreasing

Let  $x, y \in \mathbb{R}$  such that  $x < y$ . Thus,  
 $\{\zeta \leq x\} \subseteq \{\zeta \leq y\}$  ————— ①

- Recall from the measure theory.

Let  $(\Omega, \mathcal{F}, \mu)$  be a measure space.  
If  $A, B \in \mathcal{F}$  and  $A \subseteq B$  then  $\mu(A) \leq \mu(B)$

Thus,  $P(\zeta \leq x) \leq P(\zeta \leq y)$   
 $F_\zeta(x) \leq F_\zeta(y)$ .

• Right Continuity

Let  $\{x_n\}_{n \in \mathbb{N}}$  be an decreasing sequence and  $x_n \rightarrow x$  as  $n \xrightarrow{\text{lim}} \infty$ .

Since,  $\{x_n\}_n$  is decreasing sequence,  $x_1 > x_2 > \dots$

Now observe that,

$$\{\xi \leq x_1\} \supseteq \{\xi \leq x_2\} \supseteq \dots$$

Observe that  $(\{\xi \leq x_1\} \cap \{\xi \leq x_2\} \cap \dots) = \{\xi \leq x\}$

$$P(\underbrace{\{\xi \leq x_1\} \cap \{\xi \leq x_2\} \cap \dots}_{P(\xi \leq x)}) = \underset{n \xrightarrow{\text{lim}} \infty}{\lim} P(\{\xi \leq x_n\}) = \underset{n \xrightarrow{\text{lim}} \infty}{\lim} F_\xi(x_n)$$

$$P(\xi \leq x) = \underset{n \xrightarrow{\text{lim}} \infty}{\lim} F_\xi(x_n)$$

$$F_\xi(x) = \underset{n \xrightarrow{\text{lim}} \infty}{\lim} F_\xi(x_n)$$

Therefore,  $F_\xi$  is right continuous.

$$\circ n \xrightarrow{\lim_{\rightarrow \infty}} F_\xi(n)$$

First observe that

$$\{\xi < -1\} \supseteq \{\xi < -2\} \supseteq \{\xi < -3\} \supseteq \dots$$

$$\text{Further } \phi = \{\xi < -1\} \cap \{\xi < -2\} \cap \{\xi < -3\} \dots$$

$$x \xrightarrow{\lim_{\rightarrow -\infty}} F_\xi(x) = n \xrightarrow{\lim_{\rightarrow \infty}} F_\xi(-n)$$

$$= n \xrightarrow{\lim_{\rightarrow \infty}} P(\xi \leq -n)$$

$$= P(\{\xi < -1\} \cap \{\xi < -2\} \cap \dots)$$

$$= P(\phi)$$

$$= 0 \quad (\text{measure of empty set})$$

$$\bullet n \xrightarrow{\lim_{n \rightarrow \infty}} F_\xi(n)$$

First observe that,

$$\{\xi < 1\} \subseteq \{\xi < 2\} \subseteq \dots$$

$$\text{and } \{\xi < 1\} \cup \{\xi < 2\} \cup \dots = \Omega$$

$$\begin{aligned} x \xrightarrow{\lim_{n \rightarrow \infty}} F_\xi(x) &= n \xrightarrow{\lim_{n \rightarrow \infty}} F_\xi(n) \\ &= n \xrightarrow{\lim_{n \rightarrow \infty}} P(\xi < n) \\ &= P(\{\xi < 1\} \cup \{\xi < 2\} \cup \dots) \\ &= P(\Omega) \\ &= 1 - 0 = 1 \end{aligned}$$

**Definition 2.7.** If there is a Borel function  $f : \mathbb{R} \rightarrow \mathbb{R}$  such that for any Borel set  $B \subset \mathbb{R}$

$$P\{\xi \in B\} = \int_B f_\xi(x) dx,$$

then  $\xi$  is said to be a random variable with absolutely continuous distribution and  $f_\xi$  is called the **density of  $\xi$** . If there is a (finite or infinite) sequence of

pairwise distinct real numbers  $x_1, x_2, \dots$  such that for any Borel set  $B \subset \mathbb{R}$

$$P\{\xi \in B\} = \sum_{x_i \in B} P\{\xi = x_i\},$$

then  $\xi$  is said to have a discrete distribution with values  $x_1, x_2, \dots$  and mass  $P\{\xi = x_i\}$  at  $x_i$ .

**Exercise 2.5.** Suppose that  $\xi$  has a continuous distribution with density  $f$ . Show that  $f$  is continuous at  $x$ .

**Hint:** Express  $F(x)$  as an integral of  $f$ .

Suppose that  $\xi$  has a continuous distribution with density  $f_\xi$ .

$$F_\xi(x) = P(\xi < x) = \int_{-\infty}^x f_\xi(y) dy.$$

Using the fundamental theorem of calculus.

$$F'_\xi(x) = f_\xi(x)$$

Show that if  $\xi$  has discrete distribution with values  $x_1, x_2, \dots$ , then  $F_\xi$  is constant on each interval  $(s, t]$  not containing any of the  $x_i$ 's and has jumps of size  $P\{\xi = x_i\}$  at each  $x_i$ . Hint The increment  $F_\xi(t) - F_\xi(s)$  is equal to the total mass of the  $x_i$ 's that belong to the interval  $[s, t]$ .

Suppose that  $\xi$  has a continuous distribution with density  $f_\xi$ .

$$F_\xi(x) = P(\xi < x) = \int_{-\infty}^x f_\xi(y) dy.$$

Using the fundamental theorem of calculus,

$$F'_\xi(x) = f_\xi(x)$$

**Definition 2.8.** The **joint distribution** of several random variables  $\xi_1, \dots, \xi_n$  is a probability measure  $P_{\xi_1, \dots, \xi_n}$  on  $\mathbb{R}^n$  such that

$$P_{\xi_1, \dots, \xi_n}(B) = P\{\xi_1, \dots, \xi_n \in B\}$$

for any Borel set  $B$  in  $\mathbb{R}^n$ . If there is a Borel function  $f_{\xi_1, \dots, \xi_n} : \mathbb{R}^n \rightarrow \mathbb{R}$  such that

$$P\{(\xi_1, \dots, \xi_n) \in B\} = \int_B f_{\xi_1, \dots, \xi_n}(x_1, \dots, x_n) dx_1 \cdots dx_n$$

for any Borel set  $B$  in  $\mathbb{R}^n$ , then  $f_{\xi_1, \dots, \xi_n}$  is called the **joint density** of  $\xi_1, \dots, \xi_n$ .

**Definition 2.9.** A random variable  $\xi : \Omega \rightarrow \mathbb{R}$  is said to be **integrable** if

$$\int_{\Omega} |\xi| dP < \infty.$$

The integral

$$\mathbb{E}(\xi) = \int_{\Omega} \xi dP$$

exists and is called the expectation of  $\xi$ . The family of integrable random variables  $\xi : \Omega \rightarrow \mathbb{R}$  will be denoted by  $L^1$  or, in case of possible ambiguity, by  $L^1(\Omega, \mathcal{F}, P)$ .

**Example 2.3.** The **indicator function**  $\mathbf{1}_A$  of a set  $A$  is equal to 1 on  $A$  and 0 on the complement  $\Omega \setminus A$  of  $A$ . i.e.:

$$\mathbf{1}_A(a) := \begin{cases} 1 & \text{if } a \in A \\ 0 & \text{if } a \notin A \end{cases}$$

For any event  $A$ ,

$$\mathbb{E}[\mathbf{1}_A] = \int_{\Omega} \mathbf{1}_A dP = P(A)$$

we say that  $\eta : \Omega \rightarrow \mathbb{R}$  is a step function if

$$\eta = \sum_{i=1}^n \eta_i \mathbf{1}_{A_i},$$

where  $\eta_1, \dots, \eta_n$  are real numbers and  $A_1, \dots, A_n$  are pairwise disjoint events. Then,

$$\mathbb{E}[\eta] = \int_{\Omega} \eta dP = \sum_{i=1}^n \eta_i \int_{\Omega} \mathbf{1}_{A_i} dP = \sum_{i=1}^n \eta_i P(A_i)$$

**Exercise 2.6.** Show that for any Borel function  $h : \mathbb{R} \rightarrow \mathbb{R}$  such that  $h(X)$  is integrable,

$$\mathbb{E}(h(X)) = \int h(x) dP_X(x).$$

**Hint:** First verify the equality for step functions  $h : \mathbb{R} \rightarrow \mathbb{R}$ , then for non-negative ones by approximating them by step functions, and finally for arbitrary Borel functions by splitting them into positive and negative parts

**More to go ...**

## 2.3 Conditional Probability and Independence

**Definition 2.10.** For any events  $A, B \in \mathcal{F}$  such that  $P(B) \neq 0$ , the conditional probability of  $A$  given  $B$  is defined by

$$P(A | B) = \frac{P(A \cap B)}{P(B)}$$

**Exercise 2.7.** Prove the **total probability formula** for any event  $A \in \mathcal{F}$  and any sequence of pairwise disjoint events  $B_1, B_2, \dots \in \mathcal{F}$  such that  $B_1 \cup B_2 \cup \dots = \emptyset$  and  $P(B_n) \neq 0$  for any  $n$ .

**Hint:**  $A = (A \cap B_1) \cup (A \cap B_2) \cup \dots$

Let  $A \in \mathcal{F}$ . Let  $\{B_n\}_{n \in \mathbb{N}}$  be a sequence of pairwise disjoint events with  $B_n \in \mathcal{F} \quad \forall n = 1, 2, \dots$  such that  $B_1 \cup B_2 \cup \dots = \Omega$  and  $P(B_n) \neq 0$ .

By definition of conditionally probability  
 $P(A \cap B_n) = P(A|B_n) \cdot P(B_n) \quad \forall n = 1, 2, \dots$

Note that  $A = (A \cap B_1) \cup (A \cap B_2) \cup \dots$   
 $P(A) = P((A \cap B_1) \cup (A \cap B_2) \cup (A \cap B_3) \dots)$   
 $= P(A \cap B_1) + P(A \cap B_2) + P(A \cap B_3) + \dots$   
(Countable additive)  
 $= P(A|B_1) \cdot P(B_1) + P(A|B_2) \cdot P(B_2) + \dots$

**Definition 2.11.** Two events  $A, B \in \mathcal{F}$  are called **independent** if

$$P(A \cap B) = P(A)P(B).$$

In general, we say that  $n$  events  $A_1, \dots, A_n \in \mathcal{F}$  are **independent** if for any indices  $1 \leq i_1 < i_2 < \dots < i_k \leq n$ ,

$$P(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_k}) = P(A_{i_1})P(A_{i_2}) \cdots P(A_{i_k}).$$

**Exercise 2.8.** Let  $P(B) \neq 0$ . Show that  $A$  and  $B$  are independent events if and only if

$$P(A | B) = P(A).$$

**Hint:** If  $P(B) \neq 0$ , then you can divide by it.

Let  $A, B$  are events. Let  $P(B) \neq 0$ .

$A$  and  $B$  are independent  $\Leftrightarrow P(A) \cdot P(B) = P(A \cap B)$

$$\Leftrightarrow \frac{P(A) \cdot P(B)}{P(B)} = \frac{P(A \cap B)}{P(B)}$$

$$P(A) = P(A|B)$$

**Definition 2.12.** Two random variables  $\xi$  and  $\eta$  are called independent if for any Borel sets  $A, B \in \mathcal{B}(\mathbb{R})$ , the two events

$$\{\xi \in A\} \text{ and } \{\eta \in B\}$$

are independent.

We say that  $n$  random variables  $\xi_1, \dots, \xi_n$  are independent if for any Borel sets  $B_1, \dots, B_n \in \mathcal{B}(\mathbb{R})$ , the events

$$\{\xi_1 \in B_1\}, \{\xi_2 \in B_2\}, \dots, \{\xi_n \in B_n\}$$

are independent.

In general, a (finite or infinite) family of random variables is said to be independent if any finite number of random variables from this family are independent.

**Proposition 2.1.** If two integrable random variables  $\xi, \eta : \Omega \rightarrow \mathbb{R}$  are independent, then they are uncorrelated, i.e.,

$$E(\xi\eta) = E(\xi)E(\eta),$$

provided that the product  $\xi\eta$  is also integrable.

If  $\xi_1, \dots, \xi_n : \Omega \rightarrow \mathbb{R}$  are independent integrable random variables, then

$$E(\xi_1\xi_2 \cdots \xi_n) = E(\xi_1)E(\xi_2) \cdots E(\xi_n),$$

provided that the product  $\xi_1\xi_2 \cdots \xi_n$  is also integrable.

**Definition 2.13.** Two  $\sigma$ -fields  $\mathcal{G}$  and  $\mathcal{H}$  contained in  $\mathcal{F}$  are called independent if any two events

$$A \in \mathcal{G} \text{ and } B \in \mathcal{H}$$

are independent.

Similarly, any finite number of  $\sigma$ -fields  $\mathcal{G}_1, \dots, \mathcal{G}_n$  contained in  $\mathcal{F}$  are independent if any  $n$  events

$$A_1 \in \mathcal{G}_1, \dots, A_n \in \mathcal{G}_n$$

are independent.

In general, a (finite or infinite) family of  $\sigma$ -fields is said to be independent if any finite number of them are independent.

**Exercise 2.9.** Show that two random variables  $\xi$  and  $\eta$  are independent if and only if the  $\sigma$ -fields  $\sigma(\xi)$  and  $\sigma(\eta)$  generated by them are independent.

**Hint:** The events in  $\sigma(\xi)$  and  $\sigma(\eta)$  are of the form  $\{\xi \in A\}$  and  $\{\eta \in B\}$ , where  $A$  and  $B$  are Borel sets.

Let  $\xi, \eta$  be random variables. Let  $A, B$  be Borel sets in  $\mathbb{R}$ . Further  $\sigma(\xi)$  and  $\sigma(\eta)$  have events in the form  $\{\xi \in A\}$  and  $\{\eta \in B\}$

$\sigma(\xi)$  and  $\sigma(\eta)$  are independent  $\iff$   $\{\xi \in A\}$  and  $\{\eta \in B\}$  are independent (by def<sup>n</sup> 1.14 in the book.)

$\iff$   $\xi$  and  $\eta$  are independent (by def<sup>n</sup> 1.13 in book)

Sometimes it is convenient to talk of independence for a combination of random variables and  $\sigma$ -fields.

**Definition 2.14.** We say that a random variable  $\xi$  is independent of a  $\sigma$ -field  $\mathcal{G}$  if the  $\sigma$ -fields

$$\sigma(\xi) \text{ and } \mathcal{G}$$

are independent. This can be extended to any (finite or infinite) family consisting of random variables or  $\sigma$ -fields or a combination of them both. Namely,

such a family is called independent if for any finite number of random variables  $\xi_1, \dots, \xi_m$  and  $\sigma$ -fields  $\mathcal{G}_1, \dots, \mathcal{G}_n$  from this family, the  $\sigma$ -fields

$$\sigma(\xi_1), \dots, \sigma(\xi_m), \mathcal{G}_1, \dots, \mathcal{G}_n$$

are independent.

## Chapter 3

# Random Walk to Brownier Motion.

Slandered approach model stochastic dynamic in discrete time.

Let  $\eta_i$  be an random variable on a conman probability space. We often  $\Omega, \mathcal{F}, P$  assume that i.i.d. This case  $\eta_i$  is called white noise, otherwise coloured noise. Now we have definite dynamics. It will be given as discreate time dynamical systems recursively by some non linear function. We define,

$$X_{n+1} = X_n + \phi_{n+1}(X_n, \eta_{n+1}) \quad n = 0, 1, 2, \dots \quad (3.1)$$

where  $\phi_n : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}^d$  are measurable. Further, if  $X_0$  and  $\eta_0$  are all independent then  $X_n$  is called **Markov Chain**.

Now let  $\eta_i$  be i.i.d and defineda random walk

$$S_n := \sum_{i=1}^n \eta_i \quad (3.2)$$

$$S_{(n+1)} := S_n + \eta_{(n+1)} \quad (3.3)$$

We can rewrite (3.1) as

$$X_{n+1} - X_n = \phi_{n+1}(X_n, S_{n+1} - S_n) \quad n = 0, 1, \dots$$

This equation is called *Stochastic difference equations*.

**AIM:** Develop a continuous time analogous.

**Question** What to use an contionous time replacement of the random walks?

**Definition 3.1.** Let  $I$  be index set ( $I = \mathbb{N}$  or  $I = \mathbb{R}^+$ ). A collection of random variables  $(X_t)_{t \in I}$  on  $(\Omega, \mathcal{F}, P)$  is called stochastic process.

We need  $I$  to be a just totally ordered set for convention of time. If it is not an totally ordered set it is not a stochastic process but a random field.

Now we need a notation of a filtration.

**Definition 3.2.** Le  $\mathcal{F}_t$  be non-decreasing sequence of sub sigma algebras of  $\mathcal{F}$  (i.e.  $\mathcal{F}_s \subseteq \mathcal{F}_t$  for all  $s \geq t, s, t \in I$ ), then  $(\mathcal{F}_t)_{t \in I}$  is called a filtration.

Last we need the notation adaptness.

**Definition 3.3.** A stochastic process  $X_t$  is called adapted to filtration  $(\mathcal{F}_t)_t$  if  $X_t \in \mathcal{F}_t$ . i.e:  $X_t$  is measurable

**Theorem 3.1** (Central Limit Theorem). *Let  $Y_{n_i} : \Omega \rightarrow \mathbb{R}^d$  (be collection of random variables),  $1 \leq i \leq n < \infty$  be identical distributed and square integrable random variable on  $(\Omega, \mathcal{F}, P)$  such that  $Y_{n_1}, Y_{n_2}, \dots, Y_{n_n}$  are independent for all  $n \in \mathbb{N}$ . Then*

$$\left( \frac{1}{\sqrt{n}} \sum_{i=1}^n Y_{n_i} - \mathbb{E}[Y_i] \right) \xrightarrow{\mathcal{D}} N(0, C) \text{ as } n \rightarrow \infty$$

, where  $N(0, C)$  is multivariable normal distribution with covariance matrix

$$Y_{k,l} = \text{Cov}[Y_{n_i}^{(k)} - Y_{n_i}^{(l)}]$$

and  $\xrightarrow{\mathcal{D}}$  means “distribution is convergent” to ““

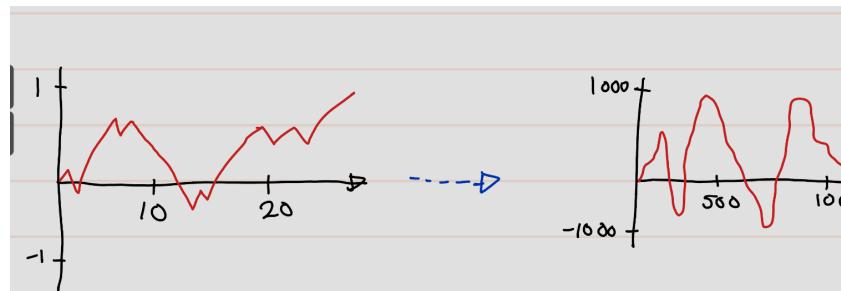
*Proof.* Omitted

□

We consider the random walk

$$S_n = \sum_{i=1}^n \eta_i$$

with  $\eta_i \in L^2(\Omega, \mathcal{F}, P)$  and normalized.(i.e.  $\mathbb{E}[\eta_i] = 0, \text{Var}[\eta_i] = 1$ )



**Plotting** (Linear Interpolation)

This gives an idea about the existence of a scaling limit. Now a question might be rising.

**Question:** What is right rescaling?

That is we try to define a rescaled random walk  $S_t^m$  (Here superscript  $m$  is for mesh size), ( $t = 0, \frac{1}{m}, \frac{2}{m}, \dots$ ) with step-size  $\frac{1}{m}$

$$S_{\frac{k}{m}}^{(m)} = c_m S_k$$

Here  $c_m$  is rescaling constant. It is difficult to correct  $c_m$ , because unless it decay so fast at the end you convert to zero or blow up whole thing and goes to infinity.

For  $t = \frac{k}{m}$  we have

$$\text{Var}[S_t^{(m)}] = c_m^2$$



## Chapter 4

# Conditional Expectation

### 4.1 Conditioning on an Event

The first and simplest case to consider is that of the conditional expectation  $\mathbf{E}(\xi|B)$  of a random variable  $\xi$  given an event  $B$ .

**Definition 4.1.** For any integrable random variable  $\xi$  and any event  $B \in \mathcal{F}$  such that  $P(B) \neq 0$ , the conditional expectation of  $\xi$  given  $B$  is defined by

$$E(\xi | B) = \frac{1}{P(B)} \int_B \xi dP.$$

**Example 4.1.** Three coins, 10p, 20p, and 50p are tossed. The values of those coins that land heads up are added to work out the total amount  $\xi$ . What is the expected total amount  $\xi$  given that two coins have landed heads up?

Let  $B$  be the event two coins landed two  
i.e.  $B := \{HHT, HTH, THH\}$  ( $H$ -head  $T$ -tail)

Need to calculate  $E(S|B)$ .

$$P(HHT) = P(HTH) = P(THH) = \frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} = \frac{1}{8}.$$

$$P(B) = P(HHT) + P(HTH) + P(THH) = 3/8$$

Further,  $S(HHT) = 10 + 20 = 30$

$$S(HTH) = 10 + 50 = 60$$

$$S(THH) = 20 + 50 = 70$$

Then by above def<sup>b</sup>

$$E(S|B) = \frac{1}{P(B)} \int_B S dP.$$

$$= \frac{1}{(3/8)} \left( \frac{30}{8} + \frac{60}{8} + \frac{70}{8} \right)$$

$$= \frac{160}{3} //$$

**Exercise 4.1.** Show that  $E(\xi | D) = E(\xi)$ .

**Hint:** The definition of  $E(\xi)$  involves an integral and so does the definition of  $E(\xi | D)$ . How are these integrals related?

$$E(\xi | \Omega) = \frac{1}{P(\Omega)} \left( \int_{\Omega} \xi dP \right) = E(\xi)$$

Note that  $P(\Omega) = 1$  and  $\int_{\Omega} \xi dP = E(\xi)$  (defn 1.9 in book)

**Exercise 4.2.** Show that if

$$\mathbf{1}_A(\omega) = \begin{cases} 1 & \text{for } \omega \in A \\ 0 & \text{for } \omega \notin A \end{cases}$$

(the indicator function of  $A$ ), then

$$E(\mathbf{1}_A | B) = P(A | B),$$

where

$$P(A | B) = \frac{P(A \cap B)}{P(B)}$$

is the conditional probability of  $A$  given  $B$ .

**Hint:** Write  $\int_B \mathbf{1}_A dP$  as  $P(A \cap B)$ .

$$\begin{aligned}
 E(I_A | B) &= \frac{1}{P(B)} \cdot \int_B I_A \, dP \\
 &= \frac{1}{P(B)} \cdot \int_{A \cap B} 1 \, dP + \int_{A^c \cap B} 0 \, dP \\
 &= \frac{1}{P(B)} \cdot |A \cap B| = P(A|B)
 \end{aligned}$$

## 4.2 Conditioning on a Discrete Random Variable

The next step towards the general definition of conditional expectation involves conditioning by a discrete random variable  $\eta$  with possible values  $y_1, y_2, \dots$  such that  $P\{\eta = y_n\} \neq 0$  for each  $n$ . Finding out the value of  $\eta$  amounts to finding out which of the events  $\{\eta = y_n\}$  has occurred or not. Conditioning by  $\eta$  should therefore be the same as conditioning by the events  $\{\eta = y_n\}$ . Because we do not know in advance which of these events will occur, we need to consider all possibilities, involving a sequence of conditional expectations

$$E(\xi | \{\eta = y_1\}), E(\xi | \{\eta = y_2\}), \dots$$

A convenient way of doing this is to construct a new discrete random variable constant and equal to  $E(\xi | \{\eta = y_n\})$  on each of the sets  $\{\eta = y_n\}$ . This leads us to the next definition.

**Definition 4.2.** Let  $\xi$  be an integrable random variable and let  $\eta$  be a discrete random variable as above. Then the conditional expectation of  $\xi$  given  $\eta$  is

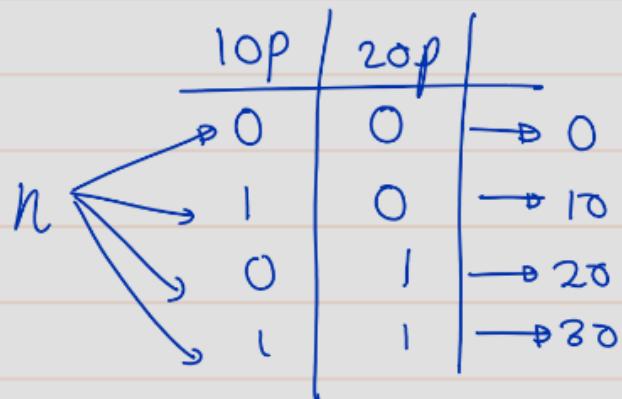
defined to be a random variable  $E(\xi | \eta)$  such that

$$E(\xi | \eta)(\omega) = E(\xi | \{\eta = y_n\}) \text{ if } \eta(\omega) = y_n$$

for any  $n = 1, 2, \dots$

**Example 4.2.** Three coins, 10p, 20p, and 50p are tossed as in Example 4.1. What is the conditional expectation  $E(\xi | \eta)$  of the total amount  $\xi$  shown by the three coins given the total amount  $\eta$  shown by the 10p and 20p coins only?

First observe that  $\eta$  is discrete random variable.  
 All the possible values for  $\eta$  is 0, 10, 20 and 30.



All possible events.

<u>50p</u>	<u>20p</u>	<u>10p</u>	<u><math>\eta</math></u>	<u><math>E</math></u>
H	H	H	30	80
H	H	T	20	70
H	T	H	10	60
H	T	T	0	50
T	H	H	20	30
T	H	T	20	20
T	T	H	10	10
T	T	T	0	0

16 :

$$\begin{aligned} E(\xi | \{\eta=0\}) &= \frac{1}{P(\{\eta=0\})} \cdot \int_{\{\eta=0\}} \xi dP \\ &= \frac{1}{(2/8)} \left[ 0 + \frac{50}{8} \right] = 25 \end{aligned}$$

$$\begin{aligned} E(\xi | \{\eta=10\}) &= \frac{1}{P(\{\eta=10\})} \cdot \int_{\{\eta=10\}} \xi dP \\ &= \frac{1}{(2/8)} \left[ \frac{60}{8} + \frac{10}{8} \right] = 35 \end{aligned}$$

$$\begin{aligned} E[\xi | \{\eta=20\}] &= \frac{1}{P(\{\eta=20\})} \cdot \int_{\{\eta=20\}} \xi dP \\ &= \frac{1}{(2/8)} \cdot \left( \frac{70}{8} + \frac{20}{8} \right) = 45 \end{aligned}$$

$$\begin{aligned} E(\xi | \{\eta=30\}) &= \frac{1}{P(\{\eta=30\})} \cdot \int_{\{\eta=30\}} \xi dP = \frac{1}{(2/8)} \left( \frac{80}{8} + \frac{30}{8} \right) \\ &= 55 \end{aligned}$$

Therefore,

$$E(\xi | \eta)(\omega) := \begin{cases} 25 & \text{if } \eta(\omega) = 0 \\ 35 & \text{if } \eta(\omega) = 10 \\ 45 & \text{if } \eta(\omega) = 20 \\ 55 & \text{if } \eta(\omega) = 30. \end{cases}$$

**Example 4.3.** Take  $\Omega = [0, 1]$  with the  $\sigma$ -field of Borel sets and  $P$  the Lebesgue measure on  $(0, 1]$ . We shall find  $E(\xi | \eta)$  for

$$\xi(x) = 2x^2, \quad \eta(x) = \begin{cases} 1 & \text{if } x \in [0, \frac{1}{3}], \\ 2 & \text{if } x \in (\frac{1}{3}, \frac{2}{3}), \\ 0 & \text{if } x \in (\frac{2}{3}, 1]. \end{cases}$$

First observe that  $\eta$  is discrete r.v.  
All possible values for  $\eta$  is 0, 1, 2.

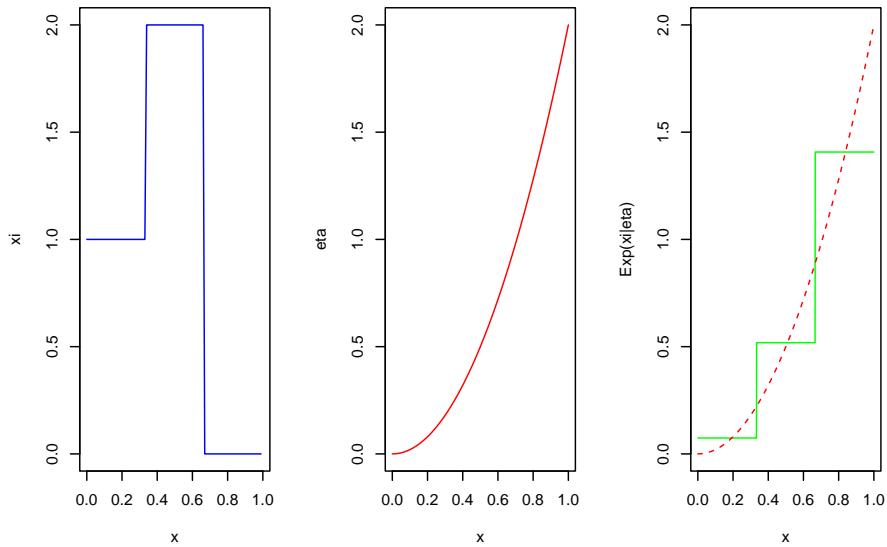
$$\begin{aligned}\{w \in \Omega \mid \eta(w) = 1\} &= \{\eta = 1\} = [0, \frac{1}{3}] \\ \{w \in \Omega \mid \eta(w) = 2\} &= \{\eta = 2\} = (\frac{1}{3}, \frac{2}{3}) \\ \{w \in \Omega \mid \eta(w) = 0\} &= \{\eta = 0\} = (\frac{2}{3}, 1]\end{aligned}$$

$\begin{aligned} &\text{If } x \in [0, \frac{1}{3}] \\ &E(S \eta)(x) \\ &= E(S \mid \{\eta=1\}) \\ &= E(S \mid [0, \frac{1}{3}]) \\ &= \frac{1}{P(\{\eta=1\})} \int_S dP \\ &= \frac{1}{(\frac{1}{3})} \int_0^{\frac{1}{3}} 2x^2 dx \\ &= \frac{1}{\frac{1}{3}} \int_0^{\frac{1}{3}} 2x^2 dx \\ &= \frac{3}{2} \cdot 2 \cdot \frac{x^3}{3} \Big _0^{\frac{1}{3}} \\ &= \frac{3}{2} \cdot 2 \cdot \frac{(\frac{1}{3})^3}{3} \\ &= \frac{2}{27} \end{aligned}$	$\begin{aligned} &\text{If } x \in (\frac{1}{3}, \frac{2}{3}) \\ &E(S \eta)(x) \\ &= E(S \mid \{\eta=2\}) \\ &= \frac{1}{(\frac{1}{3})} \int_{\frac{1}{3}}^{\frac{2}{3}} 2x^2 dx \\ &= \frac{3}{2} \cdot 2 \cdot \frac{x^3}{3} \Big _{\frac{1}{3}}^{\frac{2}{3}} \\ &= 2 \cdot \left[ \frac{8}{27} - \frac{1}{27} \right] \\ &= \frac{14}{27} \end{aligned}$	$\begin{aligned} &\text{If } x \in (\frac{2}{3}, 1] \\ &E(S \eta)(x) \\ &= E(S \mid \{\eta=0\}) \\ &= \frac{1}{(\frac{1}{3})} \int_{\frac{2}{3}}^1 2x^2 dx \\ &= \frac{3}{2} \cdot 2 \cdot \frac{x^3}{3} \Big _{\frac{2}{3}}^1 \\ &= 2 \cdot \left( 1 - \frac{8}{27} \right) \\ &= \frac{38}{27} \end{aligned}$
--	--	---

Let's summarize this above result

$$E(\xi|\eta)(\omega) \begin{cases} 2/27 & \text{if } \omega \in [0, 1/3] \\ 14/27 & \text{if } \omega \in (1/3, 2/3) \\ 38/27 & \text{if } \omega \in [2/3, 1] \end{cases}$$

The graph of  $E(\xi|\eta)$  is shown in following figure together with those of  $\xi$  and  $\eta$



**Exercise 4.3.** Show that if  $\eta$  is a constant function, then  $E(\xi | \eta)$  is constant and equal to  $E(\xi)$ .

**Hint:** The event  $\{\eta = c\}$  must be  $\emptyset$  or  $\Omega$  for any  $c \in \mathbb{R}$ .

Suppose that  $\eta$  is constant function and  
 $\eta(\omega) = c \in \mathbb{R} \quad \forall \omega \in \Omega$   
 Thus,  $\{\eta = c\} = \{\omega \in \Omega \mid \eta(\omega) = c\} = \Omega$

Therefore,  $E(\xi | \eta)(\omega) = E(\xi | \{\eta = c\}) = E(\xi | \omega)$   
 $= E(\xi) \quad (\text{Exercise 2.1 inbook})$   
 for each  $\omega \in \Omega$ .

**Exercise 4.4.** Show that

$$E(\mathbf{1}_A | B)(\omega) = \begin{cases} P(A | B) & \text{if } \omega \in B, \\ P(A | \Omega \setminus B) & \text{if } \omega \notin B \end{cases}$$

for any  $B$  such that  $0 < P(B) < 1$

**Hint** How many different values does  $\mathbf{1}_B$  take? What are the sets on which these values are taken?

Let  $B$  such that,  $P(B) \neq 1$  and  $P(B) \neq 0$   
 Let  $I_A$  an  $I_B$  be the indicator function.

$$I_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases} \quad I_B = \begin{cases} 1 & \text{if } \omega \in B \\ 0 & \text{if } \omega \notin B \end{cases}$$

First observe that,

$$\{I_B = 1\} = \{\omega \in \Omega \mid I_B(\omega) = 1\} = B \quad \text{and}$$

$$\{I_B = 0\} = \{\omega \in \Omega \mid I_B(\omega) = 0\} = \Omega \setminus B.$$

- If  $\omega \in B$  then,

$$E(I_A | I_B)(\omega) = E(I_A | \{I_B = 1\}) = E(I_A | B) = E(A|B)$$

Exercise 2.2

- If  $\omega \notin B$  then,

$$E(I_A | I_B)(\omega) = E(I_A | \{I_B = 0\}) = E(I_A | \Omega \setminus B) = E(A|\Omega \setminus B)$$

Let's summarize

$$E(I_A | I_B)(\omega) := \begin{cases} P(A|B) & \text{if } \omega \in B \\ P(A|\Omega \setminus B) & \text{if } \omega \notin B \end{cases}$$

**Exercise 4.5.** Assuming that  $\eta$  is a discrete random variable, show that

$$E(E(\xi | \eta)) = E(\xi).$$

**Hint:** Observe that

$$\int_B E(\xi | \eta) dP = \int_B \xi dP$$

for any event  $B$  on which  $\eta$  is constant. The desired equality can be obtained by covering  $\Omega$  by countably many disjoint events of this kind.

**Proposition 4.1.** *If  $\xi$  is an integrable random variable and  $\eta$  is a discrete random variable, then:*

1.  $E(\xi | \eta)$  is  $\sigma(\eta)$ -measurable;
2. For any  $A \in \sigma(\eta)$ ,

$$\int_A E(\xi | \eta) dP = \int_A \xi dP.$$

*Proof.* Suppose that  $\eta$  has pairwise distinct values  $y_1, y_2, \dots$ . Then the events

$$\{\eta = y_1\}, \{\eta = y_2\}, \dots$$

are pairwise disjoint and cover  $\Omega$ . The  $\sigma$ -field  $\sigma(\eta)$  is generated by these events; in fact, every  $A \in \sigma(\eta)$  is a countable union of sets of the form  $\{\eta = y_n\}$ . Because  $E(\xi | \eta)$  is constant on each of these sets, it must be  $\sigma(\eta)$ -measurable.

For each  $n$ , we have

$$\int_{\{\eta=y_n\}} E(\xi | \eta) dP = \int_{\{\eta=y_n\}} E(\xi | \{\eta = y_n\}) dP = \int_{\{\eta=y_n\}} \xi dP.$$

Since each  $A \in \sigma(\eta)$  is a countable union of sets of the form  $\{\eta = y_n\}$ , which are pairwise disjoint, it follows that

$$\int_A E(\xi | \eta) dP = \int_A \xi dP,$$

as required. □

### 4.3 Conditioning on an Arbitrary Random Variable

**Definition 4.3.** Let  $\xi$  be an integrable random variable and let  $\eta$  be an arbitrary random variable. Then the **conditional expectation** of  $\xi$  given  $\eta$  is defined to be a random variable  $E(\xi | \eta)$  such that:

1.  $E(\xi | \eta)$  is  $\sigma(\eta)$ -measurable;
2. For any  $A \in \sigma(\eta)$ ,

$$\int_A E(\xi | \eta) dP = \int_A \xi dP.$$

*Remark.* We can also define the **conditional probability** of an event  $A \in \mathcal{F}$  given  $(\cdot)$  by

$$P(A|\eta) = E(1_A|\eta)$$

here  $1_A$  is the indicator function of  $A$ .

Do the conditions of Definition 2.3 characterize  $E(\xi|\eta)$  uniquely? The lemma below implies that  $E(\xi|\eta)$  is defined to within equality on a set of full measure. Namely, if

$$\xi = \xi' \text{ almost sure , then } E(\xi|\eta) = E(\xi'|\eta) \text{ almost sure}$$

An event  $A$  is said to occur almost surely (a.s.) whenever  $P(A) = 1$ . The existence of  $E(\xi|\eta)$  will be discussed later in this chapter.

**Lemma 4.1.** *Let  $(\Omega, \mathcal{F}, P)$  be a probability space and let  $\mathcal{G}$  be a  $\sigma$ -field contained in  $\mathcal{F}$ . If  $\xi$  is a  $\mathcal{G}$ -measurable random variable and for any  $B \in \mathcal{G}$ ,*

$$\int_B \xi dP = 0,$$

*then  $\xi = 0$  a.s.*

# Chapter 5

## Les

### 5.1 Integrals

First we review the definitions of the Riemann integral in calculus and the Riemann–Stieltjes integral in advanced calculus.

#### 5.1.1 Riemann Integral

Let  $f$  be an bounded function defined on a finite closed interval  $[a, b]$ . Then  $f$  is called *Riemann integrable* if the following limit exists.

$$x \tag{5.1}$$

### 5.2 Random Walks

Consider a random walk starting at 0 with jumps  $h$  and  $-h$  equally likely at times  $\delta, 2\delta, \dots$ , where  $h, \delta > 0$ . More precisely, let  $\{X_n\}_{n=1}^\infty$  be a sequence of independent and identically distributed random variables with

$$P(X_j = h) = P(X_j = -h) = \frac{1}{2}$$

Let  $Y_{\delta,h}(0) = 0$

$$Y_{\delta,h}(n\delta) = X_1 + X_2 + \dots + X_n$$

For  $t > 0$  define  $Y_{\delta,h}(t)$  by linearization: (i.e: For  $n\delta < t < (n+1)\delta$ , define

$$Y_{\delta,h}(t) = \frac{(n+1)\delta - t}{\delta} Y_{\delta,h}(n\delta) + \frac{t - n\delta}{\delta} Y_{\delta,h}((n+1)\delta).$$





# Chapter 6

## Book

### The GOY MODEL

$\{U_n(t)\}_{n \geq -1}$  is defined by,

$$U_n(t) := U_{n-1}(t) + iU_{n-2}(t) \quad \forall n=1,2,3,\dots$$

with constraint  $U_0^{(t)} = U_1^{(t)} = 0$

$$\frac{dU_n}{dt} + \gamma k_n^2 U_n + ik_n \left[ \frac{1}{4} \bar{U}_{n-1} \bar{U}_{n+1} - \bar{U}_{n+1} \bar{U}_{n+2} + \frac{1}{8} \bar{U}_{n-1} \bar{U}_{n-2} \right] = \sigma_n \frac{d\beta_n}{dt}$$

where,  $\gamma$  - Viscosity ( $\gamma > 0$ )

$k_n = 2^n k_0$  - wave number ( $k_0 > 0$ )

$(\sigma_n)_{n \geq 1} \in M_n(\mathbb{R})$  - the intensities of noise

$(\beta_n)_{n \geq -1}$  - sequence of independent complex-value

$$\text{Then } \|u\|_H^2 = \langle u, u \rangle_H$$

$$\begin{aligned} &= \operatorname{Re} \sum_{n=1}^{\infty} u_n \overline{u_n} = \operatorname{Re} \sum_{n=1}^{\infty} |u_n|^2 \\ &= \sum_{n=1}^{\infty} |u_n|^2 \end{aligned}$$

Recall:  $k_n = 2^n k_0$ ;  $n \geq 1$  and  $k_0 > 1$

Introduce now the Hilbert spaces,  $D(A) \subseteq V \subseteq H$

$$V := \left\{ u \in H \mid \sum_{n=1}^{\infty} k_n^2 |u_n|^2 < \infty \right\}$$

with norm,  $\|u\|_V^2 := \sum_{n=1}^{\infty} k_n^2 |u_n|^2$  and

$$D(A) := \left\{ u \in H \mid \sum_{n=1}^{\infty} k_n^4 |u_n|^2 < \infty \right\}$$

We define linear operator,

$$\begin{array}{ccc}
 A: D(A) \subseteq H & \longrightarrow & H \\
 u_n & \longmapsto & (Au)_n = k_n^2 u_n \\
 & \uparrow & \uparrow \\
 & \text{nth component.} & (A[u])_n
 \end{array}$$

Ex: Show that  $A$  is linear. operator.

The operator  $A$  is self adjoint

Let  $u, v \in D(A)$

$$\begin{aligned}
 \langle Au, v \rangle_H &= \left\langle (k_1^2 u_1, k_2^2 u_2, \dots), (v_1, v_2, \dots) \right\rangle_H \\
 &= \operatorname{Re} \sum_{n=1}^{\infty} (k_n^2 u_n) \bar{v}_n \\
 &= \operatorname{Re} \sum_{n=1}^{\infty} u_n (\bar{k}_n^2 v_n) (k_n \in \mathbb{R}) \\
 &= \left\langle (u_1, u_2, \dots), (k_1^2 v_1, k_2^2 v_2, \dots) \right\rangle_H \\
 &= \langle u, Av \rangle_H
 \end{aligned}$$

Thus  $A$  is adjoint.

Claim: The operator  $A$  is positive definite  
 Recall:

**Definition 4.6.3.** A self adjoint operator is *strictly positive* or *positive definite* if  $(Ax, x) > 0$  for all  $x \in H$  where  $x \neq 0$ .

Let  $u \in D(A)$

$$\begin{aligned} \langle Au, u \rangle_H &= \left\langle (k_1^2 u_1, k_2^2 u_2, \dots), (u_1, u_2, \dots) \right\rangle \\ &= \sum_{n=1}^{\infty} k_n^2 u_n \bar{u}_n = \sum_{n=1}^{\infty} |k_n|^2 |u_n|^2 \\ &\geq k_0 \sum_{n=1}^{\infty} |u_n|^2 = k_0 \|u\|_H \end{aligned}$$

Now we introduce the bilinear operator.

$$B : V \times H \rightarrow H$$

$$(u, v)_n \mapsto (B(u, v))_n = \underbrace{\quad}_{\text{Def}} \quad$$

$$D_{ik_n} \left( \frac{1}{4} \bar{v}_{n-1} \bar{u}_{n+1} - \frac{1}{2} (\bar{u}_{n+1} \bar{v}_{n+2} + \bar{v}_{n+1} \bar{u}_{n+2}) + \frac{1}{8} \bar{u}_{n-1} \bar{v}_{n-2} \right).$$

Ex: prove that the operator  $B$  is bilinear.

- i.e. •  $B(\lambda u, v) = B(u, \lambda v) = \lambda (u, v) \quad \forall \lambda \in \mathbb{R}$
- $B(u+v, w) = B(u, w) + B(v, w)$
- $B(u, v+w) = B(u, v) + B(u, w)$

Note that,

$$k_n^2 |u_n|^2 \leq \sup_n (k_n^2 |u_n|^2) \quad \forall n=1,2,3,\dots$$

Thus,

$$\begin{aligned} \sum_{n=1}^{\infty} k_n^2 |u_n|^2 |\gamma_n|^2 &\leq \sum_{n=1}^{\infty} \sup_n (k_n^2 |u_n|^2) \cdot (\|\gamma_n\|^2) \\ &= \sup_n (k_n^2 |u_n|^2) \sum_{n=1}^{\infty} |\gamma_n|^2 \\ &\leq \left( \sum_{n=1}^{\infty} k_n^2 |u_n|^2 \right) \left( \sum_{n=1}^{\infty} |\gamma_n|^2 \right) \\ &= \underbrace{\|u\|_V^2}_{\|u\|_V^2 < \infty} \underbrace{\|\gamma\|_H^2}_{\|\gamma\|_H^2 < \infty} \end{aligned}$$

Ex: Prove that  $B(u, v) \in H$ , where  $(u, v) \in V \times H$

i.e:  $\sum_{n=1}^{\infty} |(B(u, v))_n|^2$

But if  $(u, v) \in V \times H$  then we can define  $B$

$$B: H \times V \longrightarrow H$$

Now let's summarize this fact.

Lemma 1  $\exists$  constant  $C > 0$  such that

$$|B(u, v)|_H \leq C \|u\|_V \|v\|_H$$

and

$$|B(u, v)|_V \leq C \|v\|_V \|u\|_H$$

for  $u$  and  $v$  are in the proper space  
(closed sets  $\Rightarrow$  compact)

We have,  $\langle B(u, v), v \rangle$