rereception ring Work for infeatry	J-1		
separable D! Not linearly separable:	Gradient computation		
map to higher dim feature space, in which it is linearly separable. Normalize datapoints is good. Terminates after	$\nabla_{w(l)} E_i = r^{(l+1)} \mathbf{z}^{(l)} , \ \nabla_{w_q^{(1)}} E_i = r_q^{(1)} \mathbf{x}$		
no more than $1/\gamma^2$ updates, where $\gamma = \min_{i=1,,n} t_i \mathbf{w}_*^T \tilde{\phi}_i$	For b: keep residual only ! (i.e $\mathbf{x} = 1$)		
Normalize features $\phi_i \to \tilde{\phi_i}$!	Momentum learning		
Update rule (when misclassified	_		
$t_i \mathbf{w}^T \tilde{\phi}_i \leq 0$): $\mathbf{w} \leftarrow \mathbf{w} + t_i \tilde{\phi}_i$ until $t_i \mathbf{w}^T \tilde{\phi}_i > 0$ for all i	$\Delta \mathbf{w}_k = \mathbf{w}_{k+1} - \mathbf{w}_k$		
· '-	$\Delta \mathbf{w}_k = -\eta (1 - \mu) \nabla_{\mathbf{w}_k} E + \mu \Delta \mathbf{w}_{k-1}$		
Stochastic Gradient Descent	$\eta = \text{learning rate e.g. } 1/k, \mu = \text{momen}$		
$\mathbf{w}_{k+1} = \mathbf{w}_k - \eta \nabla_{\mathbf{w}} E_i(\mathbf{w}_k)$	tum term		
Least square estimation	Linear Regression. LSE $E(a,b) = \frac{1}{2} \sum_{i=1}^{n} (y(x_i) - t_i)^2 \frac{\partial E}{\partial a} =$		
$1 \sum_{n=1}^{n} (1 + 1)^n$	$E(a,b) = \frac{1}{2} \sum_{i=1}^{n} (y(x_i) - t_i)^2 \frac{\partial E}{\partial a} = \sum_{i=1}^{n} (y(x_i) - t_i)^2 \frac{\partial E}{\partial a}$		
$E(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^{n} (y(\mathbf{x}_i) - t_i)^2$	$\sum_{i} (ax_{i} + b - t_{i})x_{i} = n(a\langle x^{2} \rangle + b\langle x \rangle - b\langle x \rangle)$		
1 0			
$=\frac{1}{2}\ \mathbf{\Phi}\mathbf{w}-\mathbf{t}\ ^2$	Overfit: add tikhonov regularizer		
Gradient for Squared Error (yields normal equation for least squares if set	Univariate Linear Regression $\langle x \rangle = n^{-1} \sum_i x_i$. Similar for t, tx, x^2		
to 0)	$\langle x \rangle = n$ $\sum_i x_i$. Similar for i, ix, x		
,	$y_i = wx_i + b$		
$\nabla_{\mathbf{w}} E = \sum_{i=1}^{n} \frac{\partial E}{\partial y_i} \nabla_{\mathbf{w}} y_i = \Phi^T (\underbrace{\Phi \mathbf{w} - \mathbf{t}}_{residual})$	$\rightarrow w = \frac{Cov(x,t)}{Var(x)}, b = \langle t \rangle - w \langle x \rangle$		
Multi-Layer Perceptron	Normal Equations $(\Phi^T \Phi) \mathbf{w} = \Phi^T \mathbf{t}$		
Overfitting: Split data, early stopping.	\(\frac{1}{2}\) \(\frac{1}{2}\		
Parameters: w and b. Layers: Do	$\hat{\mathbf{w}} = \operatorname*{argmin}_{w} E(\mathbf{w}) = (\mathbf{\Phi}^{T} \mathbf{\Phi})^{-1} \mathbf{\Phi}^{T} \mathbf{t}$		
not count entry but count output. Error function: not good if not bounded be-			
low, not good ig grow too fast for large	Probability. Decision Theory		
residuals. Momentum : prevent zig zag,	Probability Independence : $p(a,b) = p(a,b)$		
allow higher learning rates.	$ p(a)p(b) E[X] = E[E[X Y]] $ Sum rule: $P(X) = \sum_{Y} P(X, Y) $		
Forward pass	Product rule: $P(X,Y) = P(X Y)P(Y)$		
$a_q^l = (\mathbf{w}_q^{(l)})^T \mathbf{z}^{(l-1)} + b_q^{(l)}$	Bayes $P(B F) = \frac{P(F B)P(B)}{P(F)}$ Var(t) = E[Var(t x)] + Var(E[t x])		
$z_q^{(l)} = g(a_q^{(l)}), q = 1,,h_l$	Bayes-Optimal Classifier Do the		

best choice.

Pattern Classification and Machine Learning

Perceptron Alg. ! Work for linearly

Linear Classification

Backward pass

 $r^{(L)} = \frac{\partial E_i}{\partial a^{(L)}} = \begin{cases} a^L - t_i & \text{for } E_{sq} \\ \sigma(a^L) - \tilde{t_i} & \text{for } E_{loc} \end{cases}$

 $r_q^{(l)} = g'(a_q^{(l)}) \sum_{i=1}^{n_{l+1}} w_{jq}^{(l+1)} r_j^{(l+1)}$

Covariance $Cov(\mathbf{x}, \mathbf{y}) = E[\mathbf{x}\mathbf{y}^T] - E[\mathbf{x}]E[\mathbf{y}], \quad Cov[\mathbf{A}\mathbf{x}] = \mathbf{A}Cov[\mathbf{x}]\mathbf{A}^T,$

 $\mathbf{S} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{x_i} - \bar{\mathbf{x}}) (\mathbf{x_i} - \bar{\mathbf{x}})^T = \frac{1}{n} \mathbf{X}^T \mathbf{X}$

Bayes error: $R = P\{f(x) \neq t\}$

Bayes error: when the classifier is

Bayes under Loss function Risk:

 $f^*(\mathbf{x}) = \underset{j \in \tau}{\operatorname{argmin}} \sum_{l \in \tau} L(j, k) P(t = k | \mathbf{x})$

 $R(f) = E[L(f(\mathbf{x}), t)]$ Opt. Classifier:

Optimal discriminant:

Sample

if $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i = 0$

derive using
$$\bar{x} = n^{-1} \sum_{i} x_{i}$$
, $p_{mixture}(x) = \sum_{k=1}^{blu} P(\omega_{k}) P(x|\gamma_{k})$

MLE $\hat{p}_{1} = \operatorname{argmax}_{p_{1} \in [0,1]} P(D|p_{1})$

Maximize $\log P(D|p_{1}) : \frac{d \log P(D|p_{1})}{dp_{1}} = 0$

O or minimize $-\log P(D|p_{1})$

Gaussian

$$N(x|\mu, \sigma^{2}) = \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{-\frac{(x-\mu)^{2}}{2\sigma^{2}}}$$

Multivariate $N(\mathbf{x}|\mu, \mathbf{\Sigma}) = \frac{1}{|2\pi\mathbf{\Sigma}|^{-1/2}} e^{-\frac{1}{2}(\mathbf{x}-\mu)^{T}} \mathbf{\Sigma}^{-1}(\mathbf{x}-\mu)$

Graph of the property of the proper

Naive Bayes Classifier $\hat{P}(\mathbf{x}|t=k)\hat{P}(t=k)=(\prod_{m=1}^{M}(\hat{p}_{m}^{(k)})^{\phi_{m}(\mathbf{x})}(1-\hat{p}_{m}^{(k)})^{1-\phi_{m}(\mathbf{x})})\frac{n_{k}}{2}$ $R^* = \mathbf{E} \begin{bmatrix} \min_{j \in \tau} \sum_{r, s} L(j, k) P(t = k | \mathbf{x}) \end{bmatrix} \begin{bmatrix} \hat{p}_m^{(k)} \right)^{1 - \phi_m(\mathbf{x})} \frac{\hat{n}_k}{n} \\ \hat{p}_m^{(k)} = \frac{\sum_{i=1}^n I_{\{t_i = k\}} \phi_m(\mathbf{x}_i)}{n} \end{bmatrix}$ Probab. Models. Max. Likelihood $P(\mathbf{x}|N, t = k) = \prod_{m=1}^{M} \left(p_m^{(k)}\right)^{\phi_m(\mathbf{x})}$ Likelihood function : $\prod_{i=1}^{n} p(x_i|\gamma)$, with $\phi_m(\mathbf{x}) = \sum_{i=1}^{N} I_{\{x_i = m\}}$ Generalization. Regularization Training error: $\hat{R}_n(f) = \frac{1}{n} \sum_{i=1}^n I_{\{(f(\mathbf{x_i}) \neq t_i)\}}$ Generalization error: $R(f) = E^*[I_{\{(f(\mathbf{x})\neq t\}}]$ Tikhonov Regularization term $\frac{\nu}{2}||\mathbf{w}||^2$ $\rightarrow (\Phi^T \Phi + \nu I) \mathbf{w} = \Phi^T \mathbf{t}$ $MAP \quad \hat{p_1} = \operatorname{argmax}_{p_1} p(p_1|D) =$ $\operatorname{argmax}_{p_1} P(D|p_1)p(p_1)$ Cond. Likelihood. Logistic Reg.

 $R^* = R(f^*) = 1 - E\left[\max_{k \in \tau} P(t = k|\mathbf{x})\right] \qquad \frac{Cov(x_j, x_k)}{\sqrt{Var(x_k)Var(x_k)}} \in [-1, 1]$

 $y^*(x) = \log \frac{p(\mathbf{x}|t=1)}{p(\mathbf{x}|t=0)} + \log \frac{P(t=1)}{P(t=0)} > 0 \rightarrow \hat{y}(\hat{\mathbf{x}}) = 0 \hat{\mathbf{w}}^T \mathbf{x} - \frac{1}{2}(||\hat{\mu}_{+1}||^2 - ||\hat{\mu}_{-1}||^2)$

ML plugin discriminant

 $+\log\frac{\hat{\pi}_1}{1-\hat{\pi}_1}$ where $\hat{\pi}_1=n_1/n$ and

 $y_k^*(\mathbf{x}) = -\frac{1}{2}||\mathbf{x} - \mu_k||^2 + \log P(t = k) + C$

Conditional

Likelihood $p(\mathbf{t}|\theta) \xrightarrow{} E(t,\theta)$ Loss func-

Logistic regression $P(t|y) = \sigma(ty)$

$$\begin{split} E_{log}(\mathbf{w}) &= -\log P(\mathbf{t}|\mathbf{w}) \\ &= \sum_{i=1}^{n} \log \left(1 + e^{-t_i y_i}\right) \\ &= \sum_{i=1}^{n} -\log \sigma(t_i y_i) \end{split}$$

Likelihood

Likelihood

Bridge

 $\hat{\mathbf{w}} = \hat{\mu}_{\perp 1} - \hat{\mu}_{-1}$

Generative Modeling $p(\mathbf{x}, t \theta) = p(\mathbf{x} t, \theta)P(t \theta)$ joint max. likeli.	Maximize Criterion for dual:	with $m_k = \mu^T \mu_k$. Maximize ratio!	surrogate criterion
$p(\mathbf{x} t,\theta)P(t \theta)$ joint max. likeli. $max_{\theta}\prod_{i=1}^{n}p(\mathbf{x}_{i},t_{i} \theta)$	$\phi_D(\alpha) = \sum_i \alpha_i - \frac{1}{2} \sum_{i,j} \alpha_i t_i K_{ij} t_j \alpha_j$	$S_B = (\hat{\mu}_1 - \hat{\mu}_0)(\hat{\mu}_1 - \hat{\mu}_0)^T = \mathbf{dd}^T$	$E(0, \{0, \}) = \sum_{i=1}^{n} E(0, 0_i)$
Discriminative Modeling	$2\sum_{i=1}^{n}$	n	$E(\theta; \{Q_i\}) = \sum_{i=1}^{n} E_i(\theta; Q_i)$
$P(t \mathbf{x}, \theta) \rightarrow max_{\theta} \prod_{i=1}^{n} P(t_{i} \mathbf{x}_{i}, \theta)$ conditional maximum likelihood	subj. to $\alpha_i \in [0, C], \sum_i \alpha_i t_i = 0$	$S_W = n^{-1} \sum_{\mathbf{i}} (\mathbf{x_i} - \hat{\mu}_{\mathbf{t_i}}) (\mathbf{x_i} - \hat{\mu}_{\mathbf{t_i}})^T$	n
Multiway logistic Regression	Discriminant $y^*(\mathbf{x}) = \mathbf{w_*}^T \phi(\mathbf{x}) =$	i=1	$= \sum_{i=1}^{n} E_{Q_i}[\log P(\mathbf{x_i}, \mathbf{h_i} \theta)]$
$P(t = k \mathbf{x}) = \frac{e^{y_k^*(\mathbf{x})}}{\sum_{k} y_k^*(\mathbf{x})} = \sigma_k(\mathbf{y}^*(\mathbf{x}))$	$\sum_{i=1}^{n} \alpha_{*,i} t_i K(\mathbf{x}, \mathbf{x}_i) + b_*$	$\hat{\mu}_{FLD} = \frac{S_W^{-1} \mathbf{d}}{ S_W^{-1} \mathbf{d} }, \ \mathbf{d} = \hat{\mu}_1 - \hat{\mu}_0$	
$\sum_{\tilde{i}_{i}} e^{-\kappa}$	$b = \frac{1}{ S } \sum_{i \in S} (t_i - \tilde{y}_i)$, S = essential support vectors	11 W 11	Perfect for missing data ! Hint: $\frac{\partial \log p(x_i)}{\partial \gamma_k} = \frac{1}{p(x_i)} \frac{\partial p(x_i \omega_k)P(\omega_k)}{\partial \gamma_k}$
Soft-max mapping $\sigma_k(\nu) = e^{\nu_k - lsexp(\nu)}$	Support vectors	Total covariance $S = S_W + \alpha(1 - \alpha) \mathbf{dd}^T$ with $\alpha = \frac{n_1}{n_2}$	$= \frac{\frac{\partial \gamma_k}{\partial x_i} - \frac{1}{p(x_i)} \frac{\partial \gamma_k}{\partial \gamma_k}}{\frac{1}{p(x_i)} \frac{\partial \log p(x_i \omega_k)}{\partial \gamma_k}}$
with $lsexp(\nu) = log \sum_{\tilde{k}} e^{\nu_{\tilde{k}}}$ $\nabla_{v} lsexp(\mathbf{v}) = \sigma(\mathbf{v})$		16	$= \frac{p(x_i)}{p(x_i)} \frac{\partial \gamma_k}{\partial x_i}$
Support Vector Machines	$\begin{cases} \alpha_i = 0 & 1 - t_i y_i \le 0 \text{ not} \\ \alpha_i \in (0, C) & 1 - t_i y_i = 0 \text{ essential} \end{cases}$	LDA Generalization for multiple classes: $S = S_W + S_B$, $S_B =$	$= P(\omega_k x_i) \frac{\partial \log p(x_i \omega_k)}{\partial \gamma_k} \text{Then insert derivative already computed (right)}$
Kernel function: $K(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^T \phi(\mathbf{x}')$ Gaussian: $e^{-\frac{\tau}{2} \mathbf{x}-\mathbf{x}' ^2}$, $\tau > 0$	$\alpha_i = C \qquad 1 - t_i y_i \ge 0 \ bound$	$n^{-1} \sum_{k} n_k \mathbf{d}_k \mathbf{d}_k^T, \ \mathbf{d}_k = \hat{\mu}_k - \hat{\mu}$	part)
Polynomial: $K(\mathbf{x}, \mathbf{x}') = (\mathbf{x}^T \mathbf{x}')^T$	Model Selection and Evaluation	$\max_{U \in \Re^{d \times M}} tr(U^T S_B U) \ s.t. \ U^T S_W U = I$	Beautiful Maths
Max margin perceptron	Bias-Variance Decomposition	$U \in \Re^{d \times M}$	Cauchy-Schwarz $ \mathbf{a}^T \mathbf{b} \le \mathbf{a} \mathbf{b} $
) E[(\hat{\hat{\hat{\hat{\hat{\hat{\hat{	Unsupervised Learning	Logistic function $\sigma(v) = \frac{1}{1+e^{-v}},$ $\sigma'(v) = \sigma(v)\sigma(-v) = \sigma(v)(1-\sigma(v))$
$\max_{w,b} \left\{ \gamma_D(\mathbf{w}, b) = \min_{i=1n} \frac{t_i(\mathbf{w}^T \phi(\mathbf{x}_i) + b)}{ \mathbf{w} } \right.$	$\begin{cases} E[(y(\mathbf{x} D) - E[t x]) \mid \mathbf{x}] = \\ (E[t] + E[t] + E[t])^2 + \mathbf{x} \cdot (\mathbf{x}(t E)) \end{cases}$	Kmeans Iterate until assignment no	
w, b ($t=1n$ \mathbf{w}	$\int \underbrace{\left(E[\hat{y}(\mathbf{x} D) \mathbf{x}] - E[t \mathbf{x}]\right)^2}_{} + \underbrace{Var(\hat{y}(x D) \mathbf{x})}_{}$	longer change. Assignment step:	tanh $g(a) = \tanh(a) = \frac{e^a - e^{-a}}{e^a + e^{-a}}$
! D linearly separable !	$Bias^2$ $Variance$	$ \mathbf{x_i} - \mu_{t_i} = \min_{k=1K} \mathbf{x_i} - \mu_k $ Update prototypes:	$g(a)' = 1 - g(a)^2$
Hard margin (convex optimization problem): $\min_{w,b} \frac{1}{2} \mathbf{w} ^2$ subj. to	Ens. meth. $\hat{y}_{ens}(\mathbf{x}) = \frac{1}{L} \sum_{l=1}^{L} \hat{y}_{l}(\mathbf{x})$	$\mu_k = n_k^{-1} \sum_{i=1}^n I_{\{t_i = k\}} \mathbf{x_i}$	Trace $tr(\mathbf{A}) = \sum_{j=1}^{d} a_{jj} = \sum_{j=1}^{d} \lambda_j$
$t_i(\mathbf{w}^T \phi(\mathbf{x_i}) + b) \ge 1, i = 1n$	CV $\hat{R}_{CV}^{(M)}(D) = \frac{1}{n} \sum_{i=1}^{n} L(\hat{y}_{\nu}^{-m(i)}, t_i)$	n K	
Soft margin SVM w^2 regul term	Dimensionality Reduction	$\phi(\mathbf{t}, \mu) = \sum_{i=1}^{n} \sum_{k=1}^{n} I_{\{t_i = k\}} \mathbf{x}_i - \mu_k ^2$	$\mathbf{x}^{T}\mathbf{A}\mathbf{x} = tr(\mathbf{x}^{T}\mathbf{A}\mathbf{x}) = tr(\mathbf{A}\mathbf{x}\mathbf{x}^{T})$
1	PCA First principal component direction v is the unit norm eigen-	i=1 $k=1$	Eigs $\mathbf{A}\mathbf{v} = \lambda\mathbf{v}, \ \mathbf{A} = \mathbf{U}\boldsymbol{\Lambda}\mathbf{U}^{-1}, A $
$\min_{w,b,\xi} \frac{1}{2} \mathbf{w} ^2 + C \sum_{i=1}^{n} \xi_i$	value corresponding to the largest eigen-	Gaussian Mixture Model Soft as-	$ \lambda I = 0$
i=1	value of S. Symmetry : empirical	signement: assign to clusters with prob- abilities. If one mixture component as-	$ A = \prod_{j=1}^{d} \lambda_{j}$
subj. to $t_i(\mathbf{w}^T \phi(\mathbf{x_i}) + b) \ge 1 - \xi_i, \ \xi_i \ge 0$	mean 0, can take half of the points. $\mathbf{S} = \frac{1}{2} \sum_{i=1}^{n} (\mathbf{x}_{i} - \bar{\mathbf{x}})(\mathbf{x}_{i} - \bar{\mathbf{x}})^{T} \bar{\mathbf{x}} = \mathbf{x}$	signed to a single central datapoint, its	Positive semi-definite matrix $\mathbf{A} \in \mathbb{R}^{d \times d}$ symmetric: $\mathbf{v}^T \mathbf{A} \mathbf{v} \geq 0 \ \forall \mathbf{v} \in \mathbb{R}^{d \times d}$
$1 \dots 1^{2} \cdot \sum_{i=1}^{n} t_{i}$	$\mathbf{S} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{x}_i - \bar{\mathbf{x}}) (\mathbf{x}_i - \bar{\mathbf{x}})^{\hat{T}} \bar{\mathbf{x}} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_i, \mathbf{z} = \mathbf{U}^{\mathbf{T}} \mathbf{x} with U^T U = \mathbf{U}^T \mathbf{x}$	variance will shrink to small values : EM	$\Re^d, \mathbf{v} \neq 0$
$\equiv \min_{w,b} \frac{1}{2C} \mathbf{w} ^2 + \sum_{i=1}^{n} [1 - t_i y_i]_+$	$I_{M \times M}$	diverge with lare likelihood values.	$\mathbf{Beta}(lpha,eta)$
$E_{snm}(w,b)$	$\mathbf{u}_* = \operatorname{argmax} \mathbf{u}^T \mathbf{S} \mathbf{u}$	$p(\mathbf{x}) = \sum_{k=1}^{K} p(\mathbf{x} t=k)P(t=k)$	$p(p_1 \alpha,\beta) = \frac{1}{B(\alpha,\beta)} (p_1)^{\alpha-1} (1-p_1)^{\beta-1}$
	$\mathbf{u}: \mathbf{u} =1$	$p(\mathbf{x}) = \sum_{k=1}^{\infty} p(\mathbf{x} t=k) 1 \ (t=k)$	$P(p_1 \alpha,\beta) = \frac{1}{B(\alpha,\beta)}(p_1) \qquad (1-p_1)$
Repr. Thm $\mathbf{w}_* = \sum_{i=1}^n \alpha_{*,i} \phi(x_i)$	PC directions = eigendirections of	K	With $B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$. Mode $\frac{\alpha-1}{\alpha+\beta-2}$
	$Cov(\mathbf{x})$ Goals:maximize $Cov(z)$ / Min-	$= \sum_{k=1} N(\mathbf{x} \mu_{\mathbf{k}}, \Sigma_k) P(t=k)$	Hinge function $[x]_+ = max(x,0)$
$\rightarrow y(\mathbf{x}) = \sum_{i=1}^{n} \alpha_i K(\mathbf{x}, \mathbf{x_i}) + b$	imize $E[\hat{x} - x ^2]$ / decorrelate components of z	κ≡1	Cross-entropy, divergence
Calution Drival/Dual	! PCA doesn't depends on labels t!	Compute: $n_k = \sum_{i=1}^n P(t_i = k \mathbf{x_i})$ Update: $\pi_k = \frac{n_k}{n}$,
Solution Primal/Dual: $p_* = \min_{\mathbf{w}, b} \max_{0 < \alpha_i < C} L(\mathbf{w}, b, \alpha)$	Fischer	Update: $\pi_k = \frac{1}{n}$ $\mu_k = \frac{1}{n_k} \sum_{i=1}^n P(t_i = k \mathbf{x_i}) \mathbf{x_i}$	$D(\mathbf{q} \mathbf{p}) = \sum_{l=1}^{L} q_l \log(\frac{q_l}{p_l}) \ge 0$
d_* max-min (reversed). Weak duality	$(m_1 - m_0)^2$ $\mathbf{u}^T S_D \mathbf{u}$	Expectation Maximization E-step:	$\lim_{l \to 1} \frac{\sum_{l=1}^{l} p_l}{p_l} = 0$
$d_* \le p_*$ (strong duality iff =) $L(\mathbf{w}, b, \alpha) = \frac{1}{2} \mathbf{w} ^2 + \sum_{i=1}^n \alpha_i (1 - t_i y_i) $	$J(\mathbf{u}) = \frac{(m_1 - m_0)^2}{s_0^2 + s_1^2} = \frac{\mathbf{u}^T S_B \mathbf{u}}{\mathbf{u}^T S_W \mathbf{u}}$	Expectation Maximization E-step: $Q_i(\mathbf{h_i}) \leftarrow P(\mathbf{h_i} \mathbf{x_i}, \theta)$ M-step: maximize	
$L(\mathbf{w}, v, \alpha_i) - \frac{1}{2} \mathbf{w} + \underline{\sum}_{i=1} \alpha_i (1 - \iota_i y_i) $	0 1 01 a SW a	******	I .