Expectation, Variance, Binomial, and Poisson proofs

by Andrew Rothman

Expectation:

1) Prove E[Y + X] = E[X] + E[Y]

$$E[Y + X] = \sum_{x} \sum_{y} (x + y)P(X = x, Y = y)$$

$$= \sum_{x} \sum_{y} xP(X = x, Y = y) + \sum_{x} \sum_{y} yP(X = x, Y = y)$$

$$= \sum_{x} xP(X = x) + \sum_{y} yP(Y = y)$$

$$= E[X] + E[Y]$$

2) Prove E[a + X] = a + E[X]

$$E[a+X] = \sum_{x} (a+x)P(X=x)$$

$$= \sum_{x} aP(X=x) + \sum_{x} xP(X=x)$$

$$= (a*1) + \sum_{x} xP(X=x)$$

$$= a + E[X]$$

3) Prove E[aX] = aE[X]

$$E[aX] = \sum_{x} axP(X = x)$$
$$= a\sum_{x} xP(X = x)$$
$$= aE[X]$$

Variance Proof:

1) Prove $Var[X] = E[X^2] - (E[X])^2$

$$Var[X] = E[(X - E[X])^{2}]$$

$$= E[X^{2} - 2XE[X] + (E[X])^{2}]$$

$$= E[X^{2}] - 2E[X]E[X] + (E[X])^{2}$$

$$= E[X^{2}] - 2(E[X])^{2} + (E[X])^{2}$$

$$= E[X^{2}] - (E[X])^{2}$$

2) Prove Var[a + X] = Var[X]

$$Var[a + X] = E[(a + X - E[a + X])^{2}]$$

$$= E[(a + X - a - E[X])^{2}]$$

$$= E[(X - E[X])^{2}]$$

$$= Var[X]$$

3) Prove $Var[aX] = a^2Var[X]$

$$Var[aX] = E[(aX - E[aX])^{2}]$$

$$= E[(aX - aE[X])^{2}]$$

$$= E[a^{2}(X - E[X])^{2}]$$

$$= a^{2}E[(X - E[X])^{2}]$$

$$= a^{2}Var[X]$$

4) Prove Var[X + Y] = Var[X] + Var[Y] + 2Cov(X, Y)

$$Var[X + Y] = E[(X + Y - E[X + Y])^{2}]$$

$$= E[(X + Y - E[X] - E[Y])^{2}]$$

$$= E[XX + XY - XE[X] - XE[Y] + YX + YY - YE[X] - YE[Y] - XE[X] - YE[X] + E[X]E[X] + E[X]E[Y] - XE[Y] - YE[Y] + E[Y]E[X] + E[Y]E[Y]]$$

$$= (E[X^{2} - 2XE[X] + E[X]^{2}]) + (E[Y^{2} - 2XE[Y] + E[X]^{2}]) + 2E[XY - XE[Y] - YE[X] + E[X]E[Y]]$$

$$= E[(X - E[X])^{2}] + E[(Y - E[Y])^{2}] + 2E[(X - E[X])(Y - E[Y])]$$

$$= Var[X] + Var[Y] + 2Cov(X, Y)$$

Covariance Proof:

1) Prove Cov(X,Y) = E[XY] - E[X]E[Y]

$$Covar(X,Y) = E[(X - E[X])(Y - E[Y])]$$

$$= E[XY - XE[Y] - E[X]Y + E[X]E[Y]]$$

$$= E[XY] - 2E[X]E[Y] + [X]E[Y]$$

$$= E[XY] - E[X]E[Y]$$

2) Prove Cov(X + Y, Z) = Cov(X, Z) + Cov(Y, Z)

$$\begin{aligned} Covar(X+Y,Z) &= E[(X+Y-E[X+Y])(Z-E[Z])] \\ &= E[XZ+YZ-E[X]Z-E[Y]Z-XE[Z]-YE[Z]+E[X]E[Z]+E[Y]E[Z]] \\ &= E[XZ]+E[YZ]-2E[X]E[Z]-2E[Y]E[Z]+E[X]E[Z]+E[Y]E[Z] \\ &= E[XZ]-E[X]E[Z]+E[YZ]-E[Y]E[Z] \\ &= Cov(X,Z)+Cov(Y,Z) \end{aligned}$$

Bernoulli Distribution:

PMF:
$$f(k; p) = p^k (1-p)^{1-k}$$

Parameters: $p \in [0,1]$

Support: $k \in \{0,1\}$

1) $X \sim Bern(p)$, prove E[X] = p

$$E[X] = \sum_{x=0}^{1} x p^{x} (1-p)^{1-x}$$

$$= (0)p^{0} (1-p)^{1} + (1)p^{1} (1-p)^{0}$$

$$= 0 + p$$

$$= p$$

2) $X \sim Bern(p)$, prove Var[X] = p(1-p)

$$Var[X] = E[X^2] - (E[X])^2$$

$$E[X^{2}] = \sum_{x=0}^{1} x^{2} p^{x} (1-p)^{1-x}$$

$$= (0)p^{0} (1-p)^{1} + (1)p^{1} (1-p)^{0}$$

$$= 0 + p$$

$$= p$$

$$Var[X] = p - (p)^2$$
$$= p(1-p)$$

Binomial Distribution:

PMF:
$$f(k; n, p) = \binom{n}{k} p^k (1 - p)^{n-k}$$

Parameters: $p \in [0,1]$

$$n \in \{0, 1, 2, ..., \infty\}$$

Support: $k \in \{0,1,2,...,n\}$

1) $X \sim Binom(n, p)$, prove E[X] = np

Approach #1:

$$E[X] = \sum_{x=0}^{n} x \binom{n}{x} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=1}^{n} x \binom{n}{x} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=1}^{n} \binom{n}{x-1} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=1}^{n} n \binom{n-1}{x-1} p^{x} (1-p)^{n-x}$$

$$= np \sum_{x=1}^{n} \binom{n-1}{x-1} p^{x-1} (1-p)^{n-x}$$

$$= np \sum_{x=1}^{n} \binom{n-1}{x-1} p^{x-1} (1-p)^{(n-1)-(x-1)}$$

$$= np \sum_{j=0}^{m} \binom{m}{j} p^{j} (1-p)^{m-j} = np * (1)$$

$$= np$$

Approach #2:

Note that with $Y \sim Binom(n,p)$, $Y = \sum_{i=1}^n X$, where we have "n" iid $X \sim Bern(p)$

$$Y = \sum_{i=1}^{n} X$$

$$E[Y] = E\left[\sum_{i=1}^{n} X\right] = \sum_{i=1}^{n} E[X] = \sum_{i=1}^{n} p = np$$

2) $X \sim Binom(n, p)$, prove Var[X] = np(1 - p)

$$Var[X] = E[X^2] - (E[X])^2$$

Note that:

$$E[X^{2}] = E[X^{2}] - E[X] + E[X]$$

$$= E[X^{2} - X] + E[X]$$

$$= E[X(X - 1)] + E[X]$$

$$E[X(X-1)] = \sum_{x=0}^{n} x(x-1) {n \choose x} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} x(x-1) {n \choose x} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} {n \choose x-2} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} n(n-1) {n-2 \choose x-2} p^{x} (1-p)^{n-x}$$

$$= \sum_{x=2}^{n} n(n-1) p^{2} {n-2 \choose x-2} p^{x-2} (1-p)^{(n-2)-(x-2)}$$

$$= n(n-1) p^{2} \sum_{j=0}^{m} {m \choose j} p^{j} (1-p)^{m-j} = n(n-1) p^{2}$$

$$E[X^{2}] = E[X(X - 1)] + E[X]$$

= $n(n - 1)p^{2} + np$

$$Var[X] = E[X^{2}] - (E[X])^{2}$$

$$= n(n-1)p^{2} + np - n^{2}p^{2}$$

$$= n^{2}p^{2} - np^{2} + np - n^{2}p^{2}$$

$$= np - np^{2}$$

$$= np(1-p)$$

Approach #2:

Note that with $Y \sim Binom(n, p)$, $Y = \sum_{i=1}^{n} X$, where we have "n" iid $X \sim Bern(p)$

$$Y = \sum_{i=1}^{n} X$$

$$Var[Y] = Var\left[\sum_{i=1}^{n} X\right] = \sum_{i=1}^{n} Var[X] = \sum_{i=1}^{n} p(1-p) = np(1-p)$$

Poisson Distribution:

PMF:
$$f(k; \lambda) = \frac{\lambda^k e^{\lambda}}{k!}$$

Parameters: $\lambda \in (0, \infty)$

Support: $k \in \{0,1,2,...\}$

1) $X \sim Poisson(\lambda)$, prove $E[X] = \lambda$

$$E[X] = \sum_{x=0}^{\infty} x \frac{\lambda^x e^{\lambda}}{x!}$$

$$= \sum_{x=1}^{\infty} x \frac{\lambda^x e^{\lambda}}{x!}$$

$$= \sum_{x=1}^{\infty} \frac{\lambda^x e^{\lambda}}{(x-1)!}$$

$$= \lambda \sum_{x=1}^{\infty} \frac{\lambda^x e^{\lambda}}{(x-1)!}$$

$$= \lambda \sum_{y=0}^{\infty} \frac{\lambda^y e^{\lambda}}{y!}$$

$$= \lambda$$

2) $X \sim Poisson(\lambda)$, prove $Var[X] = \lambda$

Note that:

$$E[X^{2}] = E[X^{2}] - E[X] + E[X]$$

$$= E[X^{2} - X] + E[X]$$

$$= E[X(X - 1)] + E[X]$$

$$E[X(X-1)] = \sum_{x=0}^{\infty} x(x-1) \frac{\lambda^x e^{\lambda}}{x!}$$

$$= \sum_{x=2}^{\infty} x(x-1) \frac{\lambda^x e^{\lambda}}{x!}$$

$$= \sum_{x=2}^{\infty} \frac{\lambda^x e^{\lambda}}{(x-2)!}$$

$$= \lambda^2 \sum_{x=2}^{\infty} \frac{\lambda^{x-2} e^{\lambda}}{(x-2)!}$$

$$= \lambda^2 \sum_{y=0}^{\infty} \frac{\lambda^y e^{\lambda}}{y!}$$

$$= \lambda^2$$

$$Var[X] = E[X^2] - (E[X])^2 = E[X(X-1)] + E[X] - (E[X])^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$$

Sample mean Proof:

$$x_1, x_2, \dots, x_n \sim X$$
, with $E[X] = \mu$

1) Prove that the sample mean $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$ is an unbiased estimate of μ

$$\begin{split} E[\bar{x}] &= E\left[\frac{1}{n} \sum_{i=1}^{n} x_i\right] \\ &= \frac{1}{n} E[(x_1 + x_2 + \dots + x_n)] \\ &= \frac{1}{n} (E[x_1] + E[x_2] + \dots + E[x_n]) \\ &= \frac{1}{n} (\mu + \mu + \dots + \mu) \\ &= \frac{n\mu}{n} = \mu \end{split}$$

Sample variance Proof:

$$x_1, x_2, \dots, x_n \sim X$$
, with $Var[X] = \sigma^2$

1) Prove that $s^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2$ is an unbiased estimate of σ^2 when μ is known:

$$E[s^{2}] = E\left[\frac{1}{n}\sum_{i=1}^{n}(x_{i} - \mu)^{2}\right]$$

$$= \frac{1}{n}E[(x_{1} - \mu)^{2} + (x_{2} - \mu)^{2} + \dots + (x_{n} - \mu)^{2}]$$

$$= \frac{1}{n}(E[(x_{1} - \mu)^{2}] + E[(x_{2} - \mu)^{2}] + \dots + E[(x_{n} - \mu)^{2}])$$

$$= \frac{1}{n}(\sigma^{2} + \sigma^{2} \dots + \sigma^{2})$$

$$= \frac{n\sigma^{2}}{n} = \sigma^{2}$$

2) Prove that $s^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2$ is an unbiased estimate of σ^2 when μ is unknown:

$$\begin{split} E[s^2] &= E\left[\frac{1}{n-1}\sum_{i=1}^n(x_i - \bar{x})^2\right] \\ &= \frac{1}{n-1}E\left[\sum_{i=1}^nx_i^2 - 2x_i\bar{x} + \bar{x}^2\right] \\ &= \frac{1}{n-1}E\left[\left(\sum_{i=1}^nx_i^2\right) - \left(\sum_{i=1}^n2\bar{x}x_i\right) + \left(\sum_{i=1}^n\bar{x}^2\right)\right] \\ &= \frac{1}{n-1}E\left[\left(\sum_{i=1}^nx_i^2\right) - \left(2\bar{x}\sum_{i=1}^nx_i\right) + n\bar{x}^2\right] \\ &= \frac{1}{n-1}E\left[\left(\sum_{i=1}^nx_i^2\right) - 2n\bar{x}^2 + n\bar{x}^2\right] \\ &= \frac{1}{n-1}E\left[\left(\sum_{i=1}^nx_i^2\right) - n\bar{x}^2\right] \\ &= \frac{1}{n-1}\left(\left(\sum_{i=1}^nE[x_i^2]\right) - nE[\bar{x}^2]\right) \\ &= \frac{1}{n-1}\left(n(\sigma^2 + \mu^2) - nE[\bar{x}^2]\right) \\ &= \frac{n}{n-1}(\sigma^2 + \mu^2 - E[\bar{x}^2]) \end{split}$$

** note that:
$$Var[\bar{x}] = E[\bar{x}^2] - E[\bar{x}]^2$$

$$E[\bar{x}^2] = Var[\bar{x}] + E[\bar{x}]^2$$

$$= Var\left[\frac{1}{n}\sum_{i=1}^n x_i\right] + \mu^2$$

$$= \frac{1}{n^2} \left(\sum_{i=1}^n Var[x_i]\right) + \mu^2$$

$$= \frac{n\sigma^2}{n^2} + \mu^2 = \frac{\sigma^2}{n} + \mu^2$$

Therefore:

$$E[s^2] = \frac{n}{n-1} (\sigma^2 + \mu^2 - E[\bar{x}^2])$$

$$= \frac{n}{n-1} \left(\sigma^2 + \mu^2 - \frac{\sigma^2}{n} - \mu^2\right)$$

$$= \frac{n}{n-1} \left(\sigma^2 - \frac{\sigma^2}{n}\right)$$

$$= \frac{n}{n-1} \left(\frac{n\sigma^2 - \sigma^2}{n}\right)$$

$$= \frac{n}{n-1} \left(\frac{(n-1)\sigma^2}{n}\right) = \sigma^2$$