

# STAT604-Assignment-2

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## Assignment-2

### Part 1

#### Section 6.2 Problem#22

Solution:(a) Let , 'X' be random variable has pdf.

$$u_1 = E(x) = \int_0^1 x(\theta + 2)x^\theta dx$$

$$(\theta + 1) \int_0^1 x^{\theta+1} dx$$

$$\theta + 1 \left( \frac{x^{\theta+2}}{\theta + 2} \right)$$

$$(\theta + 1/(\theta + 2))$$

$$(\theta + 1 + 1 - 1/(\theta + 2))$$

$$1 - \frac{1}{\theta + 2}$$

$$1 - u_1 = \frac{1}{\theta + 2}$$

$$(\theta + 2) = \frac{1}{1 - u_1}$$

$$\theta = \frac{1}{(1 - u_1)^{-2}}$$

It is a moment estimate or of

$$\theta = \bar{X}$$

mean of the sample data

**Solution: (b)**

Sample

$$\bar{X} = \frac{x_1 + x_2 + x_3 + x_4 + x_5 + x_6 + x_7 + x_8 + x_9 + x_{10}}{10}$$

$$\bar{X} = \frac{0.92 + 0.79 + 0.9 + 0.65 + 0.86 + 0.47 + 0.73 + 0.97 + 0.94 + 0.77}{10}$$

$$\bar{X} = 0.8$$

The estimate of

$$\theta = \frac{1}{(1 - u_1)^{-2}} = 5 - 2 = 3$$

## Section 6.2 Problem#23

Given the probability density function (pdf) of the measurement error:  $f(x; \theta) =$

$$\frac{1}{\sqrt{2\pi\theta}} \cdot e^{-\frac{x^2}{2\theta}}$$

We want to find the maximum likelihood estimate (MLE) of

$$\theta$$

based on ( n ) independent measurements  $(x_1, x_2, x_3, \dots, x_n)$  The likely hood function

$$L(\theta)$$

is the product of the pdf for each measurement:

$$\ln(L(\theta)) = \sum_{i=1}^n \ln(f(x; \theta))$$

The log-likelihood function is given by:

$$\ln(L(\theta)) = \sum_{i=1}^N \ln \left( \frac{1}{\sqrt{2\pi\theta}} \cdot e^{-\frac{x_i^2}{2\theta}} \right)$$

Now, let's differentiate this with respect to  $\theta$  step by step:

$$\frac{d}{d\theta} \ln(L(\theta)) = \sum_{i=1}^N \frac{d}{d\theta} \ln \left( \frac{1}{\sqrt{2\pi\theta}} \cdot e^{-\frac{x_i^2}{2\theta}} \right)$$

Applying the chain rule and sum rule:

$$\frac{d}{d\theta} \ln(L(\theta)) = \sum_{i=1}^N \frac{d}{d\theta} \left( -\frac{1}{2} \ln(2\pi\theta) - \frac{x_i^2}{2\theta} \right)$$

Differentiating each term:

$$\frac{d}{d\theta} \ln(L(\theta)) = \sum_{i=1}^N \left( -\frac{1}{2\theta} + \frac{x_i^2}{2\theta^2} \right)$$

Combine terms:

$$\frac{d}{d\theta} \ln(L(\theta)) = \sum_{i=1}^N \frac{-1 + x_i^2}{2\theta^2}$$

Setting the derivative equal to zero to find the MLE:

$$\sum_{i=1}^N \frac{-1 + x_i^2}{2\theta^2} = 0$$

The equation obtained from setting the derivative of the log-likelihood equal to zero is:

$$\sum_{i=1}^N \frac{-1 + x_i^2}{2\theta^2} = 0$$

To solve for  $\theta$ , rearrange the equation:

$$\sum_{i=1}^N -1 + x_i^2 = 0$$

Now, isolate  $\theta$ :

$$\begin{aligned} \sum_{i=1}^N x_i^2 &= \sum_{i=1}^N 1 \\ \sum_{i=1}^N x_i^2 &= N \end{aligned}$$

Finally, solve for  $\theta$ :

$$\theta = \frac{\sum_{i=1}^N x_i^2}{N}$$

So, the maximum likelihood estimate (MLE) for  $\theta$  in this context is  $\frac{\sum_{i=1}^N x_i^2}{N}$ .

## Section 6.2 Problem#27

Solution(a): The probability density function (PDF) of the gamma distribution is given by:

$$f(x; a, b) = \frac{1}{\Gamma(a)b^a} x^{(a-1)} e^{-\frac{x}{b}}$$

where  $\Gamma(a)$  is the gamma function of  $a$ .

Assuming  $\Gamma(a)$  is the gamma function of  $a$ , the likelihood function for a sample of  $n$  independent observations  $X_1, X_2, \dots, X_n$  from a gamma distribution is given by:

$$\begin{aligned} L(a, b) &= f(X_1; a, b) \cdot f(X_2; a, b) \cdot \dots \cdot f(X_n; a, b) \\ &= \frac{1}{\Gamma(a)^n b^{a \cdot n}} \prod_{i=1}^n X_i^{(a-1)} e^{-\frac{X_i}{b}} \end{aligned}$$

The log-likelihood function is:

$$\ln L(a, b) = -n \ln \Gamma(a) - na \ln b + (a-1) \sum \ln(X_i) - \sum \frac{X_i}{b}$$

To find the maximum likelihood estimators of  $a$  and  $b$ , we take the partial derivatives of the log-likelihood function with respect to  $a$  and  $b$  and set them equal to zero:

$$\frac{\partial \ln L(a, b)}{\partial a} = -n\psi(a) + \sum \ln(X_i) = 0$$

where  $\psi(a)$  is the digamma function (derivative of the logarithm of the gamma function).

$$\frac{\partial \ln L(a, b)}{\partial b} = -\frac{na}{b} + \sum \frac{X_i}{b^2} = 0$$

The above equations cannot be solved explicitly. However, numerical methods can be used to find the values of  $a$  and  $b$  that maximize the log-likelihood function.

Solution(b): The probability density function (PDF) of the gamma distribution is given by:

$$f(x; a, b) = \frac{1}{\Gamma(a)b^a} x^{(a-1)} e^{-\frac{x}{b}}$$

where  $\Gamma(a)$  is the gamma function of  $a$ .

The mean of a gamma distribution is given by:

$$E(X) = a \cdot b$$

Therefore, we have:

$$a \cdot b = \mu$$

Since the Maximum Likelihood Estimators (MLE) of  $a$  and  $b$  are:

$$a = \left( \frac{n}{\psi(a)} \right)^{-1}$$

$$b = \frac{1}{n} \sum_{i=1}^n X_i$$

where  $\psi(a)$  is the gamma function, which is the first derivative of the natural logarithm of the gamma function. Substituting these values of  $a$  and  $b$  in the equation for  $\mu$ , we get:

$$\mu = a \cdot b = \left[ \left( \frac{n}{\psi(a)} \right)^{-1} \right] \left[ \frac{1}{n} \sum_{i=1}^n X_i \right] = \bar{X}$$

Here,  $\bar{X}$  represents the sample mean.

## Part 2

To find MLE for  $d$ :

### Part (i)

The probability function of  $x$  is given by:

$$P(x, \lambda) = \frac{e^{-\lambda} \cdot \lambda^x}{x!} \quad \text{for } x = 0, 1, 2, \dots$$

The likelihood function is:

$$L = e^{-nd} \cdot \frac{\lambda^{x_1} \cdot x_2 \cdot \dots \cdot x_n}{x_1! \cdot x_2! \cdot \dots \cdot x_n!}$$

So,

$$\ln L = -nd + \sum_i x_i \ln \lambda - \ln(x_1! \cdot x_2! \cdot \dots \cdot x_n!)$$

Taking the partial derivative with respect to  $d$  and equating it to 0, we get:

$$\frac{\partial}{\partial d} \ln L = -n + \frac{\sum_i x_i}{d} = 0$$

Differentiating (1) again partially with respect to  $d$ , we get:

$$\frac{\partial^2}{\partial d^2} \ln L = -\frac{\sum_i x_i}{d^2} < 0$$

This shows that equation (1) will indeed give the maximum value. Solving for  $d$ , we get:

$$d = \frac{\sum x_i}{n} = \bar{x}$$

So, we prove that the MLE for  $d$  is given by  $\bar{d} = \bar{x}$ .

### Part (ii)

To find the probability function of  $x$ , it is given by:

$$P(x, d) = \frac{e^{-\lambda} \cdot d^x}{x!}$$

The first moment about the origin is given by:

$$M_1^1 = E(x) = \sum_{x=0}^{\infty} x \cdot P(x) = \sum_{x=0}^{\infty} x \left( \frac{e^{-d} \cdot d^x}{x!} \right)$$

After simplifying, we get:

$$M_1^1 = \lambda$$

But  $\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n}$ , so we get:

$$\bar{\lambda} = (\lambda)$$

Thus, we prove that the moment estimator of  $\lambda$  is given by  $\bar{\lambda} = \bar{x}$ .

### Part (iii)

They are the same: the maximum likelihood estimate of  $d$  is given by  $\frac{\sum x_i}{n} = \bar{x}$ . The maximum likelihood of

$$\lambda$$

=

$$0 \times 4 + 1 \times 12 + 2 \times 22 + 3 \times 14 + 4 \times 9 = 112$$

Therefore,

$$\frac{112}{50} = 2.24$$