

# HW4 MATH 4000

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## 1 Exercises

1. Let  $T$  be a geometrically distributed random variable with parameter  $p$ . Using the probability generating function, which was derived in class, find the expected value and variance of  $T$ . Then, by definition

$$P\{T = n\} = (1 - p)^{n-1}p$$

and

$$1 = \sum_{n=1}^{\infty} P\{T = n\} = \sum_{n=1}^{\infty} (1 - p)^{n-1}p.$$

Recall that the expected value  $E(T)$  is given as the sum of the possible values of  $T$  multiplied by each value's respective probability. Hence

$$E(T) = \sum_{n=1}^{\infty} nP\{T = n\} = \sum_{n=1}^{\infty} n(1 - p)^{n-1}p.$$

Rearranging, we find

$$\begin{aligned} E(T) &= \sum_{n=1}^{\infty} n(1 - p)^{n-1}p \\ &= \sum_{n=1}^{\infty} (1 - p)^{n-1}p + \sum_{n=1}^{\infty} (n - 1)(1 - p)^{n-1}p \\ &= 1 + \sum_{n=1}^{\infty} (n - 1)(1 - p)^{n-1}p \\ &= 1 + (1 - 1)(1 - p)^0 + \sum_{n=2}^{\infty} (n - 1)(1 - p)^{n-1}p \\ &= 1 + 0 + \sum_{m=1}^{\infty} m(1 - p)^m p \end{aligned}$$

$$\begin{aligned}
&= 1 + (1-p) \sum_{m=1}^{\infty} m(1-p)^{m-1}p \\
&= 1 + E(T) - pE(T)
\end{aligned}$$

Hence

$$pE(T) = 1$$

and thus the expected value is  $1/p$ . Recall next that

$$V(T) = E(T^2) - (E(T))^2.$$

Since  $(E(T))^2 = 1/p^2$ , all we need to find is  $E(T^2)$ :

$$E(T^2) = \sum_{n=1}^{\infty} n^2 P\{T = n\} = \sum_{n=1}^{\infty} n^2 (1-p)^{n-1}p.$$

Similarly to finding the expected value of  $T$ , we rearrange:

$$\begin{aligned}
E(T^2) &= \sum_{n=1}^{\infty} n^2 (1-p)^{n-1}p \\
&= \sum_{n=1}^{\infty} n(n-1)(1-p)^{n-1}p + \sum_{n=1}^{\infty} n(1-p)^{n-1}p \\
&= \sum_{n=2}^{\infty} n(n-1)(1-p)^{n-1}p + E(T) \\
&= \sum_{m=1}^{\infty} m(m+1)(1-p)^m p + \frac{1}{p} \\
&= (1-p) \sum_{m=1}^{\infty} m(m+1)(1-p)^{m-1}p + \frac{1}{p}
\end{aligned}$$

Looking closely, we find that

$$\begin{aligned}
\sum_{m=1}^{\infty} m(m+1)(1-p)^{m-1}p &= \sum_{m=1}^{\infty} m^2(1-p)^{m-1}p + \sum_{m=1}^{\infty} m(1-p)^{m-1}p \\
&= E(T^2) + E(T)
\end{aligned}$$

and

$$E(T^2) = (1-p) \left( E(T^2) + \frac{1}{p} \right) + \frac{1}{p}$$

Solving for  $E(T^2)$  yields

$$E(T^2) = \frac{2-p}{p^2},$$

and thus

$$V(T) = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}.$$

2. Yosemite Sam (YS for short) uses his trusty six gun for target practice. The number of shots it takes YS to hit the bulls eye is  $(N + 2)^4$ , where  $N$  has a Poisson distribution with parameter  $\lambda = 2$ . Find the expected number of shots for YS to hit the bulls eye.

$$P\{N = n\} = e^{-\lambda} \frac{\lambda^n}{n!}$$

Calculating the expected value of  $N$ :

$$\begin{aligned} E(N) &= \sum_{n=1}^{\infty} nP\{N = n\} = \sum_{n=1}^{\infty} ne^{-\lambda} \frac{\lambda^n}{n!} \\ &= \lambda e^{-\lambda} \sum_{n=1}^{\infty} \frac{\lambda^{n-1}}{(n-1)!} \\ &= \lambda e^{-\lambda} \sum_{m=0}^{\infty} \frac{\lambda^m}{m!}, \end{aligned}$$

And since the infinite series above represents  $e^\lambda$ :

$$E(N) = \lambda e^{-\lambda} e^\lambda = \lambda$$

Thus

$$E\{\text{number of shots}\} = E((N + 2)^4) = (\lambda + 2)^4 = 256.$$

3. To pass an algebra course Roger Rabbit must complete 4 exams having no errors. The number of errors Roger makes on exams form a sequence of independent and identically distributed random variable which are Poisson distributed with parameter  $\lambda = 3$ . Find the probability Roger must take at least 6 exams to pass the course.

Let the number of errors that RR makes on exam  $i$  be given by  $X_i$ . Then for each exam:

$$P\{X_i = 0\} = e^{-3} \approx 0.0497$$

Observe the probability that RR must take at least 6 exams is equal to the probability that at least two of his first five have an error. Suppose  $N$  was the number of exams with an error out of the first five exams. Then

$$\begin{aligned} P\{N \geq 2\} &= 1 - P\{N = 1\} - P\{N = 0\} \\ &= 1 - \binom{5}{1}(e^{-3})^4(1 - e^{-3}) - (e^{-3})^5 \approx 0.99997 \end{aligned}$$

## 2 Problems

- 4.7 (a) 1,2,3,4,5,6  
 (b) 1,2,3,4,5,6  
 (c) 2,3,4,5,6,7,8,9,10,11,12  
 (d) -5,-4,-3,-2,-1,0,1,2,3,4,5

4.8 Suppose the first die has been rolled. Then there is  $\frac{1}{6}$  chance for the result to be any of the values. If a 6 has been rolled, then the second roll does not affect the variable. If a 5 has been rolled, then only a 6 affects it. If a 4, then only a 5 or 6; and so on. Hence, if  $X$  is the variable for the maximum, then

$$P\{X = i\} = \binom{2}{1} \left(\frac{1}{6}\right) \left(\frac{i}{6}\right) - \frac{1}{36}$$

There are 2 ways to choose the first die, there is  $(1/6)$  chance that the first die is  $i$ , and there are  $(i/6)$  chance that the second die is less than or equal to  $i$ . Since there is symmetry when both dice are  $i$ , we have to subtract one outcome.

Suppose that  $Y$  was the variable for the minimum. Then

$$P\{Y = i\} = P\{X = 7 - i\} = \binom{2}{1} \left(\frac{1}{6}\right) \left(\frac{7-i}{6}\right) - \frac{1}{36}$$

Since maximum is effectively the reverse of the minimum, then all we have to do is map 1 to 6, 2 to 5, and so on. This can be achieved by  $Y = 7 - X$ . Suppose that  $Z$  was the variable for the sum. Then clearly  $P\{Z = 2\} = \frac{1}{36}$ , the outcome of rolling two ones. Likewise,  $P\{Z = 12\} = \frac{1}{36}$ , the outcome of rolling two sixes. As we approach  $Z = 7$ , the most common sum, from either direction, we always increase the outcomes by one. This is because if  $a + b = Z$ , then  $(a + 1) + b = Z + 1$ , and so for  $(a + 1) \in \{2, 3, 4, 5, 6\}$  this holds. Thus

$$P\{Z = i\} = \frac{i - 1}{36}$$

if  $i \leq 7$  and

$$P\{Z = i\} = \frac{13 - i}{36}$$

otherwise.

4.13

$$P(X = 0) = (1 - 0.3)(1 - 0.6) = 0.28$$

$$P(X = 500) = (0.3)(1 - 0.6)(0.5) + (1 - 0.3)(0.6)(0.5) = 0.27$$

$$P(X = 1000) = (0.3)(0.6)(0.5)(0.5) + (0.3)(1 - 0.6)(0.5) + (1 - 0.3)(0.6)(0.5) = 0.315$$

$$P(X = 1500) = 2 * (0.3)(0.6)(0.5)(0.5) = 0.09$$

$$P(X = 1500) = (0.3)(0.6)(0.5)(0.5) = 0.045$$

4.17 (a)

$$P(X = 1) = \left(\frac{1}{2} + 0\right) - \frac{1}{4} = \frac{1}{4}$$

$$P(X = 2) = \frac{11}{12} - \left(\frac{1}{2} + \frac{1}{4}\right) = \frac{1}{6}$$

$$P(X = 3) = 1 - \frac{11}{12} = \frac{1}{12}$$

(b)

$$P\left(\frac{1}{2} < X < \frac{3}{2}\right) = F\left(\frac{3}{2}-\right) - F\left(\frac{1}{2}\right) = \frac{5}{8} - \frac{1}{8} = \frac{1}{2}$$

4.21 The expected value after selecting a teacher will be smaller, since it is effectively just the mean, whereas the expected value after selecting a student is more representative of the student's distribution (which is more wide spread).

$$E(X) = \frac{40}{148} * 40 + \frac{33}{148} * 33 + \frac{25}{148} * 25 + \frac{50}{148} * 50 \approx 39.283$$

$$E(Y) = \frac{1}{4} * 40 + \frac{1}{4} * 33 + \frac{1}{4} * 25 + \frac{1}{4} * 50 = 37$$

4.40

$$P(X \geq 4) = P(X = 4) + P(X = 5)$$

$$= \binom{5}{4} \left(\frac{1}{3}\right)^4 \left(\frac{2}{3}\right) + \left(\frac{1}{3}\right)^5 \approx 0.0452$$

4.45 Let  $X$  be the number out of  $N$  examiners who pass him. Then if  $N = 3$ ,

$$P(X = 2) = \frac{2}{3} \binom{3}{2} (0.4)^2 (0.6) + \frac{1}{3} \binom{3}{2} (0.8)^2 (0.2) \approx 0.448;$$

and if  $N = 5$ ,

$$P(X = 3) = \frac{2}{3} \binom{5}{3} (0.4)^3 (0.6)^2 + \frac{1}{3} \binom{5}{3} (0.8)^3 (0.2)^2 \approx 0.290.$$

It's better for the student to request an examination with 3 examiners.

4.60 For a Poisson random variable with parameter  $\lambda$ ,

$$P(X = i, \lambda) = e^{-\lambda} \frac{\lambda^i}{i!}.$$

Therefore, for  $X = 2$ ,

$$P(X = 2) = \frac{0.75 * P(X = 2, \lambda = 3)}{0.75 * P(X = 2, \lambda = 3) + 0.25 * P(X = 2, \lambda = 5)}$$

$$P(X = 2) \approx 0.888.$$

### 3 Theoretical Exercises

Let  $X$  be a Poisson random variable with parameter  $\lambda$ . Then

$$P(X = i) = e^{-\lambda} \frac{\lambda^i}{i!}.$$

We want to first show that  $P$  increases monotonically as long as  $i \leq \lambda$ . Recall that  $P$  increases monotonically as long as  $P(X = i) \geq P(X = i - 1)$ , or in other words,

$$\frac{P(X = i)}{P(X = i - 1)} \geq 1$$

Therefore, if  $i \leq \lambda$ , then

$$\frac{P(X = i)}{P(X = i - 1)} = \left( e^{-\lambda} \frac{\lambda^i}{i!} \right) \left( e^{\lambda} \frac{(i - 1)!}{\lambda^{i-1}} \right) = \frac{\lambda}{i} \geq \frac{\lambda}{\lambda} = 1;$$

And hence  $P$  is monotonically increasing. Suppose instead that  $i > \lambda$ . Then

$$\frac{P(X = i)}{P(X = i - 1)} = \left( e^{-\lambda} \frac{\lambda^i}{i!} \right) \left( e^{\lambda} \frac{(i - 1)!}{\lambda^{i-1}} \right) = \frac{\lambda}{i} < \frac{\lambda}{\lambda} = 1;$$

And therefore  $P$  is monotonically decreasing.