

CSE - S'75

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ASSIGNMENT #1

①

1) Baye's Classifier ↴

$$\text{Prior: } \text{Beta}(\beta_0, \beta_1) \Rightarrow \frac{\theta^{\beta_0-1} (1-\theta)^{\beta_1-1}}{B(\beta_0, \beta_1)}$$

$\alpha_0 \rightarrow \text{Heads}$

$\alpha_1 \rightarrow \text{Tails}$

$$\text{Likelihood: } \theta^{\alpha_0} (1-\theta)^{\alpha_1}$$

we know that $\xrightarrow{\text{Prior}} \xrightarrow{\text{Likelihood}}$

$$P(D|\Theta) = \frac{P(\Theta) \times P(D|\Theta)}{\int_{\Theta} P(\Theta) P(D|\Theta) d\Theta} \quad \left. \begin{array}{l} \xrightarrow{\text{Prior}} \\ \xrightarrow{\text{Likelihood}} \end{array} \right\} \Rightarrow \text{Marginal Likelihood} = P(D)$$

$$\text{Let's say } \int_{\Theta} P(\Theta) P(D|\Theta) d\Theta = C$$

$$\text{Posterior} \propto \frac{1}{C} \theta^{\alpha_0} (1-\theta)^{\alpha_1} \cdot \frac{\theta^{\beta_0-1} (1-\theta)^{\beta_1-1}}{B(\beta_0, \beta_1)} \\ \frac{1}{C \times B(\beta_0, \beta_1)} \times \theta^{\alpha_0 + \beta_0 - 1} (1-\theta)^{\alpha_1 + \beta_1 - 1}$$

$$\propto \frac{1}{C'} \times \theta^{\alpha_0 + \beta_0 - 1} (1-\theta)^{\alpha_1 + \beta_1 - 1}$$

$$P(\Theta|D) \sim \underline{\text{Beta}(\alpha_0 + \beta_0, \alpha_1 + \beta_1)}$$

$$\text{Mean} = \frac{\alpha_0 + \beta_0}{\alpha_0 + \beta_0 + \alpha_1 + \beta_1}$$

(2)

Posterior: $P(\theta) \propto P(D|\theta)$

$$P(\theta|D) = \frac{\theta^{\beta_{\eta}-1} (1-\theta)^{\beta_T-1}}{B(\beta_{\eta}, \beta_T)} \cdot \theta^{\alpha_{\eta}} (1-\theta)^{\alpha_T}$$

$$\begin{aligned} \ln P(\theta|D) &= \ln \left[\frac{\theta^{\beta_{\eta}-1} (1-\theta)^{\beta_T-1}}{B(\beta_{\eta}, \beta_T)} \right] + \ln \left[\theta^{\alpha_{\eta}} (1-\theta)^{\alpha_T} \right] \\ &= \ln \frac{1}{B(\beta_{\eta}, \beta_T)} + \ln \frac{\theta^{\beta_{\eta}-1}}{B(\beta_{\eta}, \beta_T)} + \ln (1-\theta)^{\beta_T-1} + \ln \theta^{\alpha_{\eta}} + \ln (1-\theta)^{\alpha_T} \\ &= \frac{\ln \theta}{B(\beta_{\eta}, \beta_T)} + (\beta_{\eta}-1) \ln \theta + (\beta_T-1) \ln (1-\theta) + \alpha_{\eta} \ln \theta \\ &\quad + \alpha_T \ln (1-\theta) \end{aligned}$$

$$\frac{d \ln P(\theta|D)}{d\theta} = 0 + \frac{\beta_{\eta}-1}{\theta} + \frac{\beta_T-1}{1-\theta} + \frac{\alpha_{\eta}}{\theta} + \frac{\alpha_T}{1-\theta} = 0$$

$$\frac{\beta_{\eta}-1}{\theta} - \frac{(\beta_T-1)}{1-\theta} + \frac{\alpha_{\eta}}{\theta} - \frac{\alpha_T}{1-\theta} = 0$$

$$\frac{\beta_{\eta}-1+\alpha_{\eta}}{\theta} = \frac{\beta_T+\alpha_T-1}{1-\theta}$$

$$\beta_{\eta}+\alpha_{\eta}-1 - \theta \beta_{\eta} + \theta = \theta \alpha_T + \theta \alpha_T - \theta$$

$$\alpha_{\eta}+\beta_{\eta}-1 - \theta (\alpha_{\eta}+\beta_{\eta}-1) = \theta (\alpha_T+\beta_T-1)$$

$$\hat{\theta}_{MAP} = \frac{\alpha_{\eta}+\beta_{\eta}-1}{(\alpha_{\eta}+\beta_{\eta}-1) + (\alpha_T+\beta_T-1)}$$

(3)

② Parameter Expectation →

$x_1, x_2, x_3, \dots, x_N \in R$ (i.i.d samples)

1) we know,

$$\hat{\mu}_{MLE} = \frac{1}{N} \sum_{i=1}^N x_i$$

$$E[\hat{\mu}_{MLE}] = E\left[\frac{1}{N} \sum_{i=1}^N x_i\right]$$

$$= \frac{1}{N} \sum_{i=1}^N E[x_i]$$

$$= \frac{1}{N} (\lambda \mu)$$

$= \mu \rightarrow$ It is a true value of μ and not a random variable hence unbiased.

2)

$$\hat{\sigma}_{MLE}^2 = \frac{1}{N} \sum_{i=1}^N (x_i - \hat{\mu}_{MLE})^2$$

$$E[\hat{\sigma}_{MLE}^2] = E\left[\frac{1}{N} \sum_{i=1}^N (x_i - \hat{\mu}_{MLE})^2\right]$$

$$= \frac{1}{N} E\left[\sum_{i=1}^N (x_i^2 - 2x_i \hat{\mu} + \hat{\mu}^2)\right]$$

$$= \frac{1}{N} E\left[\sum_{i=1}^N x_i^2 - 2\hat{\mu} \sum_{i=1}^N x_i + N\hat{\mu}^2\right]$$

$$= \frac{1}{N} E\left[\sum_{i=1}^N x_i^2 - 2\hat{\mu}(N\hat{\mu}) + N\hat{\mu}^2\right]$$

$$= \frac{1}{N} E\left[\sum_{i=1}^N x_i^2 - N\hat{\mu}^2\right]$$

$$= \frac{1}{N} \sum_{i=1}^N E[x_i^2] - \frac{1}{N} x_N E[\hat{\mu}^2]$$

$$\hat{\sigma}_{MLE}^2 = \frac{1}{N} \sum_{i=1}^N E[x_i^2] - E[\hat{y}^2] \quad (4)$$

Now, we know that,

$$E[x_i^2] = \sigma^2 + \mu^2 \Rightarrow \sigma^2 = E[x_i^2] - \mu^2 \quad (1)$$

$$E[\hat{y}^2] = \frac{\sigma^2}{n} + \mu^2 \quad (2)$$

$$\hat{\sigma}_{MLE}^2 = (\sigma^2 + \mu^2) - \left(\frac{\sigma^2}{n} + \mu^2 \right) \therefore \text{from } (1) \text{ & } (2)$$

$$= \sigma^2 - \frac{\sigma^2}{n} + \mu^2 - \mu^2$$

$$\boxed{\hat{\sigma}_{MLE}^2 = \frac{(n-1)}{n} \sigma^2}$$

Here we can say that $\hat{\sigma}_{MLE}^2 \neq \sigma^2$ hence it is unbiased.

About eq (2) // Additional Work to show eqn (2) //

$$\text{Var}(x_1 + x_2 + \dots + x_n) = n \cdot \sigma_x^2$$

$$\text{Var}\left(\frac{x_1 + x_2 + \dots + x_n}{n}\right) = \frac{n \cdot \sigma_x^2}{n^2} = \frac{\sigma_x^2}{n} = \sigma_y^2$$

$$\frac{\sigma_x^2}{n} = E[\bar{x}^2] - \underbrace{(E[\bar{x}])^2}_{\mu^2}$$

$$E[\bar{x}^2] = \sigma^2 + \mu^2 \quad (2)$$

③ Naive Bayes Classifier ↴

1) Total No. of independent Parameters →

$$\text{age} \Rightarrow 2 \times (3-1) = 4$$

$$\text{income} \Rightarrow 2 \times (3-1) = 4$$

$$\text{Student} \Rightarrow 2 \times (2-1) = 2$$

$$\text{Credit-ratio} \Rightarrow 2 \times (2-1) = 2$$

$$\text{buys-computer} \Rightarrow (1-1) = \frac{1}{13} \text{ parameters}$$

Here are they: (MLE Estimations)

$$\begin{aligned} P(\text{age} = \text{youth} | Y=\text{Yes}) &= \frac{2}{9} & P(\text{age} = \text{youth} | Y=\text{No}) &= 3/5 \\ P(\text{age} = \text{middle} | Y=\text{Yes}) &= 4/9 & P(\text{age} = \text{middle} | Y=\text{No}) &= 0 \\ P(\text{income} = \text{high} | Y=\text{Yes}) &= 2/9 & P(\text{income} = \text{high} | Y=\text{No}) &= 2/5 \\ P(\text{income} = \text{low} | Y=\text{Yes}) &= 3/9 & P(\text{income} = \text{low} | Y=\text{No}) &= 1/5 \\ P(\text{student} = \text{Yes} | Y=\text{Yes}) &= 6/9 & P(\text{student} = \text{Yes} | Y=\text{No}) &= 1/5 \\ P(\text{credit-ratio} = \text{fair} | Y=\text{Yes}) &= 6/9 & P(\text{credit-ratio} = \text{fair} | Y=\text{No}) &= 2/5 \\ P(\text{buys-computer}) &= 9/14 \end{aligned}$$

3) $X = (\text{youth}, \text{medium}, \text{Yes}, \text{fair})$

$$P(Y=\text{Yes} | X) = \frac{P(X | Y=\text{Yes}) P(Y=\text{Yes})}{\sum_i P(X | Y_i)}$$

$$= P(\text{age} = \text{youth} | Y=\text{Yes}) \times P(\text{income} = \text{medium} | Y=\text{Yes}) \\ \times P(\text{student} = \text{Yes} | Y=\text{Yes}) \times P(\text{cr} = \text{fair} | Y=\text{Yes}) \times P(Y=\text{Yes})$$

$$[P(\text{age} = \text{youth} | Y=\text{Yes}) \times P(\text{income} = \text{med} | Y=\text{Yes}) \times P(\text{student} = \text{Yes} | Y=\text{Yes}) \\ \times P(\text{cr} = \text{fair} | Y=\text{Yes}) \times P(Y=\text{Yes})] + [P(\text{age} = \text{youth} | Y=\text{No}) \times \\ P(\text{income} = \text{med} | Y=\text{No}) \times P(\text{student} = \text{Yes} | Y=\text{No}) \times P(\text{cr} = \text{fair} | Y=\text{No}) \times P(Y=\text{No})]$$

(6)

$$= \frac{9/14 * \frac{2}{9} * \frac{4}{9} * \frac{6}{9} * \frac{6}{9}}{\left[\frac{9}{14} * \frac{2}{9} * \frac{4}{9} * \frac{6}{9} * \frac{6}{9} \right] + \left[\frac{3}{5} * \frac{2}{5} * \frac{1}{5} * \frac{2}{5} * \frac{5}{14} \right]}$$

$$= \frac{0.0282}{0.0282 + 0.00685}$$

$$= \frac{0.0282}{0.03505} = 0.8045$$

Since $P(Y = \text{yes} | X) > 0.5$, we will classify it as
 $\underline{Y = \text{yes}}$

④ Logistic Regression ↓

+

 $(1, 0), (0, -1)$

-

 $(0, 1), (-1, 0)$

$w^{(0)} = \begin{pmatrix} 0 \\ 0 \\ 0 \end{pmatrix}$

$$\begin{aligned} P(y=1 | x^1, w^0) &= \frac{\exp(w_0 + \sum_i w_i x_i^1)}{1 + \exp(w_0 + \sum_i w_i x_i^1)} \\ &= \frac{\exp(0+0+0)}{1+\exp(0+0+0)} \Rightarrow \frac{1}{1+1} = 0.5 \end{aligned}$$

Similarly,

$P(y=1 | x^2, w^0) = 0.5$

$P(y=1 | x^3, w^0) = 0.5$

$P(y=1 | x^4, w^0) = 0.5$

$$\begin{aligned} w_0^{(1)} &= w_0^{(0)} + \eta \cdot \sum_j (y^j - \hat{P}(y=1 | x^j, w^{(0)})) \\ &= w_0^{(0)} + \eta \cdot [(1-0.5) + (1-0.5) + (0-0.5) + (0-0.5)] \\ &= w_0^{(0)} + \eta \cdot (0.5 + 0.5 - 0.5 - 0.5) \end{aligned}$$

$= 0$

$$\begin{aligned} w_1^{(1)} &= w_1^{(0)} + \eta \cdot \sum_j x_1^j (y^j - \hat{P}(y=1 | x^j, w^{(0)})) \\ &= 0 + \eta \cdot [1 \times (1-0.5) + 0 \times (1-0.5) + 0 \times (0-0.5) \\ &\quad - 1 \times (0-0.5)] \\ &= 0 + \eta \cdot [0.5 + 0 + 0 + 0.5] \\ &= \eta \end{aligned}$$

$$w_2^{(1)} = w_2^0 + \eta \cdot \sum_j x_j^{(1)} (y_j - \hat{P}(y=1 | x_j, w^{(0)})) \quad (8)$$

$$= 0 + \eta \cdot [0 \times (1-0.5) - 1(1-0.5) + 2(0-0.5) + 0(0-0.5)]$$

$$= \eta \cdot [0 + 0.5 - 0.5 + 0]$$

$$\Rightarrow -\eta$$

$$w^{(1)} = \begin{bmatrix} 0 \\ \eta \\ -\eta \end{bmatrix} \quad \left. \begin{array}{l} \text{Here, } w_0^1 = w_0^0 \\ w_1^1 > w_1^0 \\ w_2^1 < w_2^0 \end{array} \right\}$$

$$w_0^{(2)} = w_0^{(1)} + \eta \cdot \sum_j (y_j - \hat{P}(y=1 | x_j, w^{(1)}))$$

~~$\neq 0 \neq x_1 P(y=1 | x_1, w^{(1)})$~~

$$P(y=1 | x_1, w^{(1)}) = \frac{\exp(0 + 0x(\eta) + 0x(-\eta))}{1 + \exp(0 + 0x(\eta) + 0x(-\eta))} = \frac{\exp(\eta)}{1 + \exp(\eta)}$$

$$P(y=0 | x_2, w^{(1)}) = \frac{\exp(0 + 0x(\eta) - 1(-\eta))}{1 + \exp(0 + 0x(\eta) - 1(-\eta))} = \frac{\exp(\eta)}{1 + \exp(\eta)}$$

$$P(y=1 | x_3, w^{(1)}) = \frac{\exp(0 + 0x(\eta) + 1(-\eta))}{1 + \exp(0 + 0x(\eta) + 1(-\eta))} = \frac{\exp(-\eta)}{1 + \exp(-\eta)}$$

$$P(y=1 | x_4, w^{(1)}) = \frac{\exp(0 + (-1)x(\eta) + 0(-\eta))}{1 + \exp(0 + (-1)x(\eta) + 0(-\eta))} = \frac{\exp(-\eta)}{1 + \exp(-\eta)}$$

$$\begin{aligned}
 w_0^{(2)} &= w_0^{(1)} + \gamma \sum_j (y^j - \hat{p}(y=1 | x^j, w^{(1)})) \\
 &= 0 + \gamma \cdot \left[\left(1 - \frac{\exp(\eta)}{1+\exp(\eta)} \right) + \left(1 - \frac{\exp(\eta)}{1+\exp(\eta)} \right) + \left(0 - \frac{\exp(-\eta)}{1+\exp(-\eta)} \right) \right. \\
 &\quad \left. + \dots \left(0 - \frac{\exp(-\eta)}{1+\exp(-\eta)} \right) \right]
 \end{aligned}$$

$$= \gamma \left(\frac{2}{1+\exp(\eta)} - \frac{2}{1+\exp(-\eta)} \right)$$

$$= \underline{0}$$

$$\begin{aligned}
 w_1^{(2)} &= w_1^{(1)} + \gamma \sum_j x^j_1 [y^j - \hat{p}(y=1 | x^j, w^{(1)})] \\
 &= \eta + \eta \left[1x \left(1 - \frac{\exp(\eta)}{1+\exp(\eta)} \right) + 0 + 0 + (-1) \times \left(0 - \frac{\exp(-\eta)}{1+\exp(-\eta)} \right) \right] \\
 &= \eta + \eta \left[\frac{1}{1+\exp(\eta)} + \frac{\exp(-\eta)}{1+\exp(-\eta)} \right] \\
 &= \eta + \eta \left[\frac{1}{1+\exp(\eta)} + \frac{1}{\exp(\eta)} + 1 \right] \\
 &= \eta + \eta \left[\frac{2}{1+\exp(\eta)} \right] \\
 &= \eta + \frac{2\eta}{1+\exp(\eta)} \\
 &= \frac{3\eta + \eta \exp(\eta)}{1+\exp(\eta)}
 \end{aligned}$$

(10)

$$w_2^{(2)} = w_2^{(1)} + \eta \sum_j x_2^j [y^j - \hat{p}(y=1|x^j, w^{(1)})]$$

$$= -\eta + \eta \left[0 + (-1) \times \left(1 - \frac{\exp(\eta)}{1+\exp(\eta)} \right) + 1 \left(0 - \frac{\exp(-\eta)}{1+\exp(-\eta)} \right) + 0 \right]$$

$$= -\eta + \eta \left[\frac{-1}{1+\exp(\eta)} + \frac{\exp(-\eta)}{1+\exp(-\eta)} \right]$$

$$= -\eta + \eta \left[\frac{-1}{1+\exp(\eta)} - \frac{1}{1+\exp(+\eta)} \right]$$

$$= -\eta + \eta \left[\frac{-2}{1+\exp(\eta)} \right]$$

$$= -\eta - \frac{2\eta}{1+\exp(\eta)}$$

$$\Rightarrow -\eta + \frac{\eta \exp(\eta) - 2\eta}{1+\exp(\eta)}$$

$$\Rightarrow \frac{-3\eta + \eta \exp(\eta)}{1+\exp(\eta)}$$

$$w^{(2)} = \begin{bmatrix} 0 \\ \eta \left[\frac{3+\exp(\eta)}{1+\exp(\eta)} \right] \\ -\eta \left[\frac{3+\exp(\eta)}{1+\exp(\eta)} \right] \end{bmatrix}$$

Here, we can observe that,
 $w_0^0 = w_0^1 = w_0^2$

$$w_1^2 > w_1^1 > w_1^0$$

$$w_2^2 < w_2^1 < w_2^0$$

$$P(Y=1 | x^1, w^{(2)}) = \frac{\exp\left(\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}{1 + \exp\left(\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}$$

$$P(Y=1 | x^2, w^{(2)}) = \frac{\exp\left(\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}{1 + \exp\left(\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}$$

$$P(Y=1 | x^3, w^{(2)}) = \frac{\exp\left(-\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}{1 + \exp\left(-\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}$$

$$P(Y=1 | x^4, w^{(2)}) = \frac{\exp\left(-\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}{1 + \exp\left(-\eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]\right)}$$

$$w_0^{(3)} = w_0^{(2)} + \eta \sum_j (y^j - \hat{P}(Y=1 | x^j, w^{(2)}))$$

$$\begin{aligned} &= 0 + \eta \left[\left(\frac{1 - \exp\left(\frac{3 + \exp(\eta)}{1 + \exp(\eta)}\right)}{1 + \exp(\eta)} \right) + \left(\frac{1 - \exp\left(\frac{3 + \exp(\eta)}{1 + \exp(\eta)}\right)}{1 + \exp(\eta)} \right) \right. \\ &\quad \left. + \left(0 + \frac{\exp\left(\frac{3 + \exp(\eta)}{1 + \exp(\eta)}\right)}{1 + \exp(-\eta)} \right) + \left(0 + \frac{\exp\left(\frac{3 + \exp(-\eta)}{1 + \exp(-\eta)}\right)}{1 + \exp(-\eta)} \right) \right] \\ &= \underline{0} \end{aligned}$$

(12)

let's assume

$$c = \eta \left[\frac{3 + \exp(\eta)}{1 + \exp(\eta)} \right]$$

$$\begin{aligned}
 w_{(2)}^{(3)} &= w_2^{(2)} + \eta \sum_j x_2^j [y^j - \hat{P}(y=1 | x^j, w^{(2)})] \\
 &= c + \eta \left[1 \times \left(1 - \frac{\exp(c)}{1 + \exp(c)} \right) + 0 + 0 - 1 \times \left(0 - \frac{1 - \exp(-c)}{1 + \exp(-c)} \right) \right] \\
 &= c + \eta \left[\frac{1}{1 + \exp(c)} + \frac{\exp(-c)}{1 + \exp(-c)} \right] \\
 &= c + \eta \left[\frac{1}{1 + \exp(c)} + \frac{1}{1 + \exp(c)} \right] \\
 \Rightarrow & c + \eta \left[\frac{2}{1 + \exp(c)} \right] \\
 &= c + \frac{2\eta}{1 + \exp(c)}
 \end{aligned}$$

Replacing value of c

$$w_2^{(3)} = \eta \frac{(3 + \exp(\eta))}{1 + \exp(\eta)} + \frac{2\eta}{1 + \exp\left(\frac{\eta(3 + \exp(\eta))}{1 + \exp(\eta)}\right)}$$

$$\begin{aligned}
 w_{(2)}^{(3)} &= w_2^{(2)} + \eta \sum_j x_2^j [y^j - \hat{P}(y=1 | x^j, w^{(2)})] \\
 &= -c + \eta \left[0 + (-1) \times \left(1 - \frac{\exp(c)}{1 + \exp(c)} \right) + 1 \times \left(0 - \frac{1 - \exp(-c)}{1 + \exp(-c)} \right) + 0 \right]
 \end{aligned}$$

(B)

$$\begin{aligned}
 &= -c + \eta \left[\frac{-1}{1+\exp(c)} - \frac{\exp(-c)}{1+\exp(-c)} \right] \\
 &= -c + \eta \left[\frac{-1}{1+\exp(c)} - \frac{1}{1+\exp(-c)} \right] \\
 &= -c + \eta \left[\frac{-2}{1+\exp(c)} \right] \\
 &= -c - \frac{2\eta}{1+\exp(c)}
 \end{aligned}$$

Replacing value of c

$$\omega_2^{(3)} = -\eta \frac{(3+\exp(\eta))}{1+\exp(\eta)} - \frac{2\eta}{1+\exp(\frac{3+\exp(\eta)}{1+\exp(\eta)})}$$

$$\begin{aligned}
 \omega^{(3)} = & \begin{bmatrix} 0 \\ \frac{\eta(3+\exp(\eta))}{1+\exp(\eta)} + \frac{2\eta}{1+\exp(\frac{\eta(3+\exp(\eta))}{1+\exp(\eta)})} \\ -1 \left[\frac{\eta(3+\exp(\eta))}{1+\exp(\eta)} + \frac{2\eta}{1+\exp(\frac{\eta(3+\exp(\eta))}{1+\exp(\eta)})} \right] \end{bmatrix}
 \end{aligned}$$

Here, after 3 iterations, we can see that,

 ω_0 is zero for all iterations. ω_1 is increasing with each iteration and going to $+\infty$. ω_2 is decreasing with each iteration and going to $-\infty$.

$$\omega^* = \begin{bmatrix} 0 \\ +\infty \\ -\infty \end{bmatrix}$$

(14)

Now taking

$$\omega^0 = \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}$$

$$P(y=1 | x^1, \omega^0) = \frac{\exp(0+1+0)}{1+\exp(0+1+0)} = \frac{\exp(1)}{1+\exp(1)} = 0.73$$

$$P(y=1 | x^2, \omega^0) = \frac{\exp(0+0+1)}{1+\exp(0+1)} = \frac{1}{2} = 0.5$$

$$P(x=1 | x^3, \omega^0) = \frac{\exp(0+0+0)}{1+\exp(0)} = 0.5$$

$$P(y=1 | x^4, \omega^0) = \frac{\exp(0-1+0)}{1+\exp(-1)} = \frac{\exp(-1)}{1+\exp(-1)} = 0.27$$

$$\begin{aligned} \omega_0^{(1)} &= \omega_0^{(0)} + \eta \cdot \sum_j (y^j - \hat{P}(y=1 | x^j, \omega^{(0)})) \\ &= 0 + \eta \cdot [(1-0.73) + (1-0.5) + (0-0.5) + (0-0.27)] \\ &= \eta [0.27 + 0.5 - 0.5 - 0.27] \\ &= 0 \end{aligned}$$

$$\begin{aligned} \omega_1^{(1)} &= \omega_1^{(0)} + \eta \cdot \sum_j x_1^j (y^j - \hat{P}(y=1 | x^j, \omega^{(0)})) \\ &= 1 + \eta [1 \times (1-0.73) + 0 + 0 - 1(0-0.27)] \\ &= 1 + \eta [0.27 + 0.27] \end{aligned}$$

$$\begin{aligned} \omega_2^{(1)} &= \omega_2^{(0)} + \eta \sum_j x_2^j (y^j - \hat{P}(y=1 | x^j, \omega^{(0)})) \\ &= 0 + \eta [0 - 1(1-0.5) + 1(0-0.5) + 0] \\ &= -\eta \end{aligned}$$

$$w^1 = \begin{bmatrix} 0 \\ 1+0.54\eta \\ -\eta \end{bmatrix}$$

so, here we can observe that
 w_0^1 is same as w_0^0
 $w_1^1 > w_1^0$ which is increased by 0.54η
 $w_2^1 < w_2^0$

Now,

$$P(y=1 | x^1, w^{(1)}) = \frac{\exp(0 + 1 \times (1+0.54\eta) + 0)}{1 + \exp(1+0.54\eta)} = \frac{\exp(1+0.54\eta)}{1 + \exp(1+0.54\eta)}$$

$$P(y=1 | x^2, w^{(1)}) = \frac{\exp(0 + 0 + \eta)}{1 + \exp(\eta)} = \frac{\exp(\eta)}{1 + \exp(\eta)}$$

$$P(y=1 | x^3, w^{(1)}) = \frac{\exp(0 + 0 + \eta)}{1 + \exp(-\eta)} = \frac{\exp(-\eta)}{1 + \exp(-\eta)} *$$

$$P(y=1 | x^4, w^{(1)}) = \frac{\exp(0 + (-1)(1+0.54\eta) + 0)}{1 + \exp(-1-0.54\eta)} = \frac{\exp(-1-0.54\eta)}{1 + \exp(-1-0.54\eta)}$$

$$= \frac{1}{1 + \exp(1+0.54\eta)}$$

$$w_0^{(2)} = w_0^{(1)} + \eta \cdot \sum_j (y^j - \hat{P}(y=1 | x^j, w^{(1)}))$$

$$= 0 + \eta \left[\left(1 - \frac{\exp(1+0.54\eta)}{1 + \exp(1+0.54\eta)} \right) + \left(1 - \frac{\exp(\eta)}{1 + \exp(\eta)} \right) + \left(0 - \frac{\exp(-\eta)}{1 + \exp(-\eta)} \right) \right. \\ \left. + \left(0 - \frac{\exp(-1-0.54\eta)}{1 + \exp(-1-0.54\eta)} \right) \right]$$

$$= \eta \left[\frac{1}{1 + \exp(1+0.54\eta)} + \frac{1}{1 + \exp(\eta)} - \frac{1}{1 + \exp(\eta)} - \frac{1}{1 + \exp(1+0.54\eta)} \right]$$

$$= 0$$

(16)

$$w_1^{(2)} = w_1^{(1)} + \eta \sum_j x_1^j [y^j - \hat{P}(y=1|x_1^j; w^{(1)})]$$

$$= 1+0.54\eta + \eta \left[\left(1 - \frac{\exp(1+0.54\eta)}{1+\exp(1+0.54\eta)} \right) - \left(0 - \frac{1}{1+\exp(1+0.54\eta)} \right) \right]$$

$$= 1+0.54\eta + \eta \left[\frac{1}{1+\exp(1+0.54\eta)} + \frac{1}{1+\exp(1+0.54\eta)} \right]$$

$$\Rightarrow 1+0.54\eta + \frac{2\eta}{1+\exp(1+0.54\eta)}$$

$$w_2^{(2)} = w_2^{(1)} + \eta \sum_j x_2^j [y^j - \hat{P}(y=1|x_2^j; w^{(1)})]$$

$$= -\eta + \eta \left[(-1) \times \left(1 - \frac{\exp(\eta)}{1+\exp(\eta)} \right) + 1 \times \left(0 - \frac{1}{1+\exp(\eta)} \right) \right]$$

$$= -\eta + \eta \left[-\frac{1}{1+\exp(\eta)} - \frac{1}{1+\exp(\eta)} \right]$$

$$= -\eta - \frac{2\eta}{1+\exp(\eta)}$$

$$w_2^2 = \left[1+0.54\eta + \frac{2\eta}{1+\exp(1+0.54\eta)} \right] - \left[-\eta - \frac{2\eta}{1+\exp(\eta)} \right]$$

Here, we can see that
 $w_0^2 = w_0^1 = w_0^0$

$w_1^2 > w_1^1$ since $w_1^2 = w_1^1 +$ (constant)

$w_2^2 < w_2^1$ since we are subtracting
 constant from
 $\underline{w_2^1}$

(M)

Let's say.

$$C_1 = 1 + 0.54\eta + \frac{2\eta}{1 + \exp(1 + 0.54\eta)}$$

$$C_2 = -\eta - \frac{2\eta}{1 + \exp(\eta)}$$

$$P(Y=1 | X^1, w^{(2)}) = \frac{\exp(C_2)}{1 + \exp(C_1)}$$

$$P(Y=1 | X^2, w^{(2)}) = \frac{\exp(-C_2)}{1 + \exp(-C_2)} = \frac{1}{1 + \exp(C_2)}$$

$$P(Y=1 | X^3, w^{(2)}) = \frac{\exp(C_2)}{1 + \exp(C_2)}$$

$$P(Y=1 | X^4, w^{(2)}) = \frac{\exp(-C_1)}{1 + \exp(-C_1)} = \frac{1}{1 + \exp(C_1)}$$

$$w_0^{(3)} = w_0^{(2)} + \eta \sum_j (y_j - \hat{P}(Y=1 | X^j, w^{(1)}))$$

$$= 0 + \eta \left[\left(1 - \frac{\exp(C_1)}{1 + \exp(C_1)} \right) + \left(1 - \frac{1}{1 + \exp(C_2)} \right) + \left(0 - \frac{\exp(C_2)}{1 + \exp(C_2)} \right) + \left(0 - \frac{1}{1 + \exp(C_1)} \right) \right]$$

$$= 0 + \eta \left[\frac{1}{1 + \exp(C_1)} + \frac{\exp(C_2)}{1 + \exp(C_2)} - \frac{\exp(C_2)}{1 + \exp(C_2)} - \frac{1}{1 + \exp(C_1)} \right]$$

$$\underline{\underline{=0}}$$

$$\begin{aligned}
 w_1^{(3)} &= w_1^{(2)} + \eta \sum_j x_i^j [y^j - \hat{P}(y=1 | x_i^j, w^{(2)})] \\
 &= c_1 + \eta \left[1 \times \left(1 - \frac{\exp(c_1)}{1 + \exp(c_1)} \right) - 1 \times \left(1 - \frac{\exp(-c_1)}{1 + \exp(-c_1)} \right) \right] \\
 &= c_1 + \eta \left[\frac{1}{1 + \exp(c_1)} + \frac{1}{1 + \exp(-c_1)} \right] \\
 &= c_1 + \frac{2\eta}{1 + \exp(c_1)}
 \end{aligned}$$

$$\begin{aligned}
 w_2^{(3)} &= w_2^{(2)} + \eta \sum_j x_i^j [y^j - \hat{P}(y=1 | x_i^j, w^{(2)})] \\
 &= c_2 + \eta \left[(-1) \times \left(1 - \frac{\exp(-c_2)}{1 + \exp(-c_2)} \right) + \left(0 - \frac{\exp(c_2)}{1 + \exp(c_2)} \right) \right] \\
 &= c_2 + \eta \left[-\frac{\exp(c_2)}{1 + \exp(c_2)} - \frac{\exp(-c_2)}{1 + \exp(-c_2)} \right] \\
 &= c_2 - \left[\frac{2\exp(c_2)}{1 + \exp(c_2)} \right] n \\
 &= c_2 - \frac{2\eta \exp(c_2)}{1 + \exp(c_2)}
 \end{aligned}$$

$$w^3 = \begin{bmatrix} 0 \\ c_1 + \frac{2\eta}{1 + \exp(c_1)} \\ c_2 - \frac{2\eta \exp(c_2)}{1 + \exp(c_2)} \end{bmatrix}$$

Here we can see that,

$$w_0^3 = w_0^2 = w_0'$$

$w_1^3 > w_1^2 > w_1'$, since at each iteration a positive term is being added to previous weights and going to $+\infty$,

$w_2^3 < w_2^2 < w_2'$ since at each iteration a positive term is being subtracted from previous weights, and going to $-\infty$,

$$w^* = \begin{bmatrix} 0 \\ +\infty \\ -\infty \end{bmatrix}$$

So for two different w^0 , we can see that w^* is same i.e. after enough/large number of iterations both w_2 & w_3 will converge to $+\infty$ & $-\infty$ respectively.

(5) ii)

$$f: x \rightarrow y$$

$$x = \langle x_1, \dots, x_d \rangle \in \mathbb{R}^d$$

$$Y = \{0, 1\}$$

after Training

$$w = \langle w_0, w_1, \dots, w_d \rangle$$

we are given that,

$$P(x_i | Y=k) = N(\mu_{ik}, \sigma_{ik}^2) \quad (k=0, 1, i=1 \dots d)$$

No. of independent parameters will be :

$$\left. \begin{array}{l} \sigma_i^2 \text{ & } \mu_i^0 \text{ for } Y=0 \\ \sigma_i^2 \text{ & } \mu_i^1 \text{ for } Y=1 \end{array} \right\} 4 \text{ parameter for single feature}$$

for d features, it will be $4d$ & 1 parameter for $P(Y=1)$ as $P(Y=0)$ can be calculated using $1 - P(Y=1)$
So, total,

$$\boxed{4d+1}$$

ii) Translation of logistic weights into equivalent GNB parameters :-

we have,

$$P(Y=1|x) = \frac{P(Y=1) P(x|Y=1)}{P(Y=1) P(x|Y=1) + P(Y=0) P(x|Y=0)}$$

Now, dividing numerator & denominator by $P(Y=1) P(X|Y=1)$

$$= \frac{1}{1 + \frac{P(Y=0) P(X|Y=0)}{P(Y=1) P(X|Y=1)}}$$

$$= \frac{1}{1 + \exp \left[\ln \left(\frac{P(Y=0)}{P(Y=1)} \times \frac{P(X|Y=0)}{P(X|Y=1)} \right) \right]}$$

$$= \frac{1}{1 + \exp \left[\left(\ln \frac{1-\delta}{\delta} \right) + \sum_i \ln \frac{P(x_i|Y=0)}{P(x_i|Y=1)} \right]} \quad \text{--- (1)}$$

Now solving the 2nd part

we, know

$$P(x_i|Y=k) = \frac{1}{\sigma_i \sqrt{2\pi}} \exp^{-\frac{(x_i - \mu_{ik})^2}{2\sigma_i^2}}$$

$$= \sum_i \ln \left[\frac{\exp \left(-\frac{(x_i - \mu_{i0})^2}{2\sigma_i^2} \right)}{\exp \left(-\frac{(x_i - \mu_{ii})^2}{2\sigma_i^2} \right)} \right]$$

$$= \sum_i \ln \left[\exp \left(-\frac{(x_i^2 + \mu_{i0}^2 - 2x_i \mu_{i0}) + x_i^2 + \mu_{ii}^2 - 2x_i \mu_{ii}}{2\sigma_i^2} \right) \right]$$

$$= \sum_i \frac{(\mu_{ii}^2 - \mu_{i0}^2) + 2x_i (\mu_{i0} - \mu_{ii})}{2\sigma_i^2}$$

$$= \sum_i \frac{(\mu_{i0} - \mu_{i1})x_i}{\sigma_p^2} + \left(\frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_p^2} \right)$$

Substituting the above value in Eq①

$$P(Y=1|X) = \frac{1}{1 + \exp \left[\ln \left(\frac{1-\theta}{\theta} \right) + \sum_i w_0 + \underbrace{\sum_i \frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_p^2}}_{w_i} + \underbrace{\sum_i \frac{(\mu_{i0} - \mu_{i1})x_i}{\sigma_p^2}}_{w_i} \right]}$$

So from above

$$w_0 = \ln \left(\frac{1-\theta}{\theta} \right) + \sum_i \left(\frac{\mu_{i1}^2 - \mu_{i0}^2}{2\sigma_p^2} \right)$$

$$w_i = \left(\frac{\mu_{i0} - \mu_{i1}}{\sigma_p^2} \right)$$

⑤ ii) Implementation of Gaussian Naive Bayes and logistic Regression ↴

Gaussian Naïve Bayes - Pseudo Code →

Step-1: Data processing

- Pre-process the data & divide it in to two sets i.e., training set and Test set based on the given fractions.
- Separate it by labels i.e. 0 & 1 in our case.

Step-2: Build the Model

- To build the model, first we need the Baye's theorem:

$$\text{posterior} = \text{prior} * \text{likelihood} / \text{marginal evidence}$$
- Since Naïve Bayes assumes that all the features are independent, it is equivalent to calculate the posterior probability ↴

$$P(Y=1 | X_1, \dots, X_n) = \frac{P(Y=1) * P(X_1, \dots, X_n | Y=1)}{P(X_1, \dots, X_n)}$$

Since x_1, x_2, \dots, x_n are independent

$$= \frac{P(Y=1) * P(X_1 | Y=1) * \dots * P(X_n | Y=1)}{P(X_1, \dots, X_n)}$$

Since $P(x_1, \dots, x_n)$ is constant given the input

$$P(Y \neq 1 | X_1, \dots, X_n) \propto P(Y=1) * P(x_1 | Y=1) * \dots * P(x_n | Y=1)$$

c) To calculate the above eqnⁿ, we used
prior = mean(y)

for likelihood, we used Gaussian Models:

$$P(x_i|y) = \frac{1}{\sqrt{2\pi} \sigma_y^2} \exp\left(-\frac{(x_i - \bar{y})^2}{2\sigma_y^2}\right)$$

where,

\bar{y} → mean of examples in $y=1$

σ_y^2 → standard deviation of examples for $y=1$

x_i → new test example to predict

Step 3: Predict

- a) We use the test set data and calculate the prior & likelihood for each row and get the posterior for that row by multiplying the prior & likelihood.
- b) To get accuracy, we check our predicted label with the actual label.
- c) we repeat the above steps 5 times for each fraction and calculate the average for accuracy of each fraction.

Logistic Regression Pseudo code →

Step 1: Process the data

this step is exactly similar to the Naive Bayes classifier's step and I have used same function for both of them

Step 2 & Build Model

- Initialize the coefficients for each attribute/feature randomly.
- we will use Sigmoid funcn as our logistic funcn for our logistic regression model i.e,
- Estimate/update the weights / coefficients using stochastic Gradient Descent.
 - Calculate the prediction using current coefficients
 - Update the coefficients based on error in prediction i.e,

$$\frac{1}{1+e^{-x}}$$

open/coeff \hat{y}

$$\hat{y} = \text{sigmoid}(\alpha - \text{prediction}) * \text{prediction}$$

Learning rate

$$w_i^{(t+1)} \leftarrow w_i^{(t)} + \eta \frac{\partial J(w)}{\partial w_i}$$

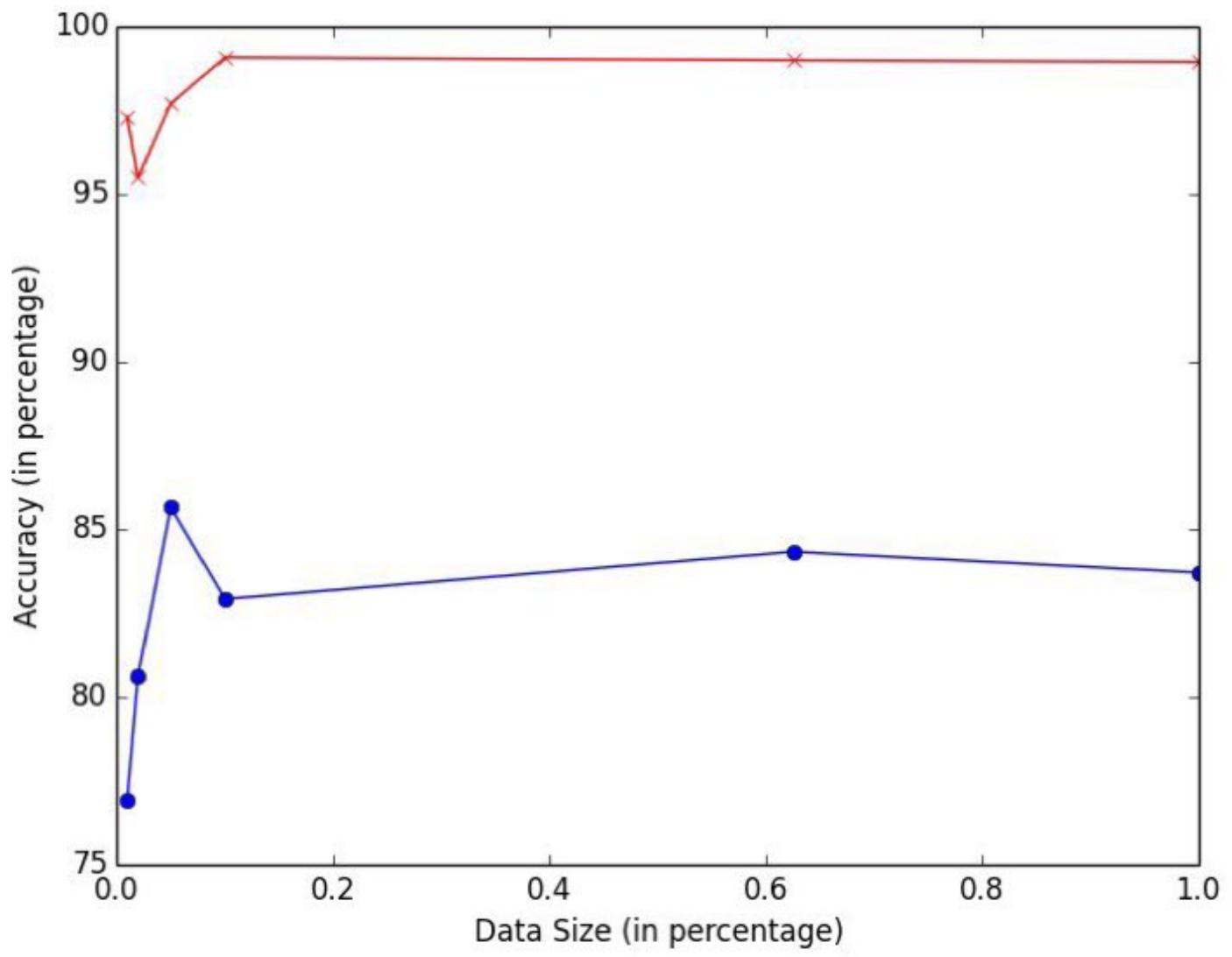
- Repeat the process

Learning Rate, $\eta > 0$

Step 3 is predict

- a) we will use the learned weights to predict the label for test input.
- b) Accuracy will be calculated by comparing the predicted and actual label for some particular row.
- c) we repeat the above steps 5 times for each fraction and calculate the average accuracy.

Plot the Learning Curve :-



Combination: 1

Attribute: 1
Original Training Set's Mean: -1.96649344171
Generated Samples' Mean: -1.86364300074
Difference: 0.102850440973

Original Training Set's SD: 1.85284431207
Generated Samples' SD: 1.86267559166
Difference: 0.00983127959427

[Attribute: 2
Original Training Set's Mean: -0.567499
Generated Samples' Mean: -0.618954999089
Difference: 0.0514559990894

Original Training Set's SD: 5.28183832332
Generated Samples' SD: 5.28684343678
Difference: 0.00500511345608

Attribute: 3
Original Training Set's Mean: 1.84226256784
Generated Samples' Mean: 2.03182337991
Difference: 0.189560812072

Original Training Set's SD: 5.13805582753
Generated Samples' SD: 4.9860273029
Difference: 0.152028524635

Attribute: 4
Original Training Set's Mean: -1.40414835477
Generated Samples' Mean: -1.44380040879
Difference: 0.0396520540194

Original Training Set's SD: 2.06651039315
Generated Samples' SD: 2.01938194917
Difference: 0.0471284439827

```
Combination: 2

Attribute: 1
Original Training Set's Mean: -1.78282045885
Generated Samples' Mean: -1.80718346018
Difference: 0.0243630013244
-----
Original Training Set's SD: 1.93454675614
Generated Samples' SD: 1.93281055487
Difference: 0.00173620126134
-----
Attribute: 2
Original Training Set's Mean: -0.793434406699
Generated Samples' Mean: -0.742865385263
Difference: 0.0505690214351
-----
Original Training Set's SD: 5.35724008266
Generated Samples' SD: 5.68345896539
Difference: 0.326218882732
-----
Attribute: 3
Original Training Set's Mean: 1.88618747368
Generated Samples' Mean: 1.45692106266
Difference: 0.429266411028
-----
Original Training Set's SD: 5.32950427606
Generated Samples' SD: 5.39534613492
Difference: 0.0658418588561
-----
Attribute: 4
Original Training Set's Mean: -1.33667266507
Generated Samples' Mean: -1.3975405891
Difference: 0.0608679240267
-----
Original Training Set's SD: 2.11839826059
Generated Samples' SD: 2.14878472877
Difference: 0.0303864681794
```

Combination: 3

Attribute: 1
Original Training Set's Mean: -1.78282045885
Generated Samples' Mean: -1.73253453385
Difference: 0.0502859250037

Original Training Set's SD: 1.93454675614
Generated Samples' SD: 2.02060426588
Difference: 0.0860575097443

Attribute: 2
Original Training Set's Mean: -0.793434406699
Generated Samples' Mean: -0.767437365155
Difference: 0.0259970415434

Original Training Set's SD: 5.35724008266
Generated Samples' SD: 5.10412861743
Difference: 0.253111465228

Attribute: 3
Original Training Set's Mean: 1.88618747368
Generated Samples' Mean: 1.61737288099
Difference: 0.26881459269

Original Training Set's SD: 5.32950427606
Generated Samples' SD: 5.69795521831
Difference: 0.368450942248

Attribute: 4
Original Training Set's Mean: -1.33667266507
Generated Samples' Mean: -1.33252189683
Difference: 0.00415076824288

Original Training Set's SD: 2.11839826059
Generated Samples' SD: 2.17718388418
Difference: 0.0587856235889

(27)

Observation → from above we can see that the difference b/w the original mean and standard deviation with the mean & standard deviation of 400 generated examples is almost same / difference is close to zero, because the generated examples belong to the same distribution, hence the difference b/w them is negligible.