STATS 202 PS2

10/10/18

1) 10.7.2

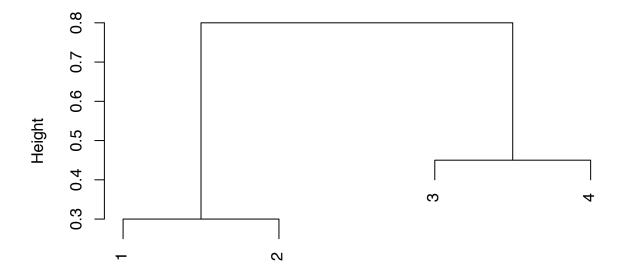
1a)

```
X = matrix(c(0, 0.3, 0.4, 0.7,
0.3, 0, 0.5, 0.8,
0.4, 0.5, 0, 0.45,
0.7, 0.8, 0.45, 0), nrow = 4)

hc.complete = hclust(as.dist(X), method="complete")

plot(hc.complete, main="Complete Linkage", xlab="", sub="")
```

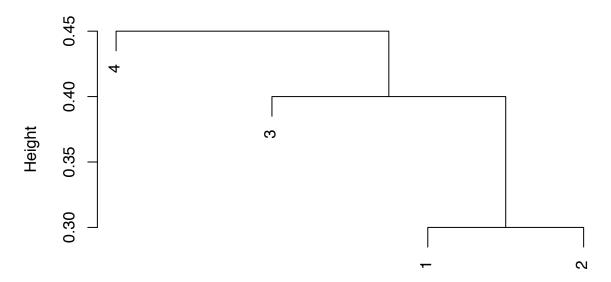
Complete Linkage



1b)

```
hc.simple = hclust(as.dist(X), method="single")
plot(hc.simple, main="Single Linkage", xlab="", sub="")
```

Single Linkage



1c)

Two clusters from 1(a): (1, 2), (3, 4)

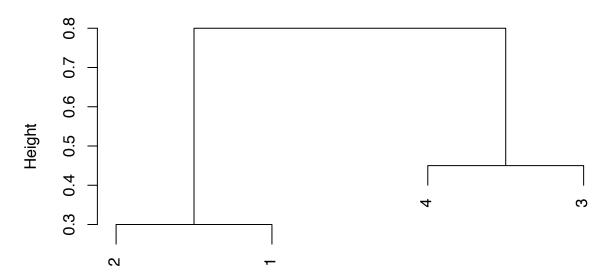
1d)

Two clusters from 1(b): (1, 2, 3), (4)

1e)

plot(hc.complete, main="Complete Linkage", xlab="", sub="", labels=c(2,1,4,3))

Complete Linkage



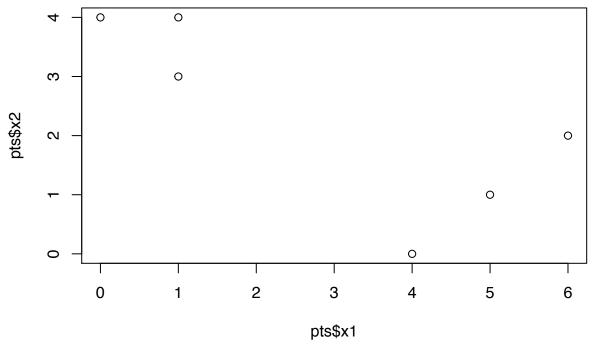
2) 10.7.3

2a)

```
set.seed(202)

pts = data.frame(
   obs = c(1:6),
   x1 = c(1,1,0,5,6,4),
   x2 = c(4,3,4,1,2,0)
)

plot(pts$x1, pts$x2)
```



2b)

```
pts$cluster = sample(c(1,2), 6, replace=TRUE)
pts
    obs x1 x2 cluster
##
        1
           4
## 1
      1
                    1
      2
                    2
## 2
         1
            3
## 3
      3 0 4
                    1
## 4
      4
         5 1
                    1
         6
            2
## 5
      5
                    1
## 6
      6 4 0
```

```
2c)
```

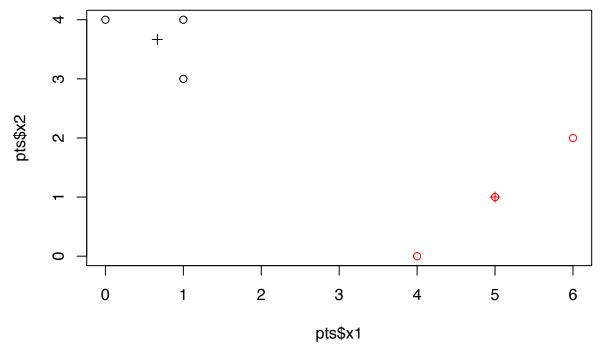
```
c1 = c(mean(pts[pts$cluster == 1, 2]), mean(pts[pts$cluster == 1, 3]))
c2 = c(mean(pts[pts$cluster == 2, 2]), mean(pts[pts$cluster == 2, 3]))
c1
## [1] 3.00 2.75
c2
## [1] 2.5 1.5
2d)
pts$d1 = ((pts$x1 - c1[1]) ^ 2) + ((pts$x2 - c1[2]) ^ 2)
pts$d2 = ((pts$x1 - c2[1]) ^ 2) + ((pts$x2 - c2[2]) ^ 2)
pts$cluster[(pts$d1 < pts$d2)] = 1</pre>
pts$cluster[(pts$d1 >= pts$d2)] = 2
pts
    obs x1 x2 cluster
                                d2
##
                           d1
                1 5.5625 8.5
## 1 1 1 4
## 2
      2 1 3
                    1 4.0625 4.5
## 3
     3 0 4
                   1 10.5625 12.5
## 4
     4 5 1
                    2 7.0625 6.5
      5 6 2
                    1 9.5625 12.5
## 5
## 6
      6 4 0
                    2 8.5625 4.5
2e)
It converges after another iteration:
c1 = c(mean(pts[pts$cluster == 1, 2]), mean(pts[pts$cluster == 1, 3]))
c2 = c(mean(pts[pts$cluster == 2, 2]), mean(pts[pts$cluster == 2, 3]))
pts$d1 = ((pts$x1 - c1[1]) ^ 2) + ((pts$x2 - c1[2]) ^ 2)
pts$d2 = ((pts$x1 - c2[1]) ^ 2) + ((pts$x2 - c2[2]) ^ 2)
pts$cluster[(pts$d1 < pts$d2)] = 1</pre>
pts$cluster[(pts$d1 >= pts$d2)] = 2
pts
##
     obs x1 x2 cluster
                           d1
                                d2
                    1 1.5625 24.5
## 1
      1 1 4
## 2
      2 1 3
                    1 1.0625 18.5
                    1 4.5625 32.5
## 3
      3 0 4
## 4
      4 5 1
                    2 14.0625 0.5
## 5
      5 6 2
                    2 17.5625 4.5
## 6
      6 4 0
                    2 14.5625 0.5
```

2f)

```
plot(pts$x1, pts$x2, col=pts$cluster)

c1 = c(mean(pts[pts$cluster == 1, 2]), mean(pts[pts$cluster == 1, 3]))
c2 = c(mean(pts[pts$cluster == 2, 2]), mean(pts[pts$cluster == 2, 3]))

points(c1[1], c1[2], col=1, pch = 3)
points(c2[1], c2[2], col=2, pch = 3)
```

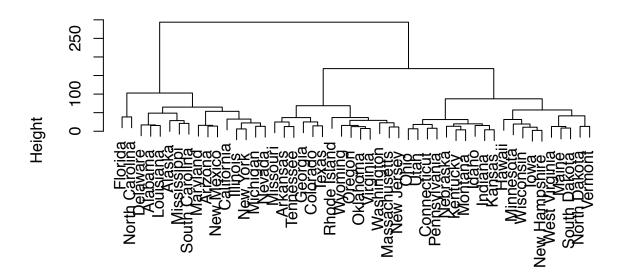


3) 10.7.9

3a)

```
hc.arrests.complete = hclust(dist(USArrests), method="complete")
plot(hc.arrests.complete, main="Complete Linkage", xlab="", sub="")
```

Complete Linkage



3b)

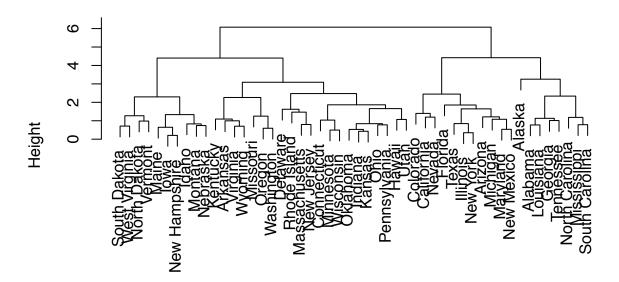
```
cutree(hc.arrests.complete, 3)
```

##	Alabama	Alaska	Arizona	Arkansas	California
##	1	1	1	2	1
##	Colorado	Connecticut	Delaware	Florida	Georgia
##	2	3	1	1	2
##	Hawaii	Idaho	Illinois	Indiana	Iowa
##	3	3	1	3	3
##	Kansas	Kentucky	Louisiana	Maine	Maryland
##	3	3	1	3	1
##	Massachusetts	Michigan	Minnesota	Mississippi	Missouri
##	2	1	3	1	2
##	Montana	Nebraska	Nevada	New Hampshire	New Jersey
##	3	3	1	3	2
##	New Mexico	New York	North Carolina	North Dakota	Ohio
##	1	1	1	3	3
##	Oklahoma	Oregon	Pennsylvania	Rhode Island	South Carolina
##	2	2	3	2	1
##	South Dakota	Tennessee	Texas	Utah	Vermont
##	3	2	2	3	3
##	Virginia	Washington	West Virginia	Wisconsin	Wyoming
##	2	2	3	3	2

3c)

```
arrests.sc = scale(USArrests)
hc.arrests.sc.complete = hclust(dist(arrests.sc), method="complete")
plot(hc.arrests.sc.complete, main="Complete Linkage", xlab="", sub="")
```

Complete Linkage



3d)

Scaling the variables has the effect of balancing the clustering branches and reducing the height of the tree; it is needed because the variables are measured in different units and have different variances. The variance for Assault is much higher than that of the other variables.

```
apply(USArrests , 2, var)

## Murder Assault UrbanPop Rape
## 18.97047 6945.16571 209.51878 87.72916
```

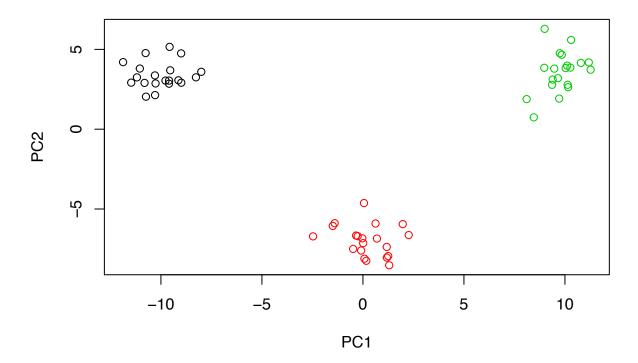
4) 10.7.10

4a)

```
X = matrix(rnorm(60*50), nrow = 60)
X[1:20] = X[1:20] + 10
X[21:40, 2] = X[21:40, 2] + 10
X[41:60, 1] = X[41:60, 1] - 10
```

4b) Perform PCA:

```
pr.out = prcomp(X)
plot(pr.out$x[,1:2], col=cbind(rep(1, 20), rep(2, 20), rep(3, 20)))
```



(4c)

Perform K-means clustering of the observations with K = 3.

```
km.out = kmeans(X, 3, nstart = 20)
table(km.out$cluster, cbind(rep(1,20), rep(2,20), rep(3,20)))
##
##
        1
           2
              3
##
     1 20
           0
              0
        0 20
              0
##
##
     3 0
           0 20
```

The clusters perform really well, and we get 3 separate clusters.

(4d) K = 2

There are two clusters, one which combines two of the original clusters.

$$(4e) K = 4$$

```
km.out = kmeans(X, 4, nstart = 20)
table(km.out$cluster, cbind(rep(1,20), rep(2,20), rep(3,20)))
##
##
        1
           2
              3
##
        0
           7
              0
##
     2
        0
           0 20
     3 20
           0
              0
##
##
     4
        0 13
              0
```

One of the original clusters splits into two new clusters while the other two remain intact.

(4f) K = 3, PCA vectors,

The clusters also perform really well, and we get 3 separate clusters as seen in the original data.

(4g) K = 3, scaled

Scaling worsens the clustering because scaling the observations down to the same stdev shrinks the distance between the observations.

5) 3.7.3

5a)

iii) For a fixed value of IQ and GPA, males earn more on average than females provided that the GPA is high enough.

Given fixed values for IQ and GPA, the effect of the coefficients on gender is $(\hat{\beta}_3 + \hat{\beta}_5 X_1)X_3$. If X_1 (GPA) is high enough, then this will be a negative effect for females, where $\hat{\beta}_3 = 35$ and $\hat{\beta}_5 = -10$, e.g if GPA = 4.0 then the net effect for females is negative.

5b)

```
50 + (20 * 4.0) + (0.07* 110) + (35*1) + (0.01 * 4.0 * 110) + (-10 * 1 * 4.0)
## [1] 137.1
5c)
```

False, the significance of a regression term has nothing to do with the size of the coefficient itself, but instead the p-value of the coefficient.

6) 3.7.10

6a)

```
library(ISLR)
fit = lm(Sales ~ Price + Urban + US, data = Carseats)
summary(fit)
##
## Call:
## lm(formula = Sales ~ Price + Urban + US, data = Carseats)
##
## Residuals:
##
      Min
                1Q Median
                                3Q
                                       Max
                           1.5786 7.0581
  -6.9206 -1.6220 -0.0564
##
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 13.043469
                           0.651012
                                    20.036
                                             < 2e-16 ***
## Price
               -0.054459
                           0.005242 - 10.389
                                             < 2e-16 ***
## UrbanYes
               -0.021916
                           0.271650
                                    -0.081
                                               0.936
                                      4.635 4.86e-06 ***
## USYes
                1.200573
                           0.259042
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.472 on 396 degrees of freedom
## Multiple R-squared: 0.2393, Adjusted R-squared: 0.2335
## F-statistic: 41.52 on 3 and 396 DF, p-value: < 2.2e-16
```

6b)

As the price increases by 1, the sales of car seats decreases by 54, and this is a significant relationship. From the regression p-values, we can see that the UrbanYes variable does not have a significant effect on sales. If the store is in the US, the sales of car seats increases by 1201, and this is a significant relationship. For a store that is not in the US and not urban, and if the price is 0, then they sell 13043 car seats.

6c)

```
Sales = 13.043469 + -0.054459 * Price + -0.021916 * UrbanYes + 1.200573 * USYes
```

6d)

We can reject the null hypothesis for Price and USYes because the p-values are below 0.05.

6e)

```
fit = lm(Sales ~ Price + US, data = Carseats)
summary(fit)
##
## Call:
## lm(formula = Sales ~ Price + US, data = Carseats)
##
## Residuals:
##
               1Q Median
                               ЗQ
      Min
                                      Max
## -6.9269 -1.6286 -0.0574 1.5766 7.0515
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 13.03079
                          0.63098 20.652 < 2e-16 ***
                          0.00523 -10.416 < 2e-16 ***
## Price
              -0.05448
## USYes
               1.19964
                          0.25846
                                   4.641 4.71e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 2.469 on 397 degrees of freedom
## Multiple R-squared: 0.2393, Adjusted R-squared: 0.2354
## F-statistic: 62.43 on 2 and 397 DF, p-value: < 2.2e-16
```

7) 3.7.14

7a)

```
set.seed(1)
x1=runif(100)
x2=0.5*x1+rnorm(100)/10
y=2+2*x1+0.3*x2+rnorm(100)
```

Linear model: $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$

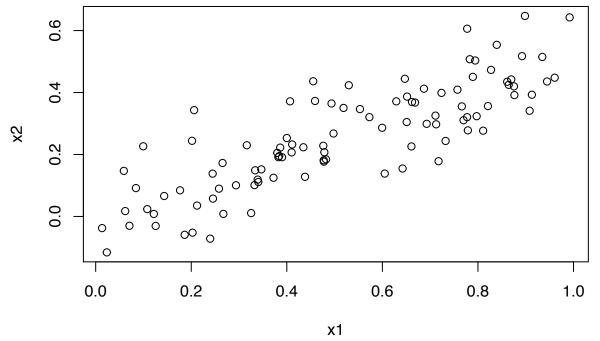
Regression coefficients: $\beta_0 = 2, \beta_1 = 2, \beta_2 = 0.3$

7b)

```
cor(x1, x2)

## [1] 0.8351212

plot(x1, x2)
```



7c)

```
fit = lm(y - x1 + x2)
summary(fit)
##
## Call:
## lm(formula = y \sim x1 + x2)
## Residuals:
                1Q Median
                                3Q
                                       Max
## -2.8311 -0.7273 -0.0537 0.6338 2.3359
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                            0.2319
                                     9.188 7.61e-15 ***
## (Intercept)
                 2.1305
## x1
                 1.4396
                            0.7212
                                     1.996
                                             0.0487 *
## x2
                 1.0097
                            1.1337
                                     0.891
                                             0.3754
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.056 on 97 degrees of freedom
## Multiple R-squared: 0.2088, Adjusted R-squared: 0.1925
```

```
## F-statistic: 12.8 on 2 and 97 DF, p-value: 1.164e-05 \hat{\beta}_0=2.1305, \hat{\beta}_1=1.4396, \hat{\beta}_2=1.0097
```

 β_0 has the closest estimate to the true model and is also reflected in the low p-value. β_1 has the next closest estimate, and we can reject the null because the p-value is less than 0.05. For β_2 we cannot reject the null because the p-value is greater than 0.05, and we can see that the value is off from the true original value.

d)

```
fit = lm(y \sim x1)
summary(fit)
##
## Call:
## lm(formula = y \sim x1)
##
## Residuals:
##
        Min
                   1Q
                        Median
                                      3Q
                                              Max
   -2.89495 -0.66874 -0.07785 0.59221
                                          2.45560
##
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                 2.1124
                             0.2307
                                       9.155 8.27e-15 ***
## x1
                  1.9759
                             0.3963
                                       4.986 2.66e-06 ***
##
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.055 on 98 degrees of freedom
## Multiple R-squared: 0.2024, Adjusted R-squared: 0.1942
## F-statistic: 24.86 on 1 and 98 DF, p-value: 2.661e-06
We can reject the null hypothesis \beta_1 = 0 because the p-value for that coefficient is less than 0.05.
e
```

```
fit = lm(y - x2)
summary(fit)
##
## Call:
## lm(formula = y \sim x2)
##
## Residuals:
##
       Min
                  1Q
                       Median
                                    3Q
                                             Max
  -2.62687 -0.75156 -0.03598 0.72383
                                        2.44890
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                 2.3899
                            0.1949
                                     12.26 < 2e-16 ***
## x2
                 2.8996
                            0.6330
                                      4.58 1.37e-05 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

```
##
## Residual standard error: 1.072 on 98 degrees of freedom
## Multiple R-squared: 0.1763, Adjusted R-squared: 0.1679
## F-statistic: 20.98 on 1 and 98 DF, p-value: 1.366e-05
```

We can reject the null hypothesis $\beta_2 = 0$ because the p-value for that coefficient is less than 0.05.

f)

The results do not contradict each other because x_1 and x_2 are collinear, and x_2 is generated directly from x_1 . Therefore they are both good predictors for y, are unreliable when combined in the linear model. Here we get x_1 to be more significant in the combined model, probably because it has the larger coefficient in the model compared to x_2 .

\mathbf{g}

```
x1=c(x1, 0.1)
x2=c(x2, 0.8)
y=c(y,6)
Model from pt (c):
fit = lm(y - x1 + x2)
summary(fit)
##
## Call:
## lm(formula = y \sim x1 + x2)
##
## Residuals:
##
       Min
                  1Q
                       Median
                                    3Q
                                             Max
## -2.73348 -0.69318 -0.05263 0.66385
                                        2.30619
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                 2.2267
                            0.2314
                                     9.624 7.91e-16 ***
## x1
                 0.5394
                            0.5922
                                     0.911 0.36458
## x2
                 2.5146
                            0.8977
                                     2.801 0.00614 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.075 on 98 degrees of freedom
## Multiple R-squared: 0.2188, Adjusted R-squared: 0.2029
## F-statistic: 13.72 on 2 and 98 DF, p-value: 5.564e-06
which.max(hatvalues(fit))
## 101
```

In this model, x2 has become the significant variable and x1 is no longer significant. We see that the new point did not decrease the R-squared value, so it is not an outlier. We can see that the new point (point #101) has high leverage.

Model from pt (d):

101

```
fit = lm(y - x1)
summary(fit)
##
## Call:
## lm(formula = y \sim x1)
##
## Residuals:
##
      Min
                1Q Median
                                3Q
                                        Max
## -2.8897 -0.6556 -0.0909 0.5682 3.5665
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
                 2.2569
                            0.2390
                                      9.445 1.78e-15 ***
## (Intercept)
                                      4.282 4.29e-05 ***
                 1.7657
                            0.4124
## x1
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.111 on 99 degrees of freedom
## Multiple R-squared: 0.1562, Adjusted R-squared: 0.1477
## F-statistic: 18.33 on 1 and 99 DF, p-value: 4.295e-05
which.max(hatvalues(fit))
## 27
## 27
We can see that the new point decreases the R-squared of the model, so it is an outlier. We can see that the
new point (point #101) does not have high leverage for this model.
Model from pt (e):
fit = lm(y \sim x2)
summary(fit)
##
## Call:
## lm(formula = y \sim x2)
##
## Residuals:
##
                  1Q
        Min
                       Median
                                     ЗQ
                                             Max
## -2.64729 -0.71021 -0.06899 0.72699
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
                            0.1912 12.264 < 2e-16 ***
## (Intercept)
                 2.3451
                 3.1190
                            0.6040
                                    5.164 1.25e-06 ***
## x2
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.074 on 99 degrees of freedom
## Multiple R-squared: 0.2122, Adjusted R-squared: 0.2042
## F-statistic: 26.66 on 1 and 99 DF, p-value: 1.253e-06
which.max(hatvalues(fit))
```

101

101

We can see that the new point does not the R-squared of the model, so it is not an outlier. We can see that the new point (point #101) has high leverage.