

HOMEWORK 6

Exercise are from Freedman.

7.A.12, 7.D (all, but turn in 2, 7, 13), 7.E (all, but turn in 6), Discussion Questions: 6, 7

Also, submit the “Class Laboratory” labelled **lab2.pdf**, answering the questions there. (It is on Canvas under “Pages/Class Laboratories”).

Problem Freedman 7A12. Suppose X, Y, Z are independent normal random variables, each having variance 1. The means are $\alpha + \beta, \alpha + 2\beta, 2\alpha + \beta$, respectively: α, β are parameters to be estimated. Show that maximum likelihood and OLS give the same estimates. Note: this won’t usually be true – the result depends on the normality assumption.

Solution to Problem Freedman 7A12. If

$$\mathbf{D} = \begin{bmatrix} 1 & 1 \\ 1 & 2 \\ 2 & 1 \end{bmatrix},$$

then

$$\begin{bmatrix} X \\ Y \\ Z \end{bmatrix} = \mathbf{D} \begin{bmatrix} \alpha \\ \beta \end{bmatrix} + \boldsymbol{\varepsilon},$$

where $\boldsymbol{\varepsilon} \sim N(\mathbf{0}, I_3)$. The OLS estimate is

$$\begin{bmatrix} \hat{\alpha} \\ \hat{\beta} \end{bmatrix} = (\mathbf{R}'\mathbf{R})^{-1}\mathbf{R}' \begin{bmatrix} X \\ Y \\ Z \end{bmatrix} = \frac{1}{11} \begin{bmatrix} 1 & -4 & 7 \\ 1 & 7 & -4 \end{bmatrix} \begin{bmatrix} X \\ Y \\ Z \end{bmatrix} = \frac{1}{11} \begin{bmatrix} X - 4Y + 7Z \\ X + 7Y - 4Z \end{bmatrix}.$$

On the other hand, we have

$$f(x, y, z; \alpha, \beta) = \frac{1}{(2\pi)^{3/2}} e^{-\frac{1}{2}[(x-(\alpha+\beta))^2 + (y-(\alpha+2\beta))^2 + (z-(2\alpha+\beta))^2]},$$

and so the log-likelihood is

$$\ell(\alpha, \beta) = -\frac{1}{2} [(X - (\alpha + \beta))^2 + (Y - (\alpha + 2\beta))^2 + (Z - (2\alpha + \beta))^2] - \frac{3}{2} \log(2\pi).$$

Thus

$$\begin{aligned} \frac{\partial \ell}{\partial \alpha} &= (X - (\alpha + \beta)) + (Y - (\alpha + 2\beta)) + 2(Z - (2\alpha + \beta)) \\ \frac{\partial \ell}{\partial \beta} &= (X - (\alpha + \beta)) + 2(Y - (\alpha + 2\beta)) + (Z - (2\alpha + \beta)) \end{aligned}$$

Setting to 0 and solving two linear equations for two unknowns yields

$$\begin{aligned} \hat{\alpha} &= \frac{1}{11}(X - 4Y + 7Z) \\ \hat{\beta} &= \frac{1}{11}(X + 7Y - 4Z) \end{aligned}$$

□

Problem Freedman 7D2. Conversely, if U is uniform on $[0, 1]$, show that $F^{-1}(U)$ has distribution function F .

Solution to Problem Freedman 7D2. Let $Y = F^{-1}(U)$. Then

$$F_Y(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(F^{-1}(U) \leq y) = \mathbb{P}(U \leq F(y)) = F_U(F(y)) = F(y),$$

since $F(u) = u$ for $u \in (0, 1)$. Thus, Y has distribution function F . \square

Problem Freedman 7D7. What is the distribution of $\log U - \log(1 - U)$, where U is uniform on $[0, 1]$?

Solution to Problem Freedman 7D7. Let $g(u) = \log(u/1 - u)$. Then

$$g'(u) = \frac{1-u}{u} \frac{1}{(1-u)^2} = \frac{1}{u(1-u)} > 0$$

for $u \in (0, 1)$; whence g is increasing in u . In particular, it is 1-1 from $(0, 1)$ to $(-\infty, \infty)$, with inverse

$$g^{-1}(x) = \frac{e^x}{1 + e^x}.$$

Thus, the cdf $F_{g(U)}$ of $g(U)$ is given by

$$F_{g(U)}(t) = \mathbb{P}(g(U) \leq t) = \mathbb{P}(U \leq g^{-1}(t)) = F_U(g^{-1}(t)) = g^{-1}(t) = \frac{e^x}{1 + e^x}.$$

\square

Problem Freedman 7D13. Show that the log likelihood for the probit model is concave, and strictly concave if X has full rank.

Solution to Problem Freedman 7D13. Let \mathbf{x}_k be the k -th row of \mathbf{X} .

$$\begin{aligned} L(\boldsymbol{\beta} \mid \mathbf{y}, \mathbf{x}) &= \prod_{k=1}^n \Phi(\mathbf{x}_k \boldsymbol{\beta})^{y_k} (1 - \Phi(\mathbf{x}_k \boldsymbol{\beta}))^{1-y_k} \\ \ell(\boldsymbol{\beta} \mid \mathbf{y}, \mathbf{x}) &= \sum_{k=1}^n y_k \log \Phi(\mathbf{x}_k \boldsymbol{\beta}) + (1 - y_k) \log(1 - \Phi(\mathbf{x}_k \boldsymbol{\beta})) \end{aligned}$$

Let $g = \log \Phi$ and $h = \log(1 - \Phi)$; we know from Exercise 12 that $g'' < 0$ and $h'' < 0$ everywhere. Then

$$\begin{aligned} \frac{\partial \ell}{\partial \beta_i}(\boldsymbol{\beta} \mid \mathbf{y}, \mathbf{x}) &= \sum_{k=1}^n y_k g'(\mathbf{x}_k \boldsymbol{\beta}) x_{k,i} + (1 - y_k) h'(\mathbf{x}_k \boldsymbol{\beta}) x_{k,i} \\ \frac{\partial^2 \ell}{\partial \beta_j \partial \beta_i}(\boldsymbol{\beta} \mid \mathbf{y}, \mathbf{x}) &= \sum_{k=1}^n y_k g''(\mathbf{x}_k \boldsymbol{\beta}) x_{k,i} x_{k,j} + (1 - y_k) h''(\mathbf{x}_k \boldsymbol{\beta}) x_{k,i} x_{k,j} \\ &= \sum_{k=1}^n x'_{i,k} (y_k g''(\mathbf{x}_k \boldsymbol{\beta}) + (1 - y_k) h''(\mathbf{x}_k \boldsymbol{\beta})) x_{k,j} \\ &= -\mathbf{X}' \mathbf{D} \mathbf{X}, \end{aligned}$$

where \mathbf{D} is the diagonal matrix with k -th diagonal entry

$$d_k = y_k (-g''(\mathbf{x}_k \boldsymbol{\beta})) + (1 - y_k) (-h''(\mathbf{x}_k \boldsymbol{\beta})).$$

Note that $d_k > 0$ since it is impossible that both y_k and $1 - y_k$ equal 0, and $-g''$ and $-h''$ are strictly positive everywhere.

Let \mathbf{u} be any p column vector. Then

$$\mathbf{u}'\mathbf{X}\mathbf{D}\mathbf{X}\mathbf{u} = (\sqrt{\mathbf{D}}\mathbf{X}\mathbf{u})'(\sqrt{\mathbf{D}}\mathbf{X}\mathbf{u}) = \|\sqrt{\mathbf{D}}\mathbf{X}\mathbf{u}\|^2 \geq 0,$$

whence $-\ell''$ is non-negative definite. This implies that ℓ is concave.

Suppose now that \mathbf{X} has full rank. Let $\mathbf{u} \neq 0$. Then $\mathbf{X}\mathbf{u} \neq 0$, since \mathbf{X} has full rank. Since \mathbf{D} is a strictly positive diagonal matrix, $\sqrt{\mathbf{D}}\mathbf{X}\mathbf{u} \neq 0$. We conclude that $-\mathbf{u}'\ell''\mathbf{u} > 0$, and $-\ell''$ is positive definite. This implies that ℓ is strictly concave. \square

Problem Freedman 7E6. Student #77 is Presbyterian, went to public school, and graduated. What does this subject contribute to the likelihood function? Write your answer using ϕ in equation (15).

Solution to Problem Freedman 7E6. We have $C_i = 0$ and $Y_i = 1$. This student contributed the factor

$$\begin{aligned} \mathbb{P}(C_i = 0, Y_i = 1 \mid X_i, IsCat_i = 0) &= \mathbb{P}(U_i < -X_i b, V_i > -X_i \beta) \\ &= \iint_{\substack{u < -X_i b \\ v > -X_i \beta}} \phi(u, v) du dv \end{aligned}$$

to the likelihood function. \square

Problem Freedman 7Dis6. Powers and Rock (1999) consider a two-equation model for the effect of coaching on SAT scores:

$$\begin{aligned} X_i &= \begin{cases} 1 & \text{if } U_i\alpha + \delta_i > 0 \\ 0 & \text{otherwise.} \end{cases} \\ Y_i &= cX_i + V_i\beta + \sigma\varepsilon_i. \end{aligned}$$

Here, $X_i = 1$ if subject i is coached, else $X_i = 0$. The response variable Y_i is subject i 's SAT score; U_i and V_i are vectors of personal characteristics for subject i , treated as data. The latent variables $(\delta_i, \varepsilon_i)$ are IID bivariate normal with mean 0, variance 1, and correlation ρ ; they are independent of the U 's and V 's. (In this problem, U and V are observable, δ and ε are latent.)

- Which parameter measures the effect of coaching? How would you estimate it?
- State the assumptions carefully (including a response schedule, if one is needed). Do you find the assumptions plausible?
- Why do Powers and Rock need two equations, and why do they need ρ ?
- Why can they assume that the disturbance terms have variance 1?

Solution to Problem Freedman 7Dis6. (a) The parameter c measures the effect of coaching. The parameters can be estimated via maximum likelihood.

- The vector (U_i, V_i) needs to be independent of $(\delta_i, \varepsilon_i)$, the variable X_i is generated from the first equation, then put in the response schedule

$$Y_{i,x} = cx + V_i\beta + \sigma\varepsilon_i.$$

Why is the effect c of coaching the same for all individuals?

- X_i may be related to unmeasured variables represented in the error term. Thus ρ allows for correlation between X_i and the error ε_i .

- (d) The parameter σ scales the error term in the second equation. Parameters are not identifiable in the second if δ_i can have arbitrary variance.

□