TECHNICAL COMMENT

PLANT ECOLOGY

Comment on "Worldwide evidence of a unimodal relationship between productivity and plant species richness"

Andrew T. Tredennick, ^{1*} Peter B. Adler, ¹ James B. Grace, ² † W. Stanley Harpole, ³ † Elizabeth T. Borer, ⁴ † Eric W. Seabloom, ⁴ † T. Michael Anderson, ⁵ Jonathan D. Bakker, ⁶ Lori A. Biederman, ⁷ Cynthia S. Brown, ⁸ Yvonne M. Buckley, ⁹ Chengjin Chu, ¹⁰ Scott L. Collins, ¹¹ Michael J. Crawley, ¹² Philip A. Fay, ¹³ Jennifer Firn, ¹⁴ Daniel S. Gruner, ¹⁵ Nicole Hagenah, ¹⁶ Yann Hautier, ¹⁷ Andy Hector, ¹⁸ Helmut Hillebrand, ¹⁹ Kevin Kirkman, ¹⁶ Johannes M. H. Knops, ²⁰ Ramesh Laungani, ²¹ Eric M. Lind, ⁴ Andrew S. MacDougall, ²² Rebecca L. McCulley, ²³ Charles E. Mitchell, ²⁴ Joslin L Moore, ²⁵ John W. Morgan, ²⁶ John L. Orrock, ²⁷ Pablo L. Peri, ²⁸ Suzanne M. Prober, ²⁹ Anita C. Risch, ³⁰ Martin Schütz, ³⁰ Karina L. Speziale, ³¹ Rachel J. Standish, ³² Lauren L. Sullivan, ⁴ Glenda M. Wardle, ³³ Ryan J. Williams, ³⁴ Louie H. Yang ³⁵

Fraser et al. (Reports, 17 July 2015, p. 302) report a unimodal relationship between productivity and species richness at regional and global scales, which they contrast with the results of Adler et al. (Reports, 23 September 2011, p. 1750). However, both data sets, when analyzed correctly, show clearly and consistently that productivity is a poor predictor of local species richness.

raser et al. (1) collected a worldwide data set to examine the relationship between productivity and species richness at global and local scales. They present their results as a direct contrast with the results of Adler et al. (2). However, their presentation obscures substantial areas of agreement, and where results between the two studies do differ, problems in Fraser et al.'s statistical analysis amplify the apparent differences.

The most important area of agreement is the low explanatory power of the "humped-back model" (HBM), in which species richness peaks at intermediate productivity and declines at low and high productivity. Fraser et al. fit a bivariate relationship between productivity and diversity that accounts for less than 1% of the observed variation in species richness in their data (Table 1, marginal R^2 s for the Fraser *et al.* data set). The same is true for an analysis of the Adler et al. data set using a generalized linear mixed model (GLMM) with a block nested within-site randomeffects structure (Table 1, marginal R^2 s for the Adler et al. data set). Thus, the analyses in both Adler et al. and Fraser et al. demonstrate that productivity is an uninformative predictor of richness for most grasslands. A combined analysis using both data sets yields similar results (Table 1).

A second point of agreement is the difficulty of inferring process from bivariate patterns. The HBM can arise through a wide array of mechanisms (3, 4), meaning that the detection of a unimodal pattern does not provide evidence for any particular mechanism.

Adler et al. argued, "[e]cologists should focus on fresh, mechanistic approaches to understanding the multivariate links between productivity and richness" (2). Fraser et al. also concluded "more work is needed to determine the underlying causal mechanisms that drive the unimodal pattern" and called for "additional efforts to understand the multivariate drivers of species richness."

The key disagreement between Fraser et al. and Adler et al. concerns the statistical significance of the quadratic term that determines the downward concavity of the richness productivity relationship. Adler et al. found little evidence for a concave-down relationship at the site scale (2% of 48 sites) [figure 2 in (2)] and at the global scale reported a significant effect but noted that it was sensitive to choices about which sites to include in the analysis [figure 3 in (2)]. In contrast, Fraser et al. found that 68% of 28 site-level relationships were significantly concave-down [figure 2A in (1)], and in a global extent regression, across all sites, the negative quadratic term had a significant, and robust, P value. However, their analysis at the site level is flawed, and the presentation of the global regression in their main figure is misleading.

The site-level regressions reported by Fraser *et al.* and displayed in their figure 2A do not include the proper random-effects structure. An important feature of the Fraser *et al.* design was explicitly selecting areas (i.e., grids) to sample across productivity gradients within sites, whereas Adler *et al.* located blocks of plots randomly with respect to local productivity gradients. To properly

reflect their sampling design, in which each "grid" of quadrats was located at one point along the within-site productivity gradient, each site-level

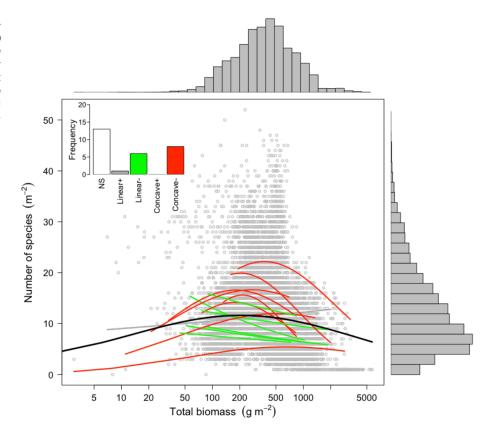
¹Department of Wildland Resources and the Ecology Center, Utah State University, 5230 Old Main, Logan, UT 84322, USA. ²U.S. Geological Survey, Wetland and Aquatic Research Center, 700 Cajundome Boulevard, Lafayette, LA 70506, USA. ³Department of Physiological Diversity, Helmholtz Center for Environmental Research – UFZ, Permoserstrasse 15, 04318 Leipzig, Germany. ⁴Department of Ecology, Evolution, and Behavior, University of Minnesota, St. Paul, MN 55108, USA. 5Department of Biology, Wake Forest University, Box 7325 Reynolda Station, Winston-Salem, NC 27109, USA. 6School of Environmental and Forest Sciences, University of Washington, 3501 NE 41st Street, Box 354115, Seattle, WA 98195, USA. ⁷Ecology, Evolution and Organismal Biology, Iowa State University, 251 Bessey Hall, Ames, IA 50010, USA. ⁸Department of Bioagricultural Sciences and Pest Management, Colorado State University, 307 University Avenue, Fort Collins, CO 80523, USA. 9School of Natural Sciences, Trinity College Dublin, University of Dublin, Zoology, Dublin 2, Ireland. 10 School of Life Sciences, Sun Yat-sen University, Xingang Xi Road 135, Guangzhou, 510275, China. ¹¹Department of Biology, University of New Mexico, Albuquerque, NM 87131, USA. ¹²Department of Biology, Imperial College London, Silwood Park, Ascot, SL5 7PY, UK. ¹³Grassland, Soil, and Water Research Laboratory, USDA-ARS, 808 East Blackland Road, Temple, TX 76502, USA. ¹⁴School of Earth, Environmental and Biological 42 Sciences, Queensland University of Technology (QUT), Gardens Point, Brisbane, Queensland, Australia, 4001. ¹⁵Department of Entomology, University of Maryland, 4112 Plant Sciences, College Park, MD 20742, USA. ¹⁶School of Life Sciences, University of KwaZulu-Natal, 1 Carbis Road, Pietermaritzburg, 3201, South Africa. ¹⁷Department of Biology, Ecology and Biodiversity group, Utrecht University, Padualaan 8, 3584 CH Utrecht, Netherlands. 18 Department of Plant Sciences, University of Oxford, South Parks Road, Oxford, OX1 3RB, UK. ¹⁹Institute for Chemistry and Biology of the Marine Environment, Carl-von-Ossietzky University Oldenburg, Schleusenstrasse 1, 26382 Wihlhemshaven, Germany. ²⁰School of Biological Sciences, University of Nebraska, 211 Manter Hall, Lincoln, NE 68588, USA. ²¹Biology Department, Doane College, 1014 Boswell Avenue, Crete, NE 68333, USA. ²²Department of Integrative Biology, University of Guelph, 50 Stone Road, Guelph, Ontario, Canada N1G 2W1. ²³Department of Plant and Soil Science, University of Kentucky, N-222D Ag Science North, Lexington, KY 40546-0091, USA. ²⁴Department of Biology, University of North Carolina at Chapel Hill, CB#3280, Chapel Hill, NC 27599, USA. ²⁵School of Biological Sciences, Monash University, Clayton Campus, Wellington Road, Clayton 3800, Victoria, Australia. ²⁶Department of Ecology, Environment and Evolution, La Trobe University, Kingsbury Drive, Bundoora 3086, Victoria, Australia. ²⁷Department of Zoology, University of Wisconsin, 430 Lincoln Drive, Madison, WI 53706, USA. ²⁸Department of Forestry, Agriculture and Water, Southern Patagonia National University-INTA-CONICET, CC 332 (CP 9400), Río Gallegos, Santa Cruz, Patagonia, Argentina. 29 Commonwealth Scientific and Industrial Research Organisation Land and Water, Private Bag 5, Wembley, WA 6913, Australia. 30 Community Ecology, Swiss Federal Institute for Forest, Snow and Landscape Research. Zuercherstrasse 111, 8903 Birmensdorf, Switzerland. ³¹Department of Ecology, INIBIOMA (CONICET-UNCO), Quintral 1250, Bariloche (8400), Rio Negro, Argentina. 32 School of Veterinary and Life Sciences, Murdoch University, Perth, Western Australia, 90 South Street, Murdoch, Western Australia 6150. ³³School of Biological Sciences, University of Sydney, Heydon-Laurence Building, A08, University of Sydney, Sydney, NSW, 2006, Australia. ³⁴Agricultural and Biosystems Engineering, Iowa State University, Ames, IA, USA. 35 Department of Entomology and Nematology, University of California, Davis, One Shields Avenue, Davis, CA 95616, USA.

*Corresponding author. E-mail: atredenn@gmail.com †Core authors that led the analysis and wrote the paper. All other authors, listed alphabetically, are Nutrient Network members and/or coauthors of Adler et al. (2011) who sign our Comment in support to show consensus among the Nutrient Network.

Table 1. Results from global-extent GLMMs for both data sets. Results from regressions with and without a quadratic effect of productivity on species richness across all sites. Both models include a random-effects structure of grid nested within site (Fraser et al.) or block nested within site (Adler et al.). Marginal and conditional R² values estimated using (7, 8). For the combined analysis, we use the same grid (or block) nested within-site random-effects structure and also include a "study" random effect.

Data set	Model type	Marginal R^2 (variance explained by fixed effects)	Conditional R ² (variance explained by fixed + random effects)	Root mean square error (in units of species number)
Fraser et al.	Linear	0.00007	0.84	8.5
Fraser et al.	Quadratic	0.009	0.84	8.3
Adler et al.	Linear	0.0007	0.79	7.7
Adler et al.	Quadratic	0.001	0.78	7.7
Combined	Linear	0.0005	0.82	8.4
Combined	Quadratic	0.003	0.82	8.3

Fig. 1. Species richness as a function of biomass production at the site level (colored lines) and at the global extent (heavy black line). These regressions are the same as presented by Fraser et al. except that we included a grid random effect for the site-level regressions, and we show the proper global extent regression line from a GLMM with grid nested within site. Nonsignificant regression fits are not plotted.



regression requires a random effect of "grid" to account for the inherent correlation among plots nested within a sampling grid. We reran the analysis of Fraser et al. with the grid random effect included (5), except for one site (6). When the proper statistical model is used, we find that only 29% of 28 site-level regressions are significantly concave-down (Fig. 1).

Fraser et al. correctly account for their sampling design at the global extent by using a GLMM with grid nested within site, as reported in their table 1. However, in their figure 2A, they plot the much more compelling fit from the statistical model without the random effects. Although still significant (P < 0.0001), the valid relationship is much weaker than the relationship presented by Fraser et al. (Fig. 1, heavy black line, and Table 1).

Despite Fraser et al.'s assertion that their results are diametrically opposed to those presented in Adler et al., the degree of concordance is impressive. In both data sets, the variance explained by the addition of a quadratic term is virtually indistinguishable from that of a linear model (Table 1). In fact, in both data sets the random effects of site and grid (block for Adler et al.) explain much more of the variation in species richness than productivity, the supposed mechanistic driver of species richness (Table 1). Furthermore, with the appropriate statistical treatment, the main difference in our results-the strength of evidence for a significant quadratic term-appears smaller.

A continued focus on this bivariate relationship hinders progress toward understanding the underlying multivariate causal relationship (4) and the development of truly predictive models. It is time to focus on effect sizes and variance explained rather than just P values. The title of Adler et al.'s paper, "Productivity is a poor predictor of plant species richness," would be a perfectly appropriate title for the Fraser et al. paper, too.

REFERENCES AND NOTES

- 1. L. H. Fraser et al., Science 349, 302-305 (2015).
- 2. P. B. Adler et al., Science 333, 1750-1753 (2011).
- 3. J. B. Grace et al., Science 335, 6075 (2012).
- 4. J. B. Grace et al., Nature 529, 10.1038/nature16524 (2016).
- 5. We used the "Ime4" package in the statistical programming environment R to fit the GLMMs at the site and global extents. Some models struggled to converge on coefficient estimates, a well-known issue with mixed-effects models. We conducted the analyses using different optimizers to make sure that our results are robust (they are), and we did our own checks of model diagnostics to make sure that the warnings could be ignored (they could). Lastly, we fit a hierarchical mixed-effects model using a Bayesian approach to make sure we obtained consistent
- results (we did). All of our analyses and results can be found on GitHub at http://github.com/atredennick/prodDiv and as release v0.1, https://github.com/atredennick/prodDiv/tag/v0.1
- 6. There are four sites, out of 28, that have only two grids. In only one case did this result in inadequate fits of the GLMM model with a "grid" random effect. We therefore fit that one site with a generalized linear model with no random effects.
- S. Nakagawa, H. Schielzeth, *Methods Ecol. Evol.* 4, 133–142 (2013).
 J. Lefcheck, R-squared for generalized linear mixed-effects models (2014); https://github.com/jslefche/rsquared.glmm

ACKNOWLEDGMENTS

We thank L. Fraser and colleagues for making their analyses and data openly available. D. Johnson, USGS, provided comments on

an earlier version of the manuscript. J.B.G. was supported by the USGS Ecosystems and Climate and Land Use Change Programs. The use of trade, firm, or product names is for descriptive purposes only and does not imply endorsement by the U.S. Government. USDA is an equal opportunity employer. We also acknowledge support from the National Science Foundation Research Coordination Network (NSF-DEB-1042132) and Long Term Ecological Research (NSF-DEB-1234162 to Cedar Creek LTER) programs, and the Institute on the Environment (DG-0001-13).

16 October 2015; accepted 17 December 2015 10.1126/science.aad6236

TECHNICAL RESPONSE

PLANT ECOLOGY

Response to Comment on "Worldwide evidence of a unimodal relationship between productivity and plant species richness"

Jason Pither, 1* Lauchlan H. Fraser, 2 Anke Jentsch, 3 Marcelo Sternberg, 4 Martin Zobel, 5 Jason Pither, Lauchian H. Fraser, Anke Jentsch, Marcelo Sternberg, Martin Zob James Cahill, Carl Beierkuhnlein, Sándor Bartha, Jonathan A. Bennett, 6 Bazartseren Boldgiv, Leslie R. Brown, Marcelo Cabido, Campetella, Gameron N. Carlyle, Stefano Chelli, Anna Mária Csergő, Sandra Diaz, Lucas Enrico, David Ensing, Anna Mária Csergő, Heath W. Garris, 2 Hugh A. L. Henry, 19 Maria Höhn, 20 John Klironomos, 1 Kadri Koorem, 5 Rachael Lawrence-Lodge, 21 Peter Manning, 22 Randall J. Mitchell, 23 Mari Moora, 5 Valério D. Pillar,²⁴ Gisela C. Stotz,⁶ Shu-ichi Sugiyama,²⁵ Szilárd Szentes,²⁶ Radnaakhand Tungalag, 10 Sainbileg Undrakhbold, Camilla Wellstein, 27 Talita Zupo 18

Tredennick et al. criticize one of our statistical analyses and emphasize the low explanatory power of models relating productivity to diversity. These criticisms do not detract from our key findings, including evidence consistent with the unimodal constraint relationship predicted by the humped-back model and evidence of scale sensitivities in the form and strength of the relationship.

redennick et al. (1), among them many contributors to the original Adler et al. study (2), argue that our findings (3) align closely with those of Adler et al. once their criticisms (described below) are addressed. This is not the case. Tredennick et al. fail to acknowledge key findings of ours that remain at odds with those of Adler et al., including (i) a significantly concave-down, global-extent relationship between productivity and richness; (ii) a significantly concave-down, global-extent quantile regression, consistent with the constraint prediction of the humped-back model (HBM); and (iii) our finding that patterns consistent with the HBM appear more evident when a broad range of productivity is sampled.

Tredennick et al. present three main criticisms of our study: (i) the analyses of the within-site productivity-diversity relationship should have included sample "grid" as a random effect, thereby accounting for our nested sampling design; (ii) our analyses focused too much on the significance of the quadratic term and not enough on the limited explanatory power of the models; and (iii) our figure 2A (3) was "misleading" and should have included a line representing the mixed-effects model for the global-extent relationship. We address each of these in turn.

(i) We agree that including "grid" as a random effect within mixed-effects models would be a reasonable approach. In our within-site analyses, we intentionally replicated the within-site analyses of Adler et al., who did not accommodate the nestedness inherent to their sampling design. In hindsight, we regret not including the results of mixed-effects models for our withinsite analyses in the supplementary materials, as we did for all other analyses. We made our data publicly available, which enabled Tredennick et al. to conduct analyses of their own, finding that 8 (29%) rather than 19 (69%) of the 28 withinsite analyses yielded a significant concave-down relationship when "grid" is included as a random effect. Crucially, these revised analyses by Tredennick et al. have no effect on the global models we presented that form the main conclusion of the study. Also, thanks to Tredennick et al. making their data and analyses publicly available, we found that the 8 sites that did exhibit a significantly concave-down relationship in their analyses encompassed a significantly larger range of productivity (on the log₁₀ scale) than the 13 sites where no association was found (permutation test on the difference in mean productivity; 9999 permutations; Z score = 2.09; P = 0.039). Moreover, the probability of detecting a concave-down relationship (i.e., significant quadratic term) over no relationship using the mixed-effects modeling approach tended to increase with increasing biomass range (logistic regression; residual deviance = 22.6 on 19 df; P = 0.078).

(ii) We recognize that regressions modeling the mean trend between productivity and richness yield limited explanatory power, and stated so in our Report (3). We suggest that Tredennick et al.'s focus on the mean trend is misplaced because, provided one samples a sufficiently broad range of productivity, the HBM predicts a constraint relationship, whereby richness is constrained to low levels at very low and very high productivity. Our study provided evidence of this, in the form of a significantly concave-down, global-extent quantile regression (both with and without random effects included). Adler et al. also tested for such a constraint relationship (without random effects) but failed to detect it, possibly because of limits to their sampling (3). For our analyses of mean trends, we focused on the form of the relationship, and hence the significance of the quadratic term, because thisnot explanatory power-lies at the heart of the debate surrounding the HBM (4, 5). Our sampling design and sampling scope allowed us to test the sensitivity of the form of the relationship to varying sampling grains and extents.

(iii) We formatted our figure 2A (3) with the objective of making it directly comparable to the results presented by Adler et al. (2), specifically their figure 2, and their global-extent regression, which was displayed in their figure 3. Adler et al. did not account for nested sampling structure in any of their analyses (i.e., using mixed-effects models), including within their global-extent analysis that yielded a significantly concave-down

¹Department of Biology, University of British Columbia, Okanagan Campus, Kelowna, BC, Canada. ²Department of Natural Resource Sciences, Thompson Rivers University, Kamloops, BC, Canada. ³Department of Disturbance Ecology, BayCEER, University of Bayreuth, Bayreuth, Germany. ⁴Department of Molecular Biology and Ecology of Plants, Tel Aviv University, Tel-Aviv, Israel. 5Department of Botany, Institute of Ecology and Earth Sciences, University of Tartu, Tartu, Estonia. 6Department of Biological Sciences, University of Alberta, Edmonton, Alberta, Canada. ⁷Department of Biogeography, Bayreuth Center of Ecology and Environmental Research (BayCEER), University of Bayreuth, Bayreuth, Germany. 8Hungarian Academy of Sciences Centre for Ecological Research, Institute of Ecology and Botany, Vácrátót, Hungary. 9School of Plant Biology, University of Western Australia, Crawley, Australia. 10 Ecology Group, Department of Biology, National University of Mongolia, Ulaanbaatar, Mongolia. 11Department of Environmental Sciences, University of South Africa, Florida, South Africa. ¹²Facultad de Ciencias Exactas, Físicas y Naturales, Universidad Nacional de Córdoba, Córdoba, Argentina. 13 School of Biosciences and Veterinary Medicine, Plant Diversity and Ecosystems Management Unit, University of Camerino, Camerino, Italy. 14 Department of Agricultural, Food and Nutritional Science, University of Alberta, Edmonton, Canada. 15 School of Natural Sciences, Trinity College Dublin, The University of Dublin, Dublin, Ireland. ¹⁶Instituto Multidisciplinario de Biología Vegetal (IMBIV), National Scientific and Technical Research Council and Facultad de Ciencias Exactas, Físicas y Naturales, Universidad Nacional de Córdoba, Córdoba, Argentina. ¹⁷Department of Biology, Queen's University, Kingston, Ontario, Canada. 18 Departamento de Botânica, UNESP -Universidade Estadual Paulista, Rio Claro, SP, Brazil. ¹⁹Department of Biology, University of Western Ontario, London, Ontario, Canada. 20 Department of Botany, Faculty of Horticultural Science, Corvinus University of Budapest, Hungary. ²¹Department of Botany, University of Otago, Dunedin, New Zealand. ²²Biodiversity and Climate Research Centre, Senckenberg Gesellschaft für Naturforschung, Germany. ²³Department of Biology, University of Akron, Akron, OH, USA. ²⁴Department of Ecology, Federal University of Rio Grande do Sul, Porto Alegre, Brazil. ²⁵Faculty of Agriculture and Life Science, Hirosaki University, Aomori, Japan. ²⁶Institute of Plant Production, Szent István University, Gödöllő, Hungary. 27 Faculty of Science and Technology, Free University of Bozen-Bolzano, Bolzano, Italy. *Corresponding author. E-mail: jason.pither@ubc.ca

relationship. We therefore opted to show our analogous regression results in figure 2A(3). We showed the results of our mixed-effects model for the global relationship in figure S1 (3).

We encourage future research to (i) explore why low species richness (per unit area) is found at the extreme ends of the productivity gradient and (ii) determine the processes that suppress species richness below its potential at intermediate levels of productivity.

REFERENCES

A. T. Tredennick *et al.*, *Science* **351**, 457 (2016).
 P. B. Adler *et al.*, *Science* **333**, 1750–1753 (2011).

3. L. H. Fraser *et al.*, *Science* **349**, 302–305 (2015). 4. J. P. Grime, *J. Environ. Manage.* **1**, 151–167 (1973). 5. L. H. Fraser, A. Jentsch, M. Sternberg, *J. Veg. Sci.* **25**, 1160–1166 (2014).

16 November 2015; accepted 17 December 2015 10.1126/science.aad8019

457-b 29 JANUARY 2016 • VOL 351 ISSUE 6272 sciencemag.org SCIENCE