

Why does the Gibbs sampler
work on hierarchical models?

Krys Łatuszyński (Warwick)

joint work with

Omiros Papaspiliopoulos (Barcelona)

Natesh Pillai (Harvard)

Gareth Roberts (Warwick)

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$$\theta \rightarrow X \rightarrow Y$$

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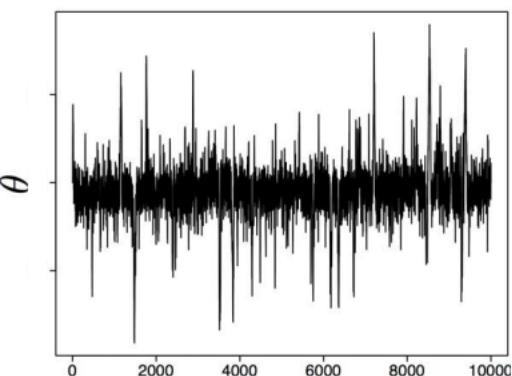
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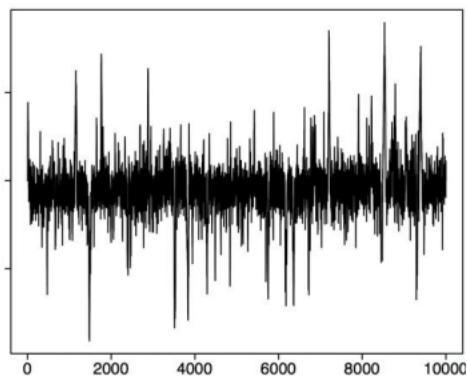
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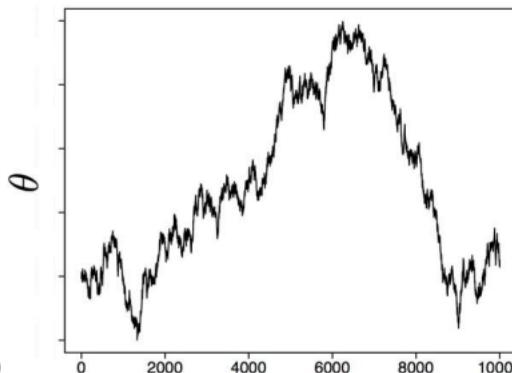
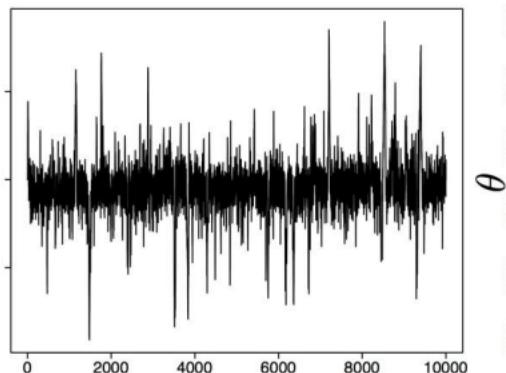


lucky \rightarrow paper
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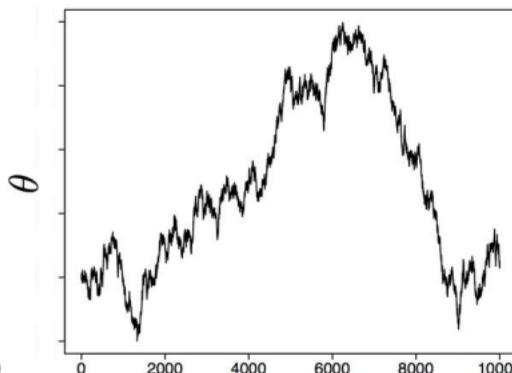
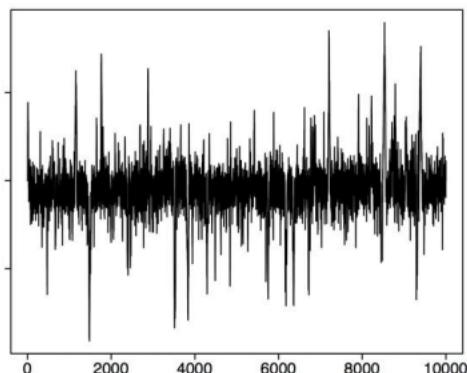


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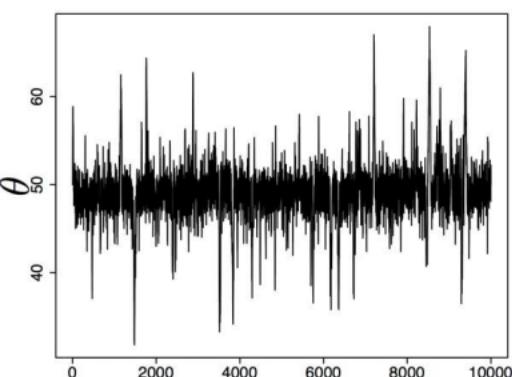
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unlucky \rightarrow no paper
NO HAPPY END!

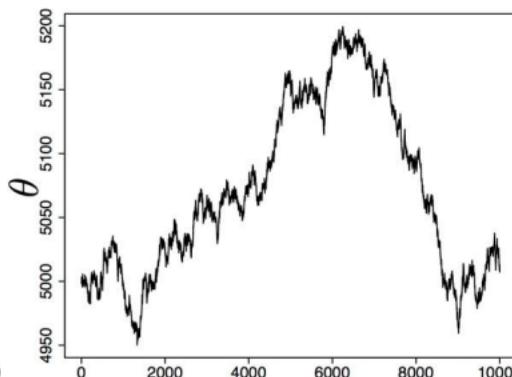
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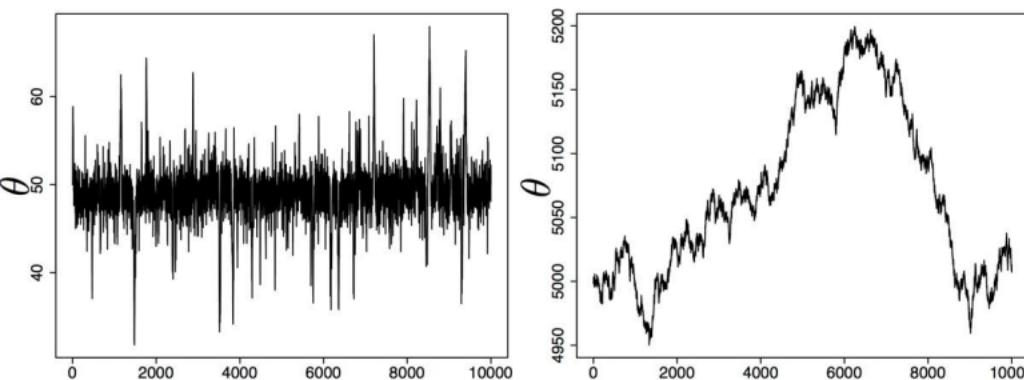


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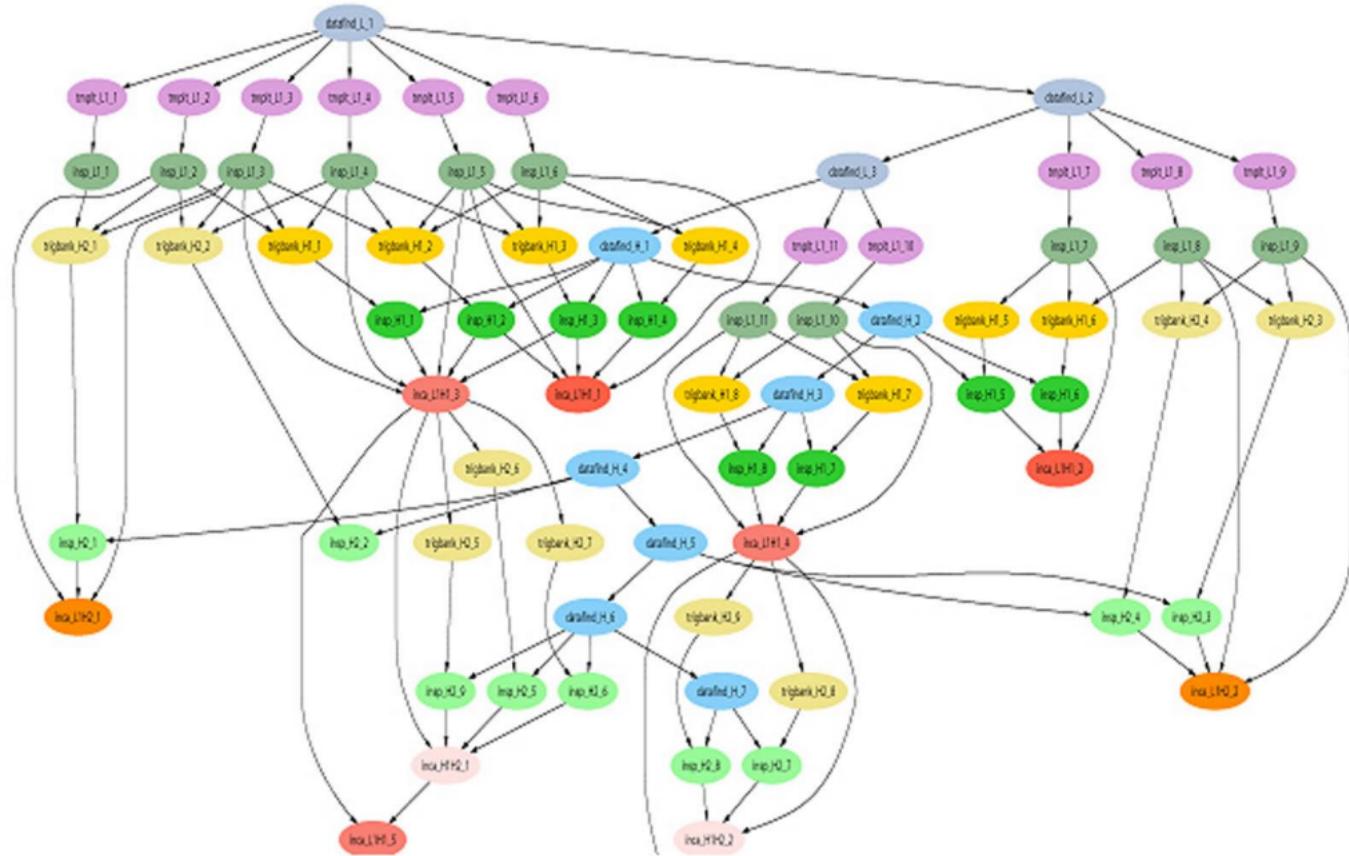


unlucky \rightarrow wrong paper (possibly)
NO HAPPY END!

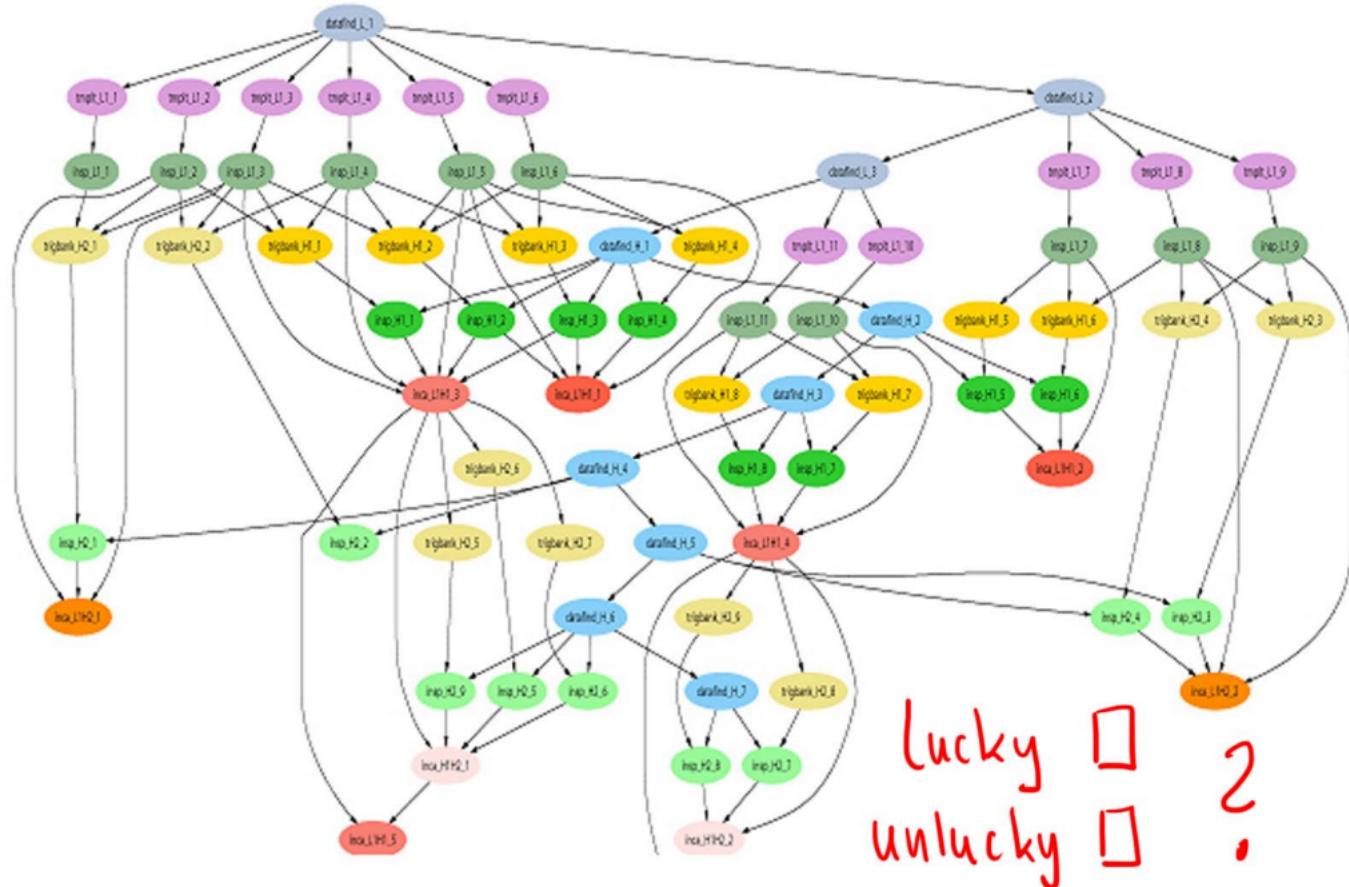
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More recently, there was a statistician ...

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Why hierarchical models?

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- Natural for model building
- Local interpretation
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 - how to tell ?

Convergence of MCMC

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the rate of convergence ≤ 1

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GOOD NEWS: for many MCMC samplers we know β EXACTLY

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GOOD NEWS: for many MCMC samplers we know β EXACTLY

THE BAD NEWS: for almost all of these $\beta = 1$

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Coarse classification:

Convergence of MCMC

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- uniformly ergodic (UE) if V bounded and $\gamma < 1$

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GE and reversibility imply that CLTs hold for all $L^2(\pi)$ functions.

GE is essentially a necessary condition for this to hold.

Consider a version of the model

$$\emptyset \rightarrow X \rightarrow Y$$

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$$\Theta \rightarrow X \rightarrow Y, \text{ where } \Theta \sim 1$$

$$X = \Theta + \tilde{X} \quad \tilde{X} \sim N(0, \sigma_x^2)$$

$$Y = X + Z \quad Z \sim N(0, \sigma_z^2)$$

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Posterior is bivariate Gaussian.

The Gibbs sampler

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The Gibbs sampler

$W = (\Theta, X)$ is a Gaussian autoregression

$$W_{t+1} = BW_t + \text{error}$$

is GE with convergence rate

$$S_C = \frac{\sigma_y^2}{\sigma_x^2 + \sigma_y^2}$$

(Roberts, Sahu 1997)

$$\Theta \sim L$$

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$$Y = X + Z$$

$$Z \sim N(0, \sigma_y^2)$$

If we reparametrize

$$\begin{matrix} \theta \\ \tilde{x} \end{matrix} \xrightarrow{\quad} x \xrightarrow{\quad} Y$$

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$$\begin{array}{ccc} \theta & \xrightarrow{\quad} & x \rightarrow Y \\ \tilde{x} & \xrightarrow{\quad} & \end{array}$$

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and take the Gibbs sampler for (θ, \tilde{X}) , then it is GE with convergence rate

$$S_{NC} = \frac{\sigma_x^2}{\sigma_y^2 + \sigma_x^2}$$

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Non-centered parametrization (vs. centered)

Heuristic: centering works well for informative data

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Non-centered parametrization (vs. centered)

Heuristic: centering works well for informative data
non-centering for non-informative data

Consider the model

$$Y = X + Z$$

$$X = \Theta + \tilde{X}$$

$$\Theta \propto 1$$

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$$Y = X + Z \leftarrow \text{Cauchy}$$

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Joint posterior \propto

$$\frac{e^{-(x-\theta)^2/2}}{1+(y-x)^2}$$

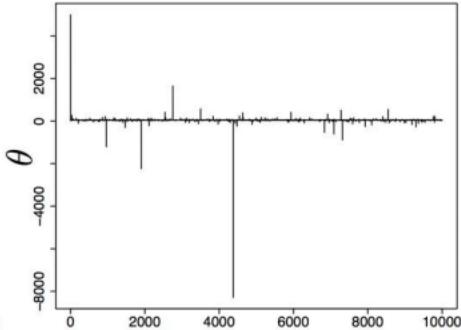
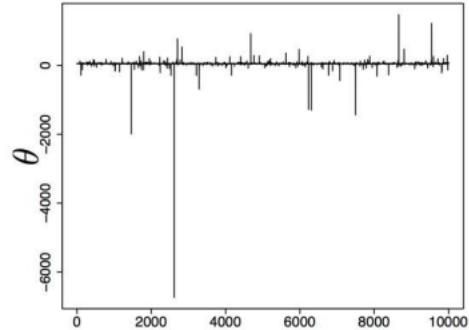
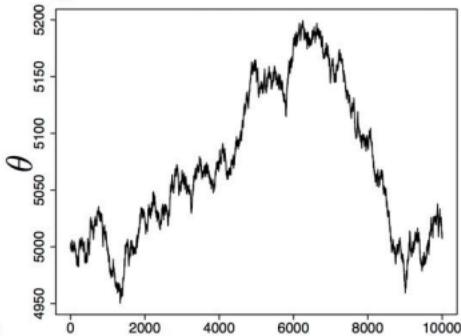
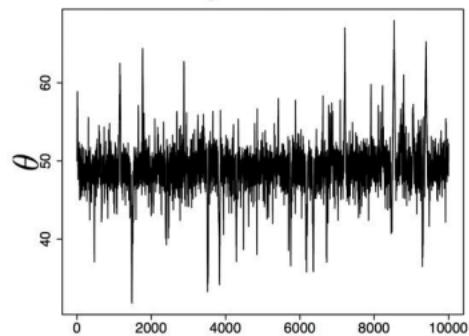
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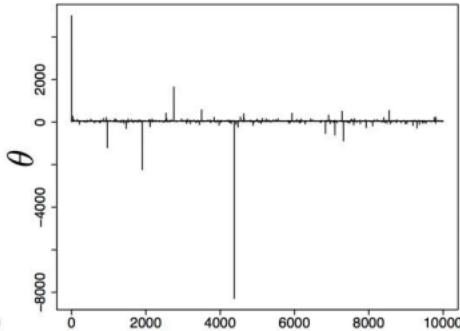
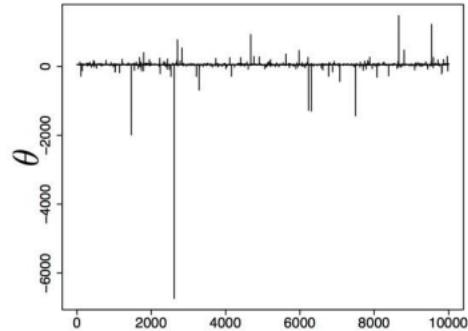
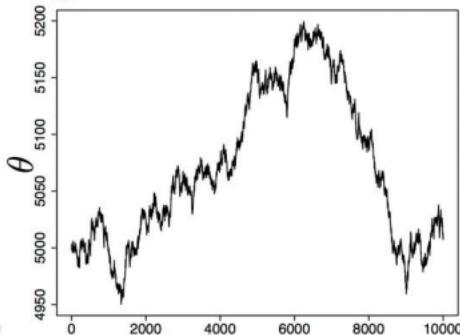
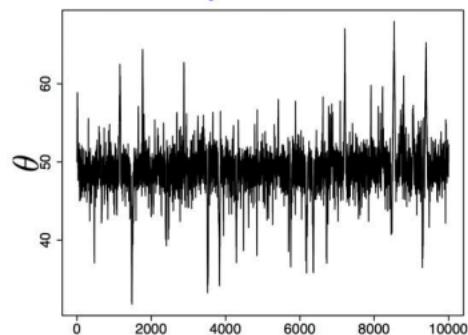
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← CENTERED
ALGORITHM
(N)



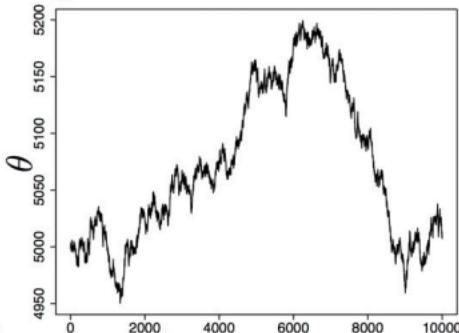
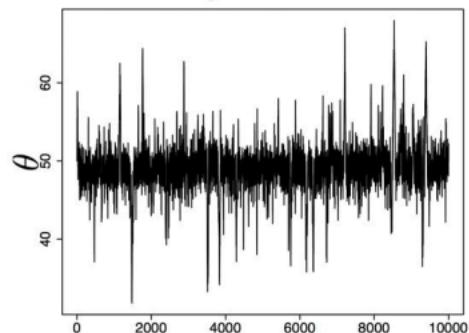
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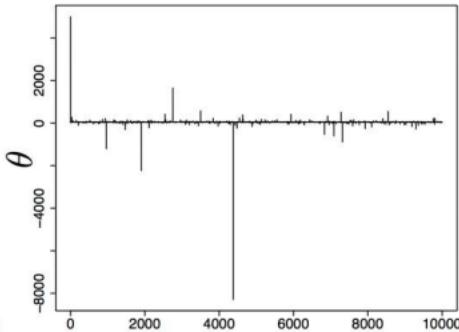
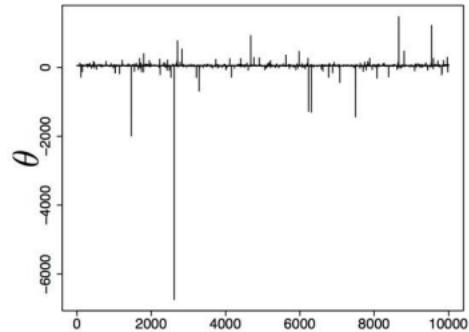
$$X = \theta + \tilde{X}^2 \leftarrow \text{Normal}$$

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← CENTERED
ALGORITHM
(N)



← NON-CENTERED
ALGORITHM
(UE)

More generally

$$Y = X + Z \quad \text{observation eqn}$$
$$X = \Theta + \tilde{X} \quad \text{hidden eqn}$$

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 $X = \theta + \tilde{X}$ hidden eqn

Error distributions for Z, \tilde{X} are

(C) Cauchy, (N) Normal, (E) Double Exponential
and (L) Light tailed $e^{-|x|^{\beta}}, \beta > 2$

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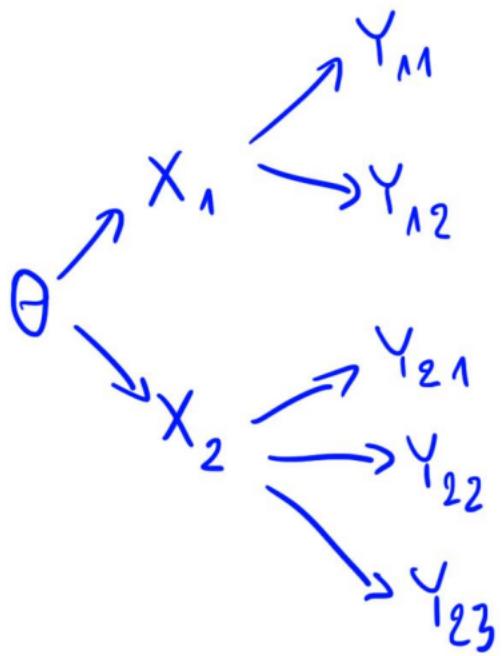
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Then the centered algorithm (θ, X) is

		Observation equation			
		C	E	G	L
Hidden eqn	C	U	U	U	U
	E	N	G/U	G	G
	G	N	G	G	G
	L	N	G	G	G

The results generalize:



Logistic regression with random effects.

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$$Y_i \sim \text{Binom}(n_i, L(x_i)) \quad 1 \leq i \leq m$$

$$x_i = \theta + z_i$$

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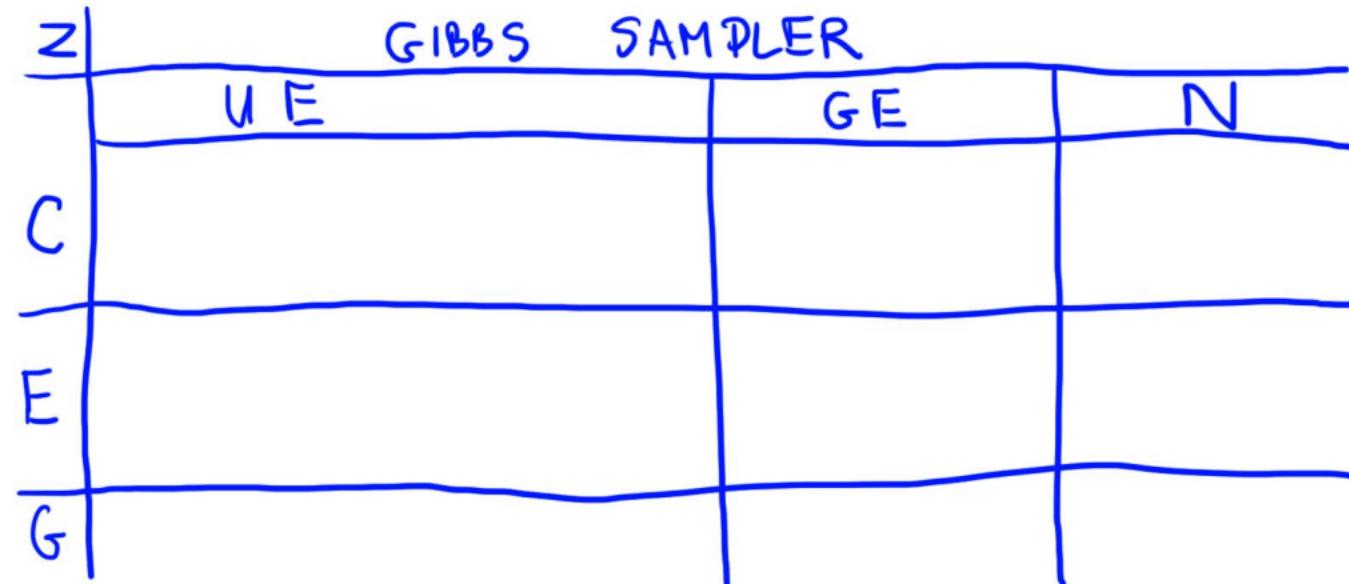
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flat prior, θ symmetric about 0, $L(x) = e^x / (1+e^x)$

Z	GIBBS SAMPLER	UE	GE	N
C		$\#\{Y_i > 0\} \geq m/2$ and $\#\{Y_i < n_i\} \geq m/2$	never	otherwise
E		$\#\{Y_i > a\} \geq m/2$ and $\#\{n_i - Y_i > a\} \geq m/2$	otherwise	never
G		never	always	never

Probit regression with random effects.

$$Y_i \sim \text{Binom}(n_i, L(x_i)) \quad 1 \leq i \leq m$$

$$x_i = \theta + z_i$$

flat prior, $\xrightarrow{\text{symmetric about 0}}$, $L(x) = \Phi(x)$

Z	GIBBS SAMPLER	UE	GE	N
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Longer Hierarchies Centered parametrization

for $\Theta_0 \rightarrow \Theta_1 \rightarrow \Theta_2 \cdots \Theta_k \rightarrow Y$

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in particular for the **RSGS**

$$E(\Theta^{(1)}) = \frac{k}{k+1} \Theta^{(0)} + \frac{1}{k+1} A \Theta^{(0)}$$

Longer hierarchies Centered parametrization

for $\Theta_0 \rightarrow \Theta_1 \rightarrow \Theta_2 \cdots \Theta_k \rightarrow Y$ with $E_i \sim N(0, \sigma_i^2)$
in particular for the **RSGS**

$$E(\Theta^{(1)}) = \frac{k}{k+1} \Theta^{(0)} + \frac{1}{k+1} A \Theta^{(0)}$$
 where

$$A = \begin{pmatrix} 0 & 1 & \dots & & \\ 1 - \rho_1 & 0 & \rho_1 & \dots & \\ 0 & 1 - \rho_2 & 0 & \rho_2 & \dots \\ \dots & \dots & \dots & \dots & \dots \\ \dots & \dots & 1 - \rho_{k-1} & 0 & \rho_{k-1} \\ & & \dots & 1 - \rho_k & 0 \end{pmatrix}$$

where $\sigma_i^2 = \sigma_i^2 / (\sigma_i^2 + \sigma_{i+1}^2)$

- The principal eigenvalue of A determines the RSGS convergence rate

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- The principal (left) eigenvector a normalised has the interpretation as the quasi-stationary vector of a Markov chain with transition matrix A and absorption from K .
- A Lyapunov drift condition is of the form $V(\theta) = \left\| \sum_{i=0}^k a_i \theta_i \right\|^2 + 1$

More general errors:

suppose $E_i \sim f_i(\cdot)$

where $f_i(x) \propto \exp\{-|x|^{\beta_i}\}$

and $2 \leq \beta_1 \leq \beta_2 \leq \dots \leq \beta_k$

Then the RSGS is GE

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Then the RSGS is GE

Remark: We generally obtain GE if the tails of the errors are lightest „close to the centre“

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Comments on proofs:

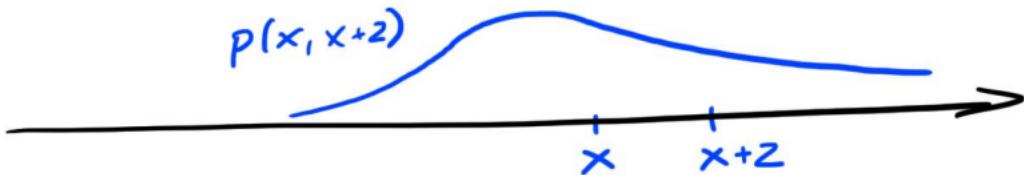
- We collapse the analysis on a bivariate Gibbs sampler, e.g $\Theta \rightarrow X \rightarrow Y$
- We conclude the properties of the bivariate Gibbs by looking at one component + Markov de-initializing processes argument.
- To deal with the one component of the bivariate Gibbs sampler, we develop a general theory of random walk like tail behaviour of Markov chains

Random walk like behaviour in the tails

P-transition kernel $p(x,y)$ - transition density

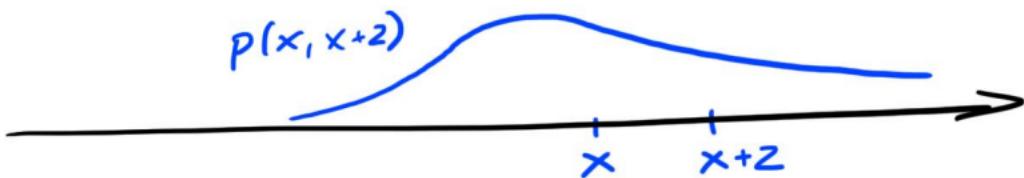
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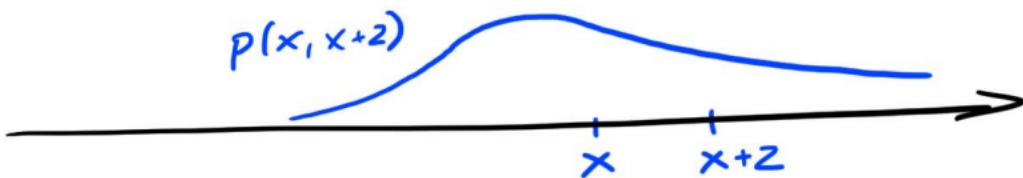
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$$\lim_{x \rightarrow \infty} p(x, x+z) =: q(z)$$

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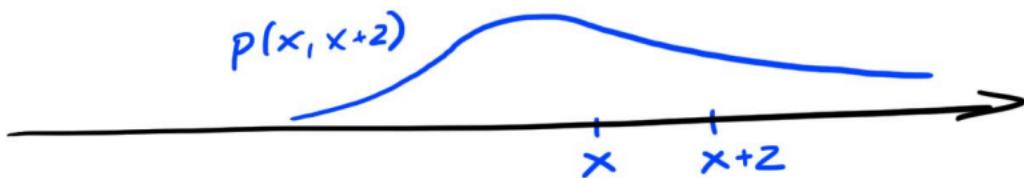
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For one component of a bivariate Gibbs sampler on a graphical model,

q typically exists!!!

The characterization via q

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Thm 1 If P is reversible and q exists,
then

$$q(z) = e^{-cz} f(z)$$

The characterization via q

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↑
constant > 0

symmetric function

The characterization via q

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or $q(z) = 0$ for all $z > 0$.

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Thm 2

$c = 0 \Leftrightarrow \pi$ has heavy tails

$\int q(z) dz > 0$. $c > 0 \Leftrightarrow \pi$ has exponential tails

+ reversibility $q(z) = 0 \Leftrightarrow \pi$ has light tails

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Thm 5 If $m_q = \int q(z) dz < 1$ (+ regularity cond)
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- There is much work still to be done!