

# Hypothesis Testing

## Small samples

Also assuming

Null

- ① •  $H_0$  is true, ②
- The sample 1 points are iid  $N(\mu_1, \sigma_1^2)$ , the sample 2 points are iid  $N(\mu_2, \sigma_2^2)$ , ③
- and the sample 1 points are independent of the sample 2 points and  $\sigma_1^2 \approx \sigma_2^2$ . ④

Alternative

P-value

Then

CI method

5 assumptions

to use →

$$K = \frac{\bar{X}_1 - \bar{X}_2 - (\#)}{S_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \sim t_{(n_1+n_2-2)}$$

Matched Pairs

Two-sample

# Hypothesis Testing

Null

Alternative

P-value

CI method

Matched Pairs

Two-sample

## Small samples

$1 - \alpha$  confidence intervals (2-sided, 1-sided upper, and 1-sided lower, respectively) for  $\mu_1 - \mu_2$  under these assumptions are of the form:

(let  $\nu = n_1 + n_2 - 2$ )

- Two-sided  $100(1 - \alpha)\%$  confidence interval for  $\mu_1 - \mu_2$

$$(\bar{x}_1 - \bar{x}_2) \pm t_{(\nu, 1-\alpha/2)} * S_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

- One-sided  $100(1 - \alpha)\%$  confidence interval for  $\mu_1 - \mu_2$  with a upper confidence bound

$$(-\infty, (\bar{x}_1 - \bar{x}_2) + t_{(\nu, 1-\alpha)} * S_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)})$$

$\uparrow \equiv$   
 $n_1 + n_2 - 2$

# Hypothesis Testing

Null

Alternative

P-value

CI method

Matched Pairs

Two-sample

## Small samples

- One-sided  $100(1 - \alpha)\%$  confidence interval for  $\mu$  with a lower confidence bound

$$\Rightarrow ((\bar{x}_1 - \bar{x}_2) - t_{(\nu, 1-\alpha)} * S_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}, +\infty)$$

$\nu = n_1 + n_2 - 2$

In general:  $\hat{L} \pm$

estimate  $\pm$  population  $\times$  SD(estimate)  
quantiles

$$\mu = 0$$

$$\mu \neq 0$$

large sample

$$\hat{x} \pm t_{(1-\alpha)/2} \frac{s}{\sqrt{n}}$$

\* Hw 99 posted (optional)

Due Thursday Dec. 12.

\* Final exam:

- Wednesday, Dec 18

9:45 - 11:45 (in-class)

- Comprehensive with focus on  
materials after quiz 3.

# Hypothesis Testing

Null

Alternative

P-value

CI method

Matched Pairs

Two-sample

## Small samples

### Example:[Springs]

The data of W. Armstrong on spring lifetimes (appearing in the book by Cox and Oakes) not only concern spring

② longevity at a 950 N/ mm<sup>2</sup> stress level but also longevity at a 900 N/ mm<sup>2</sup> stress level. ①

Let sample 1 be the 900 N/ mm<sup>2</sup> stress group and sample 2 be the 950 N/ mm<sup>2</sup> stress group.

900 N/mm <sup>2</sup> Stress	950 N/mm <sup>2</sup> Stress
216, 162, 153, 216, 225, 216, 306, 225, 243, 189	225, 171, 198, 189, 189, 135, 162, 135, 117, 162

$n_1 = 10$        $n_2 = 10$

# Hypothesis Testing

Null

Alternative

P-value

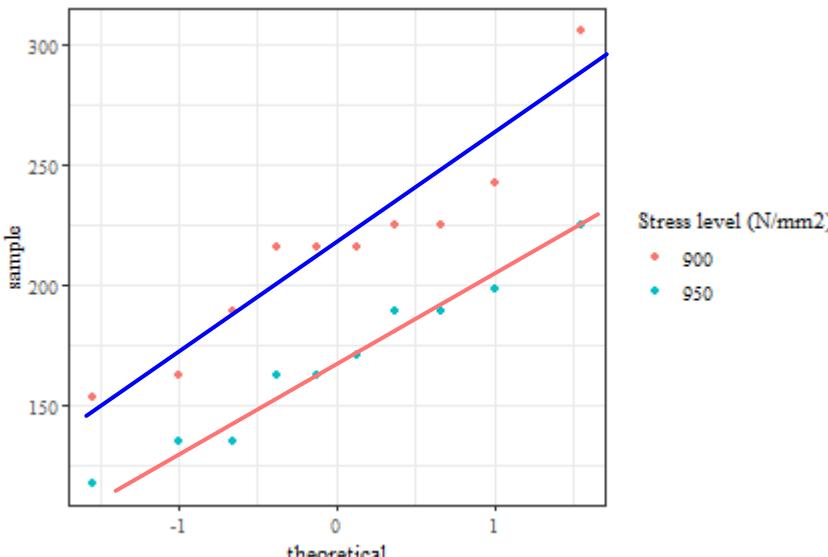
CI method

Matched Pairs

Two-sample

Small samples

Example:[Springs]



normal plots of  
Spring lifetime under two  
different stress level.

Let's do a hypothesis test to see if the sample 1 springs lasted significantly longer than the sample 2 springs. (For  $\alpha=0.05$ )

Note: however the sample sizes are small,  
the data look pretty normal.

$$\begin{cases} H_0: \mu_1 - \mu_2 = 0 \\ H_a: \mu_1 - \mu_2 > 0 \end{cases}$$

$$2, \alpha = 0.05$$

3, If we assume that  $H_0$  is true & sample 1 is iid

$\textcircled{1}$   $N(\mu_1, \sigma_1^2)$  independent from sample 2 iid  $\textcircled{2} N(\mu_2, \sigma_2^2)$

and  $\textcircled{3} \sigma_1^2 \approx \sigma_2^2$ , then the test statistic is

$$\rightarrow K = \frac{(\bar{x}_1 - \bar{x}_2) - 0}{\sqrt{s_p \frac{\sum y_{n_1}}{n_1} + \sum y_{n_2}}} \sim t_{(n_1 + n_2 - 2)}$$

$$4, \text{ calculate } \bar{x}_1 = 215.1, \quad S_1 = 42.9, \quad S_1^2 = 1840.41$$

$$\bar{x}_2 = 168.3, \quad S_2 = 33.1, \quad S_2^2 = 1095.61$$

$$\Rightarrow S_p = \sqrt{\frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1+n_2-2}} = \sqrt{\frac{(10-1)1840.41 + (10-1)1095.61}{(10+10-2)}}$$

$$\Rightarrow S_p = 38.3$$

$$\text{Then } t = \frac{(\bar{x}_1 - \bar{x}_2) - 0}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{(215.1 - 168.3)}{38.3 \sqrt{\frac{1}{10} + \frac{1}{10}}} = 2.7$$

Recall:

We need to calculate p-value now.

\* Corresponding to  $\alpha$ :  $t_{\frac{n_1+n_2-2}{}, 1-\alpha} = t_{(18, 0.95)}$

by table = 1.73

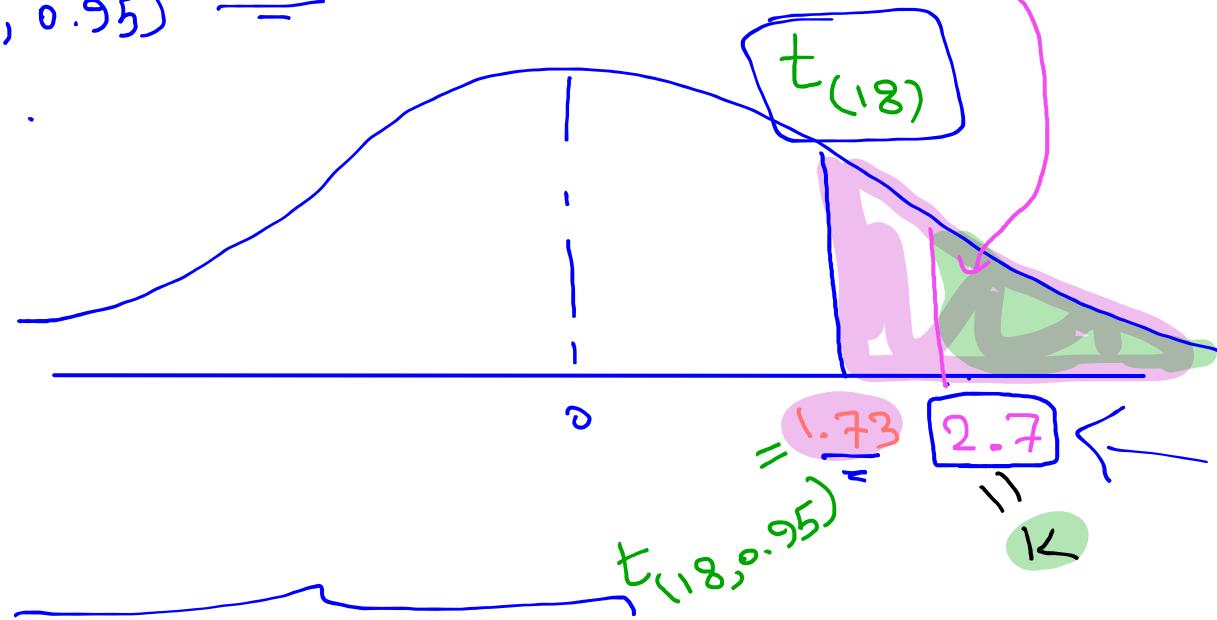
(By the methods we learned about p-value  
and  $\alpha$ , we just need to decide if  $p\text{-value} \leq \alpha$ )

→ \* Corresponding to p-value  $\equiv k = 2.7$

Now:

one sided test

$$\left. \begin{array}{l} p\text{-value} = P(T > K) = P(T > 2.7) \\ t_{(18, 0.95)} = 1.73 \end{array} \right\}$$



5. Since  $K > 1.73 (= t_{(18, 0.95)})$

(area under the curve for  $P(T > K)$  is smaller than  $\alpha$ )  $\Rightarrow p\text{-value} < \alpha \Rightarrow$  Reject  $H_0$

b) There is enough evidence to conclude  
that springs on average last longer subjected  
to  $900 \text{ N/mm}^2$  of stress than  $\underline{\underline{950 \text{ N/mm}^2}}$  of  
stress.

# Hypothesis Testing

Null

Alternative

P-value

CI method

Matched Pairs

Two-sample

## Small samples

### Example:[Stopping distance]

Suppose  $\mu_1$  and  $\mu_2$  are true mean stopping distances (in meters) at 50 mph for cars of a certain type equipped with two different types of breaking systems.

Suppose  $n_1 = n_2 = 6$ ,  $\bar{x}_1 = 115.7$ ,  $\bar{x}_2 = 129.3$ ,  $s_1 = 5.08$ , and  $s_2 = 5.38$ .

Use significance level  $\alpha = 0.01$  to test  $H_0 : \mu_1 - \mu_2 = -10$  vs.  $H_A : \mu_1 - \mu_2 < -10$ .

Construct a 2-sided 99% confidence interval for the true difference in stopping distances.

$$\left\{ \begin{array}{l} H_0: \mu_1 - \mu_2 = -10 \\ H_{\text{ar}}: \mu_1 - \mu_2 < -10 \end{array} \right.$$

$$② \alpha = 0.01$$

③ Under the assumptions that ①  $H_0$  is true and

② sample 1 is iid  $N(\mu_1, \sigma_1^2)$  ③ independent of

④ sample 2 iid  $N(\mu_2, \sigma_2^2)$  and ⑤  $\sigma_1^2 \approx \sigma_2^2$  ✓

we use test statistic

$$K = \frac{\bar{x}_1 - \bar{x}_2 - (-10)}{\sqrt{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}}$$

where  $K \sim t_{(n_1+n_2-2)}$

$$\textcircled{4} \quad S_p = \sqrt{\frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}} = \sqrt{\frac{(6-1)5.08^2 + (6-1)5.38^2}{6+6-2}}$$

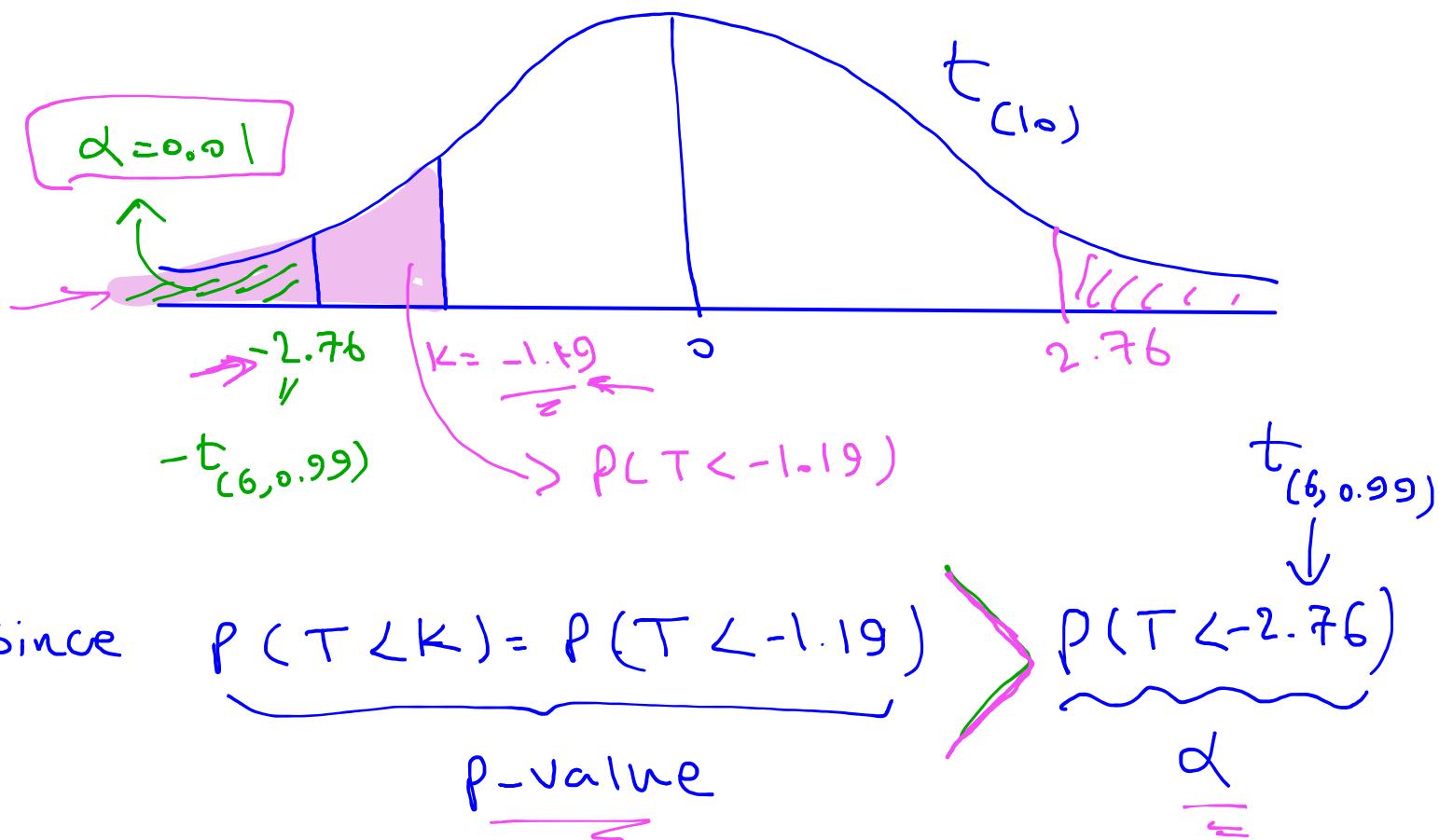
$$\rightarrow S_p = 5.23$$

$$\rightarrow K = \frac{(\bar{x}_1 - \bar{x}_2) - (-10)}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{(115.7 - 129.3) + 10}{5.23 \sqrt{\frac{1}{6} + \frac{1}{6}}} = -1.19$$

$$\left\{ \begin{array}{l} H_0: \mu_1 - \mu_2 = -10 \\ H_a: \mu_1 - \mu_2 < -10 \end{array} \right. \rightarrow p\text{-value} = P(T < K) = P(T < -1.19)$$

$T \sim t_{(n_1+n_2-2)}$

by the table of  $t$   $t_{(6+6-2, 1-\alpha)} = t_{(10, 0.99)} = 2.76$



Since  $P(T < K) = P(T < -1.19)$

we fail to reject  $H_0$ .

- ⑤ Since p-value >  $\alpha$ , we fail to reject H<sub>0</sub>
- ⑥ There is Not enough evidence to conclude  
that stopping distances for breaking system 1 are  
on average less than those of breaking system  
2 by over 10 m.
- 

Construct 2-sided 99% CI for  $\mu_1 - \mu_2$  (the true  
difference of mean stopping distance) %

using formulas

$$\rightarrow (\bar{x}_1 - \bar{x}_2) \stackrel{-}{=} t_{(n_1+n_2-2, 1-\alpha/2)} \quad \text{and} \quad (\bar{x}_1 - \bar{x}_2) \stackrel{+}{=} t_{(n_1+n_2-2, \alpha/2)}$$

$$\text{Sp} \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

$$\text{Sp} \sqrt{\frac{t_{\alpha/2}}{n_1} + \frac{t_{\alpha/2}}{n_2}}$$

$$10 \quad 1 - \frac{0.01}{2} = 0.995$$

$$= ((115.7 - 129.3) - 3.17 \cdot (5.23) \sqrt{\frac{1}{6} + \frac{1}{6}}, (115.7 - 129.3) + 3.17 \cdot (5.23) \sqrt{\frac{1}{6} + \frac{1}{0}})$$

$$= (-23.17, -4.03)$$

we are 99% confident that the true mean stopping distance of system 1 is anywhere between 23.17m to 4.03m less than that of system 2.

