## A Appendix: simulation code and model files

## A.1 R code to simulate data

```
2 ########## function to simulate survival time ################
4 # survival function is given by: S(t) = \exp(-\exp(B) \star (\exp(A \star t) - 1)) / A
_5 sim_Ti = function(n=250, alpha1, alpha2, delta=c(1,1), gamma=c(1,1)){
  # random effects
   Sigma=matrix(c(0.09, 0.09*0.16, 0.09*0.16, 0.09), nrow=2, byrow=T)
   u=rmvnorm(n,c(0,0),Sigma)
   H = matrix(rnorm(2*n), n, 2)
   W = matrix(rnorm(2*n), n, 2)
   # define the survival function
   surv = function(t) {
    if(alpha1==0 \& alpha2==0) \{res=t*exp(W%*%gamma)\}
14
    else{
      res = ((exp(alpha2*u[,1]+alpha2*t*u[,2]+alpha1*delta[1]*H[,1]+alpha1*
    delta[2]*H[,2]*t+W%*%gamma)-exp(alpha1*H[,1]*delta[1]+W%*%gamma+alpha2*u
    [,1]))/(alpha2*u[,2]+alpha1*H[,2]*delta[2]))
    return(exp(-res))
   }
20
  rnd =runif(n)
  Ti = rep(Inf,n)
  \# w = which(surv(1e8)-rnd < 0)
   w = which (1-surv(101)-rnd > 0)
   Ti[w] = sapply(w, function(j) uniroot(function(x) 1-surv(x)[j]-rnd[j],lower
    =0, upper=101) $root)
   Ti[Ti==0] = min(Ti[Ti>0])/2
   Ci = 5*rbeta(n,4,1)
   Delta = as.numeric(Ti < Ci)</pre>
   Ti2 = pmin(Ti, Ci)
   list(Ti=Ti2, event=Delta, H=H, W=W, U=u)
38 sim_longitudinal_data = function(survival_data=surdata, n=250, time=c(0, 0.25,
     0.5, 0.75, 1, 3), tau, sigma=1, beta=c(1,1), delta=c(1,1)) {
   # survival_data - data simulated from survival model
  # n - # of subjects
  # time - time points of observations
   # tau - quantile
  # sigma - scale parameter
```

```
44
   # random generation of ALD(0, sigma, p)
45
   rald = function(n, location=0, scale, p) {
46
     u = rnorm(n)
     z = rexp(n)
48
     v = scale*z
     theta = (1-2*p)/(p*(1-p))
50
     tau = sqrt(2/(p*(1-p)))
     sample = theta * z + tau * sqrt(z) * u
52
54
   time = time # at most # = length(time) observations per patient
   y = matrix(NA, nrow=n, ncol=length(time)) # wide format
   Ti = survival_data$Ti
   U = survival_data$U # random effects
   H = survival_data$H
   X = cbind(1, rnorm(n))
   count = sapply(Ti, function(x) sum(x > time)) # number of observations after
     drop-outs
62
   for (i in 1:n) {
63
     for (j in 1:count[i]) {
64
       location = beta %*% X[i, ] + delta %*% c(H[i,1], H[i,2]*time[j]) + U[i,]
      %*% c(1, time[j])
       y[i,j] = location + rald(1, scale=sigma, p=tau)
67
   }
   list(y = y, X = X, J=count)
71 }
74 ##### function to simulate multiple joint data sets #########
76 sim_multiple_data = function(N, sur_fun=sim_Ti, longi_fun=sim_longitudinal_
    data, alpha, tau) {
   # N - number of data sets to generate
   # sur_fun - function to simulate survival data
   # longi_fun - function to sumualte longitudinal data
   # alpha - association mechanism for JM
   # tau - quantile
   outdata = vector(mode='list', N)
83
   for (i in 1:N) {
     sur_data = sur_fun(alpha1=alpha[1], alpha2=alpha[2])
     longi_data = longi_fun(sur_data, tau=tau)
     outdata[[i]] = list(survival_data=sur_data, longitudinal_data=longi_data)
87
   outdata
90 }
```

## A.2 JAGS model file

```
1 model {
    zero[1] <- 0
    zero[2] \leftarrow 0
    k1 < (1-2*qt)/(qt*(1-qt))
    k2 < -2/(qt*(1-qt))
    for (i in 1:I) {
    # prior for random effects
    u[i, 1:2] ~ dmnorm(zero[], precision[,])
10
    # longitudinal process, BQR mixed model using ALD representation
11
12
      for (j in 1:J[i]) {
        er[i,j] ~ dexp(sigma)
        mu[i,j] <- u[i,1] + u[i,2]*t[j] + inprod(X[i,], beta[]) + delta[1]*H[i]
14
      ,1] + delta[2]*H[i,2]*t[j] + k1*er[i,j]
        prec[i,j] \leftarrow sigma/(k2*er[i,j])
15
        y[i,j] \sim dnorm(mu[i,j], prec[i,j])
      } #end of j loop
17
18
    # survival process, baseline hazard is set to 1
19
    A[i] \leftarrow alpha2*u[i,2] + alpha1*delta[2]*H[i,2]
    B[i] \leftarrow alpha1*delta[1]*H[i,1] + alpha2*u[i,1] + inprod(gamma, W[i,])
21
    S[i] \leftarrow \exp(-c \exp(B[i]) * (pow(exp(A[i]), Ti[i]) - 1)/A[i])
22
    h[i] <- c*exp(inprod(gamma, W[i,]) + alpha1*(delta[1]*H[i,1] + delta[2]*H[i
     ,2]*Ti[i]) + alpha2*(u[i,1] + u[i,2]*Ti[i]))
    L[i] \leftarrow pow(h[i], event[i])*S[i]/1.0E+08
24
25
26
    # zero trick
    phi[i] \leftarrow -log(L[i])
27
    zeros[i] ~ dpois(phi[i])
29
    } #end of i loop
31
32
    precision[1:2,1:2] <- inverse(Sigma[,])</pre>
      Sigma[1,1] <- 1
33
      Sigma[1,2] \leftarrow rho*sig1
34
      Sigma[2,1] \leftarrow Sigma[1,2]
35
      Sigma[2,2] \leftarrow sig1*sig1
36
    # priors for other parameters
38
    alpha1 ~ dnorm(0, 0.001)
39
    alpha2 ~ dnorm(0, 0.001)
40
    beta[1] ~ dnorm(0, 0.001)
    beta[2] ~ dnorm(0, 0.001)
42
    delta[1] ~ dnorm(0, 0.001)
43
    delta[2] ~ dnorm(0, 0.001)
44
    gamma[1] ~ dnorm(0, 0.001)
    gamma[2] ~ dnorm(0, 0.001)
    sigma ~ dgamma(0.001, 0.001)
```

```
rho ~ dunif(-1, 1)
sig1 ~ dgamma(0.01, 0.01)
c ~ dunif(0.01, 10)
for c ~ dunif(0.01, 10)
```