

Gaussian Basics

$$\begin{aligned} p(x_i | \mu, \sigma) &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x_i - \mu)^2} \\ &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{(x_i - \mu)^2}{E[(x - \mu)^2]}} \\ &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{\Delta x_i^2}{E[\Delta x^2]}} \\ &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{(x_i - E[x])^2}{E[(x - E(x))^2]}} \end{aligned}$$

$$\begin{aligned} \sigma^2 &= E[(x - \mu)^2] \\ &= \int (x - \mu)^2 p(x) dx \end{aligned}$$

$$\mu = E[x]$$

Ratio of
one squared deviation
to mean squared
deviation

→ Relative squared
deviation from the
mean squared deviation
aka. variance

Likelihood of the Linear Model

$$P(\underline{y} | \underline{X}, \mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2} (y_i - x_i \mu)^2}$$
$$= (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - x_i \mu)^2}$$

Independence
assumption

$$P(\underline{y} | \underline{X}, \underline{\mu}, \sigma^2) = \left(\frac{1}{\sqrt{2\pi} (\sigma^2 \mathbf{I})^{\frac{n}{2}}} \right) e^{-\frac{1}{2} (\underline{y} - \underline{X}\underline{\mu})^T (\sigma^2 \mathbf{I})^{-1} (\underline{y} - \underline{X}\underline{\mu})}$$

Prior

$$P(\underline{\mu}) = (2\pi\sigma_0^2)^{-\frac{1}{2}} e^{-\frac{1}{2} (\underline{\mu} - \underline{\mu}_0)^T (\sigma_0^2 \mathbf{I})^{-1} (\underline{\mu} - \underline{\mu}_0)}$$

$$P(\underline{\theta}) = (2\pi V_0)^{-\frac{1}{2}} e^{-\frac{1}{2} (\underline{\theta} - \underline{\theta}_0)^T V_0^{-1} (\underline{\theta} - \underline{\theta}_0)}$$

$$\boxed{\underline{\mu} = \underline{\theta}}$$

Posterior of $\underline{\theta}$

$$P(\underline{\theta} | \underline{y}, \underline{X}, \underline{\sigma}) \propto$$

$$P(\underline{y} | \underline{X}, \underline{\theta}, \underline{\sigma}) \cdot P(\underline{\theta})$$

$$\propto e^{-\frac{1}{2} (\underline{y} - \underline{X}\underline{\theta})^T (\sigma^2 \mathbf{I})^{-1} (\underline{y} - \underline{X}\underline{\theta})} \cdot e^{-\frac{1}{2} (\underline{\theta} - \underline{\theta}_0)^T V^{-1} (\underline{\theta} - \underline{\theta}_0)}$$

$$p(\underline{\theta} | \underline{y}, \underline{x}, \underline{\sigma}) \propto p(\underline{y} | \underline{x}, \underline{\theta}, \underline{\sigma}) \cdot p(\underline{\theta})$$

$$\propto e^{-\frac{1}{2}(\underline{y} - \underline{x}\underline{\theta})^T (\underline{\sigma}^2 \underline{I})^{-1} (\underline{y} - \underline{x}\underline{\theta})} \cdot e^{-\frac{1}{2}(\underline{\theta} - \underline{\theta}_0)^T \underline{V}_0^{-1} (\underline{\theta} - \underline{\theta}_0)}$$

$$\propto e^{-\frac{1}{2} [(\underline{y} - \underline{x}\underline{\theta})^T (\underline{\sigma}^2 \underline{I})^{-1} (\underline{y} - \underline{x}\underline{\theta}) + (\underline{\theta} - \underline{\theta}_0)^T \underline{V}_0^{-1} (\underline{\theta} - \underline{\theta}_0)]}$$

$$\propto e^{-\frac{1}{2} \left[\underbrace{\underline{y}^T (\underline{\sigma}^2 \underline{I})^{-1} \underline{y}}_{\text{const}} - \underbrace{2 \underline{y}^T (\underline{\sigma}^2 \underline{I})^{-1} \underline{x} \underline{\theta}}_{\theta} + \underbrace{\underline{\theta}^T \underline{x}^T (\underline{\sigma}^2 \underline{I})^{-1} \underline{x} \underline{\theta} + \underline{\theta}^T \underline{V}_0^{-1} \underline{\theta}}_{\theta - \text{squared}} - \underbrace{2 \underline{\theta}_0^T \underline{V}_0^{-1} \underline{\theta}}_{\theta} + \underbrace{\underline{\theta}_0^T \underline{V}_0^{-1} \underline{\theta}_0}_{\text{const}} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\text{const} + \underbrace{\underline{\theta}^T (\underline{x}^T (\underline{\sigma}^2 \underline{I})^{-1} \underline{x} + \underline{V}_0^{-1}) \underline{\theta}}_{\underline{V}_n^{-1}} - 2 (\underline{y}^T (\underline{\sigma}^2 \underline{I})^{-1} \underline{x} + \underline{\theta}_0^T \underline{V}_0^{-1}) \underline{\theta} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\text{const} + \underline{\theta}^T \underline{V}_n^{-1} \underline{\theta} - 2 \left(\frac{\underline{y}^T \underline{x}}{\underline{\sigma}^2} + \underline{\theta}_0^T \underline{V}_0^{-1} \right) \underline{\theta} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\text{const}_2 + \underline{\theta}^T \underline{V}_n^{-1} \underline{\theta} - 2 \underline{\theta}_n^T \underline{V}_n^{-1} \underline{\theta} + 2 \underline{\theta}_n^T \underline{V}_n^{-1} \underline{\theta} + \underline{\theta}_n^T \underline{V}_n^{-1} \underline{\theta}_n - 2 \left(\frac{\underline{y}^T \underline{x}}{\underline{\sigma}^2} + \underline{\theta}_0^T \underline{V}_0^{-1} \right) \underline{\theta} \right]}$$

$$\propto e^{-\frac{1}{2} \left[\text{const}_1 + (\underline{\theta} - \underline{\theta}_n)^T \underline{V}_n^{-1} (\underline{\theta} - \underline{\theta}_n) + 2 \left(\underline{\theta}_n^T \underline{V}_n^{-1} - \frac{\underline{y}^T \underline{x}}{\underline{\sigma}^2} - \underline{\theta}_0^T \underline{V}_0^{-1} \right) \underline{\theta} \right]}$$

$$\propto e^{-\frac{1}{2} [(\underline{\theta} - \underline{\theta}_n)^T \underline{V}_n^{-1} (\underline{\theta} - \underline{\theta}_n)]}$$

= 0 when
 $\underline{\theta}_n = \underline{V}_n [\underline{V}_0^{-1} \underline{\theta}_0 + \underline{x}^T \underline{y} / \underline{\sigma}^2]$