

Markov Processes



- For the concepts like the LLN and the CLT, the independence of the sequence played an important role.
- The Markov processes consider the evolution and the steady-state behavior of a stochastic process with a simple dependence model
- They are widely used in diverse applications like population studies, queueing systems and the restoration of a degraded photograph.

Conditional Independence

The Markovian model is built around the notion of conditional independence of events.

Consider three events A, B and C in (S, \mathbb{F}, P) . The joint probability of A, B and C is given by

$$P(A \cap B \cap C) = P(A)P(B \cap C | A)$$

Applying the chain rule:

$$P(A \cap B \cap C) = P(A)P(B/A)P(C/A \cap B)$$

Conditional Independence....

Given A , the events B and C are called *conditionally independent* if

$$P(B \cap C / A) = P(B / A)P(C / A) \quad (1)$$

Or

$$P(C / A \cap B) = P(C / A) \quad (2)$$

(1) and (2) are equivalent, because

$$\begin{aligned} P(C / A \cap B) &= \frac{P(A \cap B \cap C)}{P(A \cap B)} \\ &= \frac{P(A)P(B \cap C / A)}{P(A \cap B)} \\ &= \frac{P(A)P(B / A)P(C / A)}{P(A)P(B / A)} = P(C / A) \end{aligned}$$

Thus, if B and C are conditionally independent given A , the joint probability is given by

$$P(A \cap B \cap C) = P(A)P(B|A)P(C|A)$$

Example- A box contains two coins: a fair coin and one fake two-headed coin ($P(H)=1$)

A coin is chosen at random and tossed twice. Define the following events.

A= First coin toss results in an H

B= Second coin toss results in an H

C= Fair coin has been selected.

Examine if A and B are conditionally independent given C and also if A and B are (unconditionally) independent.

Solution: We have $P(A|C)=P(B|C)=1/2$. Also, given that the fair coin is selected, we have $P(A \cap B|C)=1/2 \times 1/2=1/4$.

Thus A and B are conditionally independent given C

Example- Cntd..

To find $P(A)$, $P(B)$, and $P(A \cap B)$, we use the law of total probability:

$$P(A) = P(C)P(A / C) + P(C')P(A / C')$$

$$= \frac{1}{2} \times \frac{1}{2} + \frac{1}{2} \times 1 = \frac{3}{4}$$

Similarly, $P(B) = \frac{3}{4}$

To find $P(A \cap B)$, we have

$$\begin{aligned} P(A \cap B) &= P(C)P(A \cap B / C) + P(C')P(A \cap B / C') \\ &= P(C)P(A / C)(B / C) + P(C')P(A / C')(B / C') \\ &\quad \text{(by conditional independence of A and B)} \end{aligned}$$

$$= \frac{1}{2} \times \frac{1}{2} \times \frac{1}{2} + \frac{1}{2} \times 1 \times 1 = \frac{5}{8}$$

$$\therefore P(A \cap B) \neq P(A)P(B)$$

Conditionally Independent RVs

Given X , the RVs Y and Z are called *conditionally independent* if

$$F(y, z / X = x) = F(y / X = x) F(z / X = x)$$

Equivalently, if

$$F(y / Z = z, X = x) = F(y / X = x)$$

Markov process

A random process $\{X(t), t \in \Gamma\}$ defined on (S, \mathbb{F}, P) is called a Markov process if for any sequence of time $t_1 < t_2 < \dots < t_n < t_{n+1} \in \Gamma$,

$$\begin{aligned} &P(X(t_{n+1}) \leq x | X(t_1) = x_1, X(t_2) = x_2, \dots, X(t_n) = x_n) \\ &= P(\{X(t_{n+1}) \leq x | X(t_n) = x_n\}) \end{aligned}$$

For a Markov process, given $X(t_n)$, the random variable $X(t_{n+1})$ is conditionally independent of $X(t_1), X(t_2), \dots, X(t_{n-1})$

This property is known as the *Markovian property*.

Markov Chain

$X(t_n)$ takes values from a set V called the *state space*. The elements of V are called the *states* of the process $\{X(t_n)\}$.

Suppose V is countable. Since V has one-to-one correspondence with some subset of \mathbb{Z} , we can assume V as a set consisting of integers. Thus $\{X(t_n) = i\}$ means the event that $X(t_n)$ takes the i th state. For such a process, the Markovian property can be expressed in terms of the probability mass function and the process is called a *Markov chain* (MC). A Markov chain may be a *continuous-time Markov chain (CTMC)* or a *discrete-time Markov chain (DTMC)*.

CTMC

Suppose $\{X(t)\}$ takes values from a discrete state space $V = \{0, 1, 2, \dots\}$. Then $\{X(t)\}$ is called a CTMC if for any $n \geq 1$ and

$$t_1 < t_2 < \dots < t_n < t_{n+1} \in \Gamma,$$

$$P(X(t_{n+1})=j | X(t_0)=i_0, X(t_1)=i_1, \dots, X(t_n)=i)$$

$$= P(X(t_{n+1})=j | X(t_n)=i)$$

Example Independent increment process

For any $n > 1$ and $t_0 < t_1 < \dots < t_{n+1} \in \Gamma$, we have

$$\begin{aligned} &P(X(t_{n+1})=j / X(t_0)=i_0, X(t_1)=i_1, \dots, X(t_n)=i) \\ &= P(X(t_{n+1})-X(t_n)=j-i / X(t_0)=i_0, X(t_1)=i_1, \dots, X(t_n)=i) \\ &= P(X(t_{n+1})-X(t_n)=j-i) \quad (\text{Using the independent increment property}) \end{aligned}$$

Similarly,

$$\begin{aligned} &P(X(t_{n+1})=j / X(t_n)=i) \\ &= P((X(t_{n+1})-X(t_n))=j-i / X(t_n)=i) \\ &= P(X(t_{n+1})-X(t_n)=j-i) \\ &\therefore P(X(t_{n+1})=j / X(t_0)=i_0, X(t_1)=i_1, \dots, X(t_n)=i) = P(X(t_{n+1})=j / X(t_n)=i) \end{aligned}$$

Thus $\{X(t)\}$ is a CTMC.

DTMC

Consider a *discrete-time random process* $\{X_n, n \geq 0\}$ taking values from a countable set V . $\{X_n, n \geq 0\}$ is said to be a DTMC if

$$P(X_{n+1} = j | X_0 = i_0, X_1 = i_1, \dots, X_n = i) = P(X_{n+1} = j | X_n = i).$$

Example

Suppose $\{X_n, n \geq 0\}$ is a sequence of iid and integer-valued random variables. Then

$$\begin{aligned} P(X_{n+1} = j | X_0 = i_0, X_1 = i_1, \dots, X_n = i) &= P(X_{n+1} = j) \\ &= P(X_{n+1} = j / X_n = i) \end{aligned}$$

Therefore, a sequence of integer –valued iid random variables is trivially a DTMC.

Example

Suppose $\{Z_n, n \geq 0\}$ is a sequence of iid and integer-valued random variables and $X_n = \sum_{i=0}^n Z_i$. Then

$\{X_n, n \geq 0\}$ is an MC.

Solution:

We have
$$X_{n+1} = \sum_{i=0}^{n+1} Z_i = \sum_{i=0}^n Z_i + Z_{n+1} = X_n + Z_{n+1}$$

$$\begin{aligned} \therefore P(X_{n+1} = j \mid X_n = i) &= \frac{P(X_{n+1} = j, X_n = i)}{P(X_n = i)} \\ &= \frac{P(Z_{n+1} = j - i, X_n = i)}{P(X_n = i)} \\ &= \frac{P(Z_{n+1} = j - i)P(X_n = i)}{P(X_n = i)} \\ &= P(Z_{n+1} = j - i) \end{aligned}$$

Arguing in the similar manner, we can show that $P(X_{n+1} = j \mid X_0 = i_0, X_1 = i_1, \dots, X_n = i) = P(Z_n = j - i)$

$$\therefore P(X_{n+1} = j \mid X_0 = i_0, X_1 = i_1, \dots, X_n = i) = P(X_{n+1} = j \mid X_n = i)$$

$\therefore \{X_n, n \geq 0\}$ is an MC.

Example contd..

When Z_n takes values from $\{-1, 1\}$, then $\left\{X_n = \sum_{i=0}^n Z_i, n \geq 0\right\}$ is the simple random walk process. Thus the simple random walk process is an MC.

To Summarise

- Conditional Independence

$$P(B \cap C / A) = P(B / A)P(C / A)$$

- Markov Process

For a Markov process $\{X(t), t \in \Gamma\}$

$$\begin{aligned} &P(X(t_{n+1}) \leq x | X(t_1) = x_1, X(t_2) = x_2, \dots, X(t_n) = x_n) \\ &= P(\{X(t_{n+1}) \leq x | X(t_n) = x_n\}) \end{aligned}$$

- $X(t)$ takes values from a state space V whose elements are called states.

To Summarise..

➤ CTMC

For a CTMC $\{X(t), t \geq 0\}$

$$P(X(t_{n+1})=j | X(t_0)=i_0, X(t_1)=i_1, \dots, X(t_n)=i) \\ = P(X(t_{n+1})=j | X(t_n)=i)$$

➤ DTMC

For a DTMC $\{X_n, n \geq 0\}$ taking values from a countable set V .

$$P(X_{n+1}=j | X_0=i_0, X_1=i_1, \dots, X_n=i) = P(X_{n+1}=j | X_n=i)$$

Thank You