

ALLEVIATING THE COST OF BAYESIAN CALIBRATION WITH EMBEDDED MODEL ERROR

(USING GAUSSIAN PROCESS SURROGATES AND MULTILEVEL MCMC)



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OVERVIEW



- Motivating example: Bayesian calibration for interatomic potentials
- Review two model error frameworks
 - Additive model error
 - Embedded model error

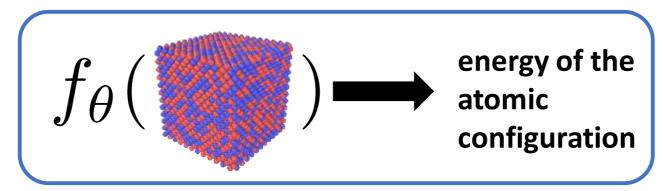
- Computational strategy for Gaussian embedded model error
 - Likelihood approximation
 - Gaussian Process (GP) surrogates with Gaussian inputs
 - Sparse GPs & multilevel MCMC for inference
- Case study: calibrating an EAM potential with embedded model error

MOTIVATING EXAMPLE: CALIBRATING INTERATOMIC POTENTIALS



Design of advanced energy storage devices often requires materials simulations at the *atomistic* level.

For practical scale systems, density functional theory (DFT) based simulations are untenable, motivating the use of empirical methods involving **interatomic potentials**.

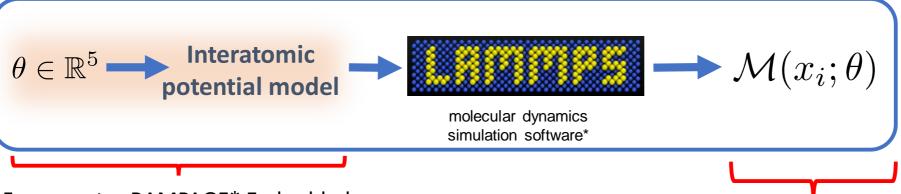


Given their empirical nature, **effective calibration of such models** requires careful consideration of the **model error***.

^{*}Hegde, A., Weiss, E., Windl, W., Najm, H., & Safta, C. "Bayesian calibration of interatomic potentials for binary alloys." *Computational Materials Science* 214 (2022): 111660.

MOTIVATING EXAMPLE: CALIBRATING INTERATOMIC POTENTIALS





5-parameter RAMPAGE* Embedded Atom Method (EAM) potential.

$$E_{i} = \frac{1}{2} \sum_{j \neq i} V_{\alpha\beta}(r_{ij}) + F_{\alpha} \left(\sum_{j \neq i} \rho_{\alpha\beta}(r_{ij}) \right)$$

$$V_{AB}(r) = D \left(e^{-2\alpha(r - r_{eq})} - 2e^{-\alpha(r - r_{eq})} \right)$$

$$\rho_{BA}(r) = r^{6} \left(e^{-S_{A}r} + 2^{9}e^{-2S_{A}r} \right)$$

$$\rho_{AB}(r) = r^{6} \left(e^{-S_{B}r} + 2^{9}e^{-2S_{B}r} \right)$$

*L. Ward, A. Agrawal, K.M. Flores, and W. Windl. Rapid production of accurate embedded-atom method potentials for metal alloys. (2012). arXiv:cond-mat.mtrl-sci/1209.0619

Model output: **102** quantities of interest (QoIs) total. 17 compositions ranging from **3%** Au to **97%** Au For each composition, 6 physical properties:

- lattice parameter
- mixing enthalpy
- elastic constants C11, C12, C44
- bulk modulus

Indexing the 102 different QOIs:

finite composition space $x \in \mathcal{X}$ physical properties $i \in \{\text{lat}, \text{mix}, \text{C}_{11}, \text{C}_{12}, \text{C}_{44}, \text{bulk}\}$

BAYESIAN CALIBRATION WITH ADDITIVE MODEL ERROR



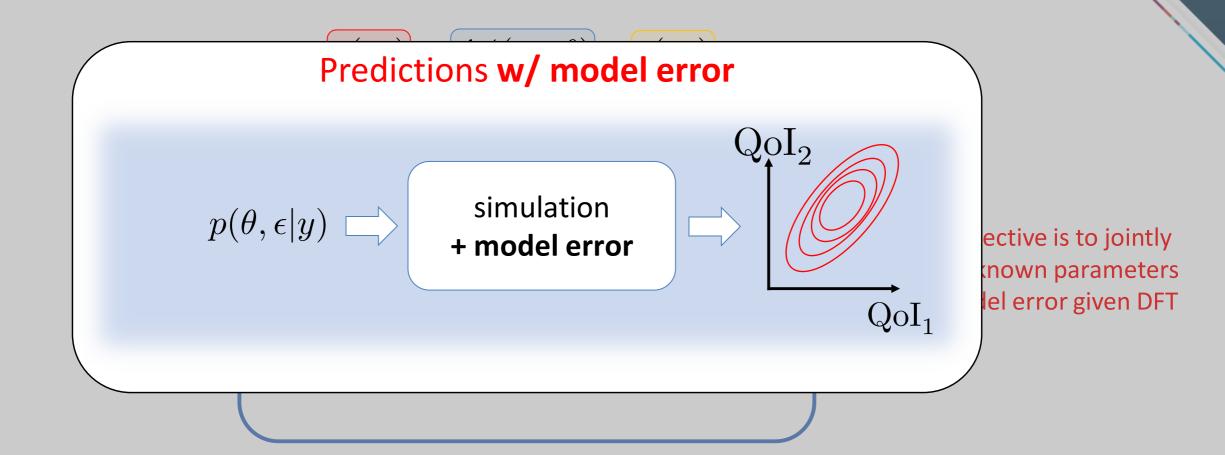
$$y(x_i) = \mathcal{M}(x_i; \theta) + \epsilon(x_i)$$
DFT data simulation model error

$$p(\theta, \epsilon | y) = \frac{p(y | \theta, \epsilon)p(\theta, \epsilon)}{p(y)}$$

Principal objective is to jointly infer the unknown parameters and the model error given DFT data.

BAYESIAN CALIBRATION WITH ADDITIVE MODEL ERROR

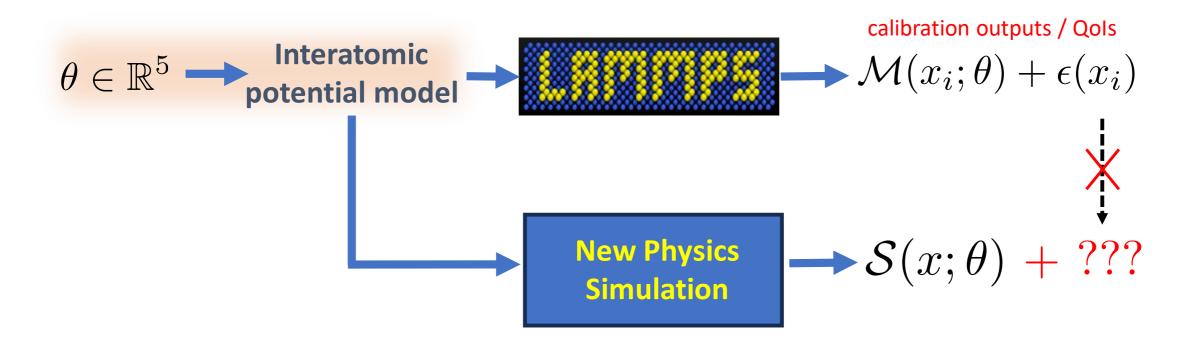




Kennedy, M. C., & O'Hagan, A. (2001). Bayesian calibration of computer models. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 63(3), 425-464.

TRANSFERABILITY TO NEW PREDICTION DOMAINS

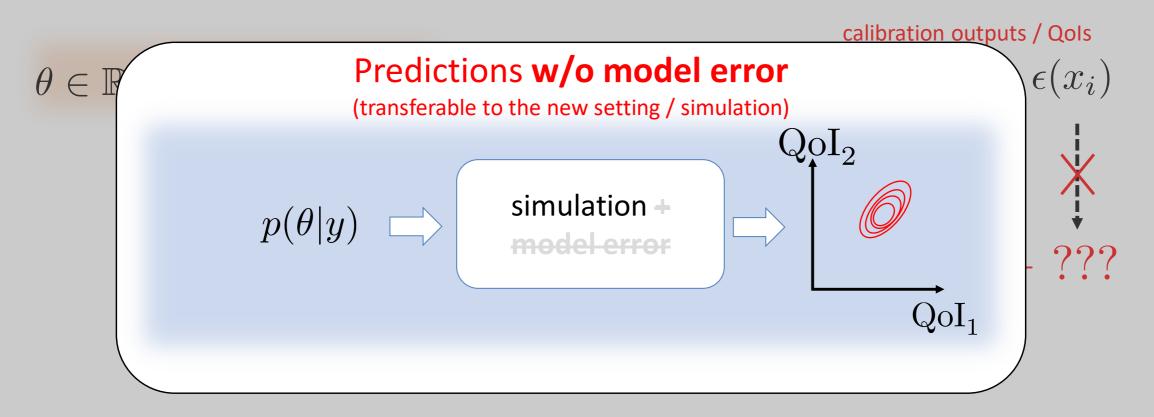




An additive error model **does not always** transfer to new prediction domains, particularly when the new prediction QoIs differ from the calibration QoIs

TRANSFERABILITY TO NEW PREDICTION DOMAINS



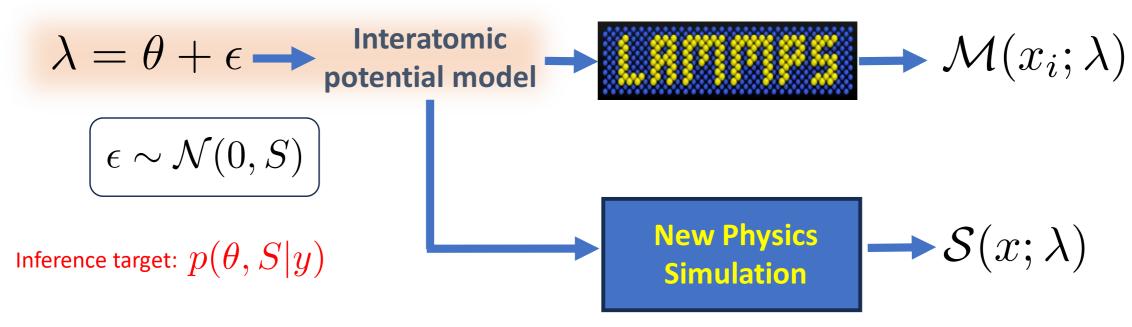


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EMBEDDED MODEL ERROR: MULTIVARIATE GAUSSIAN



Simplified form of embedded model error -- model error added to the model parameters



- Full uncertainty representation model error entirely <u>transferable</u> to new prediction domains
- Predictions constrained by the model predictive distributions supported only by values realizable by the model.
 - predictions with uncertainties satisfy the model constraints (e.g., conservation laws), but **also** inherit the biases of the model.

ABC-BASED LIKELIHOOD APPROXIMATION



ABC = Approximate Bayesian Computation

Challenge: the corresponding likelihood does not have a closed form.

Workaround: use an ABC-based approximate likelihood strategy*

$$p_{\mathrm{ABC}}(y|x,\theta,S) = \frac{1}{\nu\sqrt{2\pi}} \prod_{x,i}^{N} \exp\left(-\frac{\frac{\mathsf{nominal error}}{(\mu(x_i) - y(x_i))^2 + (\sigma(x_i) - \gamma|\mu(x_i) - y(x_i)|)^2}}{2\nu^2}\right)$$

$$\mu(x_i) = \mathbb{E}_{\lambda \sim \mathcal{N}(\theta, S)}[\mathcal{M}(x_i; \lambda)]$$

$$\sigma(x_i) = \sqrt{\mathbb{E}_{\lambda \sim \mathcal{N}(\theta, S)}[(\mathcal{M}(x_i; \lambda) - \mu(x_i))^2]}$$

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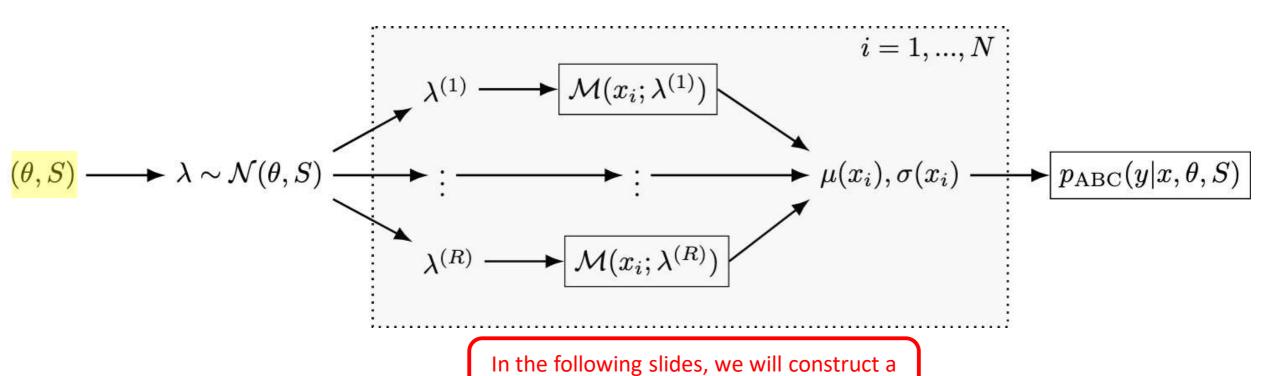
Expensive! Many model evaluations (samples) required per likelihood evaluation.

$$\mu(x_i) = \mathbb{E}_{\lambda \sim \mathcal{N}(\theta, S)}[\mathcal{M}(x_i; \lambda)]$$

$$\sigma(x_i) = \sqrt{\mathbb{E}_{\lambda \sim \mathcal{N}(\theta, S)}[(\mathcal{M}(x_i; \lambda) - \mu(x_i))^2]}$$

COST OF A SINGLE ABC-LIKELIHOOD EVALUATION





surrogate encompassing this boxed region.

GP SURROGATES WITH GAUSSIAN INPUTS



Computational expense of the simulation often necessitates the use of cheaper-to-evaluate surrogate models of the θ -to- $\mathcal{M}(x_i;\theta)$ map.

When the surrogate models are Gaussian processes (GPs) with Gaussian/RBF kernels,

$$k(\theta, \theta'; \nu) = \sigma_k^2 \exp\left(-\frac{1}{2}(\theta - \theta')^T \Gamma^{-1}(\theta - \theta')\right)$$

The GP prediction equations at a new parameter value are Gaussian with mean:

training data: $\{(\theta^j, q^j)\}_{i=1}^M$

$$m_{\star}(\theta) = m(\theta) + K(\theta, \Theta)\beta$$
 with $\beta = (K(\Theta, \Theta) + \sigma_n^2 I)^{-1}(q - m(\Theta))$

 σ_n^2 is from the GP likelihood (additive noise)

and variance:

$$\sigma_{\star}^{2}(\theta) = \sigma_{k}^{2} - K(\theta, \Theta)(K(\Theta, \Theta) + \sigma_{n}^{2}I)^{-1}K(\Theta, \theta)$$

GP SURROGATES WITH GAUSSIAN INPUTS



Surrogate model:
$$\mathcal{M}(x_i; \lambda) \longrightarrow \mathcal{N}(m_{\star}(\lambda), \sigma_{\star}^2(\lambda))$$

with embedding: $\lambda \sim \mathcal{N}(\theta, S)$

induces a pushforward density on the surrogate output

If the embedded model error is Gaussian, then we can analytically* compute the quantities that appeared in the ABC-likelihood,

$$m_{x_i}(\theta, S) = \mu(x_i) = \beta^T \ell$$

$$v_{x_i}(\theta, S) = \sigma^2(x_i) = \sigma_k^2 + \sigma_n^2 - \operatorname{trace}((K(\Theta, \Theta) + \sigma_n^2 I)^{-1} L)$$
$$+ \operatorname{trace}((L - \ell \ell^T) \beta \beta^T)$$

where $\ell \in \mathbb{R}^M$ and $L \in \mathbb{R}^{M \times M}$ require $\mathcal{O}(M)$ and $\mathcal{O}(M^2)$ computations (with respect to the surrogate model training data).

^{*}Candela, J. Q., Girard, A., Larsen, J., & Rasmussen, C. E. (2003, April). Propagation of uncertainty in Bayesian kernel models-application to multiple-step ahead forecasting. In 2003 IEEE International Conference on Acoustics, Speech, and Signal Processing, 2003. Proceedings. (ICASSP'03). (Vol. 2, pp. II-701). IEEE.

REDUCTION THROUGH SPARSE GPS AND MULTI-LEVEL MCMC

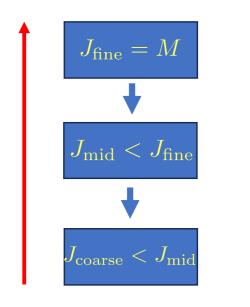


Challenge: each ABC-likelihood evaluation requires N surrogate evaluation, where each surrogate costs $\mathcal{O}(M^2)$ to <u>evaluate</u>.

Strategy: Build reduced-cost surrogate models via inducing point (pseudo data) approximations to GPs.

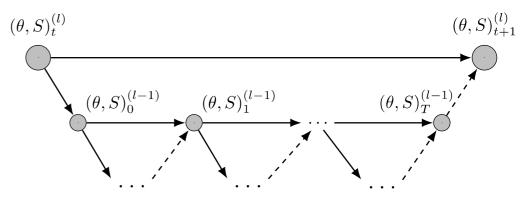
• Evaluation cost $\mathcal{O}(J^2)$ where J is the number of inducing points.

increasing fidelity, increasing evaluation cost



leads to likelihoods of varying fidelity

use multi-level MCMC algorithms for inference, e.g., multi-level delayed-acceptance (MLDA)*



*Lykkegaard, M. B., Dodwell, T. J., Fox, C., Mingas, G., & Scheichl, R. (2023). Multilevel delayed acceptance MCMC. SIAM/ASA Journal on Uncertainty Quantification, 11(1), 1-30.

CASE STUDY: CALIBRATING POTENTIALS WITH EMBEDDED ERROR



Case 1: Embedding model error into the 5 parameters *individually* v.s.

Case 2: Embedding into all 5 parameters *simultaneously* (includes all correlations between model parameters)

Case 1 Setup

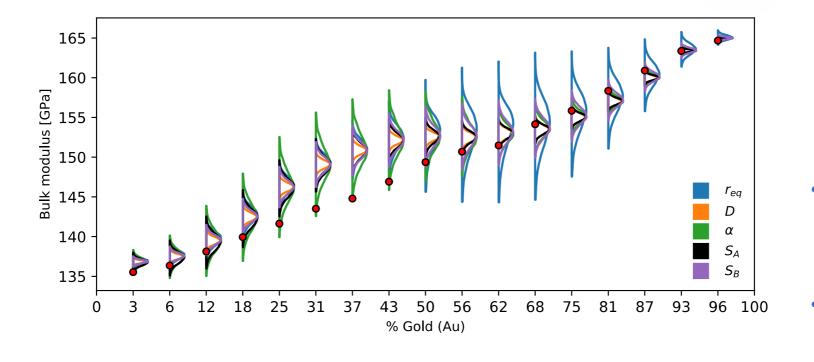
- model parameters + univariate embedding
- 102 Qols
- ABC-Likelihood w/ GP surrogates
- Adaptive MCMC
 - resulted in low autocorrelations, effective sample size > 2000

Case 2 Setup

- model parameters + multivariate embedding
- 102 Qols
- ABC-Likelihood w/ GP surrogates
- Adaptive MCMC: very high autocorrelations, even after > 1 week of sampling
 - Sparse GPs + Multilevel MCMC (MLDA)
 - 3 levels (w/ 500 and 1000 inducing points)
 - Very low autocorrelations, effective sample size > 2000



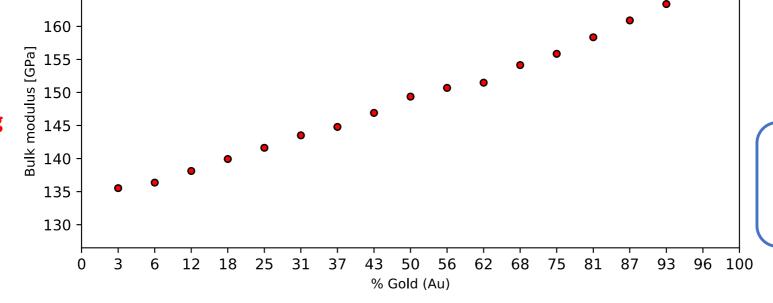
Individual embedding



- Embeddingdependent predictive capabilities
- Compositiondependent trends
- Full embedding generally provides better coverage of the data

Full embedding

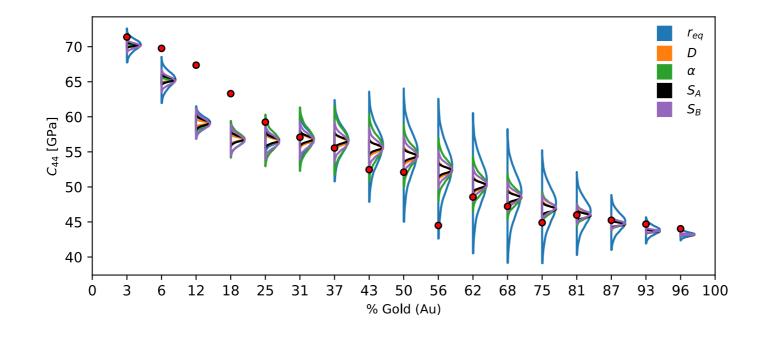
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In comparison with the additive model error, this uncertainty is fully transferable.

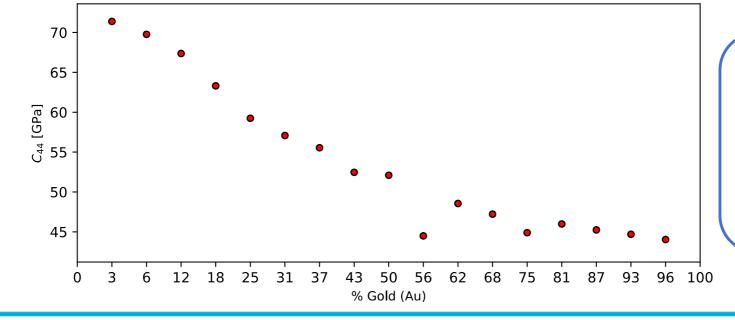


Individual embedding



- Embeddingdependent predictive capabilities
- Compositiondependent trends





Even with the full embedding, the model struggles to predict the DFT data at low compositions gold – reflects limitations of the underlying potential model.

SUMMARY



- UQ & calibration for interatomic potential models
- Embedded model error (Gaussian embedding)
 - ABC-likelihoods & GP surrogates
 - Further reduction through sparse GPs & multilevel MCMC
- Case study: predictive utility of embedded model error
 - Full uncertainty representation new predictions account for the calibrated model error
 - Depicted uncertainties are fully transferable to new prediction settings
 - Predictions constrained by the model predictions with uncertainties satisfy the model constraints, and in doing so also inherit the biases of the model.
 - Embedding-dependence and composition-dependence in predictions
 - Inability to predict certain Qols within uncertainty
 - Combined, these findings provide <u>new and useful diagnostic information</u> that was not present in our prior work.

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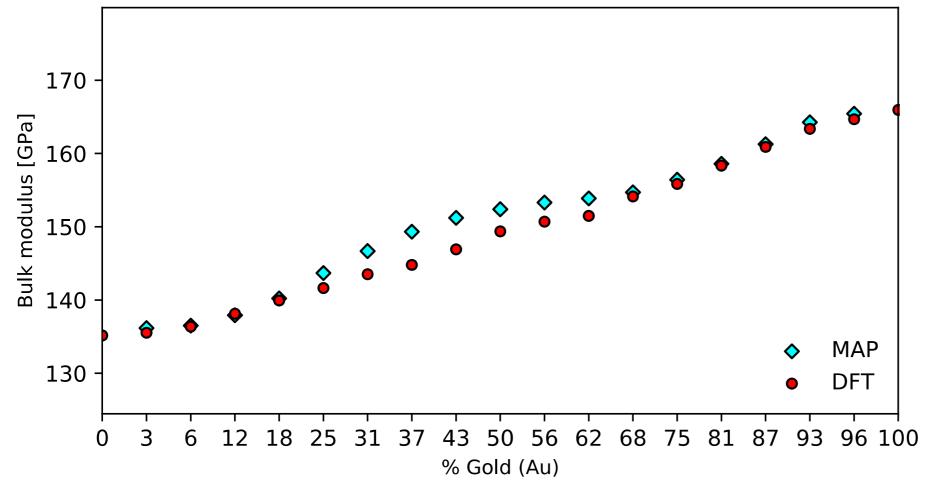


APPENDIX

ILLUSTRATION: PREDICTIONS WITH AND WITHOUT MODEL ERROR



With (left, grey) and without (right, blue) a simple form of additive error (iid Gaussian)



Notably,

- predictions w/o model error under-represent uncertainty.
- predictions with model error provide coverage of the DFT data.

In *new* prediction settings – where additive model error cannot be transferred -- the uncertainty will be considerably under-represented.

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$$\ell_i = \sigma_k^2 \frac{1}{\sqrt{\det(\Gamma_A^{-1}S + I)}} \exp\left(-\frac{1}{2}((\lambda_{\mathcal{A}^c} - \theta_{\mathcal{A}^c}^i)^T \Gamma_{\mathcal{A}^c}^{-1}(\lambda_{\mathcal{A}^c} - \theta_{\mathcal{A}^c}^i)\right) \exp\left(-\frac{1}{2}(\theta - \theta_{\mathcal{A}}^i)^T (S + \Lambda_{\mathcal{A}})^{-1}(\theta - \theta_{\mathcal{A}}^i)\right)$$

$$L_{ij} = k(\lambda|_{\lambda_{\mathcal{A}}=\theta}, \theta^{i})k(\lambda|_{\lambda_{\mathcal{A}}=\theta}, \theta^{j})$$

$$\cdot \frac{1}{\sqrt{\det(2\Gamma_{\mathcal{A}}^{-1}S + I)}} \exp\left(2\left(\theta - \frac{1}{2}(\theta_{\mathcal{A}}^{i} + \theta_{\mathcal{A}}^{j})\right)^{T} \Gamma_{\mathcal{A}}^{-1}(2\Gamma_{\mathcal{A}}^{-1} + S^{-1})^{-1}\Gamma_{\mathcal{A}}^{-1}\left(\theta - \frac{1}{2}(\theta_{\mathcal{A}}^{i} + \theta_{\mathcal{A}}^{j})\right)\right).$$

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