

Model Specification

We'll be working today with the wage2 dataset, which includes monthly wages of male earners along with a variety of characteristics. We'll be attempting to estimate some fairly standard wage models, but we'll also try to answer the most vexing question for many students: what variables should I put in my model?

The most important answer to that question is to use theory. Theory and previous results are our only guide—the data simply can't tell you by themselves what belongs in the model and what doesn't. However, we can use a combination of theory and applied data analysis to come up with a model that fits the data well and says something interesting about theory.

Missing Data

Before we start with all that, let's talk again about how Stata handles missing data. Let's assume that we want to estimate several nested models, first with hours, education and age, then the same model with mother's education, then the same model with father's education, then a final model with all variables. Our results look like this:

```
.
. reg lwage hours educ age
```

Source	SS	df	MS			
Model	21.7514568	3	7.25048559	Number of obs =	935	
Residual	143.904838	931	.15457018	F(3, 931) =	46.91	
Total	165.656294	934	.177362199	Prob > F =	0.0000	
				R-squared =	0.1313	
				Adj R-squared =	0.1285	
				Root MSE =	.39315	

	lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hours		-.0047011	.0017887	-2.63	0.009	-.0082115	-.0011906
educ		.0616404	.0058814	10.48	0.000	.0500981	.0731827
age		.0227339	.0041411	5.49	0.000	.0146069	.0308608
_cons		5.403279	.1732026	31.20	0.000	5.063366	5.743191


```
.
. reg lwage hours educ age meduc
```

Source	SS	df	MS			
Model	22.8514162	4	5.71285406	Number of obs =	857	
Residual	126.509635	852	.148485487	F(4, 852) =	38.47	
Total	149.361051	856	.174487209	Prob > F =	0.0000	
				R-squared =	0.1530	
				Adj R-squared =	0.1490	
				Root MSE =	.38534	

	lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hours							
educ							
age							
meduc							
_cons							

hours		-.0058052	.0018374	-3.16	0.002	-.0094115	-.0021989
educ		.0525597	.0064521	8.15	0.000	.0398957	.0652236
age		.0243798	.0042747	5.70	0.000	.0159896	.03277
meduc		.0184424	.0049725	3.71	0.000	.0086826	.0282022
_cons		5.33402	.1776844	30.02	0.000	4.98527	5.682771

```
.
. reg lwage hours educ age feduc
```

Source		SS	df	MS		Number of obs =	741
Model		20.2139719	4	5.05349299		F(4, 736) =	34.17
Residual		108.836202	736	.147875274		Prob > F	= 0.0000
Total		129.050173	740	.174392126		R-squared	= 0.1566
						Adj R-squared	= 0.1521
						Root MSE	= .38455

lwage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours		-.007041	.0019639	-3.59	0.000	-.0108964 -.0031855
educ		.0475258	.0070026	6.79	0.000	.0337785 .0612732
age		.0262759	.0045806	5.74	0.000	.0172834 .0352685
feduc		.0172076	.0047569	3.62	0.000	.0078689 .0265462
_cons		5.421121	.1897058	28.58	0.000	5.048692 5.79355

```
.
. reg lwage hours educ age feduc meduc
```

Source		SS	df	MS		Number of obs =	722
Model		20.519095	5	4.10381899		F(5, 716) =	27.64
Residual		106.292836	716	.148453682		Prob > F	= 0.0000
Total		126.811931	721	.1758834		R-squared	= 0.1618
						Adj R-squared	= 0.1560
						Root MSE	= .3853

lwage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours		-.0071408	.0019821	-3.60	0.000	-.0110321 -.0032495
educ		.046006	.0072322	6.36	0.000	.0318071 .0602049
age		.0249456	.0046705	5.34	0.000	.0157762 .034115
feduc		.0114239	.0055571	2.06	0.040	.0005138 .022334
meduc		.0132414	.0063094	2.10	0.036	.0008543 .0256285
_cons		5.404781	.1929204	28.02	0.000	5.026024 5.783539

The results are extremely problematic because each set of results is on a different sample! The first set has 857 observations, the second 741, and down to 722 for the final one. Stata performs casewise deletion when running regressions, and doesn't adjust unless you tell it to. In this case none of the standard tests of model fit are relevant, because it's not the same sample.

The solution is to use the `e(sample)` command to limit the sample to the relevant analysis sample. First, run the model that restricts the data the most (has the most missing data), then limit subsequent models using the statement `if e(sample)==1`.

Selecting Variables: Stepwise Regression OR A Cautionary Tale of Woe

When selecting variables for a model, students are sometimes tempted by the dark side of stepwise regression, which is a step on the path toward the greater evil that is data mining. I will illustrate why this is a bad idea. The basic idea with stepwise regression is to eliminate variables from the model one at a time—if the variable is not significant, it gets dropped. However, this method is very sensitive to the overall group of variables used, essentially just pushing decisions one step back, and then using an arbitrary non-theoretical standard for variable inclusion. There is no good theoretical reason to use this procedure.

Selecting Variables: RESET test

One question that comes up frequently is whether one or more variables ought to be expressed as quadratic or higher-order polynomials in the equation. The RESET test can help with this problem. Specifying the RESET test without any options means that Stata will fit the model with the second, third and fourth powers of \hat{y} . Specifying the option `rhs` will use powers of the individual regressors.

In Stata, we would run:

```
.
. reg lwage hours age educ

      Source |      SS      df      MS      Number of obs =      935
-----+-----+-----+-----+-----+-----+-----+-----
      Model | 21.7514568      3   7.25048559   F( 3,  931) =    46.91
      Residual | 143.904838    931   .15457018   Prob > F      =    0.0000
-----+-----+-----+-----+-----+-----+-----
      Total | 165.656294    934   .177362199   R-squared     =    0.1313
                                           Adj R-squared =    0.1285
                                           Root MSE     =    .39315

-----+-----+-----+-----+-----+-----+-----+-----
      lwage |      Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----+-----+-----+-----+-----+-----+-----
      hours | -.0047011   .0017887     -2.63   0.009   -.0082115   -.0011906
       age | .0227339   .0041411      5.49   0.000    .0146069    .0308608
       educ | .0616404   .0058814     10.48   0.000    .0500981    .0731827
       _cons | 5.403279   .1732026     31.20   0.000    5.063366    5.743191
-----+-----+-----+-----+-----+-----+-----+-----

.
. estat ovtest

Ramsey RESET test using powers of the fitted values of lwage
Ho: model has no omitted variables
      F(3, 928) =      0.32
      Prob > F =      0.8089

.
. estat ovtest, rhs

Ramsey RESET test using powers of the independent variables
Ho: model has no omitted variables
      F(9, 922) =      2.12
      Prob > F =      0.0255
```

The result of the first test is not significant, but the result of the second test is. This indicates that we might want to include some additional powers of the right hand variables. Let's begin by introducing a quadratic function of age:

```
. gen agesq=age^2
.
. label var agesq "Age squared"
.
. reg lwage hours educ age agesq
```

Source	SS	df	MS	Number of obs =	935
Model	21.7551592	4	5.43878981	F(4, 930) =	35.15
Residual	143.901135	930	.154732403	Prob > F =	0.0000
				R-squared =	0.1313
				Adj R-squared =	0.1276
Total	165.656294	934	.177362199	Root MSE =	.39336

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours	-.0047	.0017897	-2.63	0.009	-.0082123 -.0011876
educ	.0615585	.0059083	10.42	0.000	.0499634 .0731536
age	.0388675	.1043805	0.37	0.710	-.1659811 .2437162
agesq	-.0002425	.0015678	-0.15	0.877	-.0033193 .0028343
_cons	5.138356	1.721371	2.99	0.003	1.760134 8.516578


```
.
. test age agesq
( 1) age = 0
( 2) agesq = 0

F( 2, 930) = 15.07
Prob > F = 0.0000
```

The two terms for age are jointly significant, but it looks like we could safely exclude age squared from the model without any loss of model fit.

Now let's try education squared:

```
. gen educsq=educ^2
.
. la var educsq "Education squared"
.
. reg lwage hours age educ educsq
```

Source	SS	df	MS	Number of obs =	935
Model	22.3576551	4	5.58941378	F(4, 930) =	36.27
Residual	143.298639	930	.154084558	Prob > F =	0.0000
				R-squared =	0.1350
				Adj R-squared =	0.1312
Total	165.656294	934	.177362199	Root MSE =	.39254

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours	-.004608	.0017866	-2.58	0.010	-.0081141 -.0011018
age	.0243563	.0042147	5.78	0.000	.0160848 .0326278
educ	.2161619	.0781252	2.77	0.006	.0628397 .369484
educsq	-.0054912	.0027685	-1.98	0.048	-.0109245 -.000058
_cons	4.286926	.5887928	7.28	0.000	3.13141 5.442443

This does result in a statistically significant increase in model fit. The way I would prefer approaching this problem is to fully specify the model, then restrict it appropriately, like so:

```
. /*Preferred method */
.
. reg lwage hours age agesq educ educsq
```

Source	SS	df	MS	Number of obs =	935
Model	22.3674226	5	4.47348452	F(5, 929) =	29.00
Residual	143.288872	929	.154239905	Prob > F =	0.0000
Total	165.656294	934	.177362199	R-squared =	0.1350
				Adj R-squared =	0.1304
				Root MSE =	.39273

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours	-.0046056	.0017875	-2.58	0.010	-.0081135 -.0010976
age	.050602	.1043806	0.48	0.628	-.154247 .255451
agesq	-.0003944	.0015672	-0.25	0.801	-.0034699 .0026812
educ	.2169837	.0782328	2.77	0.006	.0634502 .3705171
educsq	-.0055252	.0027732	-1.99	0.047	-.0109676 -.0000828
_cons	3.849225	1.836393	2.10	0.036	.2452658 7.453184


```
.
. test age agesq

( 1) age = 0
( 2) agesq = 0

F( 2, 929) = 16.71
Prob > F = 0.0000

.
. test educ educsq

( 1) educ = 0
( 2) educsq = 0

F( 2, 929) = 56.44
Prob > F = 0.0000
```

Selecting Variables: Non-Nested Models

In many situations, models are based on competing hypotheses, and so they don't nest within one another. Let's say we have one model that posits education as the key to wages, another that posits iq as the key to wages. To test whether one is better than the other, we use the Davidson-Mackinnon test:

```
. reg lwage hours iq
```

Source	SS	df	MS	Number of obs =	935
Model	17.2420918	2	8.62104588	F(2, 932) =	54.14
Residual	148.414203	932	.159242707	Prob > F =	0.0000
				R-squared =	0.1041

```

-----+-----
Total | 165.656294 934 .177362199
Adj R-squared = 0.1022
Root MSE = .39905

-----+-----
lwage | Coef. Std. Err. t P>|t| [95% Conf. Interval]
-----+-----
hours | -.0041302 .0018124 -2.28 0.023 -.007687 -.0005734
iq | .0089535 .0008698 10.29 0.000 .0072465 .0106606
_cons | 6.053607 .1150416 52.62 0.000 5.827837 6.279378
-----+-----

.
. reg lwage hours educ

Source | SS df MS Number of obs = 935
-----+-----
Model | 17.0930188 2 8.54650938 F( 2, 932) = 53.62
Residual | 148.563276 932 .159402656 Prob > F = 0.0000
-----+-----
Total | 165.656294 934 .177362199 R-squared = 0.1032
Adj R-squared = 0.1013
Root MSE = .39925

-----+-----
lwage | Coef. Std. Err. t P>|t| [95% Conf. Interval]
-----+-----
hours | -.0044454 .0018159 -2.45 0.015 -.0080091 -.0008817
educ | .0611697 .005972 10.24 0.000 .0494497 .0728898
_cons | 6.150426 .108791 56.53 0.000 5.936923 6.36393
-----+-----

.
. nnest lwage hours iq (hours educ)

M1 : Y = a + Xb with X = [hours iq]
M2 : Y = a + Zg with Z = [hours educ]

J test for non-nested models

H0 : M1 t(931) 5.89476
H1 : M2 p-val 0.00000

H0 : M2 t(931) 5.97647
H1 : M1 p-val 0.00000

Cox-Pesaran test for non-nested models

H0 : M1 N(0,1) -8.87744
H1 : M2 p-val 0.00000

H0 : M2 N(0,1) -9.05449
H1 : M1 p-val 0.00000

.

```

The results of this test indicate that it would be better to include both of these models, in a sort of “super” model.

Interactions

Interactions can be difficult to understand, but they are key to getting a handle on sometimes important moderating variables in an analysis.

Interactions with two binary variables

Let's say we're interested in whether marriage affects wages differently for black and white men. The specification of an interaction between the two binary variables of white and married would look like this:

```
. reg lwage hours age educ i.black#i.married iq meduc south urban
```

Source	SS	df	MS	Number of obs = 857		
Model	38.9381574	10	3.89381574	F(10, 846) = 29.83		
Residual	110.422894	846	.130523515	Prob > F = 0.0000		
Total	149.361051	856	.174487209	R-squared = 0.2607		
				Adj R-squared = 0.2520		
				Root MSE = .36128		

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hours	-.0063301	.0017273	-3.66	0.000	-.0097204	-.0029398
age	.0220385	.0040456	5.45	0.000	.0140979	.0299792
educ	.0363765	.0068217	5.33	0.000	.0229871	.0497659
black#i.married						
0 1	.1817023	.0438419	4.14	0.000	.0956506	.267754
1 0	-.2750052	.1036788	-2.65	0.008	-.478503	-.0715073
1 1	.0597192	.0605138	0.99	0.324	-.0590557	.178494
iq	.0039649	.0010438	3.80	0.000	.0019162	.0060137
meduc	.0104632	.0047967	2.18	0.029	.0010483	.0198781
south	-.0729291	.0276262	-2.64	0.008	-.1271531	-.0187051
urban	.179466	.0280946	6.39	0.000	.1243228	.2346092
_cons	5.084363	.1866763	27.24	0.000	4.717961	5.450766

Interactions with a binary and continuous variable

Let's say we're interested in whether education affects wages differently for black and white men. The specification of an interaction between a binary variable and a continuous variable would look like this:

```
. reg lwage hours age educ black i.black#c.educ married iq meduc south urban
```

Source	SS	df	MS	Number of obs = 857		
Model	38.7974052	10	3.87974052	F(10, 846) = 29.69		
Residual	110.563646	846	.130689889	Prob > F = 0.0000		
Total	149.361051	856	.174487209	R-squared = 0.2598		
				Adj R-squared = 0.2510		
				Root MSE = .36151		

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
hours	-.0063045	.00173	-3.64	0.000	-.0097001	-.002909
age	.0215072	.0040455	5.32	0.000	.0135669	.0294476
educ	.0378404	.0069883	5.41	0.000	.0241239	.0515569
black	.1034702	.2766783	0.37	0.709	-.4395862	.6465265
black#c.educ						
1	-.0196117	.0215854	-0.91	0.364	-.0619789	.0227554

married		.2051596	.0402851	5.09	0.000	.1260892	.28423
iq		.0040191	.0010442	3.85	0.000	.0019695	.0060687
meduc		.0102747	.0047964	2.14	0.032	.0008604	.0196889
south		-.069586	.0276172	-2.52	0.012	-.1237922	-.0153798
urban		.179133	.0281112	6.37	0.000	.1239572	.2343088
_cons		5.05532	.187864	26.91	0.000	4.686586	5.424055

Interactions with two continous variables

Finally let's say we think that education will affect your wages differently depending on your age. The specification of an interaction between two continous variables would look like this:

```
. reg lwage hours age educ c.age#c.educ black married iq meduc south urban
```

Source		SS	df	MS		Number of obs =	857
Model		39.1769054	10	3.91769054		F(10, 846) =	30.08
Residual		110.184146	846	.130241307		Prob > F =	0.0000
Total		149.361051	856	.174487209		R-squared =	0.2623
						Adj R-squared =	0.2536
						Root MSE =	.36089

lwage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
hours		-.0066151	.0017294	-3.83	0.000	-.0100095 -.0032208
age		-.0280105	.0260076	-1.08	0.282	-.0790575 .0230364
educ		-.0876756	.0645406	-1.36	0.175	-.2143541 .0390029
c.age#c.educ		.0037019	.0019136	1.93	0.053	-.0000542 .0074579
black		-.1466181	.0430343	-3.41	0.001	-.2310847 -.0621515
married		.2063299	.0402134	5.13	0.000	.1274003 .2852596
iq		.0040003	.0010423	3.84	0.000	.0019545 .0060461
meduc		.0100804	.0047844	2.11	0.035	.0006897 .019471
south		-.06698	.0276104	-2.43	0.015	-.121173 -.012787
urban		.182844	.02813	6.50	0.000	.1276312 .2380569
_cons		6.747921	.8848456	7.63	0.000	5.011171 8.484672

The do file has extensive examples of how to use margins to create nice plots of various types of interactions. We'll go over these in class.

Not-so-quick Exercise

In pairs, I would like you to estimate the best possible model using the wage2 dataset. Think about model specification and functional form, with an eye toward possible non-linearities and other issues. Generate a do file that walks through your process of identifying the best model. Generate a fancy graph that shows the predictions made by your model.