# Problem 42

## Gamma Ray Arrivals

### Estimating the Parameter of Poisson Arrivals with Variable Time Lengths

data <- read.csv("../Excel Comma/Chapter 8/gamma-ray.csv")</pre>

We will assume that gamma ray arrivals follow a Poisson process with parameter  $\lambda$  arrivals per second. Under this assumption, let the 100 observations be indexed as i = 1, 2, ...n. Let  $t_i$  and  $X_i$  be the length of the *i*th time interval and the number of observed gamma rays during the *i*th time interval, respectively. I will assume that the time intervals are disjoint.

According to the model,  $X_i$  follows a Poisson distribution with parameter  $\lambda t_i$ . Due to the assumption that the time intervals are disjoint, the  $X_i$ s are independent of each other.

The distribution of  $X_i$  is thus

$$P(X_i = k_i | \lambda, t_i) = \frac{(\lambda t_i)^{k_i} e^{-\lambda t_i}}{k_i!}$$

The likelihood function of  $\lambda$  is thus

$$lik(\lambda|\mathbf{t}, \mathbf{X}|) = \frac{(\lambda t_i)^{\sum_{i=1}^{n} X_i} e^{-\lambda \sum_{i=1}^{n} t_i}}{\prod_{i=1}^{n} (X_i!)}$$

The log likelihood is thus

$$l(\lambda) = \ln(\lambda) \sum_{i=1}^{n} X_i + \sum_{i=1}^{n} X_i \ln(t_i) - \lambda \sum_{i=1}^{n} t_i - \sum_{i=1}^{n} \ln(X_i!)$$

The maximum likelihood estimator satisfies the first-order condition

$$0 = \frac{1}{\hat{\lambda}} \sum_{i=1}^{n} X_i - \sum_{i=1}^{n} t_i$$

Which after some rearranging, yields the maximum likelihood estimator

$$\hat{\lambda} = \frac{\sum_{i=1}^{n} X_i}{\sum_{i=1}^{n} t_i}$$

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Setting all  $t_i$ 's to 1 yields the usual estimator for  $\lambda$  in an iid sample from a Poisson distribution.

lambdaHat <- sum(data\$count) / sum(data\$seconds)</pre>

Thus  $\hat{\lambda} = 0.0038809$ .

#### The Standard Error of the Estimate

Define  $S = \sum_{i=1}^{n} X_i$  and  $T = \sum_{i=1}^{n} t_i$ . Then the estimator  $\hat{\lambda}$  can be rewritten as

$$\hat{\lambda} = \frac{S}{T}$$

Since S is the sum of independent Poisson variables with parameter  $\lambda t_i$ , S follows a Poisson distribution with parameter  $\lambda T$ . Thus  $Var(S) = \lambda T$ . Since  $\hat{\lambda} = \frac{S}{T}$ ,  $Var(\hat{\lambda}) = \frac{\lambda}{T}$ , and  $SD(\hat{\lambda}) = \sqrt{\frac{\lambda}{T}}$ . This gives us an estimator

$$s_{\hat{\lambda}} = \sqrt{\frac{\hat{\lambda}}{T}}$$

s\_lambdaHat <- sqrt(lambdaHat / sum(data\$seconds))</pre>

This produces  $s_{\hat{\lambda}} = 4.9689212 \times 10^{-4}$ .

#### Posterior Distribution Using an Improper Gamma

Assume that the prior distribution of  $\lambda$  follows an improper gamma distribution,  $f_{\Lambda}(\lambda) \propto \lambda^{-1}$ . Then the posterior distribution is proportional to

$$f_{\Lambda}(\lambda|\mathbf{X}, \mathbf{t}) \propto \lambda^{\sum_{i=1}^{n} x_i - 1} \prod_{i=1}^{n} t_i^{x_i} \frac{e^{-\lambda \sum_{i=1}^{n} t_i}}{\prod_{i=1}^{n} x_i} \propto \lambda^{\sum_{i=1}^{n} x_i - 1} e^{-\lambda \sum_{i=1}^{n} t_i}$$

Which is a gamma distribution with parameters  $\sum_{i=1}^{n} x_i$  and  $\sum_{i=1}^{n} t_i$ .