

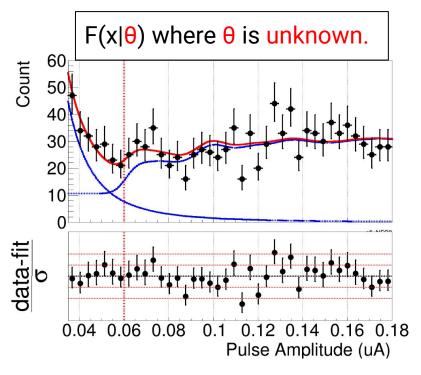
Statistics for CAKE

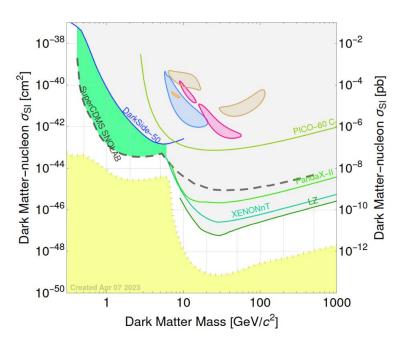
Ata Sattari



Why are we here?

- 1. Discuss how to fit a model to data.
- 2. Learn how to perform statistical tests. (Upper limits,)



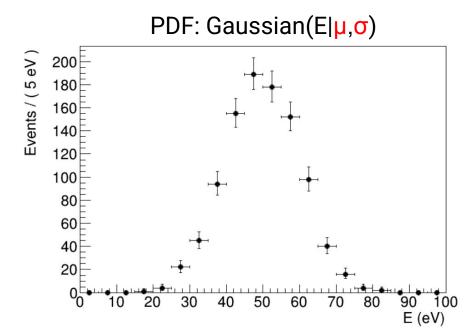






General statistics

What is the probability to observe this data?



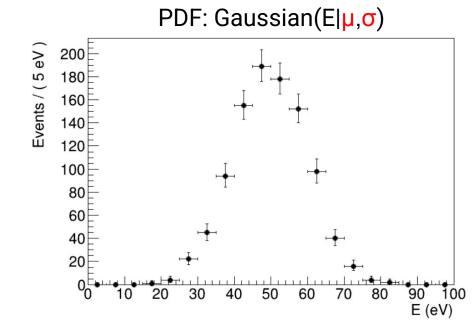
What is the probability to observe this data?

Multiply probabilities to observe events:

$$L(data|\vec{\theta}) = \prod_{Events} PDF(E_i|\vec{\theta}) dE$$

Shape only fit

(Unbinned likelihood)



What is the probability to observe this data?

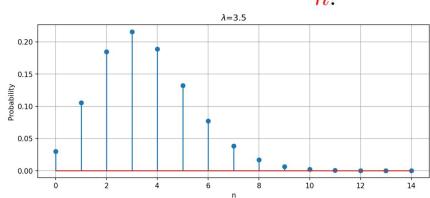
PDE: Cauccian(E|µ,o)

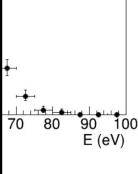
$$L(data|\vec{\theta}) =$$

Multiply p

What is the probability to see \mathbf{n} events when we expect λ ? $L(data|\vec{\theta}) = \mathbf{n} : \text{Integer}$ $\lambda : \text{Positive real (average number of event)}$

$$Poisson(\mathbf{n}, \lambda) = \frac{\lambda^{\mathbf{n}} e^{-\lambda}}{n!}$$





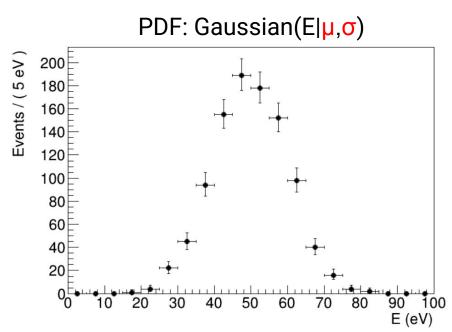
Likelihood definition - Extended unbinned likelihood

What is the probability to observe this data?

Multiply probabilities to observe events:

$$L(data|\vec{\theta}) = \prod_{Events} PDF(E_i|\vec{\theta}) dE \cdot Poisson(N_{total}|\lambda(\vec{\theta}))$$

Shape and event count fit (Extended unbinned likelihood)

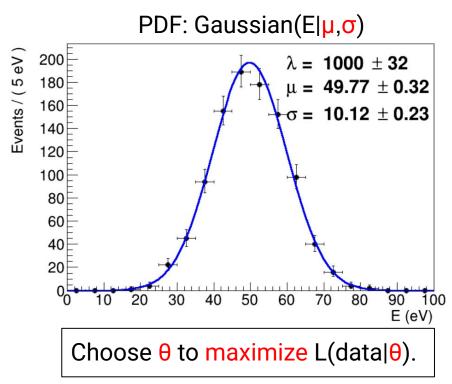


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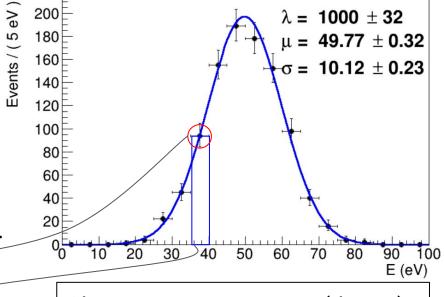
Shape and event count fit (Extended unbinned likelihood)

Alternatively, we can multiply bin probabilities.

$$L(data|\vec{\theta}) = \prod_{Bins} Poisson(n_i|\lambda_i(\vec{\theta}))$$

(Extended binned likelihood)





Choose θ to maximize L(data $|\theta$).

What is the

Multiply prol

$$L(data|\vec{\theta}) =$$

Alternativel

$$L(data|\vec{\theta}) = \prod_{Bi}$$

(Exte

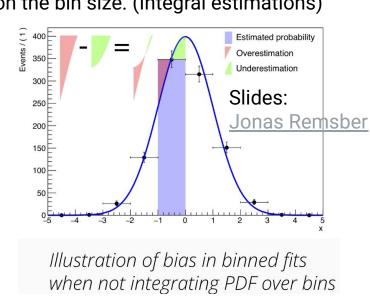
Features of binned likelihood:

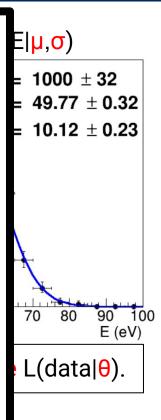
- + Less time consuming minimization. (100 events per bin?)
- + Numerically more stable.
- May be biased based on the bin size. (Integral estimations)

Typically:

Unbinned for low statistic and analytical PDFs.

Binned for high statistic and models from MC simulation.





Change θ to maximize L(data $|\theta$).

 θ_{max} gives the best fit (point estimate).

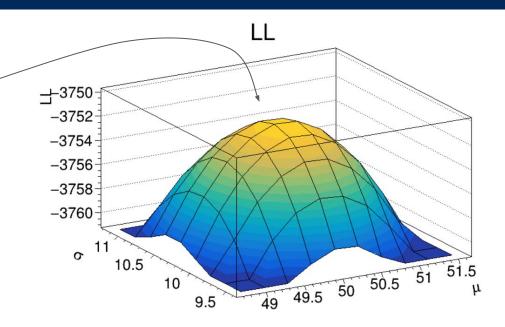
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How to do this?

1- For numerical stability use Log(L).

$$\prod_i o \sum_i$$



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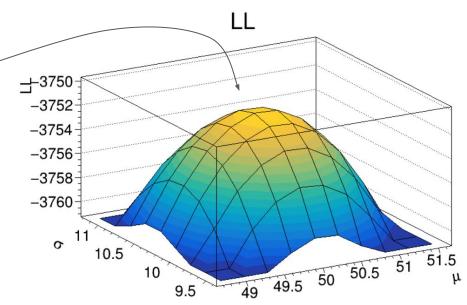
How to do this?

1- For numerical stability use Log(L).

$$\prod_i o \sum_i$$

2- Impossible to evaluate LL on a grid.

2 parameters 200² evaluations! 200 points each (What about 100s of parameters?)



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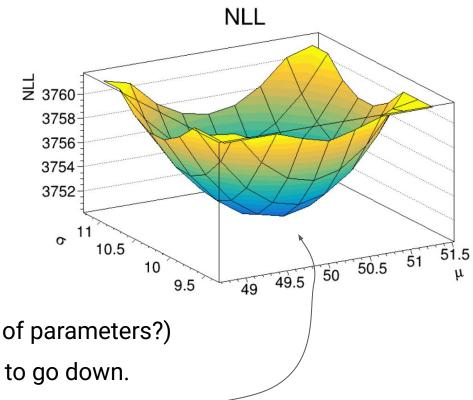
How to do this?

1- For numerical stability use Log(L).

$$\prod_i o \sum_i$$

- 2- Impossible to evaluate LL on a grid.
 - 2 parameters 200² evaluations! (What about 100s of parameters?)
- 3- Flip LL (negative LL) and use derivatives to go down.

ROOT minimizes the NLL (θ_{min}) .



Frequentist confidence (uncertainty)

What is uncertainty on θ_{min} ?

Frequentist confidence (uncertainty)

What is uncertainty on θ_{\min} ?

Imagine there are many replicas of the data to fit.

$$\begin{cases} L(\text{data}_{1}|\vec{\theta}) \to \vec{\theta}_{min}^{1} \\ L(\text{data}_{2}|\vec{\theta}) \to \vec{\theta}_{min}^{2} \\ \vdots \end{cases}$$

Frequentist confidence (Neyman construction)

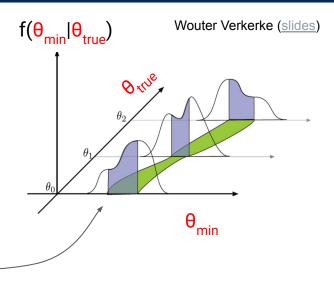
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 θ_{true} can be anything and is unknown.

How does θ_{min} distribute for a θ_{true} ?



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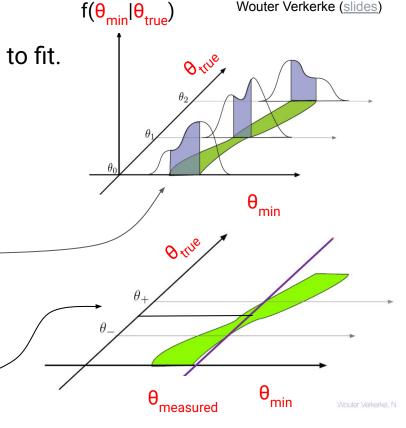
 θ_{true} can be anything and is unknown.

How does θ_{min} distribute for a θ_{true} ?

The confidence (uncertainty) range shows

how often θ_{true} is within $\theta_{measured}$.

$$1 \sigma \sim 68\%$$
 $2 \sigma \sim 95\%$



Frequentist confidence (Neyman construction)

What Neyman construction is intuitive but has issues.

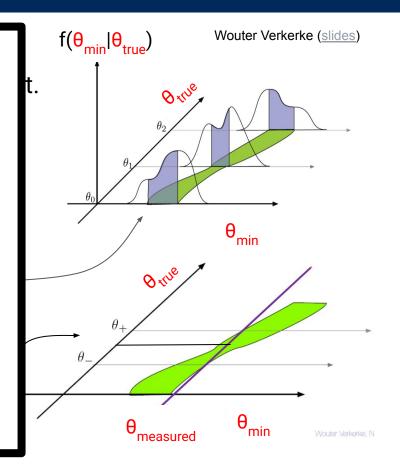
For instance finding $f(\theta_{min}|\theta_{true})$.

What to do?

Use likelihood ratio test

The c

How

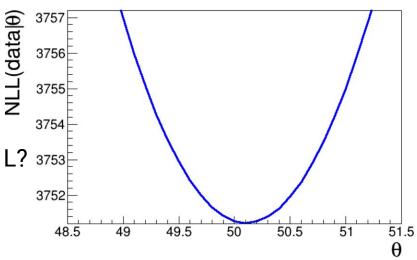


Confidence interval from likelihood ratio

Value of NLL(data $|\theta$):

- θ_{\min} gives the minimum.
- As θ deviates NLL increases.

How much θ changes for a significant change in L? 3753



Likelihood ratio test to measure the confidence interval

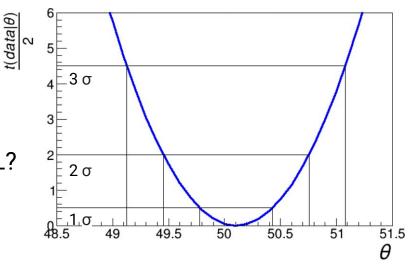
Value of NLL(data $|\theta$):

- θ_{min} gives the minimum.
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How much θ changes for a significant change in L?

Decide with the Likelihood Ratio (LR) test:

$$t(data|\theta) = -2\log \frac{L(data|\theta)}{L(data|\hat{\theta})}\bigg|_{\hat{\theta}: \text{Best fit}} = \left(\frac{\theta - \hat{\theta}}{\sigma}\right)^2\bigg|_{data \to \infty}$$



Papers sometime show Hessian errors. (An approximation!)

Likelihood ratio test to measure the confidence interval

Value

•

•

How

Decid

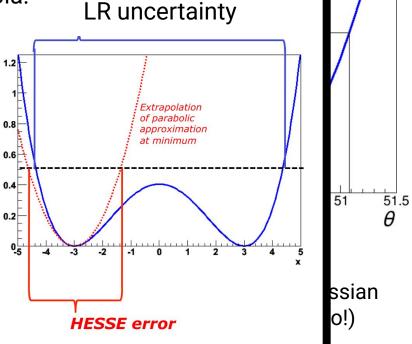
t(data

At lower statistic LR may not give a parabola.



Still the threshold of 1 for $t(data|\theta)$ gives a good 68% coverage.

Why?



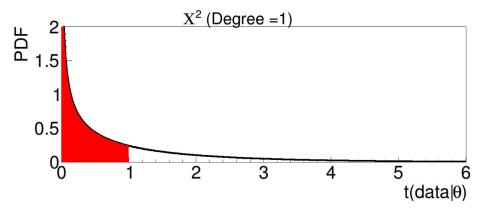
Distribution of $t(data|\theta)$

What is the distribution of $t(data|\theta)$?

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 $t(data|\theta)$ follows a chi² distribution, independent of θ !



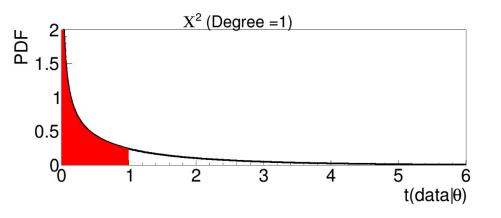
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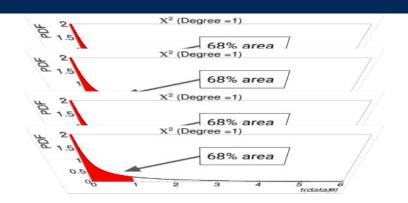
What is the distribution of $t(data|\theta)$?

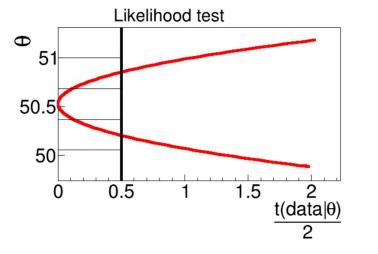
 $t(data|\theta)$ follows a chi² distribution, independent of θ !

You find a value for t at each θ .

A value above 1 is an unlikely outcome.

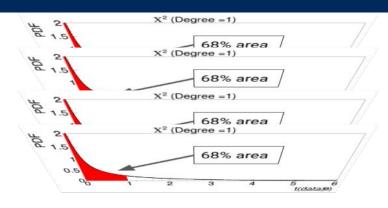


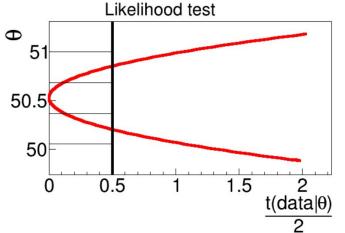




Distribution of t(data|θ)

It is always good to do more validation, especially with small data. More on this later! What if the likelihood has multiple parameters?





Nuisance parameters

<u>Definition</u>: parameters that will impact our results/change our PDF but are not the parameter of interest we care about fitting.

Typically they are measured or computed separately and have some uncertainty associated with them.

Examples: resolution, efficiency, calibration constants, isotope activity...

Two ways of including these in your likelihood, depending on your school of thought.

Notation in next few slides assumes μ is the POI and θ the nuisance parameter/s.

Nuisance parameters

<u>Definition</u>: parameters that will impact our results/change our PDF but are not the

parameter of intere

associated with the

Examples: resolution

Two ways of includ

We want to measure the parameter of Typically they are n interest. (fraction of DM events in data)

> We also have NPs that control the shape of the fit model.

How to propagate the uncertainty of NPs Notation in next few into the measurement?

Profile likelihood ratio

uncertainty

chool of thought.

e parameter/s.

Nuisance parameters

Frequentist:

Construct your likelihood with two data sets so we include the probability of observing data_x given μ and θ : $P(\mathrm{data}_x|\mu,\theta)$ and the probability of observing our other "nuisance" data_y (e.g., calibration data set) given θ $P(\mathrm{data}_y|\theta)$

Total likelihood is then:
$$\mathcal{L}(\mu, \theta) = P(\text{data}_x | \mu, \theta) P(\text{data}_y | \theta)$$

Bayesian:

Adjust your priors given the information you have on θ (nominally some mean value θ_0 and uncertainty $\Delta\theta$) given your previous measurements. Typically takes Poisson or Gaussian form:

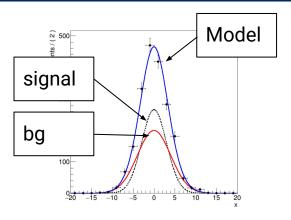
$$pdf(\mu, \theta) = P(data_x | \mu, \theta) Gauss(\theta | \theta_0, \Delta \theta)$$

Model: Gaussian-1 (signal) + Gaussian-2 (bg).

 μ_{Interest} : Ratio of sig/bg θ_{NP} : Background width.

Signal width is constant

We want to measure $\mu_{Interest}$.



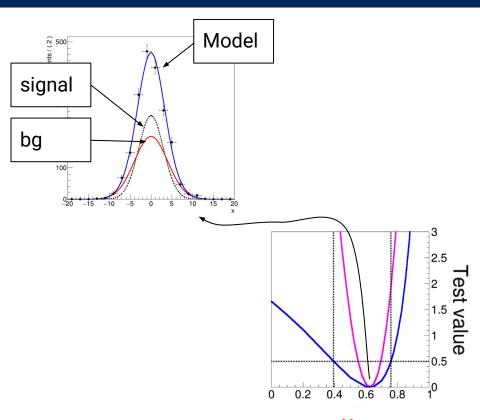
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We want to measure $\mu_{Interest}$.

LR quantifies how fast the model becomes inconsistent with data.



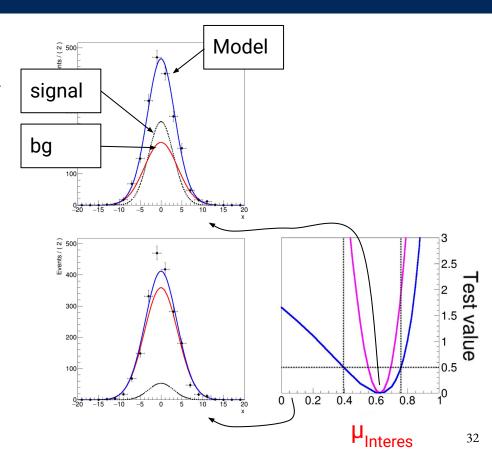
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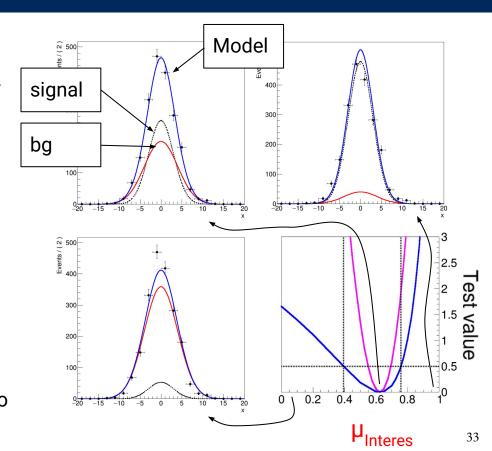
Signal width is constant

We want to measure $\mu_{Interest}$.

LR quantifies how fast the model becomes inconsistent with data.

 $\boldsymbol{\theta}_{NP}$ can compensate for the change in $\boldsymbol{\mu}_{Interest}.$

For small signal ($\mu_{Interest}$), make bg (θ_{NP}) pointier to keep the model consistent (smaller LR).



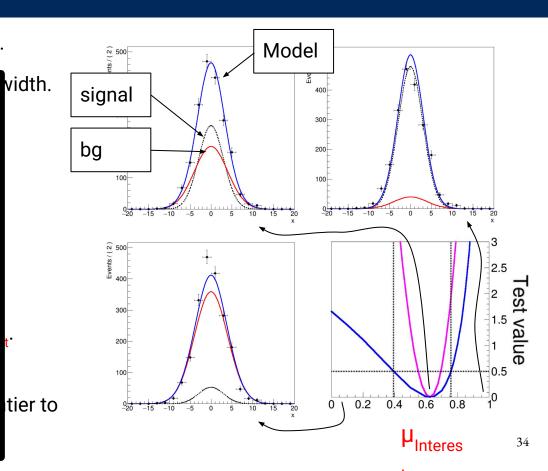
Model: Gaussian-1 (signal) + Gaussian-2 (bg).

This is the effect of NPs.

NPs have uncertainties. Let them float as a function of $\mu_{Interest}$ to keep the model consistent with data.

This is called Profile likelihood.

PL grows slower, uncertainty becomes wider.



PL propagates the uncertainty of NPs.

Best fit of θ for given μ

$$Log\left(\frac{L(\mu)}{L(\hat{\mu})}\right) \to Log\left(\frac{L(\mu, \hat{\hat{\theta}}(\mu))}{L(\hat{\mu}, \hat{\theta})}\right)$$

Global best fit values

PL propagates the uncertainty of NPs.

Best fit of
$$\theta$$
 for given μ

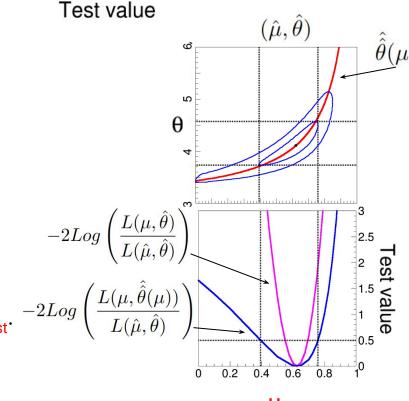
$$Log\left(\frac{L(\mu)}{L(\hat{\mu})}\right) \to Log\left(\frac{L(\mu, |\hat{\hat{\theta}}(\mu)|)}{L(\hat{\mu}, \hat{\theta})}\right)$$

Global best fit values

At each μ_{Interest} repeat the fit with floating θ_{NP} .

Changes in θ_{NP} compensates for the change in $\mu_{Interest}.$

Test reaches 0.5 slower -> Wider uncertainty.



From likelihood to profile likelihood (PL)

PL propagates the uncertainty of NPs.

Best fit of
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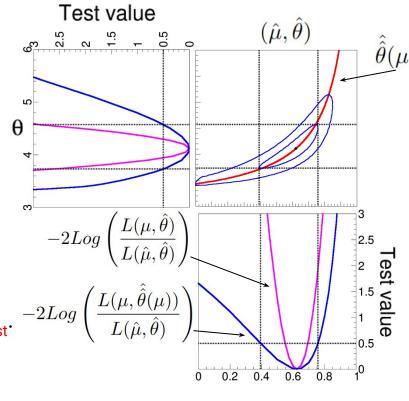
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From likelihood to profile likelihood (PL)

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Best fit of 6

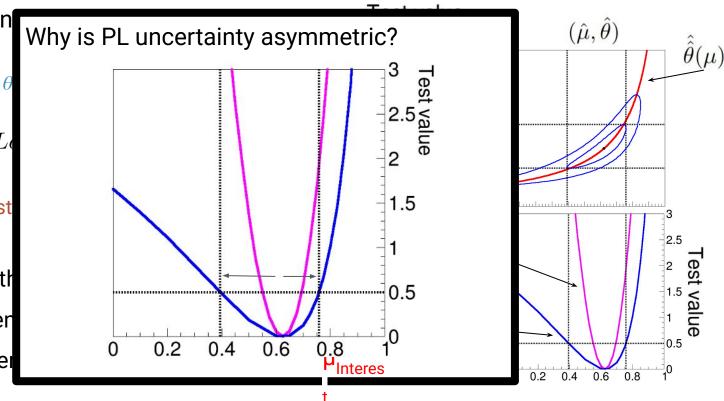
 $Log\left(\frac{L(\mu)}{L(\hat{\mu})}\right) \to L$

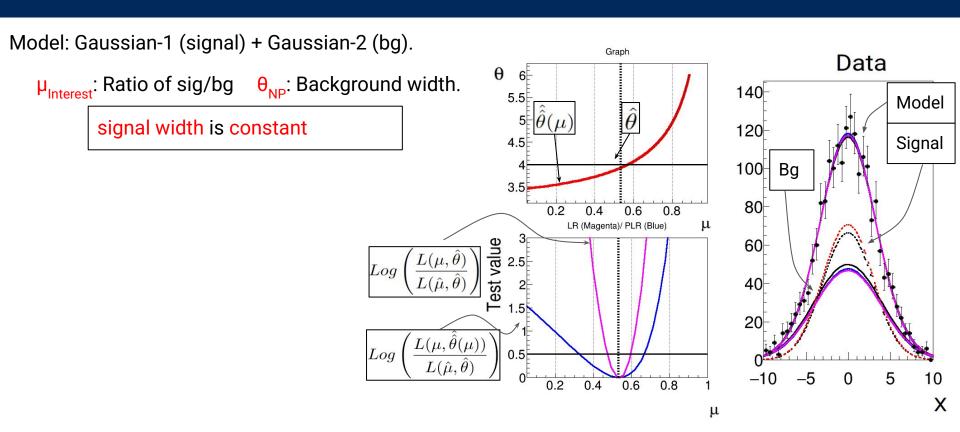
Global best

At each µ_{Interest} repeat the

Changes in θ_{NP} comper

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Model: Gaussian-1 (signal) + Gaussian-2 (bg).

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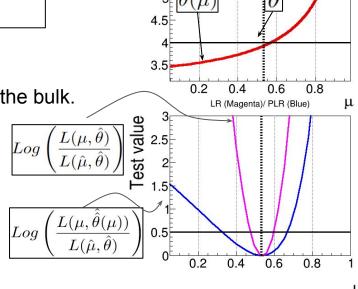
signal width is constant

Low μ :

Weak signal.

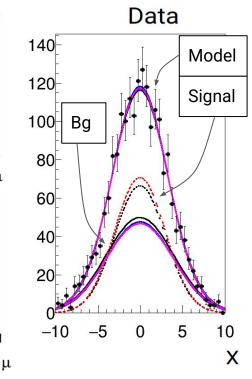
bg becomes narrower to model the bulk.

- It compensates the signal
- PLR grows slower than LR.



θ

Graph



Model: Gaussian-1 (signal) + Gaussian-2 (bg).

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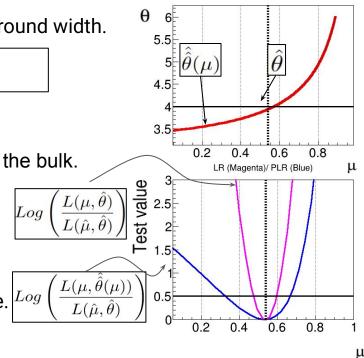
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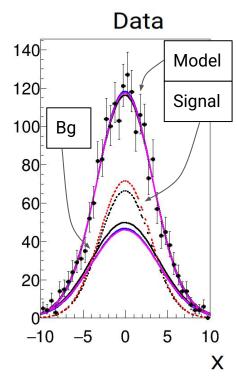
PLR grows slower than LR.

High µ:

- bg is depleted.
- ullet Tail cannot be modeled anymore. $|^{Log}$
- Model becomes inconsistent faster.
- PLR and LR increase parallel



Graph



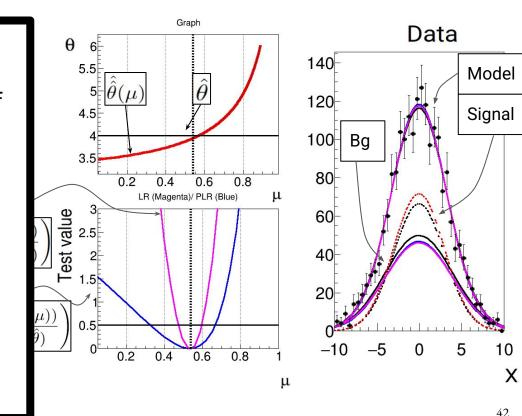
Model: Gaussian-1 (signal) + Gaussian-2 (hg)
Bottom line:

- 1. NPs increase the uncertainty if there is correlation.
- Try to understand strong correlations and the PLR response.

Low

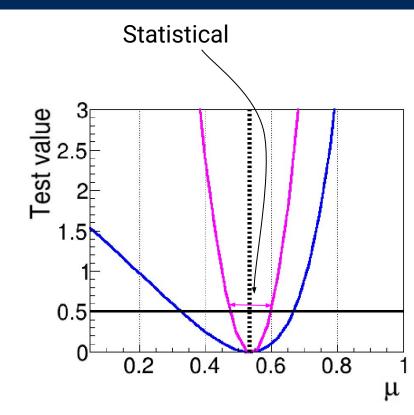
High

3. Use PLR to propagate uncertainties.



For fits with many floating parameters (5, 100, 1000, ...):

- 1. Fix all parameter to the best fit. $(\hat{\mu}, \hat{\theta})$
- 2. Do LR on μ to find the Statistical uncertainty.



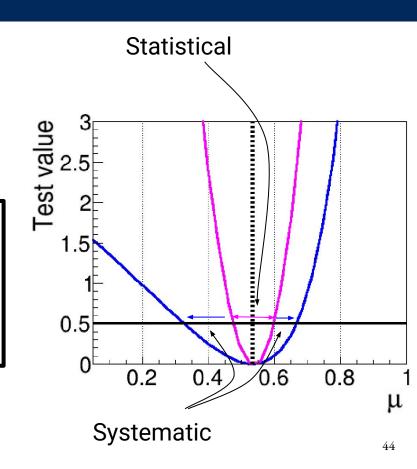
For fits with many floating parameters (5, 100, 1000, ...):

- 1. Fix all parameter to the best fit. $(\hat{\mu}, \hat{\theta})$
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- 3. Do PLR with one NP at a time for systematics.

Statistical uncertainty is dominant?

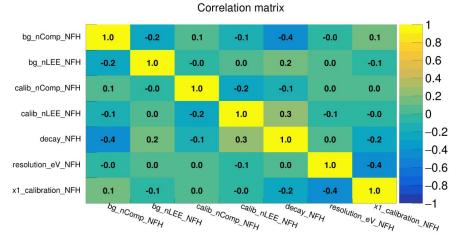
Deep breath, your life is easy ...

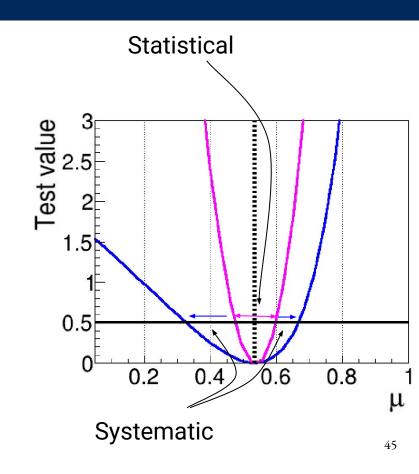
Assuming you considered all systematics;)

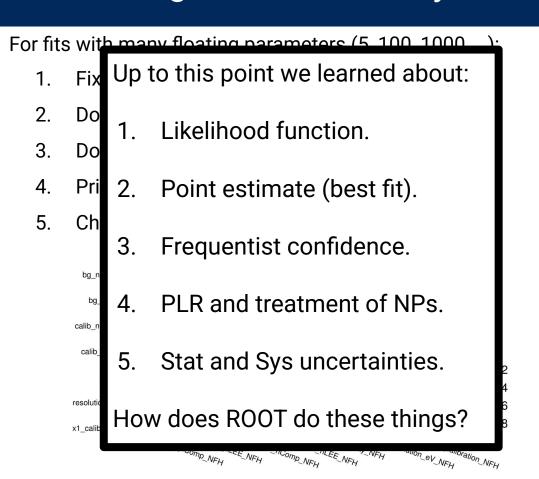


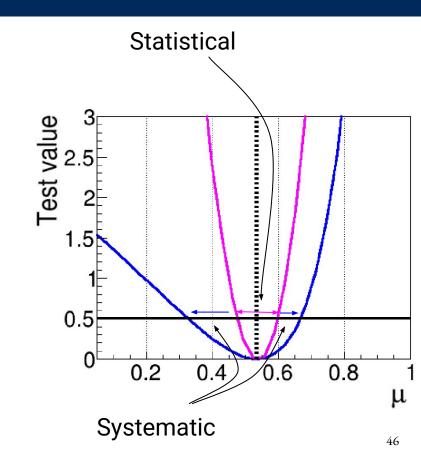
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- 3. Do PLR with one NP at a time for systematics.
- 4. Prioritize to understand strong correlations.
- 5. Check ranking plots. (More on this later)









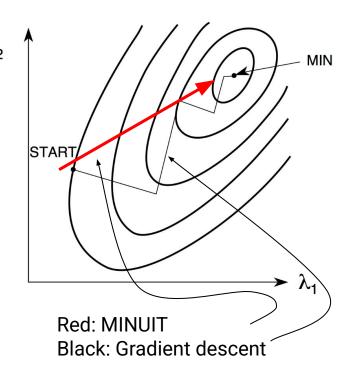
PLR diagnosis - (1) Understanding the minimizer (MINUIT)

MINUIT is the minimization package behind RooFit

It checks the gradient, and its change to find the direction to min. λ_2

$$f(\vec{\theta_0} + \delta \vec{\theta}) = f(\vec{\theta}) + \Delta f^T \theta + \frac{1}{2} \delta \vec{\theta}^T H \delta \vec{\theta} + \mathcal{F}$$

$$\nabla f(\vec{\theta_0} + \delta \vec{\theta}) = \nabla f(\vec{\theta_0}) + H \delta \vec{\theta} + \mathcal{F}$$



PLR diagnosis - (1) Understanding the minimizer (MINUIT)

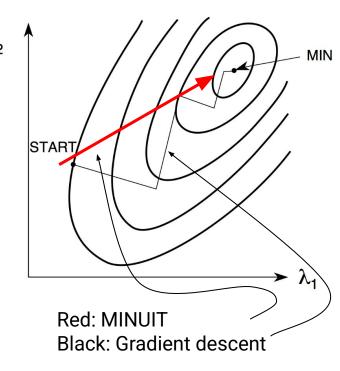
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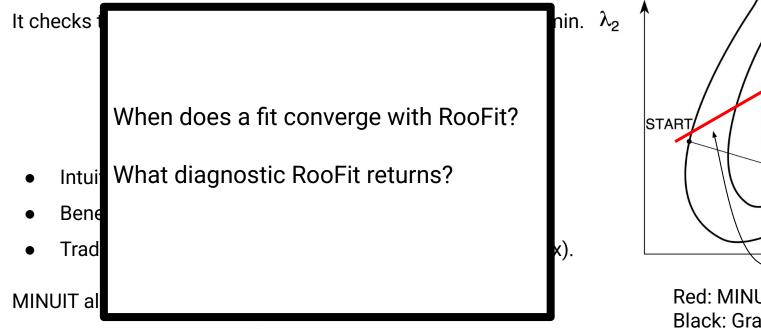
$$\nabla f(\vec{\theta_0} + \delta \vec{\theta}) = \nabla f(\vec{\theta_0}) + H \delta \vec{\theta} + \mathcal{F}$$

- Intuition: Take a big step if gradient is stationary.
- Benefit: Fast convergence.
- Tradeoff: Requires the Hessian (second derivative matrix).

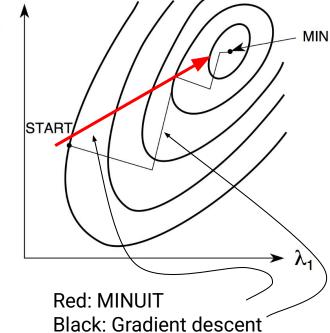


PLR diagnosis - (1) Understanding the minimizer (MINUIT)

MINUIT is the minimization package behind RooFit



Check "EDM = $\frac{1}{2} \cdot \nabla f^T H^{-1} \nabla f < 0.001$ " per step.

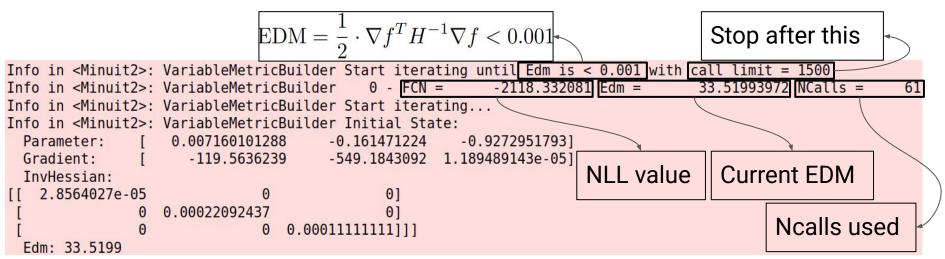


PLR diagnosis - (2) Convergence and output (MINUIT)

- 1- Initial minimization to EDM< 0.001 using approximate calculation of the Hessian (H).
- 2- Depending on the "strategy code":
 - 0 -> claim convergence.
 - 2 -> Find exact H, and continue the minimization if EDM>0.001
 - 1 (default) -> If approximate and exact H are close terminate. Continue with strategy 2 otherwise.

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More info: TMinuit reference - MINUIT manual

PLR diagnosis - (3) what may go wrong

Extreme correlations

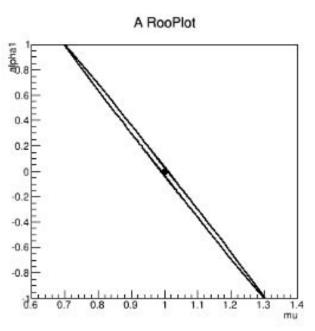
- Hessian and inverse Hessian calculation problems.
- Global minimum becomes like a valley instead of a point.
- Not well-defined minimum or uncertainty.
- HESSE fails when ratio of weakest-to-strongest eigenvalue < 10-6

Output:

```
Warning in <Minuit2>: VariableMetricBuilder Matrix not pos.def, gdel = 0.506583 > 0
Warning in <Minuit2>: MnPosDef non-positive diagonal element in covariance matrix[ 9 ] = -0.0231748
Warning in <Minuit2>: MnPosDef non-positive diagonal element in covariance matrix[ 10 ] = -0.270261
Warning in <Minuit2>: MnPosDef non-positive diagonal element in covariance matrix[ 12 ] = -21.0648
Warning in <Minuit2>: MnPosDef non-positive diagonal element in covariance matrix[ 16 ] = -0.0247647
Warning in <Minuit2>: MnPosDef Added to diagonal of Error matrix a value 21.5648
Warning in <Minuit2>: MnPosDef Matrix forced pos-def by adding to diagonal 1.38155
```

Solution:

Reparameterize or simplify the model.



 $\rho = -0.9995$

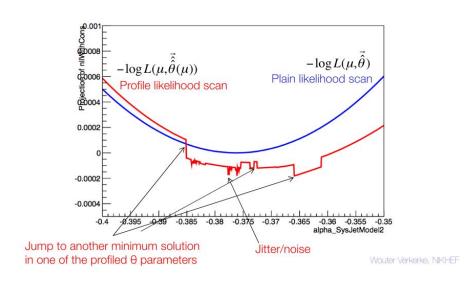
PLR diagnosis - (3) what may go wrong

Numerical instabilities

- Numerical integral.
- Numerical smearing.
- Very high statistic and complex model.
- Multiple minimums close the global minimum.

Solution:

- Increasing the numerical integral precession.
- Define Analytical integrals for PDF.
- Case by case debugging.
- Also talk to people with more experience.







Back ups

Useful references

- Asymptotic formulae for likelihood-based tests of new physics (Cowan, Cranmer, Gross, Vitells) https://arxiv.org/pdf/1007.1727.pdf
- PDG stats chapter
- Advances statistics (Verkerke)
 https://www.precision.hep.phy.cam.ac.uk/wp-content/people/mitov/lectures/GraduateLectures/Advanced-Statistics-Verkerke.pdf
- Dealing with uncertainty/errors: https://www.nikhef.nl/~ivov/Talks/2013_03_21_DESY_PoissonError.pdf
- RooFit tutorials available in c++ and python (in regular files and notebook format): https://root.cern/doc/master/group_tutorial_roofit.html
- RooStats tutorials available in c++ and python (in regular files and notebook format): https://root.cern/doc/master/group_tutorial_roostats.html
- Diagnostics guide from ATLAS:
 https://statisticalmethods.web.cern.ch/StatisticalMethods/recommendations/rec_diagnostics_checks/#investigating-simple-fits
- Introductory ROOT/RooFit/RooStats tutorials: https://root.cern/get_started/courses/