Appendix 4

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"Consistent ecosystem functional response across precipitation extremes in a sagebrush steppe"

PeerI

Section A4.1 Conducting our analysis with NDVI

- 6 The first step of our analysis is to convert NDVI from field-based radiometric measurements to
- aboveground net primary productivity. However, this conversion is not perfect and the associated
- 8 uncertainty is sometimes high (Table A1-2). Therefore, we also assessed ecosystem functional
- 9 response across treatments using NDVI as the response variable rather than ANPP. The model
- 10 essentially the same as the ANPP model, except we use a Beta likelihood for NDVI since its values
- range from 0 to 1, not including true 0s or 1s (see Stan code, below).
- Our NDVI-model results are very similar to the ANPP-model results (Figure A4-1). The only
- difference is there is a higher probability that the drought slope offset is less than 1 (90%
- probability). However, these negative offsets do not result in significantly different slopes once
- applied to the control slope, just as in the ANPP-model.

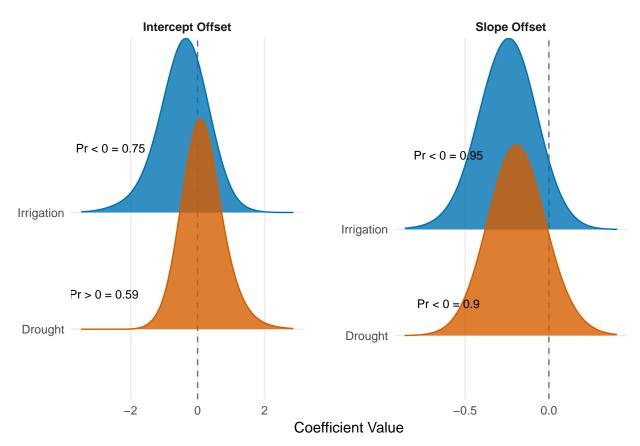


Figure A4-1 Posterior distributions for the effects of drought and irrigation on the intercept and slope of the NDVI-soil moisture relationship. Treatment effects show the difference between the coefficients estimated in the treated plots and the control plots. Probabilities (Pr </> 0 =) for each coefficient indicate the one-tailed probability that the coefficient is less than or greater than zero, depending on whether the median of the distribution is less than or greater than zero. The posterior densities were smoothed for visual clarity by increasing kernel bandwidth by a factor of five.

¹⁶ Section A4.1.1 Stan Code

```
data {
  int<lower=0> Npreds;
                         # number of covariates, including intercept
  int<lower=0> Npreds2;
                            # number of random effect covariates
 int<lower=0> Nplots;
                            # number of plots
 int<lower=0> Ntreats;
                            # number of treatments
 int<lower=0> Nobs;
                            # number of observations
                       # number of years
 int<lower=0> Nyears;
 vector[Nobs] y;
                            # vector of observations
 row_vector[Npreds] x[Nobs]; # design matrix for fixed effects
 row_vector[Npreds2] z[Nobs]; # simple design matrix for random effects
 int plot_id[Nobs];
                            # vector of plot ids
 int treat_id[Nobs];
                            # vector of treatment ids
 int year_id[Nobs];  # vector of year ids
}
parameters {
   vector[Npreds] beta;
                                     # overall coefficients
   vector[Nyears] year_off;  # vector of year effects
   cholesky_factor_corr[Npreds2] L_u; # cholesky factor of plot random effect corr matrix
 vector[Npreds2] beta_plot[Nplots]; # plot level random effects
   vector<lower=0>[Npreds2] sigma_u; # plot random effect std. dev.
                                     # treatment-level observation std. dev.
   real<lower=0> sd_y;
   real<lower=0> sigma_year;
                                   # year std. dev. hyperprior
   real<lower=0> phi;
                                     # dispersion parameter
}
transformed parameters {
 vector[Nobs] A;
                           # parameter for beta distn
 vector[Nobs] B;
                      # parameter for beta distn
 vector[Nobs] yhat;
                                 # vector of expected values
 vector[Npreds2] u[Nplots]; # transformed plot random effects
 matrix[Npreds2,Npreds2] Sigma_u; # plot ranef cov matrix
 Sigma_u = diag_pre_multiply(sigma_u, L_u); # cholesky factor for plot-level covariance matri
 for(j in 1:Nplots)
   u[j] = Sigma_u * beta_plot[j]; # plot random intercepts and slopes
```

```
# regression model for expected values (one for each plot-year)
  for (i in 1:Nobs)
    yhat[i] = inv_logit(x[i]*beta + z[i]*u[plot_id[i]] + year_off[year_id[i]]);
 A = yhat * phi;
 B = (1.0 - yhat) * phi;
}
model {
  #### PRIORS
 phi \sim cauchy(0, 5);
  sigma_year ~ cauchy(0,2.5);
  sd_y \sim cauchy(0,2.5);
 year_off ~ normal(0, sigma_year); # priors on year effects, shared variance
                                    # priors on treatment coefficients
 beta \sim normal(0,5);
 L_u ~ lkj_corr_cholesky(2.0); # prior on the cholesky factor which controls the
                                    # correlation between plot level treatment effects
  sigma_u ~ cauchy(0,2.5);
 for(i in 1:Nplots)
        beta_plot[i] ~ normal(0,1); # plot-level coefficients for intercept and slope
   #### LIKELIHOOD
 y ~ beta(A, B); # observations vary according to beta distribution
}
generated quantities{
  corr_matrix[Npreds2] R = multiply_lower_tri_self_transpose(L_u);
 cov_matrix[Npreds2] V = quad_form_diag(R,sigma_u);
}
```