# **Robust Regression**

유제진

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# QUANTILE REGRESSION

#### Quantile

► r-th Quantile of random variable Y with Cdf F<sub>y</sub>

$$Q_Y(\tau) = F_y^{-1}(\tau) = Inf\{y : F_Y(y) \ge \tau\}, \tau \in (0,1)$$

► Here, we define the loss function as

$$\rho_{\tau}(y) = y * (\tau - I(y < 0))$$
 where  $I$ : Indicator Function

Also known as "Check Function".

# QUANTILE REGRESSION

#### Quantile

► CDF Function : function returns probabilities of *X* being smaller than or equal to some value *x*.

$$Pr(X \le x) = F(x)$$

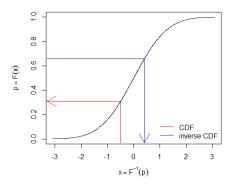
► The inverse of CDF = Quantile Function

$$F^{-1}(p) = x$$

where x would return some value *P*.

Check function

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**Gummbel Distribution** 

$$F(x) = e^{-e^{-x}} \iff F^{-1}(p) = -ln(-lnp)$$

# QUANTILE REGRESSION

#### **Generalized Inverse Distribution Function**

- ► Not every function has an inverse.
- ► So condition "Monotonically increasing" is needed.
- ► Also, to be a function, one to one condition need to be satisfied.
- ► CDF satisfy this condition.
- But what happens when it comes to discrete random variable CDF?
- ► Discrete random variable CDF is not continuous, but increasing.

# QUANTILE REGRESSION

#### **Generalized Inverse Distribution Function**

► Generalized Inverse Distribution only requires non-decreasing condition.

$$Q_Y(\tau) = F_y^{-1}(\tau) = Inf\{y : F_Y(y) \ge \tau\}, \tau \in (0,1)$$

- ▶ For given probability value  $\tau$ ,
- ▶ Look for some y that results in F(y) returning value greater or equal then  $\tau$ .
- ▶ But since there could be multiple values of *y* that meet this condition.
  - e.g.  $F(y) \ge 0$  is true for any y
- ► So use Infimum to take the smallest among *y*.

#### **Derivation of Check Function**

- ► Find "Location" *x*\* relative to a distribution or set of data *F*.
- mean

Check function

$$L_F(\bar{x}) = \int_R (x - \bar{x})^2 dF(x)$$

mean minimized the expected squared residual.

- $\blacktriangleright$  *L*<sub>F</sub> is for loss function determined by *F*.
- $\blacktriangleright$  To show  $x^*$  minimized any function begins with,
- Demonstrating the function's value does not decrease when  $x^*$  is changed by a little bit.
- ► Such a value is called a critical point of the function.

Check function

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#### Derivation of Check Function

 $\blacktriangleright$  What kind of loss function  $\Lambda$  would result in a percentile  $F^{-1}(\alpha)$  being a critical point?

$$L_F(F^{-1}(\alpha)) = \int_R \Lambda(x - F^{-1}(\alpha)) dF(x)$$
$$= \int_0^1 \Lambda(F^{-1}(u) - F^{-1}(\alpha)) du$$

► For this to be critical point, derivative must be zero.

Check function

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# **Derivation of Check Function**

$$0 = L'_F(x^*) = L_F(F^{-1}(\alpha)) = -\int_0^1 \Lambda'(F^{-1}(u) - F^{-1}(\alpha)) du$$
$$= -\int_0^\alpha \Lambda'(F^{-1}(u) - F^{-1}(\alpha)) du - \int_0^1 \Lambda'(F^{-1}(u) - F^{-1}(\alpha)) du$$

- ▶ On the left hand side, argument of  $\Lambda'$  is negative.
- ▶ On the right hand side, argument of  $\Lambda'$  is positive.
- ▶ Other than that, we have little control over the values of these integrals.
- Because F could be any distribution function.

# QUANTILE REGRESSION

#### **Derivation of Check Function**

- ▶ Now, make  $\Lambda'$  depend only on the sign of its argument, otherwise it must be constant.
- This implies Λ will be piecewise linear, potentially with different slopes to the left and right of zero.
- Moreover, Rescaling  $\Lambda$  by a constant will not change its properties.
- ► So may feel free to set the left hand slope to -1.

# OUANTILE REGRESSION

#### **Derivation of Check Function**

- Let  $\tau > 0$  be the right hand slope.
- ► Then final equation simplifies to

$$0 = \alpha - \tau(1 - \alpha) \Longrightarrow \tau = \frac{\alpha}{1 - \alpha}$$

To conclude,

$$\Lambda(x) = \begin{cases} \frac{-x}{\alpha} & \Longrightarrow \Lambda(x) = \begin{cases} -(1-\alpha)x & x \le 0\\ \alpha x & x \ge 0 \end{cases}$$

#### Drawing Check function

#### Check Function

```
▶ rho <- function(u) {u * (tau - ifelse(u <</p>
 0,1,0))
 tau < .25; curve (rho, -2, 2, lty=1, lwd=3)
 tau \leftarrow .50; curve(rho,
 -2,2,1ty=2,col="blue",add=T,lwd=3)
 tau <- .90; curve(rho,
 -2, 2, 1tv=3, col="red", add=T, 1wd=3)
 abline (v=0, lty=5, col="gray", lwd=3)
  legend("bottomleft",c(".25",".5",".9"),
  lty=1:3, col=c("black", "blue", "red"), cex=.6)
```

Check function

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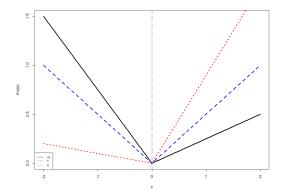


Figure 1: Loss of check function

#### Median

► Consider simple  $\{y_1 \cdots y_n\}$ 

$$\min_{\mu} \{ \sum_{i=1}^{n} |y_i - \mu| \} \iff \min_{\mu, a, b} \{ \sum_{i=1}^{n} a_i + b_i \}$$

subject to 
$$a_i, b_i \ge 0$$
 and  $y_i - \mu = a_i - b_i \quad \forall 1, \dots, n$ 

- ► Turn into Linear Programming Problem.
- ► To illustrate, consider a lognormal sample.

#### Median

► Pre-built Function

```
n=101
set.seed(32420)
y = rlnorm(n)
median(y)
[1] 1.03435
```

#### Median

► Solve by Linear Programming

#### Quantile

Now Change the equation into quantile regression,

$$\min_{\mu,a,b} \{ \sum_{i=1}^{n} \tau a_i + (1-\tau)b_i \}$$

subject to 
$$a_i, b_i \ge 0$$
 and  $y_i - [\beta_0^{\tau} + \beta_1^{\tau} x_i] = a_i - b_i \quad \forall 1, \dots, n$ 

### **Ouantile**

► Pre-built Function

```
t.au = .3
quantile (y, tau)
30% 0.5775248
```

► Solve by Linear Programming

```
A1 = cbind(diag(2*n), 0)
A2 = cbind(diag(n), -diag(n), 1)
r = lp("min", c(rep(tau,n), rep(1-tau,n), 0),
        rbind(A1, A2), c(rep(">=", 2*n),
        rep("=", n)), c(rep(0,2*n), y))
tail(r$solution,1)
30% 0.5775248
```

Building OR

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#### **Ouantile**

- Other than above programming, there are several more approaches.
- ▶ If interested, check out more on this page. Quantile Regression Computation: From the Inside and the Outside
- ► Instead building your own function, we can use R package "quantreg".

# **Basic Setting**

Import necessary packages.

```
library (MASS)
library (quantreg)
set.seed(32420)
```

- Same seed will be used repeatedly.
- ▶  $n_1 = 1000$  and  $n_2 = 200$  for this trial.

#### Generate Data Sample

▶ Generate *n*<sub>1</sub> of Good Observation.

```
data1 \leftarrow mvrnorm(n=n1, mu=c(0, 0), Sigma =
matrix(c(1,0.8,0.8,1),ncol=2))
```

▶ Generate *n*<sub>2</sub> of Bad Observation.

```
data2 \leftarrow mvrnorm(n=n2, mu=c(1.5, -1.5), Sigma
= .2*diag(c(1,1)))
```

#### Generate Data Sample

Bind the Data and turn it into data frame.

```
data <- rbind(data1, data2)</pre>
data <- data.frame(data)</pre>
```

Distinguish the Good and Bad using Indicator vector.

```
names(data) <- c("X", "Y")</pre>
ind <- c(rep(1,n1), rep(2,n2))
```

#### **Drawing Scatter Plot**

Draw plot of the generated Data Sample.

```
plot (Y \sim X, data, pch=c("x", "o") [ind],
     col=c("black", "red")[ind],
     main=paste("N1 = ", n1, "N2 = ", n2)
```

# OUTLIER PROTECTION ON MEDIAN REGRESSION Data Plot

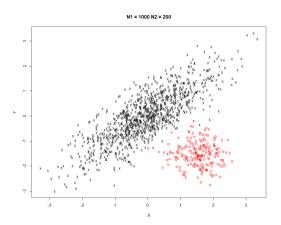


Figure 2: Scatter Plot

# **Data Plot**

▶ Fit and Draw Quantile Regression of  $\tau$ =0.5.

```
r1 <- rg(Y \sim X, data=data, tau=0.5)
abline(r1)
summary (r1)
```

See Summary if you need.

# OUTLIER PROTECTION ON MEDIAN REGRESSION Drawing Quantile Regression Plot

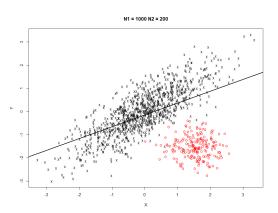


Figure 3: Plot with QR

### **Comparison Data Plot**

► Draw Ordinary Least Square.

```
abline(lm(Y\sim X, data), lty=2, col="red")
```

► Draw Ordinary Least Square on Good.

► Draw Topleft Index.

```
legend("topleft",c("L1","ols","ols on good"),
    inset=0.02, lty=c(1,2,1),
    col=c("black","red","blue"),cex=.9)
```

# OUTLIER PROTECTION ON MEDIAN REGRESSION Data Plot

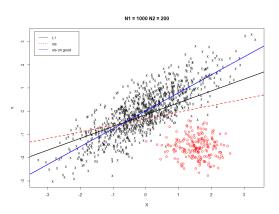


Figure 4: Final Comparison Plot

# **Comparison Data Plot**

- ► Produce 4 simulation plot with difference amount of bad observation.
- ▶ By observing several graph, we can guess how much outlier can be bear-ed on median regression.
- ▶ In this simulation, 1000 good observations could bear till 371 bad observations.

### Merge into One Function

```
gresult <- function(n1, n2) {</pre>
 data \leftarrow mvrnorm(n=n1, mu=c(0,0),
           Sigma = matrix(c(1,0.8,0.8,1),ncol=2))
 data <- rbind(data, mvrnorm(n=n2, mu=c(1.5, -1.5),
           Sigma = .2*diag(c(1,1)))
 data <- data.frame(data) names(data) <- c("X", "Y")
 ind <-c(rep(1,n1), rep(2,n2))
 plot(Y\sim X, data, pch=c("x", "o")[ind], col=c("black", "red")[ind],
       main=paste("N1 =",n1,"N2 =",n2))
 summary(r1 <- rg(Y\simX,data=data,tau=0.5))
 abline(r1)
 abline (lm(Y \sim X, data), ltv=2, col="red")
 abline (lm(Y \sim X, data, subset= 1:n1), lty=1, col='blue')
 legend("topleft",c("L1", "ols", "ols on good"),
           inset=0.02, ltv=c(1,2,1),
           col=c("black", "red", "blue"), cex=.9)}
```

# OUTLIER PROTECTION ON MEDIAN REGRESSION Data Plot

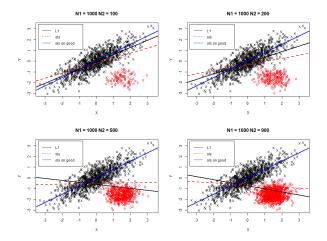


Figure 5: Final Comparison Plot



# OUTLIER PROTECTION ON MEDIAN REGRESSION Data Plot

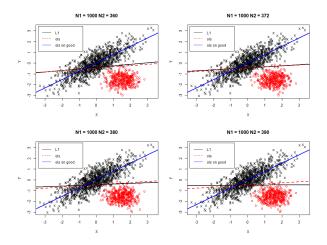


Figure 6: Final Comparison Plot



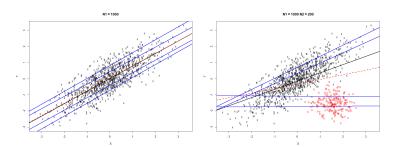
### Heteroscedasticity

- Simply to check, Heteroscedasticity, check out the Quantile Regression line.
- If they look like having different slope, probably there will be Heteroscedasticity.

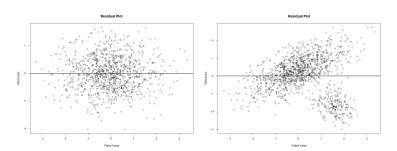
#### **Predictive Interval**

- ► For different quantiles, there will be different predictive intervals obviously.
- ► By plotting quantiles, the accuracy of OLS interval can be checked.

#### Heteroscedasticity



#### Heteroscedasticity - Residual Plot



# Predictive Intervals

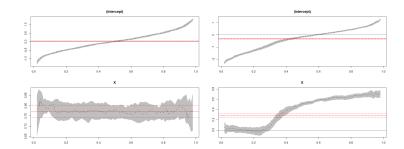
► Summary Plot of the Intercept and Coefficient.

```
plot (summary (rq (Y \sim X, datal,
        tau=2:98/100))
```

- ► The horizontal line is the OLS estimate.
- ▶ Dashed lines for confidence interval for OLS estimate.

### OUTLIER PROTECTION ON MEDIAN REGRESSION

#### **Predictive Intervals**



### QUANTILE REGRESSION ON ENGEL DATA

#### Engel's Law

- ► As income rises, the proportion of income spent on food falls, even if absolute expenditure on food rises.
- ► Empirical Law based on observation.
- ► This example shows expenditures on food as a function of income for 19th century Belgian households.

#### **Plot Data**

► Engel Data is pre-included in the R package 'quantreg'.

```
data(engel)
```

► Draw the Scatter plot.

▶ Draw the OLS on the plot.

```
abline(lm(foodexp~income,engel))
```

# QUANTILE REGRESSION ON ENGEL DATA Data Plot

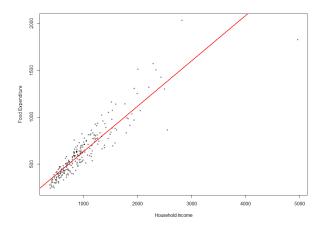


Figure 10: Scatter Plot and OLS



## OUANTILE REGRESSION ON ENGEL DATA

#### **Plot Quantile Regression**

Draw Median Regression on the plot.

```
abline(rq(foodexp~income, engel, tau=.5),
col="blue")
```

► Draw the 10-20-75-90 Quantile Regression on the plot.

```
taus \leftarrow c(.1,.25,.75,.90)
for( i in 1:length(taus)){
abline(rq(foodexp~income, engel, tau=taus[i]),)
       col="gray")}
```

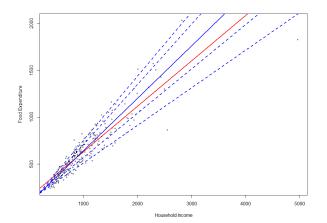


Figure 11: Quantile Regression



### QUANTILE REGRESSION ON ENGEL DATA

#### **Plot Summary**

► Summary Plot of the Intercept and Coefficient.

- ► The horizontal line is the OLS estimate.
- ▶ Dashed lines for confidence interval for OLS estimate.

# QUANTILE REGRESSION ON ENGEL DATA Data Plot

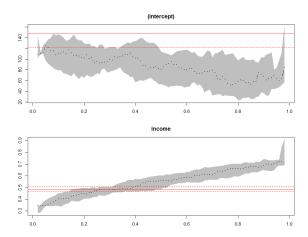


Figure 12: Summary Plot of Quantile Regression

# QUANTILE REGRESSION ON ENGEL DATA Log Transformation Case of Engel Data

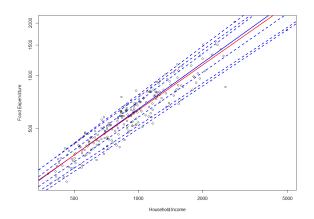


Figure 13: Summary Plot of Log10 Quantile Regression



### OUANTILE REGRESSION ON WAGE DATA

#### Wage Data from ISLR

- ► R package 'ISLR' contains only data, for use of the book "Introduction to Statistical Learning with applications in R".
- ► Load Wage Data from the package 'ISLR'.

```
library (ISLR)
data(Wage)
```

► Rest of procedure will be same as the previous case.

# QUANTILE REGRESSION ON WAGE DATA Data Plot

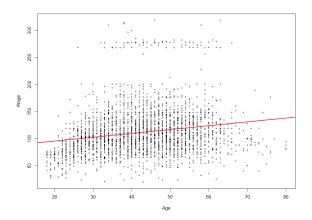


Figure 14: Scatter Plot and OLS



# QUANTILE REGRESSION ON WAGE DATA Data Plot

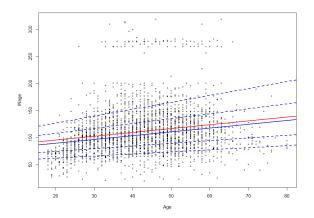


Figure 15: Quantile Regression



# QUANTILE REGRESSION ON WAGE DATA Data Plot

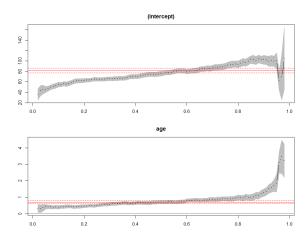


Figure 16: Summary Plot of Quantile Regression

## QUANTILE REGRESSION ON WAGE DATA Log Transformation Case of Wage Data

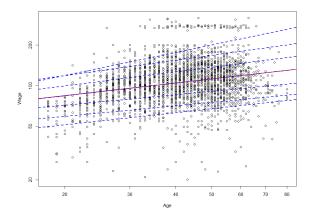


Figure 17: Summary Plot of Log10 Quantile Regression