

An abstract network diagram with numerous small, colored nodes (blue, green, yellow, pink) connected by thin, light gray lines, forming a complex web-like structure that serves as a background for the text.

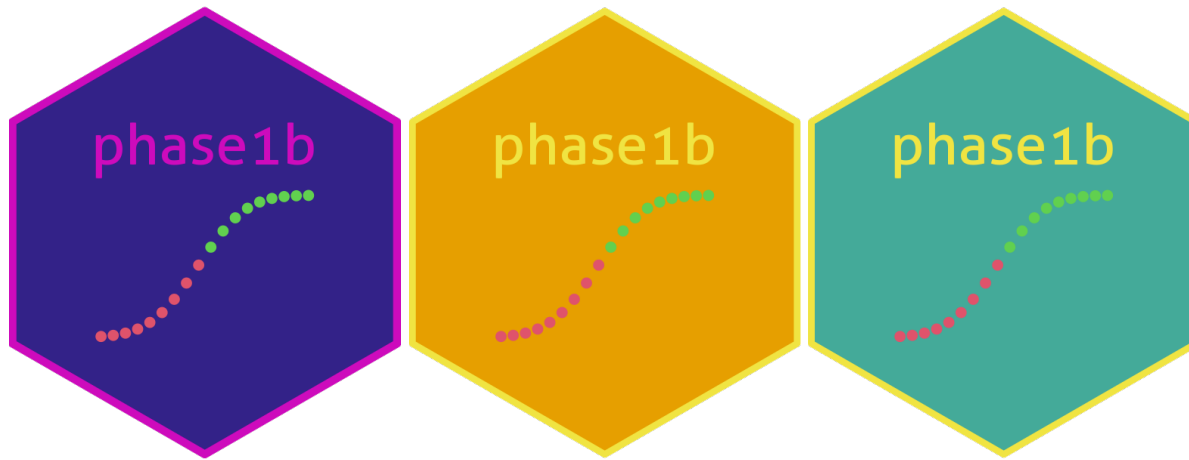
A Bayesian approach to decision making in early development clinical trials: An R solution

Audrey Yeo

Welwyn Statistics CoP, 2024

This presentation has ALT text and as much as possible, uses colour-blind friendly palettes

Early oncology trials and why **phase1b?**



Prior and Posterior of Beta Distribution for response rate

- Conjugate Prior is $f(\pi)$, where $\pi \sim \text{Beta}(\alpha, \beta)$, same family of distribution of Posterior (see below)
- We know the mean response rate (RR) is :

$$\pi = \frac{\alpha}{\alpha + \beta}$$

- Likelihood is $f(x|\pi)$, where $x \sim \text{Binomial}(x, n)$
- The updated Posterior $f(\pi|x)$ is again a *Beta* distribution (same family as prior) :

$$\pi | x \sim \text{Beta}(\alpha + x, \beta + n - x)$$

where x is the number of responders of current trials

History and how to install :

- 2015 : Started as a need in Roche's early development group, package development led by Daniel Sabanés Bové in 2015.
- 2023 : Refactoring, Renaming, adding Unit and Integration tests as current State-of-Art Software Engineering practice.
- 100% written in R and Open Source.
- website : genentech.github.io/phase1b/

```
1 library(devtools)
2 devtools::install_github("https://github.com/Genentech/phase1b")
3 library(phase1b)
```

Use case:

A single arm novel therapeutic with an assumed control response rate is at most 60%

Example	Interim	Final
Responders	16	23
n	23	40
Response rate	69.57 %	57.5 %
Posterior probability*	ask phase1b	ask phase1b
Predictive posterior probability*	ask phase1b	-
Decision to develop molecule further : Go/Stop/Grey Zone	ask phase1b	ask phase1b

* Posterior Probability : $P(\pi > 60\%|\alpha + x, \beta + n - x)$

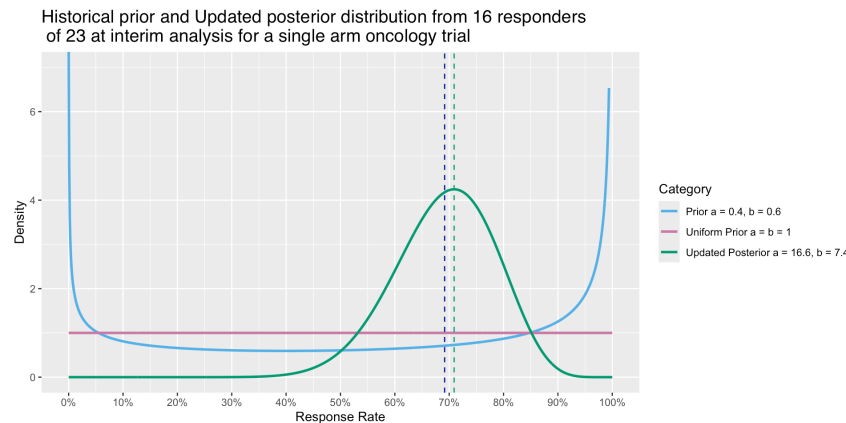
* Predictive Posterior Probability : $P(\text{success or failure at final})$

Updating the Posterior

- Using the formula for the mean, where $\alpha = 0.6$, $\beta = 0.4$ and at interim $x = 16$, $n = 23$:

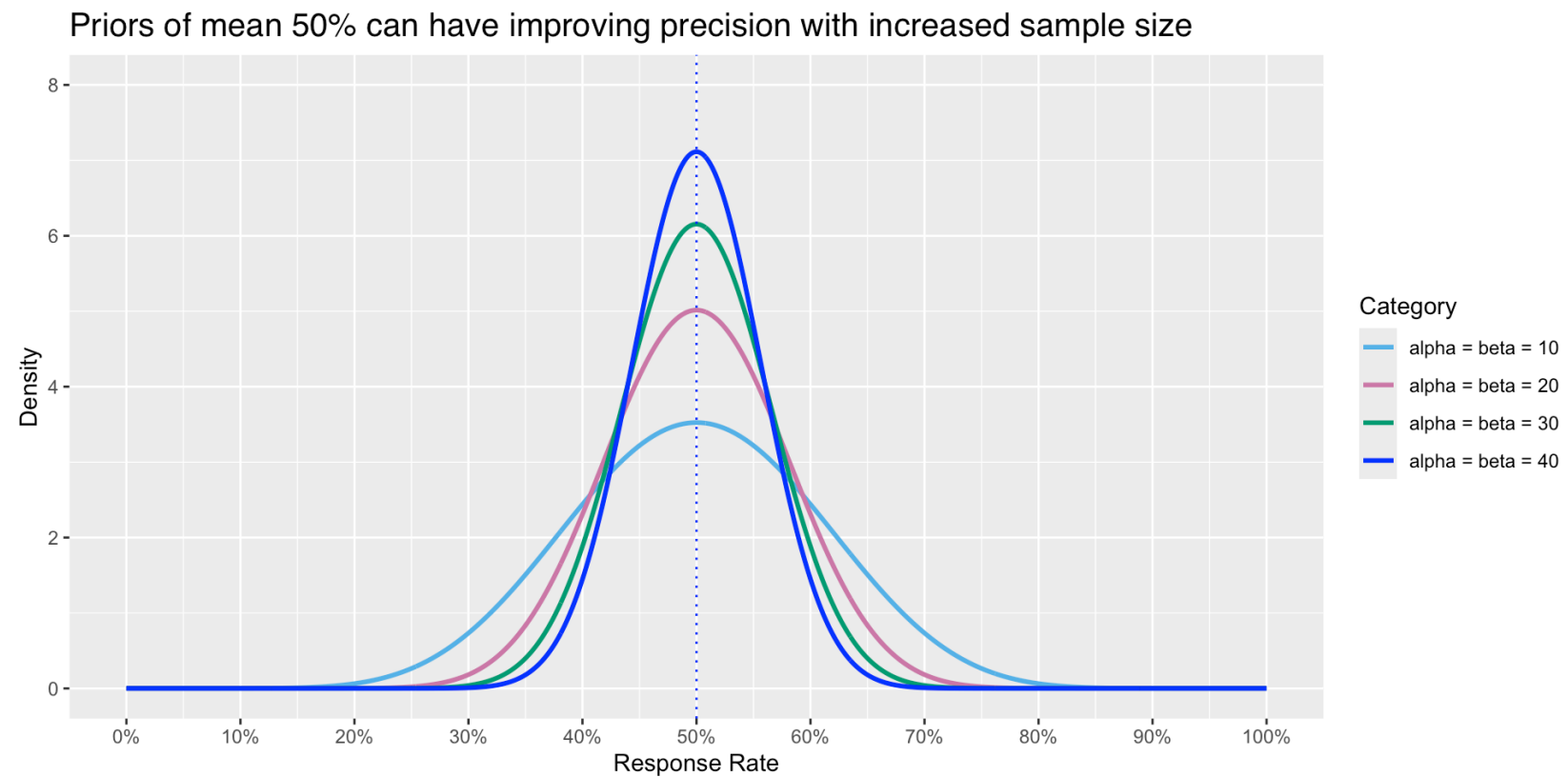
$$\pi = \frac{\alpha}{\alpha + \beta} = \frac{\alpha_{updated}}{\alpha_{updated} + \beta_{updated}} = \frac{16.6}{16.6 + 7.4} \approx 69.17\%$$

$$mode(\pi) = \frac{\alpha_{updated} - 1}{\alpha_{updated} + \beta_{updated} - 2} = \frac{16.6 - 1}{16.6 + 7.4 - 2} \approx 70.90\%$$



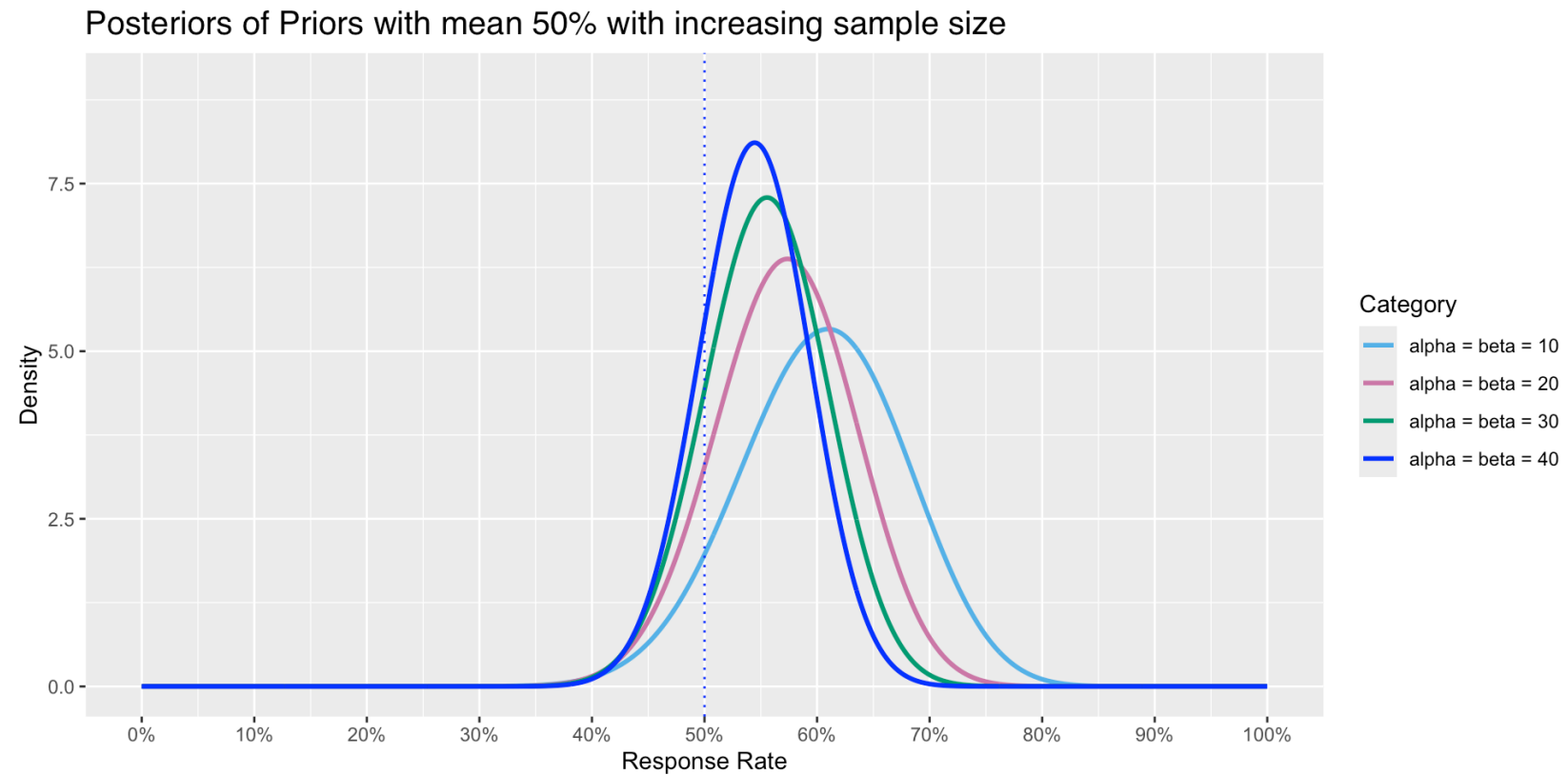
A variety of Priors

- To illustrate how density of Prior changes with increased sample size even though mean is the same



A variety of Posteriors

- To illustrate how density of Posterior changes with increased sample size even though mean is the same



Terminology

- A “look” is a stop
- A stop is when a rule is applied
- The rule is specified for Go, Stop or Evaluate (Gray zone)
- If the rule is met, the result is Go, Stop or Evaluate (Gray zone)
- Rules are applied at interim or final
- Go = Success = Efficacious
- Stop = Failure = Futile
- Rule = Criteria, e.g. Go rule is a Go Criteria

postprob() example (Lee & Liu, 2008)

Example	Interim
Responders	16
n	23
Response rate	69.57 %
Standard of Care Response rate	60 %
Posterior probability	postprob() call from phase1b

```
1 postprob(x = 16, n = 23, p = 0.60, par = c(0.6, 0.4))
```

```
[1] 0.8359808
```

Posterior Probability

- Interim trial is efficacious if posterior probability exceeds 70% or $P(RR \geq 60 \% \mid \text{data}) > 70\%$

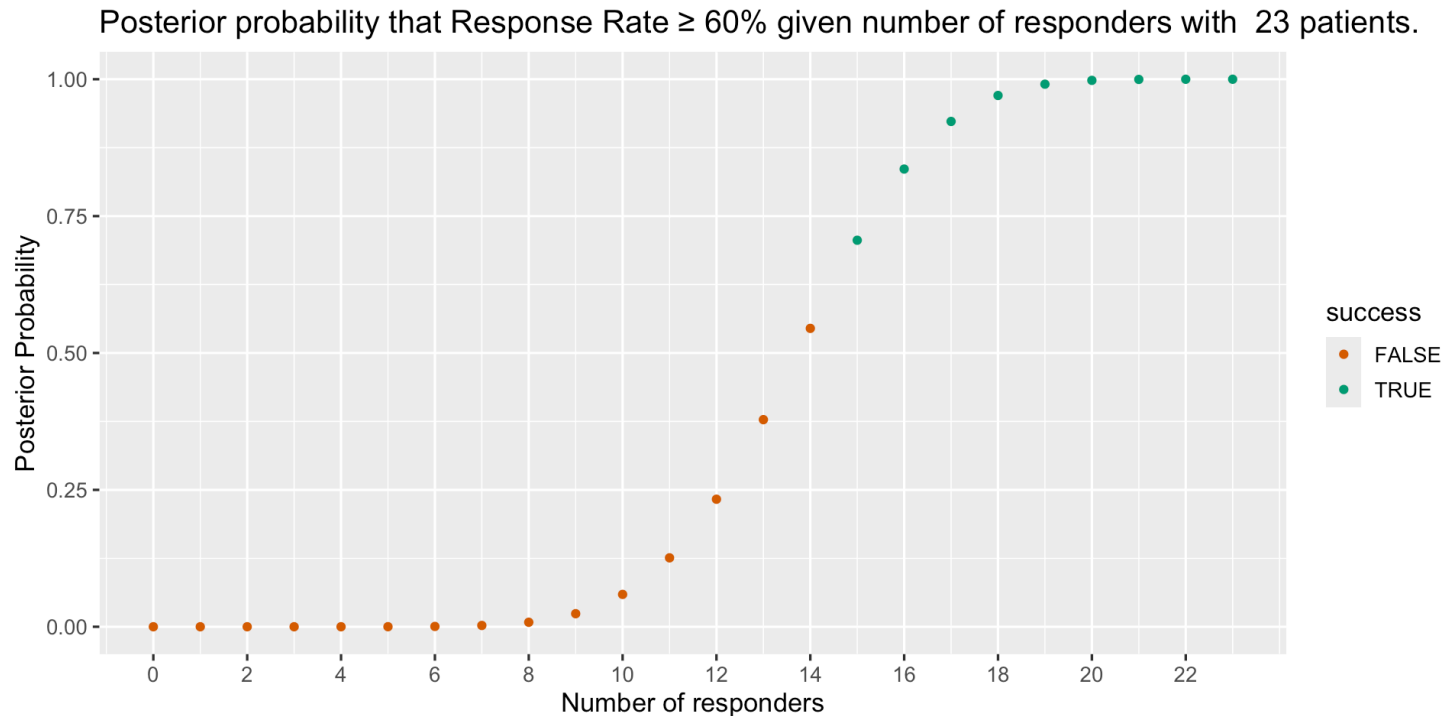


Figure 1

Beta prior mixture

- `phase1b` facilitates the flexibility of using various priors and its respective weightings:

```
Prior is P_E ~ sum(weights * beta(parE[, 1], parE[, 2]))
```

```
1 a = postprob(x = 16,  
2           n = 23,  
3           p = 0.60,  
4           parE = c(0.6, 0.4), weights = 1)  
5  
6 b = postprob(x = 16,  
7           n = 23,  
8           p = 0.60,  
9           parE = c(2, 4), weights = 1)  
10  
11 0.5*(a + b)  
12  
13 postprob(x = 16,  
14         n = 23,  
15         p = 0.60,  
16         parE = rbind(c(0.6, 0.4), c(2, 4)), weights = c(0.5, 0.5))
```

- Posterior formulation :

$$f(\pi|x) \propto \pi^x (1 - \pi)^{n-x} \sum_{j=1}^k w_j \frac{1}{B(\alpha_j, \beta_j)} \pi^{\alpha_j-1} (1 - \pi)^{\beta_j-1}$$

predprob() example (Lee & Liu, 2008)

Example	Interim
Responders	16
n	23
Response rate	69.57 %
Standard of Care Response rate	60 %
Predictive Posterior probability predprob() call from phase1b	

```
1 control = 0.6
2 confidence_sixty = 0.7
3 result <- predprob(
4   x = 16, n = 23, Nmax = 40, p = control, thetaT = confidence_sixty,
5   parE = c(0.6, 0.4)
6 )
7 result$result
```

```
[1] 0.8211011
```

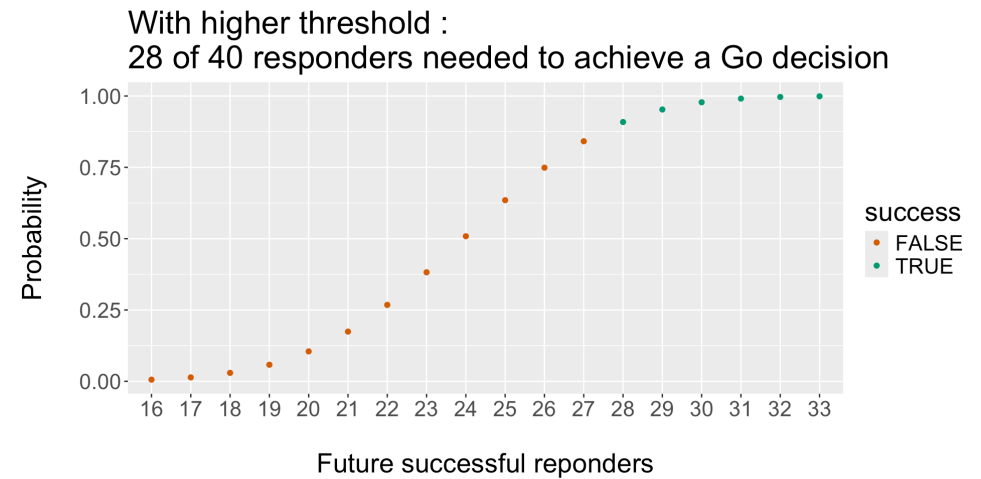
```
1 confidence_ninety = 0.9
2 result_high_thetaT <- predprob(
3   x = 16, n = 23, Nmax = 40, p = control, thetaT = confidence_ninety,
4   parE = c(0.6, 0.4)
5 )
6 result_high_thetaT$result
```

```
[1] 0.5655589
```

Predictive Posterior Probability



$$(a) P(\pi > 0.6 | data) > 70\%$$



$$(b) P(\pi > 0.6 | data) > 90\%$$

Figure 2: Predictive Posterior CDF for different Efficacy Rules

Operating Characteristics

Operating Characteristics : threshold for Success (and failure):

- Efficacy criteria, e.g. we would stop for Efficacy if :

$$\Pr(RR > p_1) > t_U$$

- Futility criteria, eg. we would stop for Futility if :

$$\Pr(RR < p_0) > t_L$$

Rules and Operating characteristics. A use case for **ocPostprob()**:

- Look for Efficacy: Go if $P(\pi > 60\% | data) > 90\%$
- Look for Futility: Stop if $P(\pi < 60\% | data) > 70\%$
- Prior of treatment arm $Beta(0.6, 0.4)$.

```
1 set.seed(2025)
2 res <- ocPostprob(
3   nnE = c(23, 40), truep = 0.60, p0 = 0.60, p1 = 0.69, tL = 0.70, tU = 0.90, parE = c(0.6, 0.4),
4   sim = 500, wiggle = TRUE, nnF = c(23, 40)
5 )
6 res$oc
```

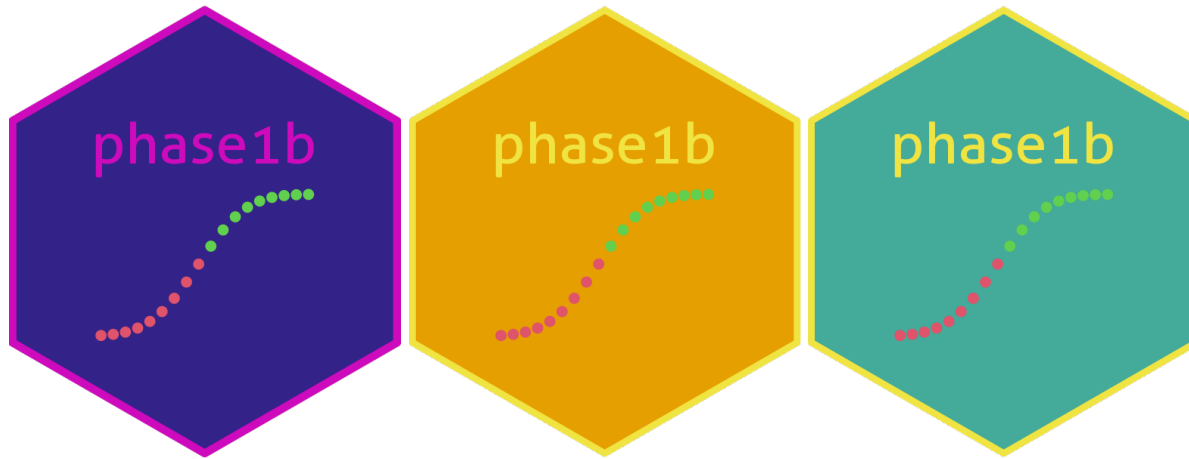
	ExpectedN	PrStopEarly	PrEarlyEff	PrEarlyFut	PrEfficacy	PrFutility	PrGrayZone
1	35.182	0.294	0.014	0.28	0.018	0.4	0.582

Expanded features

.... and wiggle room!

	SOC uncertainty	single-arm	two-arm	simulation	plotting	boundaries
postprob		✓				
postprobDist	✓	✓				
predprob		✓				
predprobDist	✓	✓				
ocPostprob		✓		✓		
ocPostprobDist	✓	✓		✓		
ocPredprob		✓		✓		
ocPredprobDist	✓	✓		✓		
ocRctPostprobDist	✓	✓	✓	✓		
ocRctPredprobDist	✓	✓	✓	✓		
plotBeta				✓	✓	
plotDecision					✓	
plotOc					✓	
plotBounds					✓	
boundsPostprob						✓
boundsPredprob						✓

Concluding remarks



- Big thank you to Daniel Sabanés Bové for mentorship. Data Science Acceleration colleagues Isaac Gravestock, John Kirkpatrick, Craig Gower-Paige et al who collaborated and supported.
- Extension to other therapeutic areas that use response rate as endpoint if beta priors are appropriate
- Open [issues here](#). Contact [me](#) to collaborate.
- Statistical Engineering Journey, PDD Tech Talk 16th (EU), 19th June (SH/SSF) for more on the [phase1b](#) journey

Some references

- Thall P F, Simon R (1994), Practical Guidelines for Phase IIB Clinical Trials, Biometrics, 50, 337-349
- Lee J J, Liu D D (2008), A Predictive probability design for phase II cancer clinical trials, 5(2), 93-106, Clinical Trials
- Yeo, A T, Sabanés Bové D, Elze M, Pourmohamad T, Zhu J, Lymp J, Teterina A (2024). Phase1b : Calculations for decisions on Phase 1b clinical trials. R package version 1.0.0, <https://genentech.github.io/phase1b>
- Inclusive Speaker Orientation [Linux Foundation](#)
- Zeileis, Fisher, Hornik, Ihaka, McWhite, Murrell, Stauffer, Wilke (2020). colorspace: A Toolbox for Manipulating and Assessing Colors and Palettes. Journal of Statistical Software.
- [Code for this presentation](#)

Backupslides

Standard of Care Distribution when unknown

Using the approach by Thall and Simon (Biometrics, 1994), this evaluates the posterior probability of achieving superior response rate in the treatment group **E** compared to standard of care **S**.

- The desired improvement is denoted as **delta**. There are two options in using **delta**. The absolute case when **relativeDelta = FALSE** and relative as when **relativeDelta = TRUE**.

Desired improvement **delta**, two approaches:

1. The absolute case is when we define an absolute delta, greater than P_S , the response rate of the standard of care or control or S group such that the posterior is $\Pr(P_E > P_S + \text{delta} \mid \text{data})$.
2. In the relative case, we suppose that the treatment group's response rate is assumed to be greater than $P_S + (1 - P_S) * \text{delta}$ such that the posterior is $\Pr(P_E > P_S + (1 - P_S) * \text{delta} \mid \text{data})$.

User facing functions with this feature include

- `postprobDist()`
- `predprobDist`
- `ocPostprobDist()`
- `ocPredprobDist()`

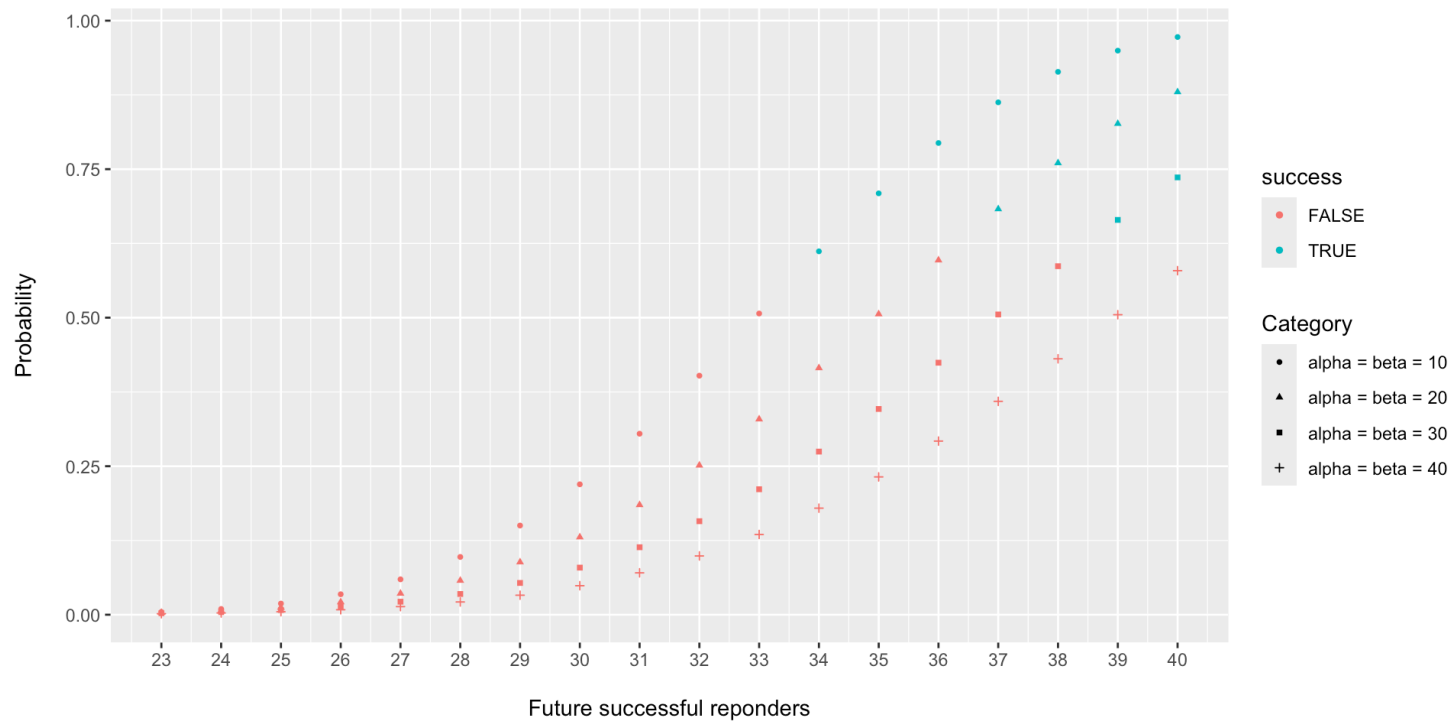
Does CDF improve with increased “data availability” of priors ?

Here we have an interim result of 16 and 23 patients. (input for Likelihood)

We have a threshold of 60% success. Our standard of care is 60%. Our priors have a mean of 50%.

Does CDF improve with increased “data availability” of priors ?

No.



Predictive posterior probability for Decision 1 :

The criteria for Decision 1 for Interim looks are :

- interim GO : $P(\text{ success at final}) > \phi_U$
- interim STOP : $P(\text{ success at final}) < \phi_L$

The criteria for Decision 1 for Final looks are:

- Final GO : $P(\text{ response rate } > p_0 \mid \text{ data}) > t_T$
- Final STOP : $P(\text{ response rate } > p_0 \mid \text{ data}) < t_T$

Predictive posterior probability for Decision 2:

The criteria for Decision 2 for Interim looks are :

- Interim GO : $P(\text{success at final}) > \phi_U$
- Interim STOP : $P(\text{failure at final}) > \phi_{Fu}$

The criteria for Decision 2 for Futility looks are :

- Final GO : $P(\text{response rate} > p_0) > t_T$
- Final STOP : $P(\text{response rate} < p_1) > t_F$