

# Bayesian Inference for Brain Activity from Dual-Resolution Functional Magnetic Resonance Imaging

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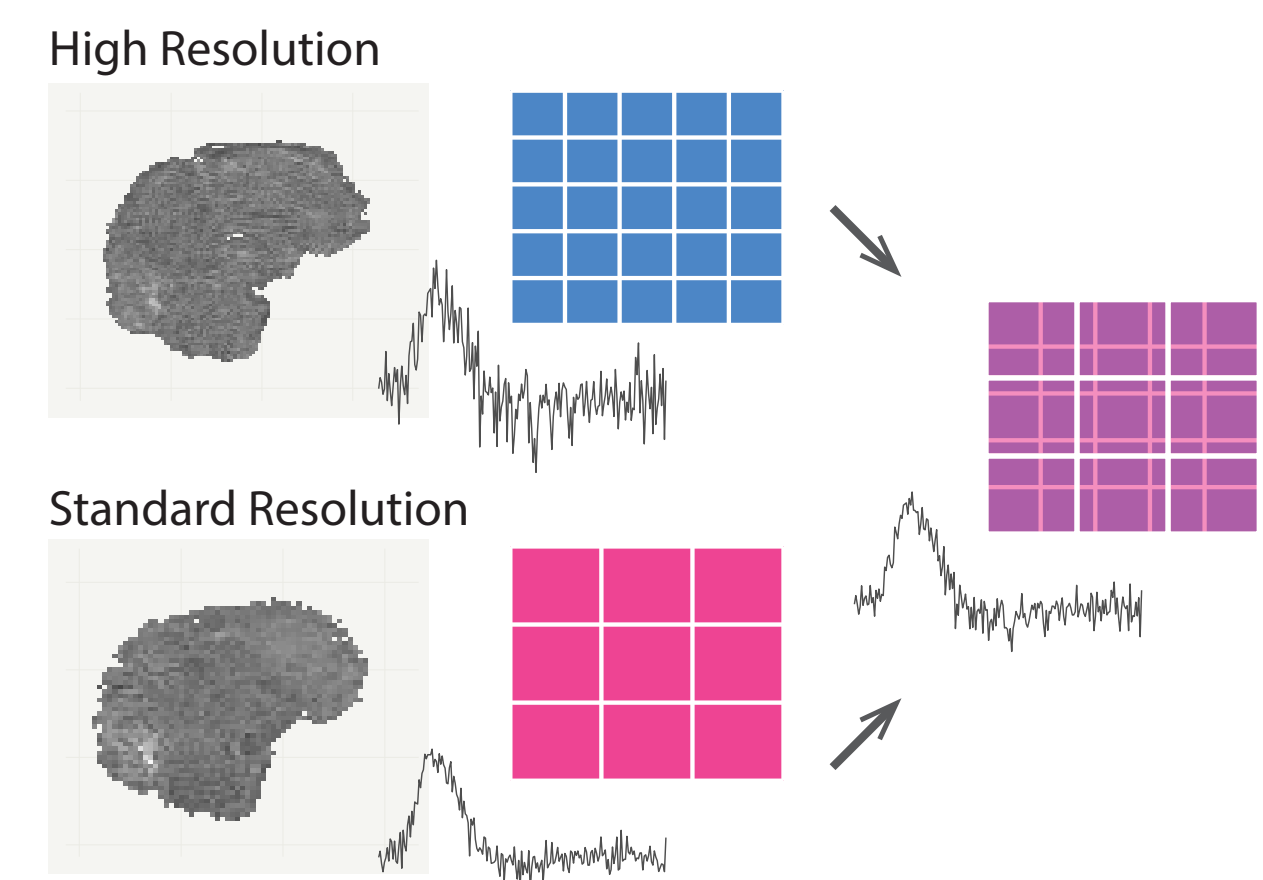


## Introduction

- **Personalized medicine:** need for patient-specific functional mapping. Surgeons want to minimize damage to healthy, functional tissue
- Variations in functional neuroanatomy [e.g. 1]
- MR imaging methods used to aid presurgical planning and neuronavigation

## Why multiple resolutions?

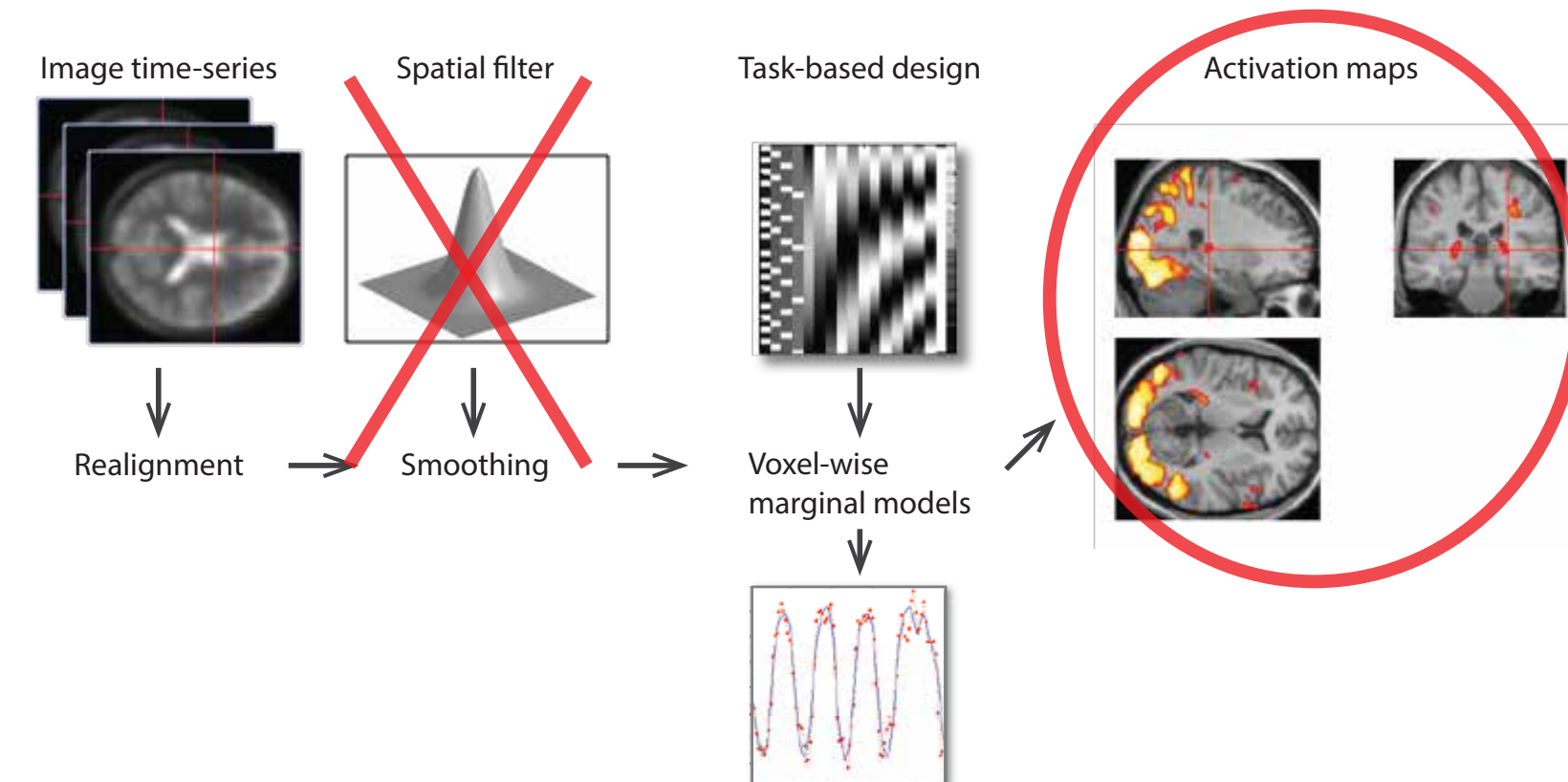
- Surgery requires high spatial precision
- Signal-to-noise ratio (SNR) decreases as spatial resolution increases



**Fig 1:** Unique blessing/curse of data collected at “high” and “standard” resolutions ( $1.8 \times 1.8 \times 2.3 \text{ mm}^3$  and  $3 \times 3 \times 3.45 \text{ mm}^3$  voxels; voxel-volumetric pixel)

**Goal:** to leverage spatial precision from high resolution data and SNR from standard resolution data

- Inference at high resolution voxel locations
- Utilize massive spatial information to conduct within-patient functional mapping
- Fully Bayesian inference in data of this size can be computationally prohibitive



**Fig 2:** Typical fMRI pipeline. We received and used unsmoothed  $z$ -statistic maps as data, treating these as noisy summaries of spatial activation. Adapted from: <https://www.fil.ion.ucl.ac.uk/spm/course/slides11/>

## Proposed Model for Dual-Resolution fMRI

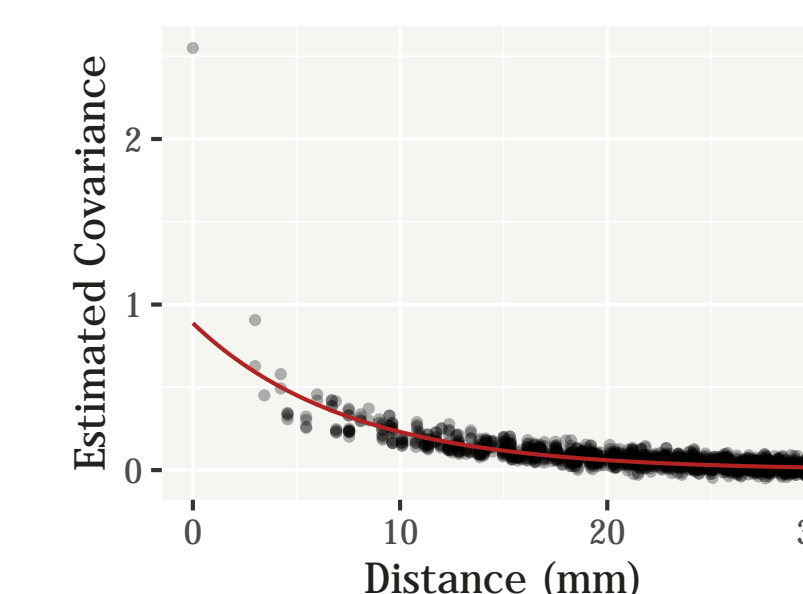
$$y_h(\mathbf{v}_h) \sim \mathcal{N}(\mu(\mathbf{v}_h), \sigma_h^2) \quad y_s(\mathbf{v}_s) \sim \mathcal{N}(\mu(\mathbf{v}_s), \sigma_s^2)$$

$$\pi(\sigma_h^2, \sigma_s^2) \propto \sigma_h^{-2} \sigma_s^{-2} \mathbf{1}(\sigma_h^{-2} < \sigma_s^{-2})$$

( $h$ —high and  $s$ —standard resolution)

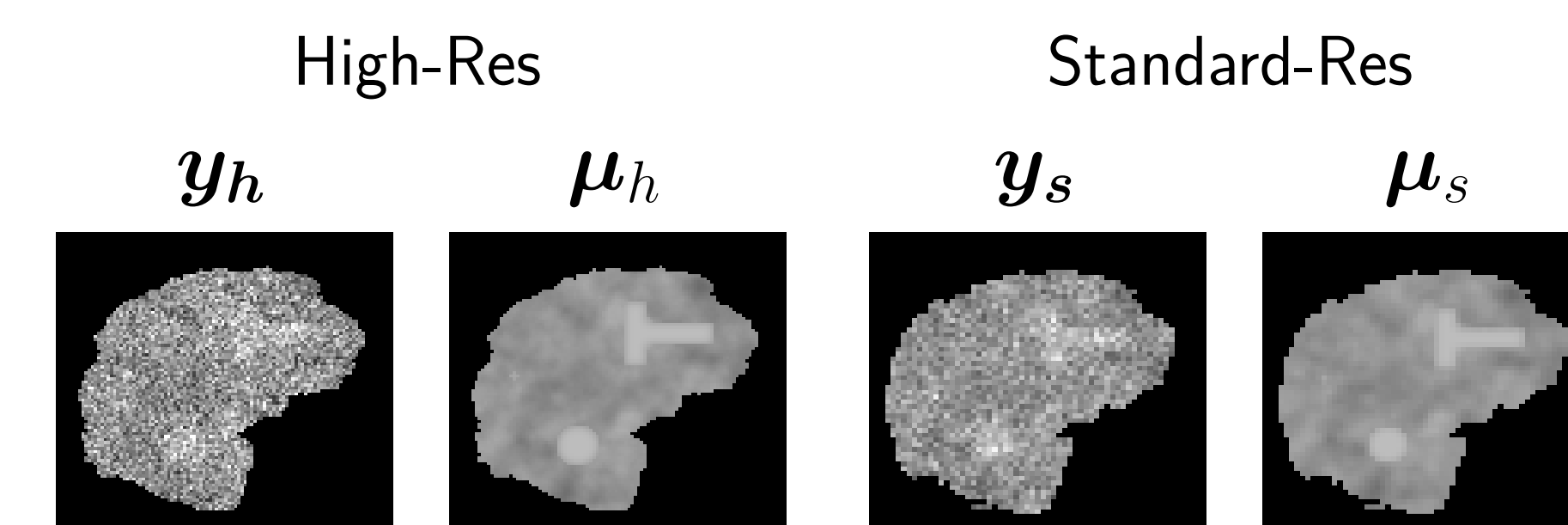
$$\mu(\mathbf{v}) \sim \mathcal{GP}(0, K) \quad K(\mathbf{v}, \mathbf{v}') = \tau^2 \exp(-\psi \|\mathbf{v} - \mathbf{v}'\|^\nu), \quad \tau^2, \psi > 0; \nu \in (0, 2]$$

- Correlation between voxels and images induced by Gaussian Process prior on  $\mu(\cdot)$
- Covariance parameters estimated via minimum contrast

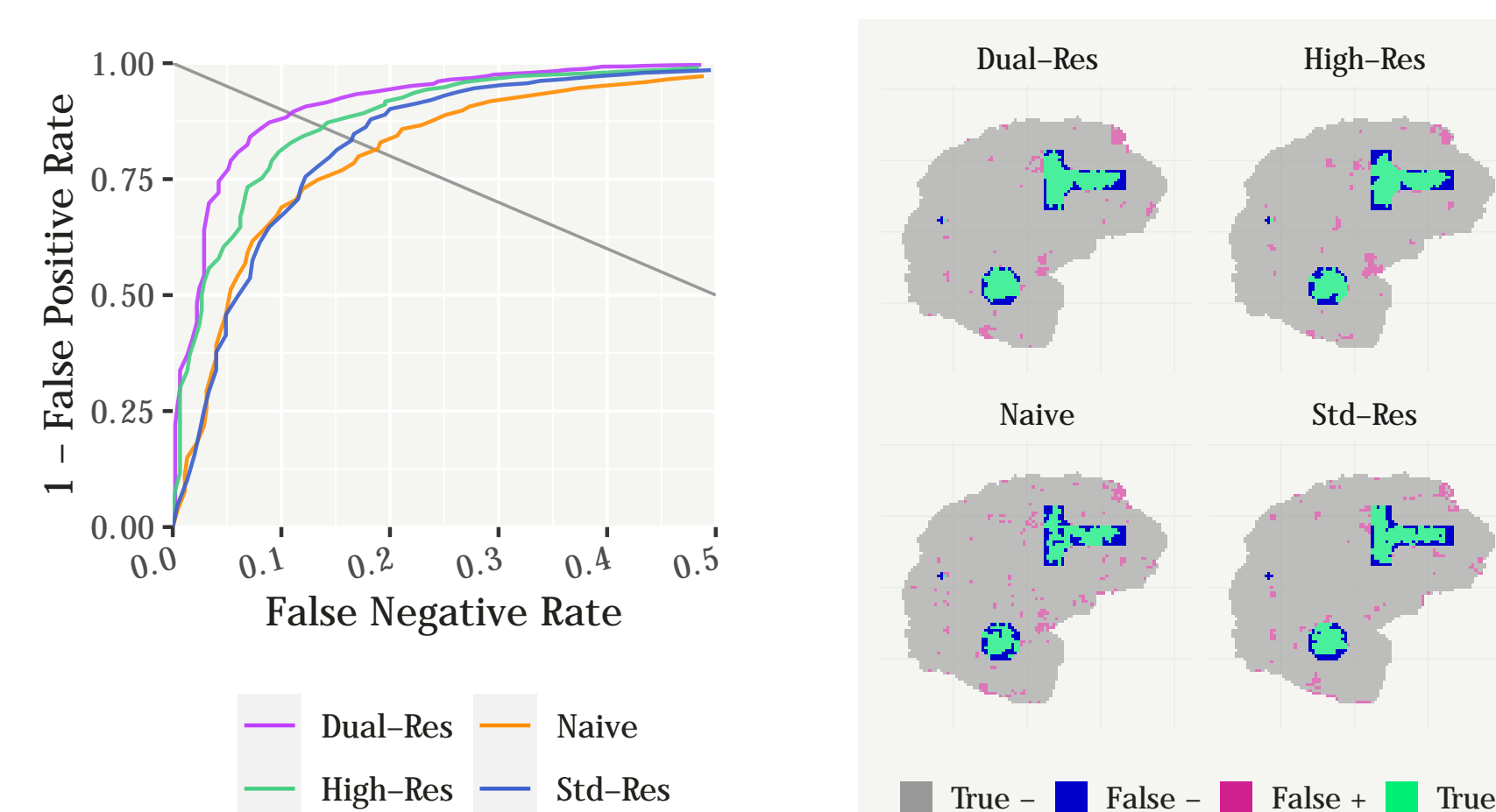


**Fig 3:** Empirical covariogram and fitted curve with exponential correlation function

## Illustrative Example: Embedded Signal in 2D Slices

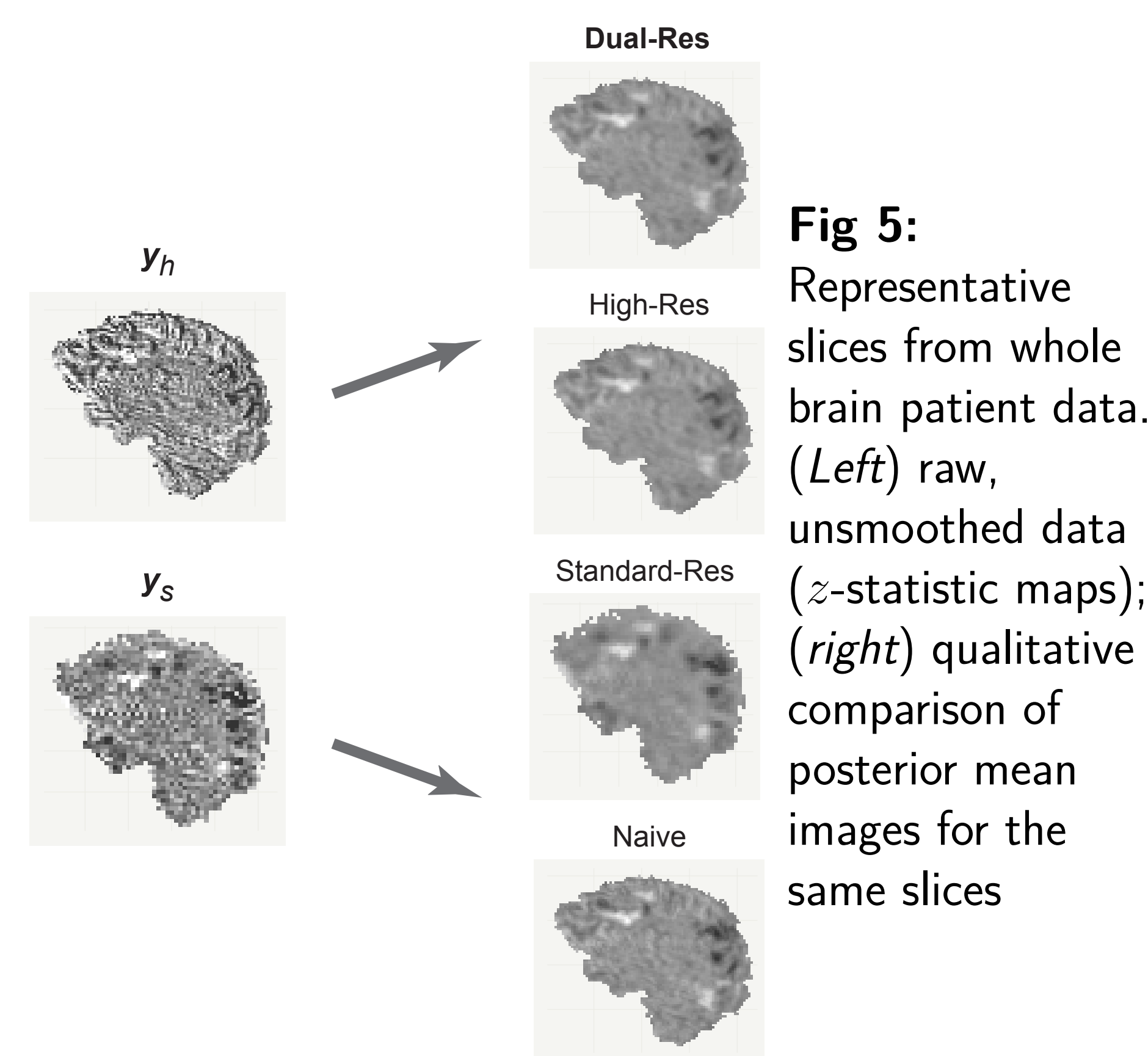


- Compare with single resolution methods
- **Naive** data combination:  $\bar{\mathbf{y}}_{hs} = (\mathbf{y}_h + \mathbf{W}^\top \mathbf{y}_s)/2$   
(Where  $\mathbf{W}^\top \mathbf{y}_s$  is an ordinary kriging interpolation of  $\mathbf{y}_s$ )

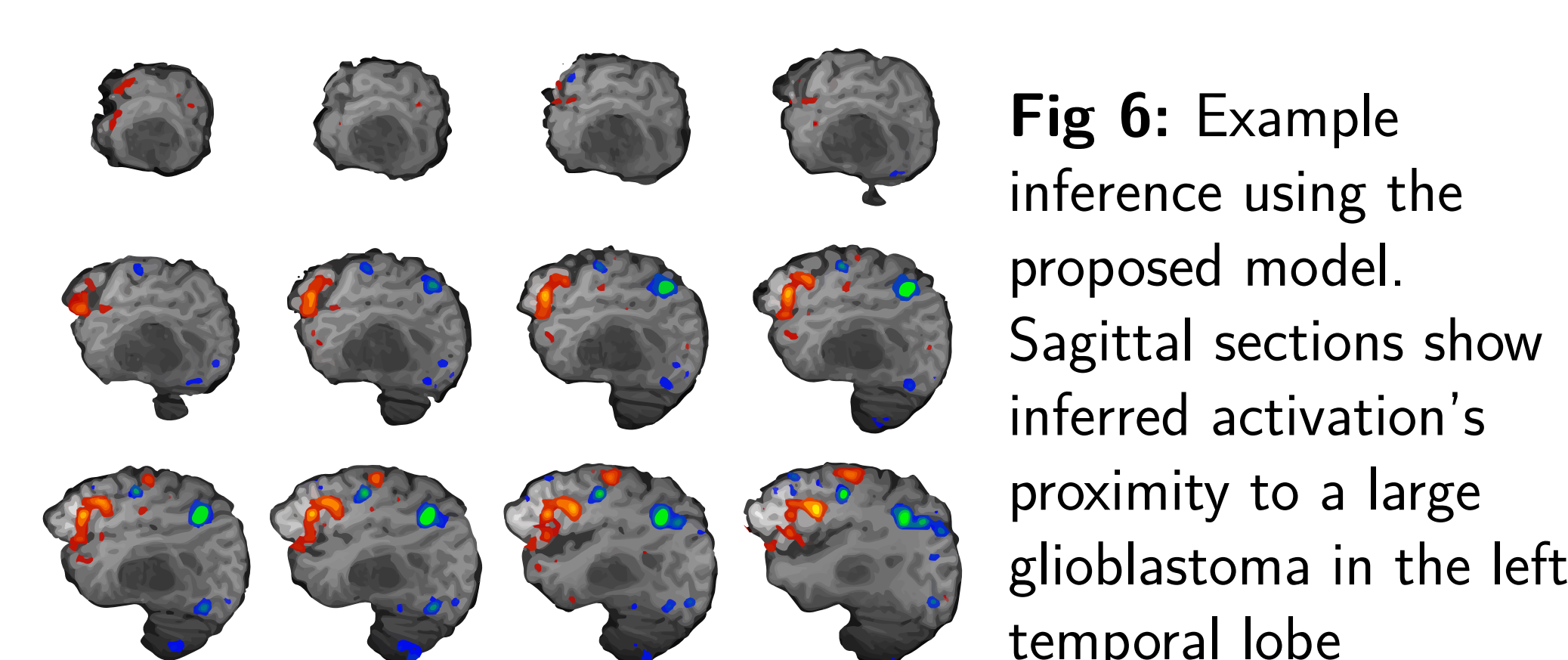


**Fig 4:** (Left) ROC curves show proposed method's improvement over single resolution methods. (Right) Panels show example inferential errors

## Patient Data Analysis



**Fig 5:** Representative slices from whole brain patient data. (Left) raw, unsmoothed data ( $z$ -statistic maps); (right) qualitative comparison of posterior mean images for the same slices



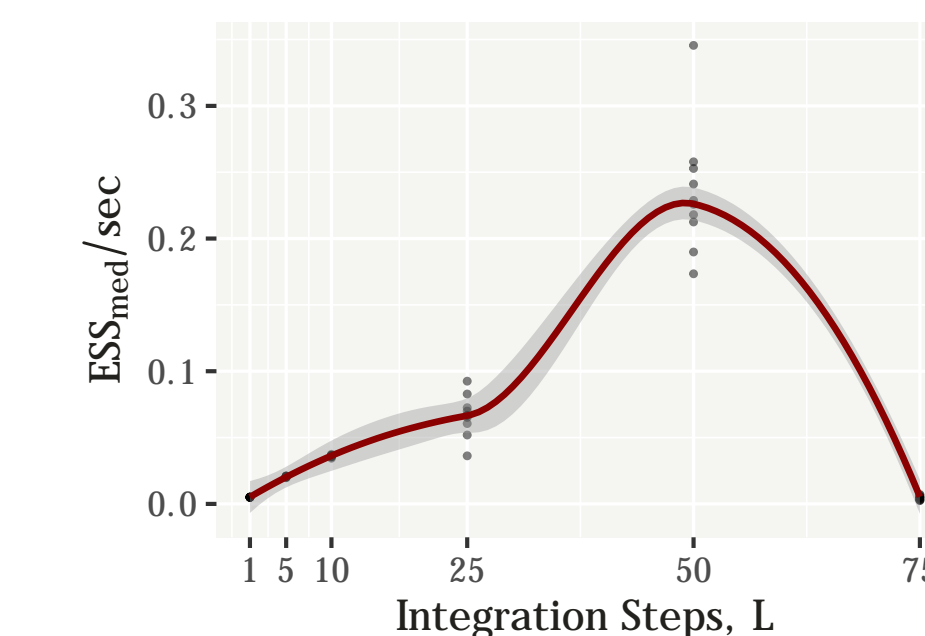
**Fig 6:** Example inference using the proposed model. Sagittal sections show inferred activation's proximity to a large glioblastoma in the left temporal lobe

## Scanning Session Description

- Patient scanned while performing a word reading task (30 sec on/off block design)
- Two resolutions collected over separate sessions  
(High-Res:  $120 \times 120 \times 62$  grid size; Std-Res:  $64 \times 64 \times 48$  grid size)
- Siemens 3 T scanner; 32 channel head coil

## Posterior Inference

- Sampling via Riemann manifold HMC [2]
- Sparse approximation with high resolution voxel locations used as the “inducing set” to maximize spatial precision
- Computational burden alleviated with circulant matrix embedding [5, 6]



**Fig 7:** Computation time in analysis of real patient data.  $ESS_{med}$  denotes median effective sample size across all voxels

- Differently penalize false negative and false positive errors. Adapted from [4, 3]

$$L(\mathbf{m}, \mathbf{d}) = \sum_i \frac{-f(m_i)d_i - [1 - f(m_i)](1 - d_i)}{\text{Gains for correct decisions} + k_1 f(m_i)(1 - d_i) + k_2 [1 - f(m_i)]d_i + td_i}$$

Penalties for false -/+

- $k_1, k_2, t$  tunable parameters, e.g.  $t$  penalizes the overall number of discoveries
- $m_i = |\mathbb{E}(\mu_{hi}|\mathbf{y}_h, \mathbf{y}_s)| / \sqrt{\text{var}(\mu_{hi}|\mathbf{y}_h, \mathbf{y}_s)}$
- $f(\cdot)$  some monotone function of signal strength: e.g.  $f(m_i) = m_i / \max_i m_i$
- Proxy measure for  $\pi(\text{“voxel } i \text{ is active”}|\mathbf{y}_h, \mathbf{y}_s)$
- Collaborating radiologist's advice: penalize false negatives  $11 \times$  more heavily than false positives

## References

- [1] Belliveau, JW et al. *Science* 254.5032 (1991), pp. 716–719.
- [2] Girolami, Mark and Calderhead, Ben. *JRSS B* 73.2 (2011), pp. 123–214.
- [3] Liu, Zhuqing et al. *Bayesian Analysis* 11.2 (2016), p. 599.
- [4] Muller, Peter, Parmigiani, Giovanni, and Rice, Kenneth. (2006).
- [5] Rue, Havard and Held, Leonhard. *Chapman and Hall/CRC*, 2005.
- [6] Wood, Andrew TA and Chan, Grace. *J Comp Graph Stat* 3.4 (1994), pp. 409–432.

## Contact & Software

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