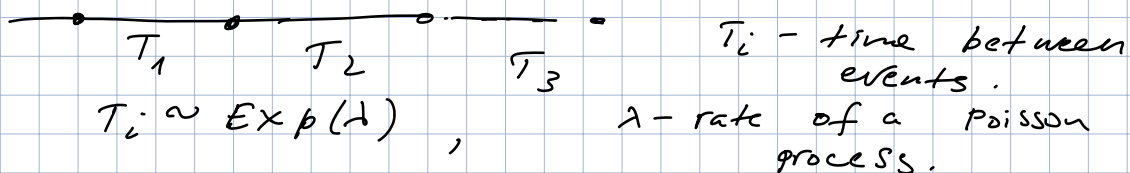


Renewal theory and its applications.

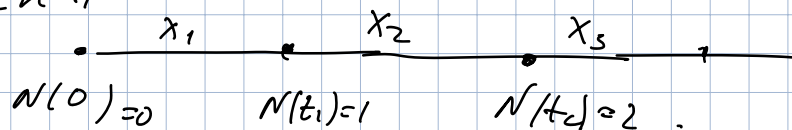
$\{N(t), t \geq 0\}$ - Poisson process.



Generalization.

Let $\{N(t), t \geq 0\}$ be a counting process.

Let X_n be the time between $(n-1)$ and n^{th} event.



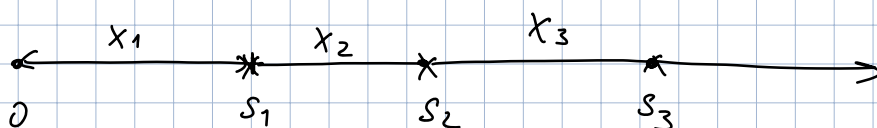
Def. If the sequence of nonnegative r.v. $\{X_1, X_2, \dots\}$ is iid then the counting process $\{N(t), t \geq 0\}$ is said to be a renewal process.

Let F be a distribution of X_i

Ex. We have an infinite supply of lightbulbs.

We use a single lightbulb at a time, and when it fails we immediately replace it with a new one.

$\{N(t), t \geq 0\}$ is a renewal process, $N(t)$ - # of lightbulbs had fail by time t .



Let x_1, x_2, \dots be interarrival times

$$s_0 = 0 \quad s_n = \sum_{i=1}^n x_i$$

$$s_1 = x_1, \quad s_2 = x_1 + x_2, \quad \dots$$

s_n - the time of n^{th} renewal.

For the distribution F we assume:

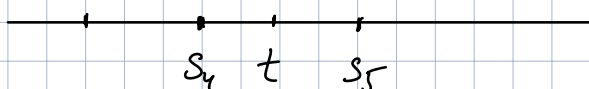
$$F(0) = P(X_n = 0) < 1$$

$$\mu = EX_n \quad n \geq 1 \quad \text{mean time between successful renewals}$$

X_n is not identically 0 $\mu > 0$.

$N(t)$ be ∞ for a finite t ?

$$N(t) = \max \{n : s_n \leq t\}$$



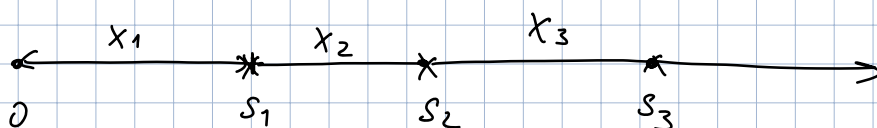
$$\text{Let } s_4 \leq t \quad s_5 > t$$

Hence, 4^{th} renewal had occurred by time t
 5^{th} \rightarrow after time t

By strong law of large numbers

$$\frac{s_n}{n} \rightarrow \mu \quad n \rightarrow \infty \quad \text{with probability 1.}$$

$$\text{Since } \mu > 0 \Rightarrow s_n \rightarrow \infty \quad n \rightarrow \infty$$



Let x_1, x_2, \dots be interarrival times

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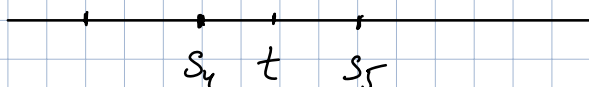
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$$\text{Since } \mu > 0 \Rightarrow s_n \rightarrow \infty \quad n \rightarrow \infty$$

Thus S_n can be $\leq t$ for at most a finite number of values of $n \Rightarrow$

$N(t) = \max \{n : S_n \leq t\} < \infty \quad \forall t$ with probability 1.

$$N(\infty) = \lim_{t \rightarrow \infty} N(t) = \infty$$

This follows since the only way in which $N(\infty)$, the total # of renewal that occur, can be finite is for one of the interarrival times to be infinite.

$$\begin{aligned} \Rightarrow P\{N(\infty) < \infty\} &= P(X_n = \infty \text{ for some } n) = \\ &= P\left(\bigcup_{n=1}^{\infty} X_n = \infty\right) \leq \sum_{n=1}^{\infty} P(X_n = \infty) = 0. \end{aligned}$$

Convergence with probability 1
(almost surely).

Consider a sequence of r.v.s X_1, X_2, \dots
all defined on the same probability space Σ
For every $s \in \Sigma$ we obtain a sample
sequence (sequence of numbers)
 $X_1(s), \dots, X_n(s), \dots$

A sequence X_1, X_2, X_3, \dots of r.v.s is said to
converge to a r.v. X with probability 1
(almost surely) if

$$P \left\{ \omega: \lim_{n \rightarrow \infty} X_n(s) = X(s) \right\} = 1$$

This means that the set of sample paths
that converge to $X(s)$ in sense of
sequence convergence has probability 1.

Equivalently X_1, \dots, X_n, \dots converges
with probability 1 if for $\forall \epsilon > 0$

$$\lim_{m \rightarrow \infty} P \left\{ |X_n - X| < \epsilon \quad \forall n \geq m \right\} = 1$$

(Strong Law of Large numbers)

Weak law of large numbers.

Convergence in probability.

A sequence of r.v. $X_1 \dots X_n \dots$ converges to

r.v. X in probability if $\forall \epsilon > 0$

$$\lim_{n \rightarrow \infty} P\{|X_n - X| < \epsilon\} = 1.$$

converg. with p. 1 \Rightarrow convergence
in probability.

\Leftarrow