# Aula 15 – Anuidades Contínuas

Danilo Machado Pires

<u>danilo.pires@unifal-mg.edu.br</u>

Leonardo Henrique Costa

<u>Leonardo.costa@unifal-mg.edu.br</u>

https://atuaria.github.io/portalhalley

- ➤ Se imaginarmos que numa anuidade fracionada o número de frações cresce infinitamente, passamos a ter o que se pode designar por uma anuidade contínua.
  - > Pagamentos por hora, por minuto, por segundo,...etc.
  - $\triangleright$  Infinitos pagamentos ao longo do ano,  $m \to \infty$
- ➤ Importante notar que:

$$\ddot{a}_{\chi}^{(m)} = \frac{1}{m} + a_{\chi}^{(m)}$$

$$\lim_{m \to \infty} \ddot{a}_{\chi}^{(m)} = \lim_{m \to \infty} \left( \frac{1}{m} + a_{\chi}^{(m)} \right)$$

$$\bar{\ddot{a}}_{\chi} = \bar{a}_{\chi}$$

Comecemos por calcular o valor atuarial de uma anuidade continua, considerando uma taxa de capitalização constante:

$$e^{\delta} = 1 + i$$

$$v = (1 + i)^{-1}$$

$$e^{\delta} = \frac{1}{v}$$

 $\triangleright$  Considere uma anuidade com duração n pagamentos fracionados em m partes a taxa de rentabilidade i.

$$\ddot{a}_{\bar{n}|}^{(m)} = \frac{1}{m} + \frac{1}{m} v^{\frac{1}{m}} + \frac{1}{m} v^{\left(\frac{1}{m}\right)^2} + \frac{1}{m} v^{\left(\frac{1}{m}\right)^3} \dots + \frac{1}{m} v^{\left(\frac{1}{m}\right)^{mn-1}} = \frac{1}{m} \left( \frac{1 - \left(v^{\frac{1}{m}}\right)^{mn}}{1 - v^{\frac{1}{m}}} \right)$$

Cada ano "n" tem "'m" partes, assim  $m \times n$  atualizações

 $\triangleright$  Pensemos no que ocorre quando  $m \to \infty$ .

$$\lim_{m \to \infty} \ddot{a}_{\bar{n}|}^{(m)} = \lim_{m \to \infty} \frac{1}{m} \left( \frac{1 - v^n}{1 - v^{\bar{m}}} \right) = (1 - v^n) \lim_{m \to \infty} \left( \frac{\frac{1}{\bar{m}}}{1 - v^{\bar{m}}} \right)$$

Por L' Hopital: 
$$\lim_{m\to\infty} \left(\frac{f(m)}{g(m)}\right) = \lim_{m\to\infty} \left(\frac{f'(m)}{g'(m)}\right)$$
, então:

$$f(m) = \frac{1}{m} \quad \mapsto \quad f'(m) = -\frac{1}{m^2}$$

$$g(m) = 1 - v^{\frac{1}{m}} \mapsto g'(m) = -e^{\frac{1}{m}\ln(v)} \left( -\frac{\ln(v)}{m^2} \right) = \frac{v^{\frac{1}{m}}(\ln(v))}{m^2}$$

$$v^{\frac{1}{m}} = e^{\frac{1}{m}\ln(v)}$$

$$\lim_{m \to \infty} \ddot{a}_{\bar{n}|}^{(m)} = (1 - v^n) \lim_{m \to \infty} \left( -\frac{1}{v^{\frac{1}{m}}(\ln(v))} \right) = \frac{(1 - v^n)}{\ln(v)} \lim_{m \to \infty} \left( -\frac{1}{v^{\frac{1}{m}}} \right)$$

$$\lim_{m \to \infty} \ddot{a}_{\bar{n}|}^{(m)} = -\frac{(1 - v^n)}{\ln(v)} \lim_{m \to \infty} \left(\frac{1}{v^m}\right) = -\frac{(1 - v^n)}{\ln(v)} \left(\frac{1}{v^0}\right)$$

$$\lim_{m \to \infty} \ddot{a}_{\bar{n}|}^{(m)} = -\frac{(1 - v^n)}{\ln(v)} = -\frac{(1 - v^n)}{\ln(e^{-\delta})} = \frac{(1 - v^n)}{\delta}$$
$$\bar{\ddot{a}}_{\bar{n}|} = \frac{(1 - v^n)}{\delta}$$

> Assim para um T aleatório :

$$\overline{\ddot{a}}_{|\overline{T}|} = \frac{\left(1 - e^{-\delta T}\right)}{\delta} = \overline{a}_{|\overline{T}|}$$

P Que é o valor presente de um fluxo contínuo de pagamentos entre [0,t].

> O <u>valor presente atuarial contínuo</u> de anuidades vitalícias por ser calculada por:

$$\bar{a}_{x} = \int_{0}^{\infty} \frac{\left(1 - e^{-\delta t}\right)}{\delta} f_{T(x)}(t) dt$$

$$\bar{a}_x = \int_0^\infty \frac{\left(1 - e^{-\delta t}\right)}{\delta} t p_x \mu_{x+t} dt$$

Considerando que :

$$\bar{a}_{\bar{t}|} = \frac{(1-e^{-\delta t})}{\delta}$$
  $f_T(t) = {}_t p_x \mu_{x+t}$ 

 $\blacktriangleright$  a variância do valor presente de um fluxo contínuo de pagamentos em [0,t] à taxa de 1 real por ano, com juros  $\delta$  .

$$var(\bar{a}_{\bar{T}|}) = var\left[\frac{(1 - e^{-\delta T})}{\delta}\right]$$

$$var(\bar{a}_{\bar{T}|}) = \frac{var(1 - e^{-\delta T})}{\delta^{2}}$$

$$var(\bar{a}_{\bar{T}|}) = \frac{var(e^{-\delta T})}{\delta^{2}} \qquad e^{-\delta T} \equiv v^{T}$$

$$var(\bar{a}_{\bar{T}|}) = \frac{{}^{2}\bar{A}_{x} - (\bar{A}_{x})^{2}}{\delta^{2}}$$

$$var(v^{T}) = \int_{0}^{n} w^{t} f_{T}(t) dt - \left[\int_{0}^{n} v^{t} f_{T}(t) dt\right]^{2}$$

> Exemplo 20

Suponha que:

$$S_{T(0)}(t) = 1 - (1 - e^{-\alpha(t)})$$

Usando a taxa de juros  $\delta$ , calcule a esperança e variância de  $\bar{a}_{\bar{T}|}$  considerando uma pessoa de idade x.

$$\bar{a}_x = \int_0^\infty \bar{a}_{\bar{t}|\ t} \, p_x \mu_{x+t} dt$$

$$var(\bar{a}_{\bar{t}|}) = \frac{{}^2\bar{A}_x - (\bar{A}_x)^2}{\delta^2} = \frac{\int_0^\infty e^{-2\delta t} \, _t p_x \mu_{x+t} dt - \left(\int_0^\infty e^{-\delta t} \, _t p_x \mu_{x+t} dt\right)^2}{\delta^2}$$

#### Exemplo 20

$$\succ$$
 (i)

$$\overline{a}_{\overline{T}|} = \frac{\left(1 - e^{-\delta T}\right)}{\delta}$$

> (ii)

$$S_{T(0)}(t) = 1 - (1 - e^{-\alpha(t)})$$

$$S_{T(x)}(t) = P(T > t + x | T > x) = \frac{1 - \left(1 - e^{-\alpha(x+t)}\right)}{1 - \left(1 - e^{-\alpha x}\right)} = \frac{e^{-\alpha(x+t)}}{e^{-\alpha x}}$$

$$P(T > t + x | T > x) = {}_{t}p_{x} = \begin{cases} e^{-\alpha t}, & t > 0 \\ 1, & c.c \end{cases}$$

$$\succ$$
 (i)

$$\overline{a}_{\overline{T}|} = \frac{\left(1 - e^{-\delta T}\right)}{\delta}$$

> (ii)

$$P(T > t + x | T > x) = {}_{t}p_{x} = \begin{cases} e^{-\alpha t}, & t > 0 \\ 1, & c. c \end{cases}$$

> (iii)

$$\mu_{x+t} = -\frac{s'(x+t)}{s(x+t)} = \frac{f(x+t)}{1 - F(x+t)} = -\frac{\alpha e^{-\alpha(t+x)}}{1 - (1 - e^{-\alpha(x+t)})}$$

$$\mu_{x+t} = \frac{\alpha e^{-\alpha(t+x)}}{e^{-\alpha(x+t)}} = \alpha$$

#### > Exemplo 20

$$\bar{a}_{x} = \int_{0}^{\infty} \bar{a}_{\bar{t}|\ t} p_{x} \mu_{x+t} dt = \int_{0}^{\infty} \frac{(1 - e^{-\delta t})e^{-\alpha t} \alpha}{\delta} dt$$

$$\bar{a}_{x} = \frac{\alpha}{\delta} \int_{0}^{\infty} e^{-\alpha t} - e^{-t(\delta + \alpha)} dt$$

$$\bar{a}_{x} = \frac{\alpha}{\delta} \left[ -\frac{1}{\alpha e^{\alpha t}} + \frac{1}{(\delta + \alpha)e^{t(\delta + \alpha)}} \right]_{0}^{\infty}$$

$$\bar{a}_{x} = \frac{\alpha}{\delta} \left[ -\frac{1}{\alpha e^{\alpha t}} + \frac{1}{(\delta + \alpha)e^{t(\delta + \alpha)}} \right]_{0}^{\infty} = \frac{\alpha}{\delta} \left[ -\frac{1}{(\delta + \alpha)} + \frac{1}{\alpha} \right]$$

$$\overline{a}_{x} = \frac{1}{\delta + \alpha} \qquad \qquad \mu_{x+t} = \alpha$$

#### > Exemplo 20

$$var(\bar{a}_{\bar{t}}|) = \frac{var(e^{-\delta t})}{\delta^2} = \frac{{}^2\bar{A}_{\chi} - (\bar{A}_{\chi})^2}{\delta^2}$$

$$\bar{A}_{\chi} = \int_0^{\infty} e^{-\delta t} {}_t p_{\chi} \mu_{\chi+t} dt = \int_0^{\infty} e^{-t\delta} e^{-\alpha t} \alpha dt$$

$$\bar{A}_{\chi} = \alpha \left[ -\frac{1}{(\delta + \alpha)e^{t(\delta + \alpha)}} \right]_0^{\infty} = \frac{\alpha}{\delta + \alpha}$$

$${}^2\bar{A}_{\chi} = \int_0^{\infty} e^{-2\delta t} {}_t p_{\chi} \mu_{\chi+t} dt = \int_0^{\infty} e^{-t2\delta} e^{-\alpha t} \alpha dt$$

$${}^2\bar{A}_{\chi} = \alpha \left[ -\frac{1}{(2\delta + \alpha)e^{t(2\delta + \alpha)}} \right]_0^{\infty} = \frac{\alpha}{2\delta + \alpha}$$

$$var(\bar{a}_{\bar{t}|}) = \frac{1}{\delta^2} \left[ \frac{\alpha}{2\delta + \alpha} - \left( \frac{\alpha}{\delta + \alpha} \right)^2 \right] = \frac{\alpha}{(2\delta + \alpha)(\delta + \alpha)^2}$$

 $\triangleright$  Qual a probabilidade de que o valor presente de uma anuidade exceda o valor presente esperado, para o caso do tempo de vida adicional ser exponencial com parâmetro  $\alpha$ ?

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(\frac{\left(1 - e^{-\delta T}\right)}{\delta} > \frac{1}{\delta + \alpha}\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(-e^{-\delta T} > \frac{\delta}{\delta + \alpha} - 1\right) = P\left(e^{-\delta T} < \frac{\alpha}{\delta + \alpha}\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(-\delta T < \ln\left(\frac{\alpha}{\delta + \alpha}\right)\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(T > -\frac{1}{\delta}\ln\left(\frac{\alpha}{\delta + \alpha}\right)\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_x) = P\left(T > -\frac{1}{\delta}\ln\left(\frac{\alpha}{\delta + \alpha}\right)\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = e^{-\alpha \left(-\frac{1}{\delta}\ln\left(\frac{\alpha}{\delta + \alpha}\right)\right)} = \left(\frac{\alpha}{\alpha + \delta}\right)^{\frac{\alpha}{\delta}}$$

Exemplo 21

Com  $T(x) \sim \exp(\alpha)$ , então  $\alpha = \mu_{x+t}$ , logo:

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(T > -\frac{1}{\delta} \ln\left(\frac{\mu_{x+t}}{\delta + \mu_{x+t}}\right)\right)$$

$$P(\bar{a}_{\bar{T}|} > \bar{a}_{x}) = P\left(T > -\frac{1}{\delta}\ln\left(\frac{\alpha}{\delta + \alpha}\right)\right) = e^{-\alpha\left(-\frac{1}{\delta}\ln\left(\frac{\alpha}{\delta + \alpha}\right)\right)} = \left(\frac{\alpha}{\alpha + \delta}\right)^{\frac{\alpha}{\delta}}$$

Exemplo 21

$$P\left(\overline{a}_{\overline{T}|} > \overline{a}_{x}\right) = P\left(T > -\frac{1}{\delta}\ln\left(\frac{\mu_{x+t}}{\delta + \mu_{x+t}}\right)\right) = e^{-\mu_{x+t}\left(-\frac{1}{\delta}\ln\left(\frac{\mu_{x+t}}{\delta + \mu_{x+t}}\right)\right)} = \left(\frac{\mu_{x+t}}{\delta + \mu_{x+t}}\right)^{\frac{\mu_{x+t}}{\delta}}$$

 $\triangleright$  Considerando  $\delta = 0.10$  e a força de mortalidade igual a 0.016

$$P(\bar{a}_{\bar{T}|} > \bar{a}_x) = \left(\frac{0,016}{0,016 + 0,10}\right)^{\frac{0,016}{0,1}} = 0,7283$$

 $\triangleright$  Considerando  $\delta = 0.01 \,\mathrm{e}\,\mu_{x+t} = 0.033$ :

$$P(\bar{a}_{\bar{T}|} > \bar{a}_x) = \left(\frac{0,033}{0,033 + 0,10}\right)^{\frac{0,033}{0,1}} = 0,4174$$

Qual a probabilidade de que o valor presente de uma anuidade seja menor que um dador valor "Pr"?

$$F(Pr) = P(\bar{a}_{\bar{T}|} \leq \Pi)$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P\left(\frac{1 - e^{-\delta T}}{\delta} \leq \Pi\right)$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P(1 - e^{-\delta T} \leq \delta \Pi)$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P(-e^{-\delta T} \leq \delta \Pi - 1) = P(e^{-\delta T} \geq 1 - \delta \Pi)$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P[-\delta T \geq \ln(1 - \delta \Pi)]$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P\left[-T \geq \frac{\ln(1 - \delta \Pi)}{\delta}\right]$$

$$P(\bar{a}_{\bar{T}|} \leq \Pi) = P\left[T \leq -\frac{\ln(1 - \delta \Pi)}{\delta}\right] = F_T\left(-\frac{\ln(1 - \delta \Pi)}{\delta}\right)$$

> O valor presente atuarial contínuo de vitalícia por ser calculada por:

$$\bar{a}_{x} = \int_{0}^{\infty} \frac{\left(1 - e^{-\delta t}\right)}{\delta} t p_{x} \mu_{x+t} dt$$

$$\bar{a}_{x} = \int_{0}^{\infty} e^{-\delta t} \,_{t} p_{x} dt$$

> Exemplo 22

Suponha que :

$$S_{T(0)}(t) = 1 - (1 - e^{-\alpha(t)})$$

Usando a taxa de juros  $\delta$ , calcule  $\bar{a}_{\bar{T}|}$  considerando uma pessoa de idade x. Utilize  $\bar{a}_x = \int_0^\infty e^{-\delta t} \ _t p_x dt$ .

$$S_{T(0)}(t) = 1 - (1 - e^{-\alpha(t)})$$

$$S_{T(x)}(t) = P(T > t + x | T > x) = \frac{1 - (1 - e^{-\alpha(x+t)})}{1 - (1 - e^{-\alpha x})} = \frac{e^{-\alpha(x+t)}}{e^{-\alpha x}}$$

$$P(T>t+x|T>x)={}_{t}p_{x}=\begin{cases}e^{-\alpha t}, & t>0\\ 1, & c.c\end{cases}$$

$$\overline{a}_{x} = \int_{0}^{\infty} e^{-\delta t} p_{x} dt = \int_{0}^{\infty} e^{-\delta t} e^{-\alpha t} dt$$

$$\overline{a}_{x} = \int_{0}^{\infty} e^{-t(\delta + \alpha)} dt$$

$$\overline{a}_{x} = \left[ -\frac{1}{(\delta + \alpha)e^{t(\delta + \alpha)}} \right]_{0}^{\infty}$$

$$\overline{a}_{\chi} = \frac{1}{(\delta + \alpha)}$$

 $\triangleright$  Prêmio puro único para anuidades contínuas em um período de cobertura n.

$$Y = \begin{cases} \bar{a}_{\bar{T}|} & se \ 0 \le T < n \\ \bar{a}_{\bar{n}|} & se \ T \ge n \end{cases}$$

$$\bar{a}_{x:\bar{n}|} = E(Y) = \int_0^n \bar{a}_{\bar{t}|\ t} p_x \mu_{x+t} dt + \int_n^\infty \bar{a}_{\bar{n}|\ t} p_x \mu_{x+t} dt$$

$$\overline{a}_{x:\overline{n}|} = \int_0^n \overline{a}_{\overline{t}|t} p_x \mu_{x+t} dt + \overline{a}_{\overline{n}|n} p_x$$

$$\overline{a}_{x:\overline{n}|} = \int_0^n e^{-\delta t} \, _t p_x dt$$