

Penalized regression for left-truncated & right-censored survival data
McGough, Incerri, Lyalin, Copping,
Narasimhan, Tibshirani (2021)
"Statistics in Medicine"

A summary
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Rationale: For some inclusion criteria
for time to survival data, one sometimes
include only patients with a genomic
test in a personalized health care context
(often the case for oncology medicine that
this context is applied). Plus with electronic
health records (aka, just lots of variables)
we have high dimensional data (because they are)

(Problem → Solution proposed)

High dimensionality → Penalized regression
such as lasso (l1) and
ridge (l2), elastic net
regression, smoothly clipped
absolute deviation
Partial likelihood at
time t : $\prod_{i=1}^n \frac{e^{x_i^T \beta}}{\sum e^{x_i^T \beta}}$

Left truncation
at time t

Overfitting → cross validation

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Things to avoid \rightarrow Preventative Measure in Analysis

"immortal time bias" \rightarrow
because observed patients
cannot die prior to
entering the study.

(that's a hard one
to avoid when
"letting" patients
into a survival analysis
study)

Poor prediction accuracy
limited $+$ generalizability

\rightarrow Don't over fit
add regularization
penalty to constrain
the size of B ,
reduce complexity
of model

avoid p predictors $>$
 n sample ($p < n$
is desirable)

Simulation study (methods)

1) Data generation of from Weibull.

T_i = latent survival time

U_i = Right censoring

V_i = study entry time

observed survival time = $Y_i = \min(T_i, U_i)$

R censored = $T_i > U_i$

L truncated = $V_i > Y_i$

2) Simulation of 10 binary predictors to an $n \times p$ matrix called X .

(11 predictors are taken from real-world dataset already, s.t. $p=21$)

$$\begin{pmatrix} \tilde{x}_{i,1}^* \\ \vdots \\ \tilde{x}_{i,p}^* \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_1 \\ \vdots \\ \mu_p \end{pmatrix}, \Sigma \right)$$

} multivariate probit approach

T_i , latent survival is taken from a Weibull survival function with density

$$f(t) = a m t^{a-1} e^{-m t^a}$$

where $m = \exp(\alpha + x^T \beta)$

s.t. $T_i \sim \text{Weibull}(\alpha, m_i)$

choice because when using Nelson-Aalen estimator of $\hat{Q}(t)$, cumulative hazard, overall survival is monotonically decreasing, >0 Weibull can capture baseline hazard

U_i , right censoring time is taken also from a Weibull distribution

V_i , study entry times is generated in 2 parts, first is the bernoulli random variable $d_i \sim \text{Bernoulli}(\pi)$ to input if V has status 1 or 0.

$V \sim \text{lognormal}(\mu, \sigma^2)$, median = 1 year
not sure rational for choice
maybe convenience from Normal distribution
as inverse cdf and pdf can be "easily" computed
and (2) skewed (see mean, median chosen)

3) Cross validation with 10 folds
where test data is 25% and training is 75%
from ~4500 observations of real-world data

3a) repeated 200 times

3b) from small and large p scenarios
 $p=21$ $p=1011$

4) Find λ that minimized partial likelihood deviance in (3)

5) plot calibration curves of survival probability of those with and without calibration curves. for small and large p scenario

6) Use C-index to compare fitted models with and without left truncation adjustment. the \uparrow C-index, the better the model is to discriminate prognosis (survival probability) between patients

7) with penalised Cox proportional model, compare hazard ratios between models that were and were not adjusted for left truncation

Results (1 example)

| Model | Covariate size | C-index | |
|-------------|----------------|------------|---------------|
| | | adjustment | no adjustment |
| Cox | small | 0.649 | 0.580 |
| | | 0.648 | 0.580 |
| Cox (lasso) | small | | |
| | | 0.625 | 0.556 |
| Cox | large | | |
| | | 0.663 | 0.600 |
| Cox (lasso) | large | | |

Annotations:
 - Red arrows point from 0.649 to 0.648 and from 0.580 to 0.580.
 - "not better" is written in red next to the arrow from 0.649 to 0.648.
 - "no diff" is written in red next to the arrow from 0.580 to 0.580.
 - Red arrows point from 0.625 to 0.663 and from 0.556 to 0.600.
 - "better" is written in red next to the arrow from 0.625 to 0.663.
 - "better" is written in red next to the arrow from 0.556 to 0.600.
 - The values 0.648, 0.580, 0.663, and 0.600 are circled in yellow.

Points of Discussion

- generally c-index is higher in model that does not adjust for left truncation? maybe because of bias in test = training sets. therefore careful understanding of the data generation process is crucial !
- this approach opens questions for future specification, interpretation and evaluation of prognostic survival models
- Calibration curves > c-index, maybe latter is more difficult to differentiate between models
- define risk set appropriately, difficult with Ⓛ truncated data ∴ when and how patients came to be included in the data
- potential selection bias because only stage IV diagnoses was included eg temporal selection bias, immortal time bias. Can be mitigated by studying the associations between left truncated time and survival time.

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