

UQ for Credit Risk Management

Deep Evidence Regression approach for Loss Given Default

Paper Presentation, CSE 8803-IUQ

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Agenda

- **Introduction**

- Credit Risk Management and Loss Given Default
- UQ for Loss Given Default
- Evidential learning

- **Approach**

- Deep Evidence Regression
 - Pros
 - Limitations
- Improved Deep Evidence Regression
 - Intuition
 - Improvement made

- **Results**

- Simulated data
- Real Data

- **Limitations**

What is Credit Risk ?

- Credit Risk Measures:

- *Probability of Default*: Creditor unable to comply with his commitments
- *Loss given Default*: Relative fraction of the outstanding amount is lost
- *Exposure to Default*: Outstanding amt at the time of default

- Credit Risk Instruments:

- Fixed Income assets like Corporate Bonds, sovereign bonds from specific countries
- *Loans*
- Mortgages etc.



UQ for Credit Risk Management?

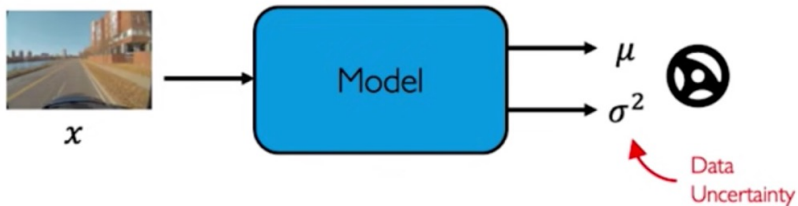
- Credit Default is a very rare event
 - Data on credit default is very limited
 - In such cases UQ is important because we are not sure of our prediction
 - Quantifying Uncertainty also helps with regulatory angle
- Quantifying uncertainty of machine learning methods for LGD *
 - Deep evidential regression as a UQ framework
 - Bond loss given defaults using Moody's Default and Recovery Database (*not public*)
 - Article also analyses the ratio between aleatoric and epistemic uncertainty
 - Also invokes XAI techniques like ALE plots for regulatory requirements

[*Max Nagl et al \(2022\)](#)

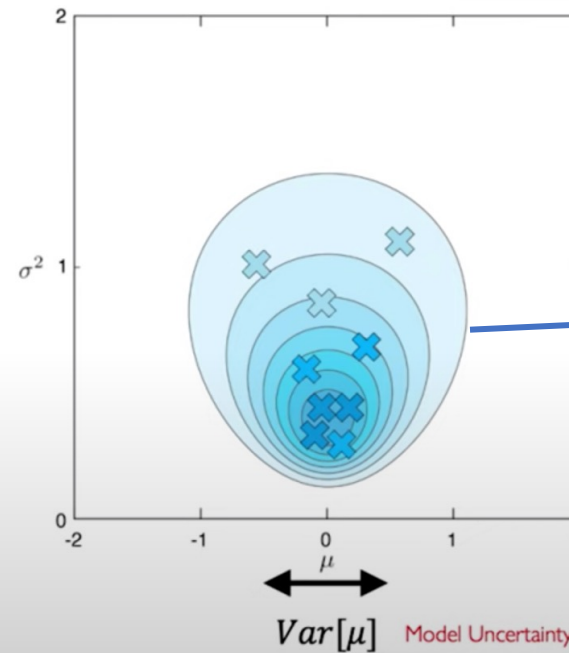
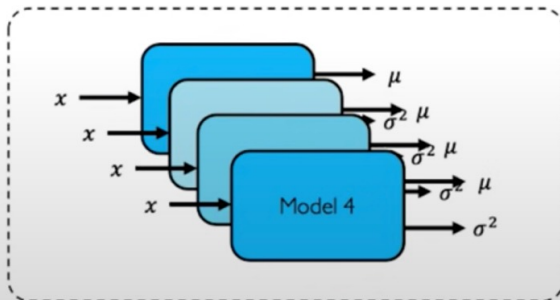
Deep Evidential learning

Model parameters as generated from higher order evidential distribution

- Deep Evidence Classification
- Deep Evidence Regression



Ensemble



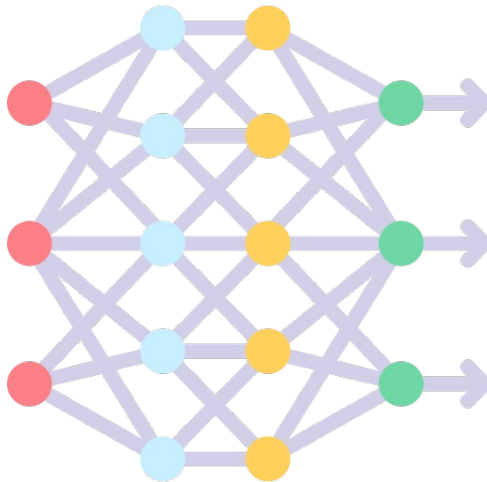
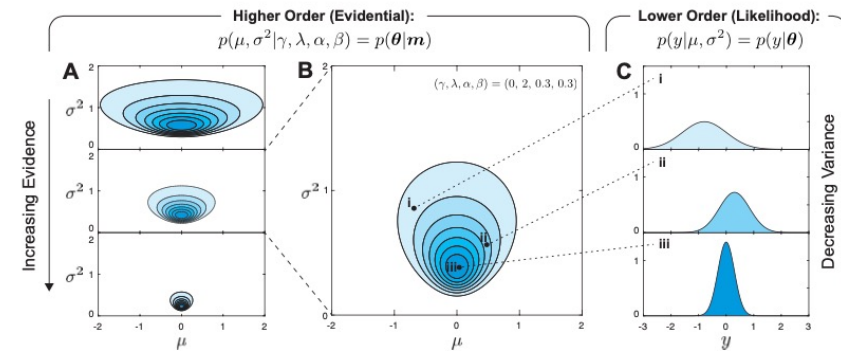
Higher Order
Evidential Distribution

Deep Evidence Regression

- Assumes target variable (y) is **Normally distributed** with unknown parameters
- Different Prior distributions are then placed on these unknown parameters

$$(y_1, \dots, y_N) \sim \mathcal{N}(\mu, \sigma^2)$$

$$\mu \sim \mathcal{N}(\gamma, \sigma^2 v^{-1}) \quad \sigma^2 \sim \Gamma^{-1}(\alpha, \beta)$$



NN output = $\mathbf{m} = \{\alpha, \beta, v, \gamma\}$

NN Loss = $L_{\{NLL\}} + c * L_{\{regularisation\}}$

$$p(y_i | \mathbf{m}) = \frac{p(y_i | \boldsymbol{\theta}, \mathbf{m}) p(\boldsymbol{\theta} | \mathbf{m})}{p(\boldsymbol{\theta} | y_i, \mathbf{m})} = \int_{\sigma^2=0}^{\infty} \int_{\mu=-\infty}^{\infty} p(y_i | \mu, \sigma^2) p(\mu, \sigma^2 | \mathbf{m}) d\mu d\sigma^2$$

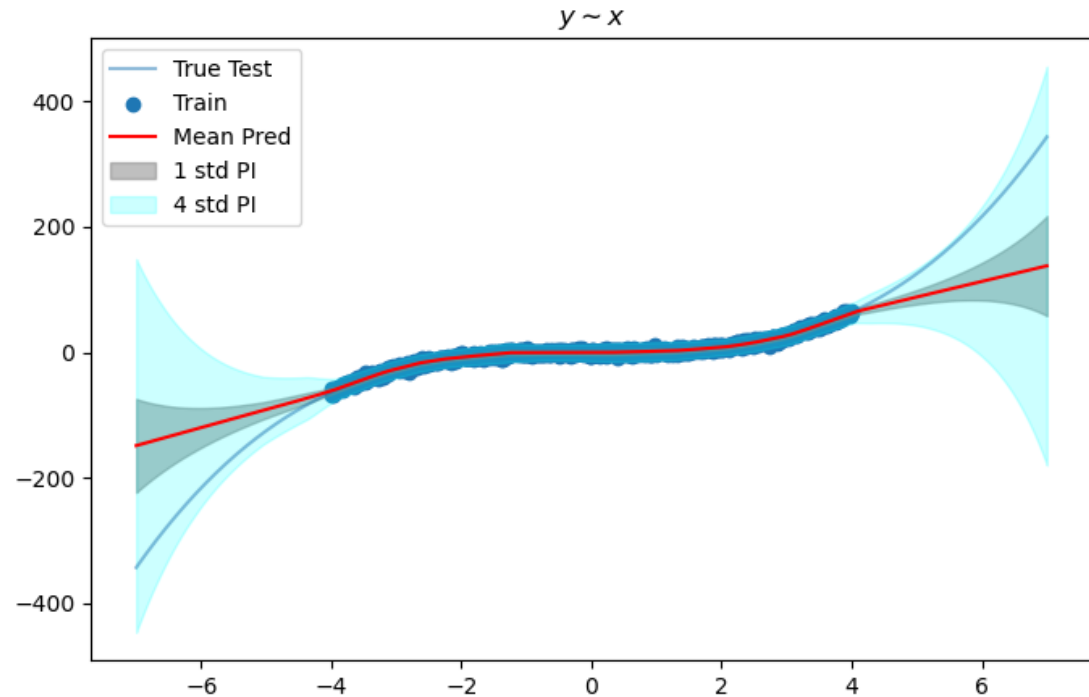
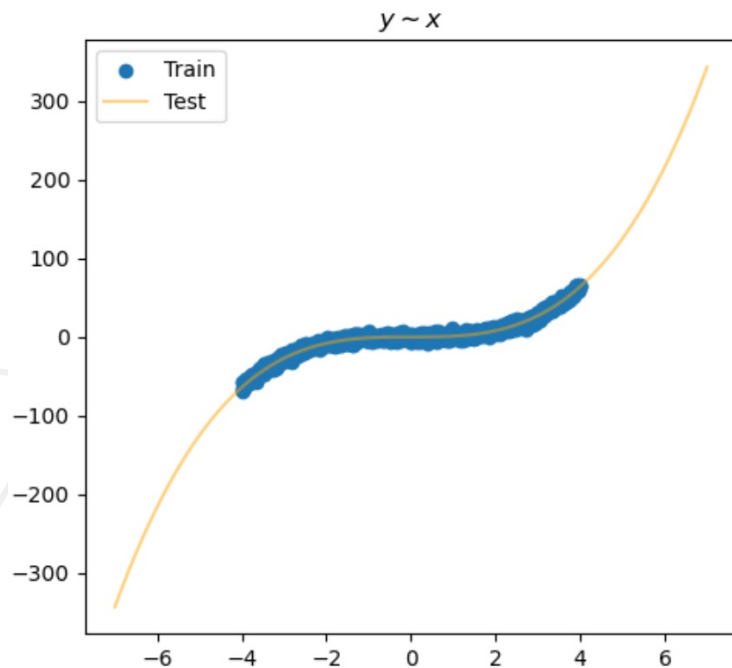
$$\underbrace{\mathbb{E}[\mu] = \gamma}_{\text{prediction}}$$

$$\underbrace{\mathbb{E}[\sigma^2] = \frac{\beta}{\alpha-1}}_{\text{aleatoric}}$$

$$\underbrace{\text{Var}[\mu] = \frac{\beta}{v(\alpha-1)}}_{\text{epistemic}}$$

Deep Evidence Regression: Pros

- Allows analytical computation of uncertainty for OOD Test data
- Example:
 - Data generated by $y = x^3 + \epsilon, \epsilon \sim N(0, \sigma^2)$
 - Train set, $x \in [-4, 4]$, Test set, $x \in [-7, 7]$



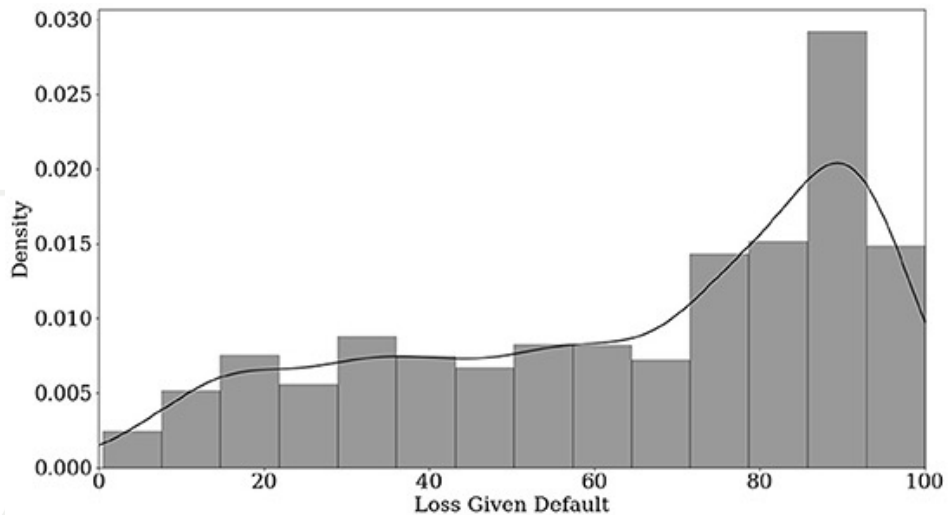
[*example recreated from Amini et al \(NeurIPS 2020\)](#)

Deep Evidence Regression: Cons

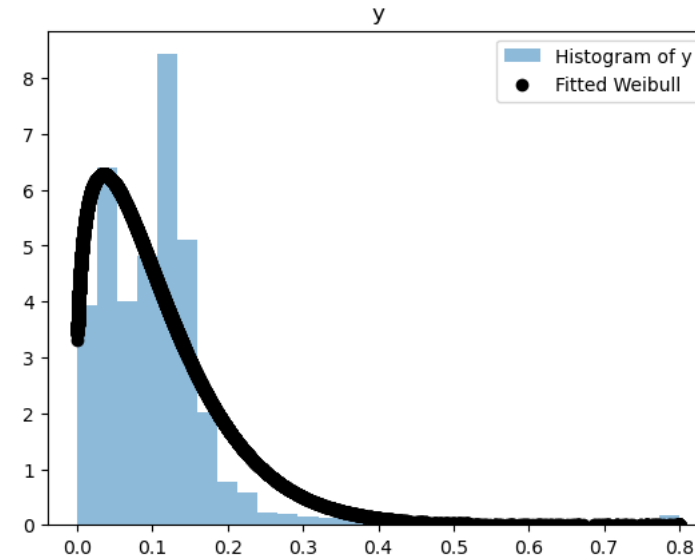
- Assumes target variable (y) is **Normally distributed** with unknown parameters
- Not suited for Loss Given Default and other credit risk measures

$$(y_1, \dots, y_N) \sim \mathcal{N}(\mu, \sigma^2) \longrightarrow \text{Not true for LGD}$$
$$\mu \sim \mathcal{N}(\gamma, \sigma^2 v^{-1}) \quad \sigma^2 \sim \Gamma^{-1}(\alpha, \beta).$$

Loss given defaults for Bonds



Recovery Rate (Home Loans)



[*image from Max Nagl et al \(2022\)](#)

Improved Deep Evidence Regression: Intuition

- What if we assume Target variable comes from Weibull Distribution
 - $y_i \sim Weibull(\lambda, k)$, k is shape parameter, λ is scale parameter
 - $f(y_i|\lambda, k) = \frac{k}{\lambda} (\frac{y_i}{\lambda})^{k-1} \exp(-(\frac{y_i}{\lambda})^k)$
- What are the prior distributions in that case ?
 - If we assume both k and λ are unknown
 - $k \sim \Gamma^{-1}(\alpha_1, \beta_1)$, $\lambda \sim \Gamma^{-1}(\alpha_2, \beta_2)$
 - The computation of Log Likelihood $p(y_i|\alpha_1, \beta_1, \alpha_2, \beta_2)$ is intractable

Improved Deep Evidence Regression

- Instead assume:

- k or shape is known, and λ scale is unknown
 - k can be estimated from training data (Assumption)

- Let $\theta = \lambda^k \sim \Gamma^{-1}(\alpha, \beta)$

- $p(y_i | \theta) = \frac{k}{\theta} y_i^{k-1} \exp(-\frac{y_i^k}{\theta})$

- $p(y_i | \alpha, \beta) = \int \frac{k}{\theta} y_i^{k-1} \exp(-\frac{y_i^k}{\theta}) * \frac{\beta^\alpha}{\Gamma(\alpha)} \frac{1}{\theta^{\alpha+1}} \exp(-\frac{\beta}{\theta}) d\theta$

- $p(y_i | \alpha, \beta) = \frac{\alpha k y_i^{k-1} \beta^\alpha}{(y_i^k + \beta)^{\alpha+1}}$

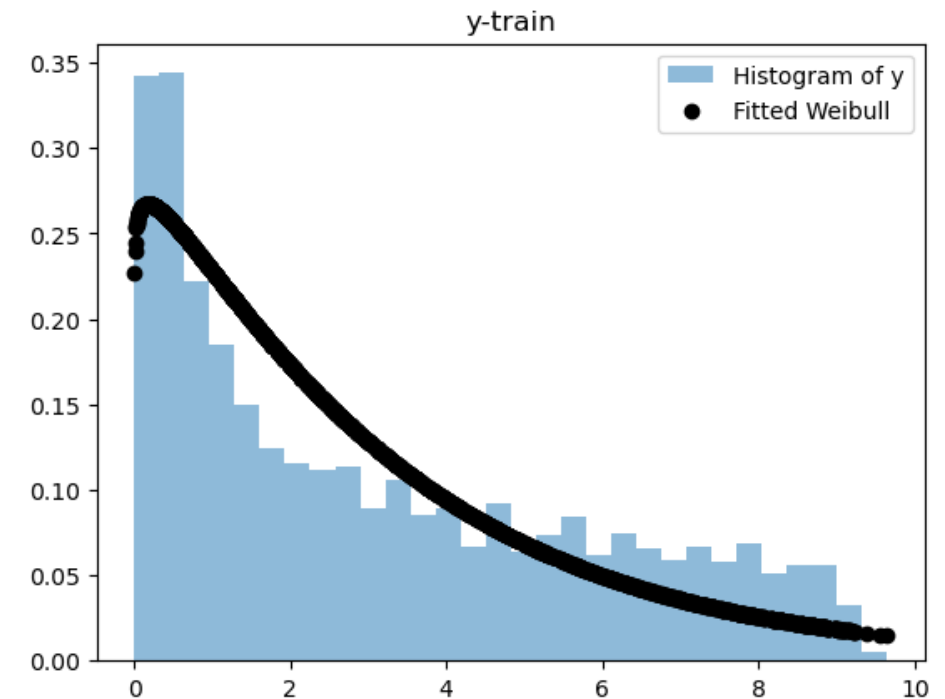
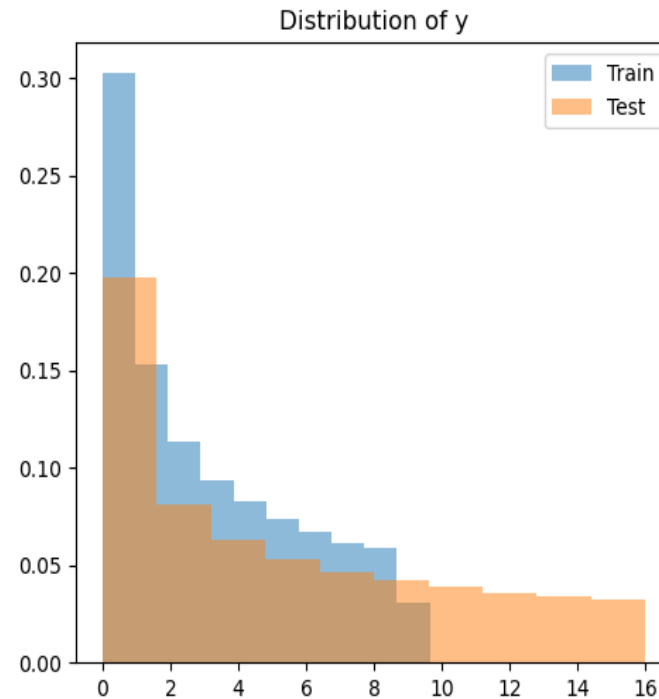
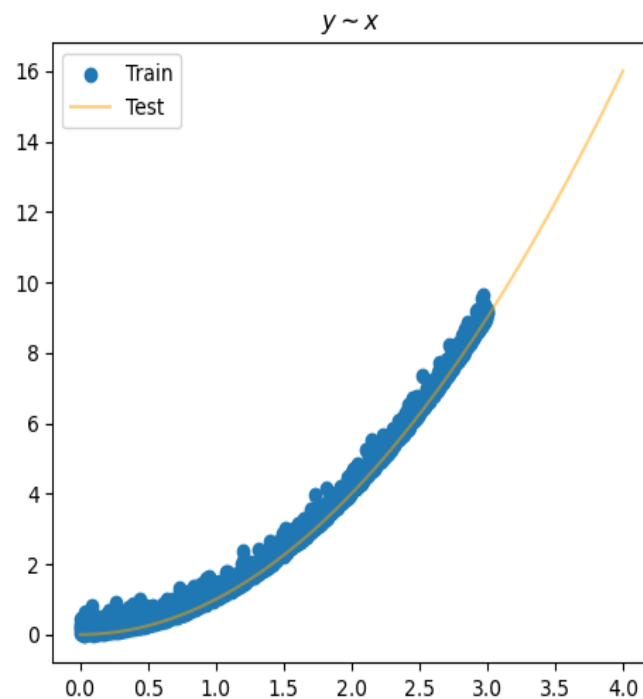
Log Likelihood of y given
evidential parameters

Mean
Prediction

- $Z = E[y_i | \alpha, \beta] = \Gamma\left(1 + \frac{1}{k}\right) E(\lambda | \alpha, \beta) = \Gamma\left(1 + \frac{1}{k}\right) \frac{1}{\Gamma(\alpha)} \Gamma\left(\alpha - \frac{1}{k}\right) \beta^{1/k}$

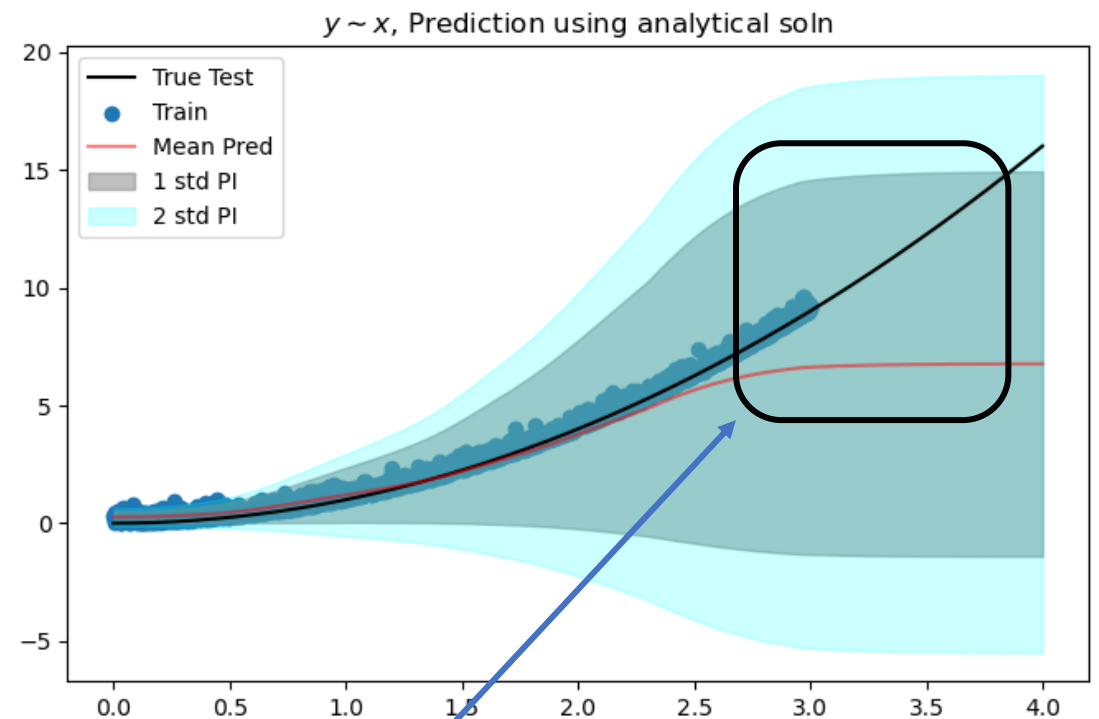
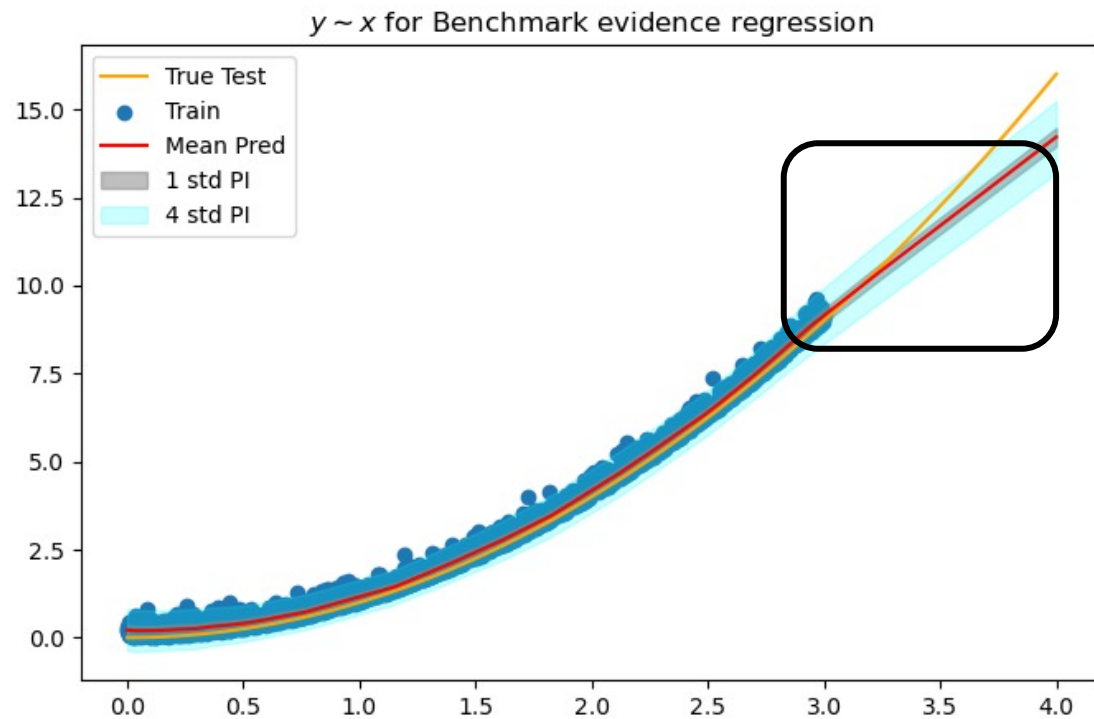
Results: Simulated Data

- Data generated by $y = x^2 + \epsilon, \epsilon \sim Weibull(k = 1.2, \lambda = 0.2)$
 - Train set, $x \in [0, 3]$
 - Test set, $x \in [0, 4]$
 - K from Weibull fit on train data = 1.05



Results: Simulated Data

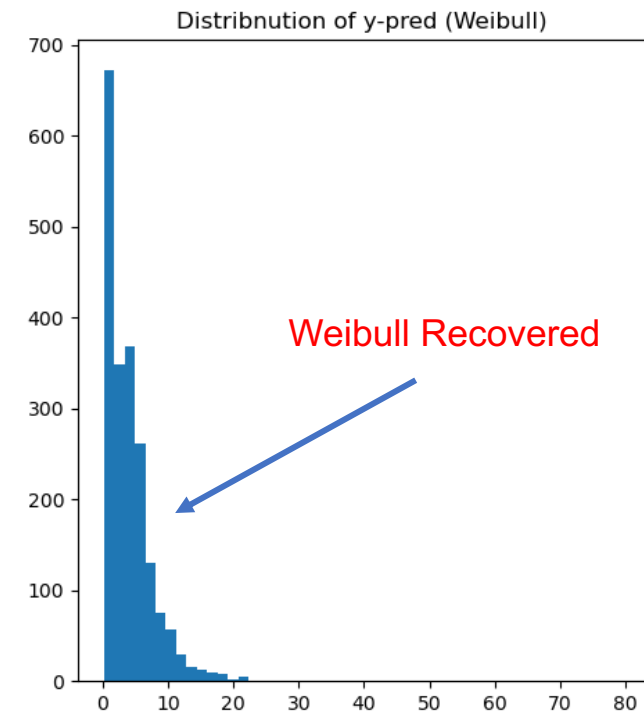
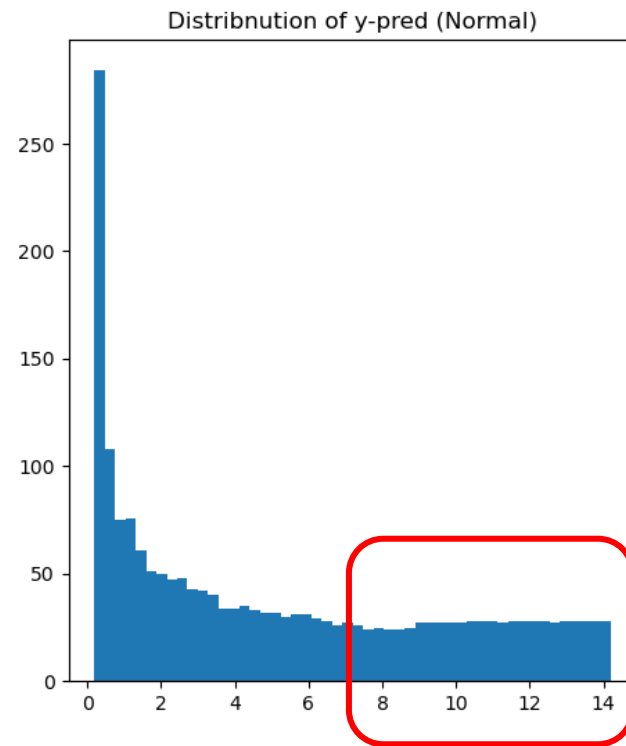
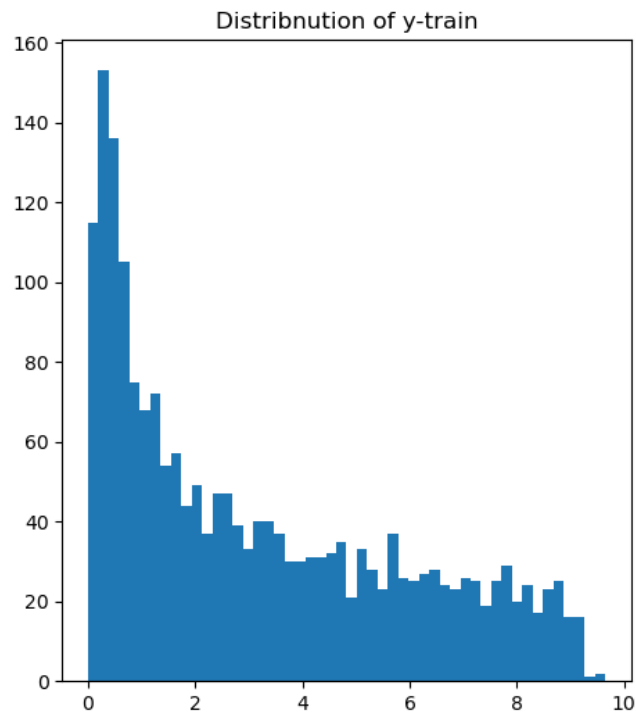
- For Weibull data Normal Evidence not able to capture Uncertainty for OOD
- Improved version captures uncertainty but not able to capture true function



Uncertainty increases
in test data

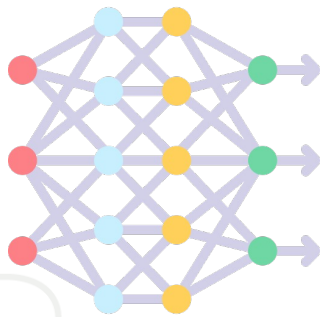
Results: Simulated Data

- Improved version also better at recovering original distribution of target

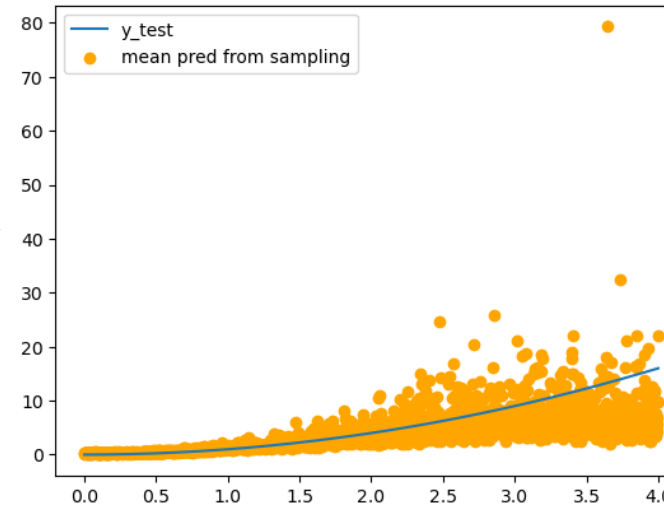


Results: Evaluation of analytical calculations

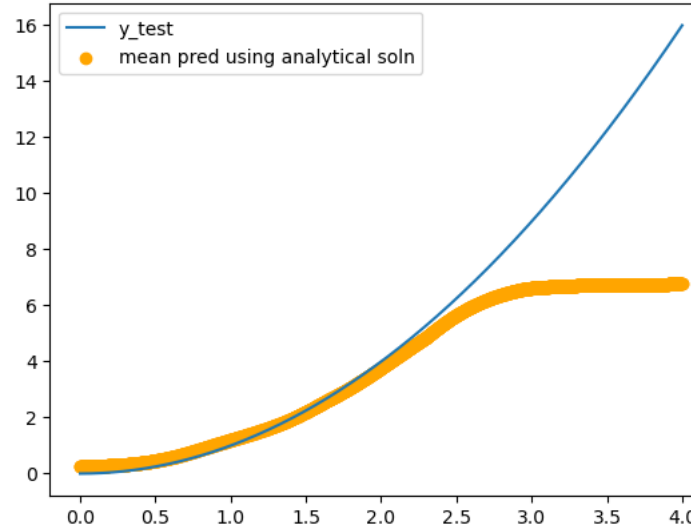
- Results from analytical computation match with those generated from sampling



$NN\ output$
 $= \{\alpha, \beta\}$



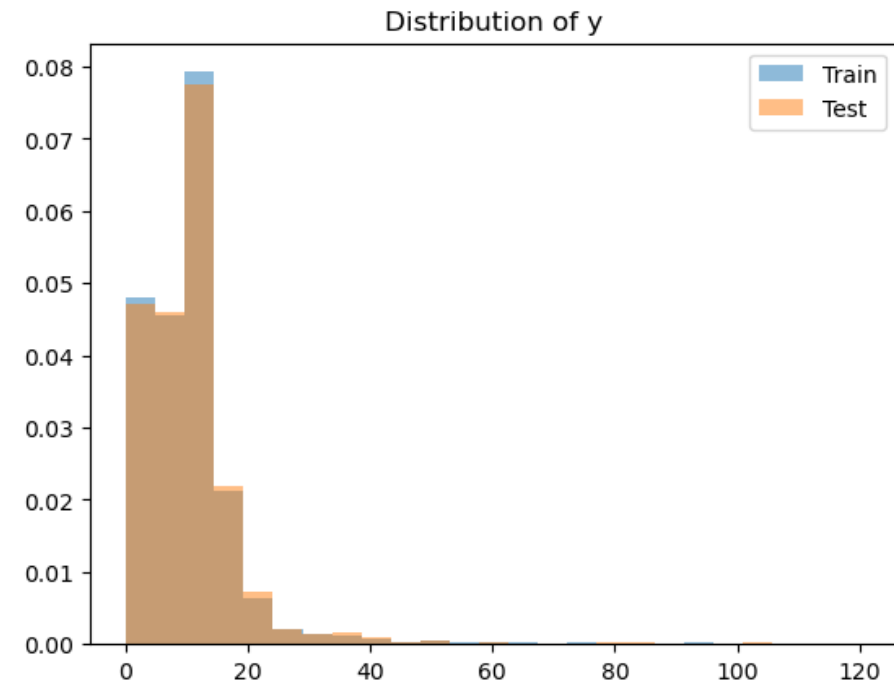
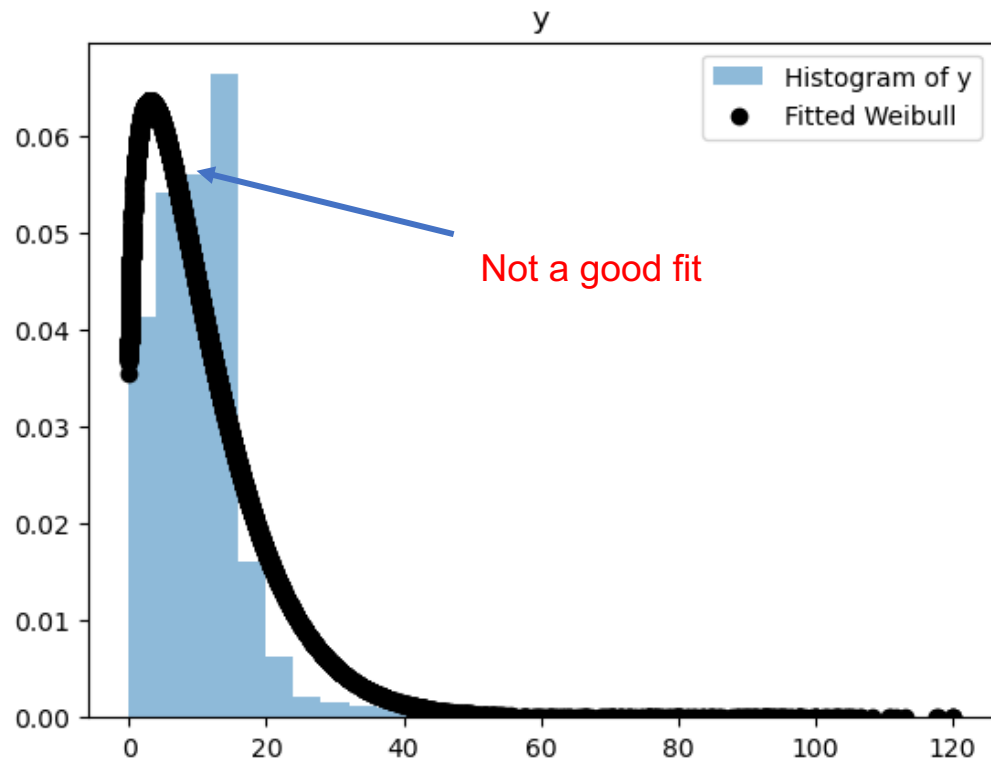
$$\begin{aligned}\theta &\sim \Gamma^{-1}(\alpha, \beta) \\ \lambda &= \theta^{1/k} \\ y_i &\sim \text{Weibull}(k, \lambda)\end{aligned}$$



$$\begin{aligned}Z &= E[y_i|\alpha, \beta] = \Gamma\left(1 + \frac{1}{k}\right) E(\lambda|\alpha, \beta) = \Gamma\left(1 + \frac{1}{k}\right) \frac{1}{\Gamma(\alpha)} \\ &\quad \Gamma\left(\alpha - \frac{1}{k}\right) \beta^{1/k}\end{aligned}$$

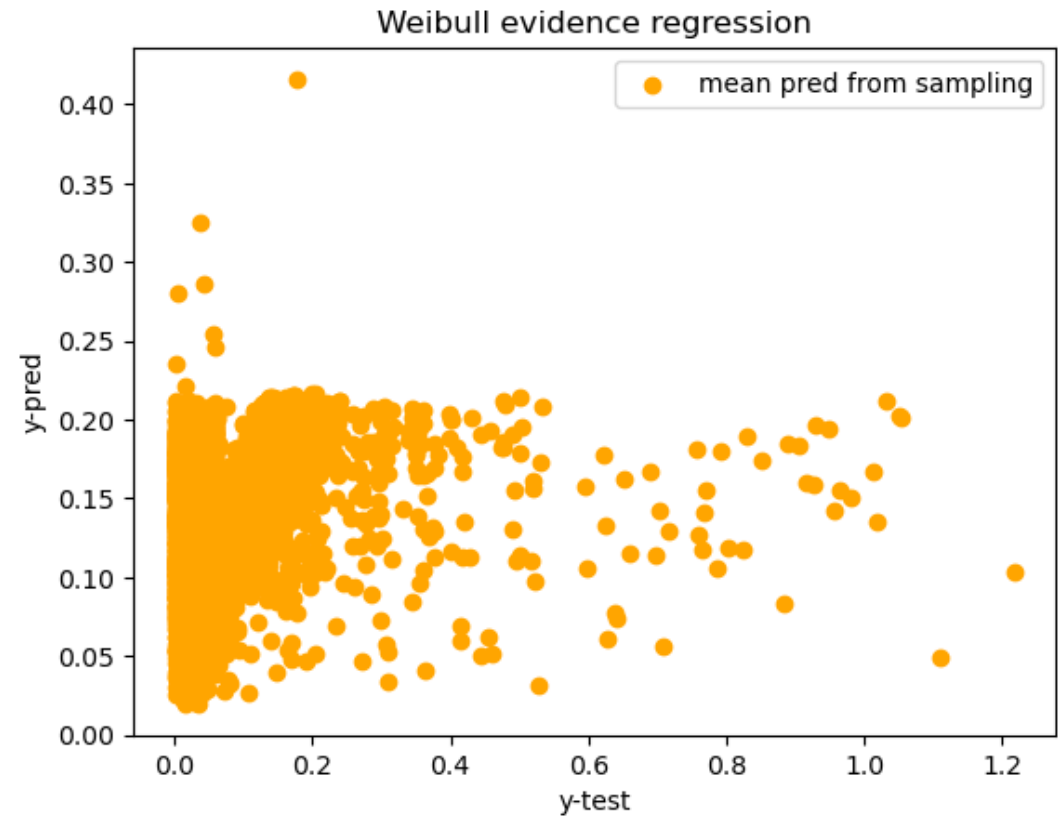
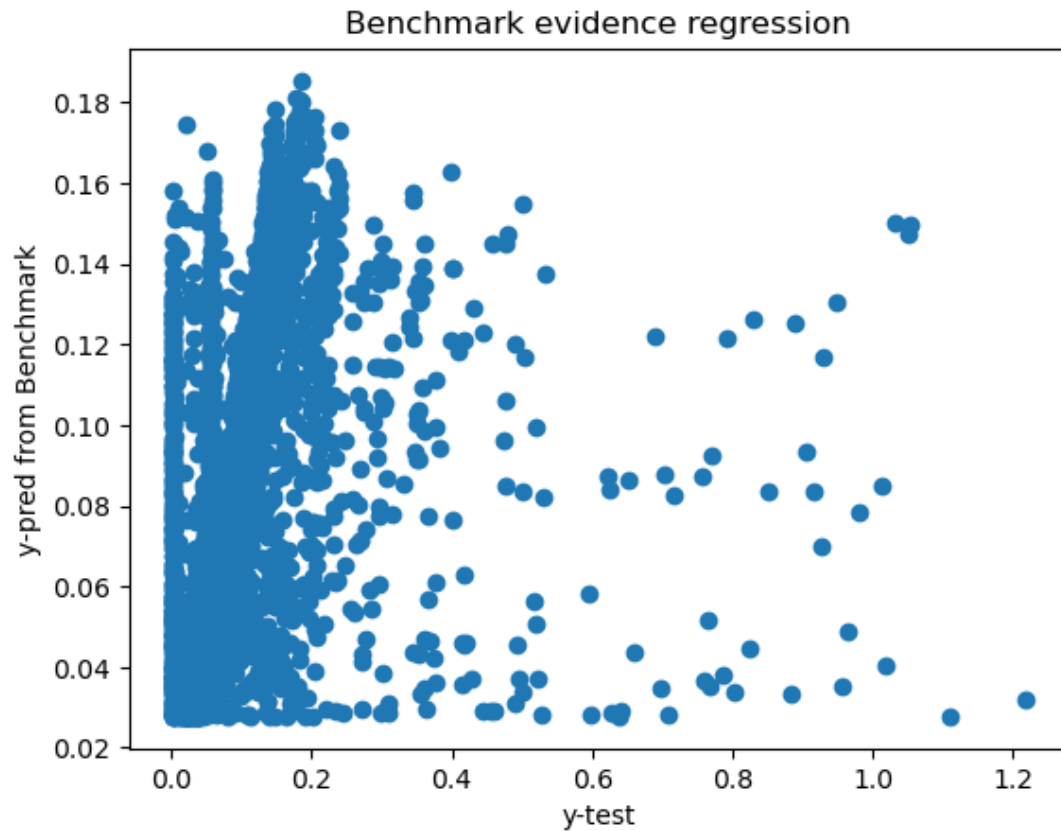
Results: Real Data

- Mortgage data from Peer to Peer lending (2007-20011)
 - Loss given default unavailable in data
 - Recovery rate used as proxy = (recoveries made/origination amount)
 - 43 'x' variables in data: time since loan, dti, joint, delinquency etc
 - ~23k rows in data



Results: Real Data

- Mortgage data from Peer to Peer lending (2007-20011)



Limitations

- Unsure of Variance calculation
- Both models very sensitive to Regularization cost
- Improved might only work for Weibull targets
- Improved requires more deep network
 - Benchmark had 4 outputs from NN – more flexible
 - Improved has 2 outputs so less flexibility

Discussion/Questions

- The presentation evolved into a very fruitful discussion, with key questions on:
 - ***How is Deep Evidential learning helpful:*** We discussed how with Evidential learning one can quantify uncertainty without sampling or keeping multiple copies of the network.
 - ***How is log-likelihood calculated:*** We summarized that log likelihood is calculated by finding joint probability of parameters of target variables given prior(evidential) parameters, and then marginalizing over the target variables parameters.
 - ***Why do we need Weibull distribution:*** We contextualized the use of Weibull distribution, showing insufficiency of normal distribution to model LGD. Furthermore, recent back collapse of SVB bank was shared as an example.
 - ***Why can we not keep both Weibull parameters unknown:*** It was noted that with both the parameters unknown, the log likelihood integral becomes intractable.
 - ***Why choose the NIG prior after reparametrizing Weibull:*** Following from the last discussion we present that this NIG prior is more amenable for analytical calculations, as the NIG lies in the conjugate family of Weibull.

Appendix: Log Likelihood Calculation

2.2. Learning Log-Likelihood. Hence we can define likelihood of y_i given the higher order evidential parameters α, β can be defined as :

$$\begin{aligned}
 (2.3) \quad Lik &= p(y_i|\alpha, \beta) = \int_{\theta} p(y_i|\theta, k) p(\theta|\alpha, \beta) d\theta \\
 &\quad \text{Now given } \lambda, k > 0 \implies \theta > 0 \\
 p(y_i|\alpha, \beta) &= \int_{\theta=0}^{\infty} \left(\frac{k}{\theta} y_i^{k-1} \exp(-y_i^k/\theta) \right) \left(\frac{\beta^\alpha}{\Gamma(\alpha)} \frac{1}{\theta^{\alpha+1}} \exp(-\frac{\beta}{\theta}) \right) d\theta \\
 &= k y_i^{k-1} \frac{\beta^\alpha}{\Gamma(\alpha)} \int_{\theta=0}^{\infty} \frac{1}{\theta^{\alpha+2}} \exp(-\frac{y_i^k + \beta}{\theta}) d\theta
 \end{aligned}$$

$$\begin{aligned}
 &\text{Now let} \\
 \Omega &= k y_i^{k-1} \frac{\beta^\alpha}{\Gamma(\alpha)} \quad \text{and,} \\
 a &= y_i^k + \beta \\
 \implies p(y_i|\alpha, \beta) &= \Omega \int_{\theta=0}^{\infty} \frac{1}{\theta^{\alpha+2}} \exp(-\frac{a}{\theta}) d\theta \\
 &\quad \text{Substituting } t = 1/\theta \\
 \implies p(y_i|\alpha, \beta) &= \Omega \int_{t=\infty}^0 t^{\alpha+2} \exp(-at) (-dt * t^{-2}) \\
 &= \Omega \int_{t=0}^{\infty} t^\alpha \exp(-at) dt
 \end{aligned}$$

$$\begin{aligned}
 &\text{Now we know from [10]} \\
 \int_0^\infty x^n \exp(-ax) dx &= \frac{\Gamma(1+n)}{a^{1+n}} \\
 \implies p(y_i|\alpha, \beta) &= \Omega \frac{\Gamma(1+\alpha)}{a^{1+\alpha}} \\
 &\quad \text{substituting back } \Omega \text{ and } a \\
 \implies Lik &= p(y_i|\alpha, \beta) = k y_i^{k-1} \frac{\beta^\alpha}{\Gamma(\alpha)} \frac{\Gamma(1+\alpha)}{(y_i^k + \beta)^{1+\alpha}} \\
 &= \frac{\alpha k y_i^{k-1} \beta^\alpha}{(y_i^k + \beta)^{\alpha+1}}
 \end{aligned}$$

Hence the log-likelihood for i'th observation is defined as:

$$(2.6) \quad Log - Lik_i = L_i^{lik} = \log \alpha_i + \log k + (k-1) \log y_i + \alpha_i \log \beta_i - (\alpha_i + 1)(y_i^k + \beta_i)$$

Appendix: Mean and variance calculation

2.3.1. Mean Prediction.

We define the mean prediction as

$$Z = E[y_i | \alpha, \beta]$$

Now given $y_i \sim Weibull(k, \lambda)$

$$\begin{aligned} 115 \quad (2.7) \quad Z &= E[\lambda * \Gamma(1 + \frac{1}{k})] = E(\lambda) * \Gamma(1 + \frac{1}{k}) \quad (k \text{ is known}) \\ E[\lambda] &= \int_{\lambda} \lambda p(\lambda) d\lambda \end{aligned}$$

116 Hence to solve for mean prediction we need to find pdf $p(\lambda)$. Because we know $\theta = \lambda^k \sim$
117 $\Gamma^{-1}(\alpha, \beta)$, we can use change of variable to find pdf of λ [9].

$$\begin{aligned} 118 \quad (2.8) \quad p(\lambda | \alpha, \beta) &= p_{\theta}(\lambda^k) * \left| \frac{d\lambda^k}{d\lambda} \right| \\ &= \frac{\beta^{\alpha}}{\Gamma(\alpha)} \left(\frac{1}{\lambda^k} \right)^{\alpha+1} \exp\left(-\frac{\beta}{\lambda^k}\right) * |k\lambda^{k-1}| \\ &= \frac{\beta^{\alpha}}{\Gamma(\alpha)} \left(\frac{1}{\lambda^k} \right)^{\alpha+1} \exp\left(-\frac{\beta}{\lambda^k}\right) * k\lambda^{k-1} \quad (\text{given } \lambda, k > 0) \end{aligned}$$

Hence,

2.9)

$$\begin{aligned} E[\lambda] &= \int_{\lambda} \lambda \frac{\beta^{\alpha}}{\Gamma(\alpha)} \left(\frac{1}{\lambda^k} \right)^{\alpha+1} \exp\left(-\frac{\beta}{\lambda^k}\right) * k\lambda^{k-1} d\lambda \\ &= \frac{k\beta^{\alpha}}{\Gamma(\alpha)} \int_{\lambda=0}^{\infty} \frac{1}{\lambda^{k\alpha+k-k}} \exp\left(\frac{\beta}{\lambda^2}\right) d\lambda \end{aligned}$$

Substituting $t = 1/\lambda$, we get:

$$dt = -1/\lambda^2 d\lambda,$$

$$E[\lambda] = \frac{k\beta^{\alpha}}{\Gamma(\alpha)} \int_{t=0}^{\infty} t^{k\alpha-2} \exp(-\beta t^k) dt$$

By table of integrals at [10]

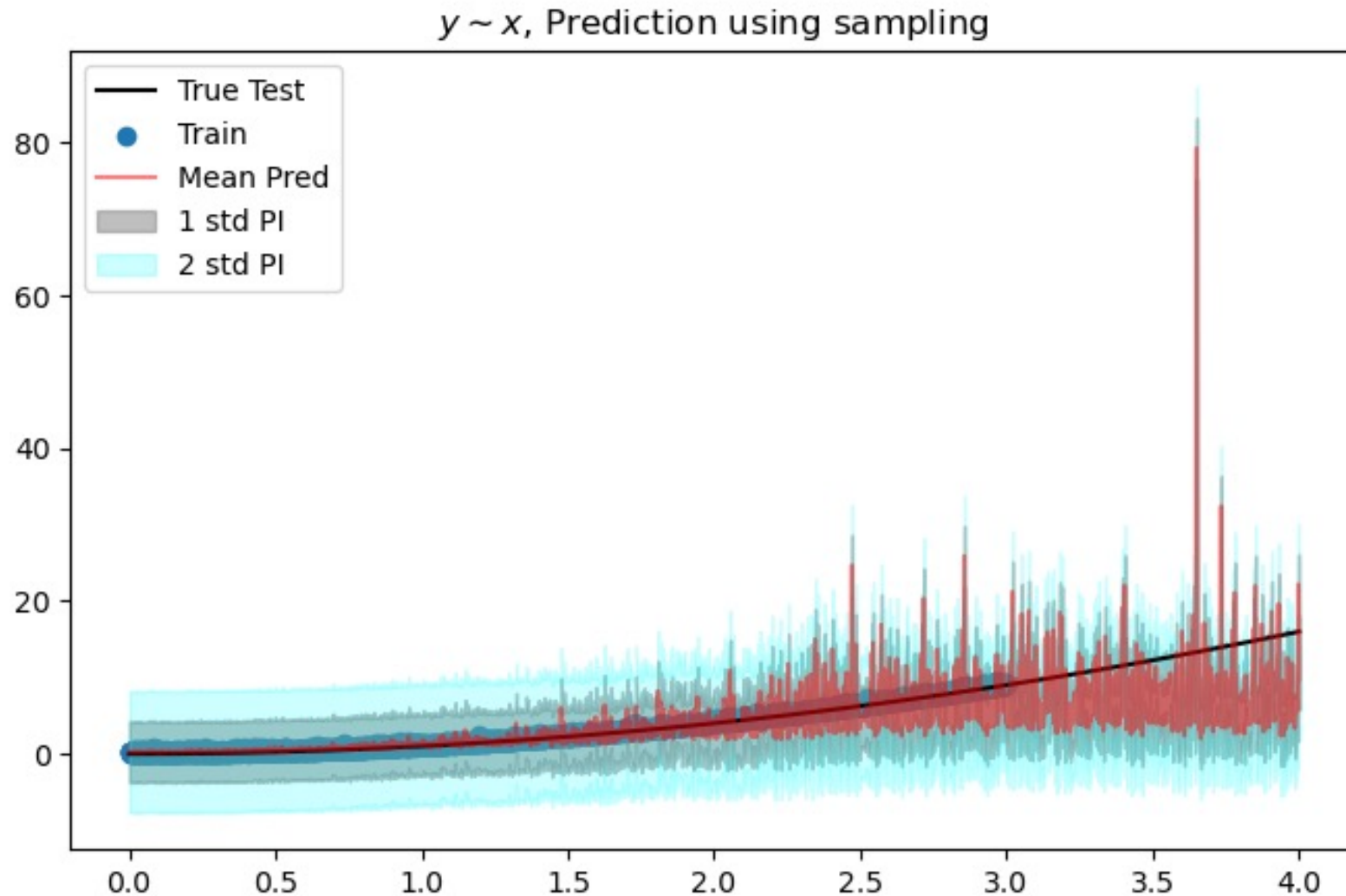
$$\begin{aligned} \int_0^{\infty} y^m e^{-by^k} dy &= \frac{\Gamma\left(\frac{m+1}{k}\right)}{kb^{(m+1)/k}} \\ \Rightarrow E[\lambda] &= \frac{k\beta^{\alpha}}{\Gamma(\alpha)} * \Gamma\left(\frac{k\alpha-1}{k}\right) * \frac{1}{k} * \frac{1}{\beta^{\frac{k\alpha-1}{k}}} \end{aligned}$$

Hence we get the mean prediction as:

2.10)

$$\begin{aligned} Z &= E[y_i | \alpha, \beta] = E(\lambda) * \Gamma(1 + \frac{1}{k}) \\ &= \frac{k\beta^{\alpha}}{\Gamma(\alpha)} * \Gamma\left(\frac{k\alpha-1}{k}\right) * \frac{1}{k} * \frac{1}{\beta^{\frac{k\alpha-1}{k}}} * \Gamma(1 + \frac{1}{k}) \\ &= \Gamma(1 + \frac{1}{k}) \frac{1}{\Gamma(\alpha)} \Gamma\left(\alpha - \frac{1}{k}\right) * \beta^{1/k} \end{aligned}$$

Appendix: Simulated data PI from sampling



Appendix Results: Real Data

- Mortgage data from Peer to Peer lending (2007-20011)

	Relative MSE
Benchmark	16.3%
Improved	9.7%

Appendix: Simulated data PI from sampling

```
Index(['Unnamed: 0', 'id', 'member_id', 'loan_amnt', 'funded_amnt',  
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'delinq_2yrs', 'earliest_cr_line', 'inq_last_6mths', 'mths_since_last_delinq',  
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'collection_recovery_fee', 'last_pymnt_d', 'last_pymnt_amnt', 'next_pymnt_d',  
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'open_acc_6m', 'open_il_6m', 'open_il_12m', 'open_il_24m', 'mths_since_rcnt_il',  
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