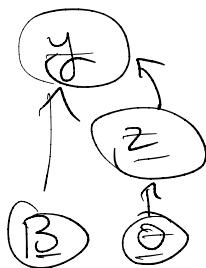


## Hierarchical models :

Baillargeon (1996) :



Data :  $[y | z, \beta]$

process :  $[z | \theta]$

parameters :  $[\beta, \theta]$

Account for data-level variation while modeling latent process.

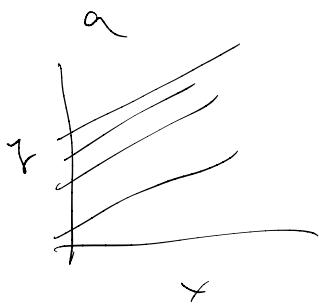
Also called :

1.) multi-level models

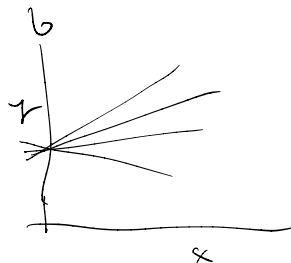
2.) mixed effects models  
    ↳ fixed and random effects

3.) state-space models,  
especially when temporal

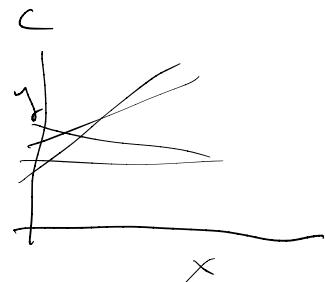
## Regression Scenarios :



$$\beta_{0i} + \beta_{1i}x_{ij}$$



$$\beta_0 + \beta_{1i}x_{ij}$$



$$\beta_{0i} + \beta_{1i}x_{ij}$$

It is common to consider varying coefficients  
as "random" effects:

For example

a.)  $i = 1, \dots, n$        $j = 1, \dots, J$

$$y_{ij} \sim N(\beta_{0i} + \beta_1 x_{ij}, \sigma^2) \quad ] \text{ data}$$

$$\beta_{0i} \sim N(\mu_0, \sigma_{00}^2) \quad ] \text{ process}$$

$$\mu_0 \sim N(\mu_{00}, \sigma_{000}^2) \quad ] \text{ parameters}$$

$$\sigma_{00}^2 \sim Ig(g_0, r_0)$$

$$\beta_1 \sim N(\mu_1, \sigma_1^2)$$

$$\sigma^2 \sim Ig(g_1, r_1) \quad ]$$

This is called the "random intercepts" model

b.)

$$y_{ij} \sim N(\beta_{0i} + \beta_1 x_{ij}, \sigma^2)$$

$$\beta_1 \sim N(\mu_1, \sigma_1^2)$$

$$\mu_1 \sim N(\mu_{10}, \sigma_{10}^2)$$

$$\sigma_1^2 \sim Ig(g_1, r_1)$$

$$\beta_{0i} \sim N(\mu_0, \sigma_{00}^2)$$

$$\sigma^2 \sim Ig(g, r)$$

This called the "random slopes" model

C.)

$$y_{ij} \sim N(\beta_0 + \beta_i x_{ij}, \sigma^2) \quad ] \text{ data}$$

$$\beta_i \sim N(\mu, \Sigma), \quad \Sigma = \sigma_\beta^2 \begin{bmatrix} 1 & p \\ p & 1 \end{bmatrix} \quad ] \text{ process}$$

$$\mu \sim N(\mu_0, \Sigma_0)$$

$$\sigma_\beta^2 \sim Ig(g_\beta, r_\beta)$$

$$p \sim Unif(-1, 1)$$

$$\sigma^2 \sim Ig(g, r)$$

All coefficients random!  
(no "fixed" effects)

Note: data model can be written as:

$$y_i \sim N(X_i \beta_i, \sigma^2 I)$$

Posterior:

$$[\beta_1, \dots, \beta_n, \mu, \sigma_\beta^2, p, \sigma^2 | y_1, \dots, y_n] \propto$$

$$\left( \prod_{i=1}^n [y_i | \beta_i, \sigma^2] [\beta_i | \mu, \sigma_\beta^2, p] \right) [\mu] [\sigma_\beta^2] [p] [\sigma^2]$$

Full-conditional Distributions:

$$[\beta_i | \cdot] \propto [y_i | \beta_i, \sigma^2] [\beta_i | \mu, \sigma_\beta^2, p] \quad \begin{array}{l} \text{(from} \\ \text{before)} \end{array}$$

$$A = X_i' ( \sigma^2 I )^{-1} X_i + \Sigma^{-1}$$

$$b = X_i' ( \sigma^2 I )^{-1} y_i + \Sigma^{-1} g$$

for  $i = 1, \dots, n$

$$\begin{aligned}
 & [M \cdot J] \propto \left( \prod_{i=1}^n [B_i | \mu, \sigma_B^2, p] \right) [M] \\
 & \propto \exp \left\{ -\frac{1}{2} \sum_i (B_i - \mu)' \Sigma' (B_i - \mu) \right\} \exp \left\{ \frac{1}{2} (\mu - \mu_0)' \Sigma_0' (\mu - \mu_0) \right\} \\
 & \propto \exp \left\{ -\frac{1}{2} \underbrace{\left( -2 \left( \sum_{i=1}^n B_i (\Sigma' + \mu' \Sigma_0') M + \mu' (n \Sigma' + \Sigma_0' M) \right) \right)}_{b'} \right. \\
 & \quad \left. + \underbrace{A} \right\} \\
 & = N(A^{-1} b', A')
 \end{aligned}$$

$$\begin{aligned}
 & [\sigma_B^2 | \cdot] \propto \left( \prod_{i=1}^n [B_i | \mu, \sigma_B^2, p] \right) [\sigma_B^2], \quad R = \begin{bmatrix} 1 & p \\ p & 1 \end{bmatrix} \\
 & \propto (\sigma_B^2)^n \exp \left\{ -\frac{1}{\sigma_B^2} \sum_i \underbrace{(B_i - \mu)' R (B_i - \mu)}_2 \right\} (\sigma_B^2)^{\frac{n(n+1)}{2}} \exp \left\{ -\frac{1}{\sigma_B^2 r_B} \right\} \\
 & \propto (\sigma_B^2)^{\underbrace{(n+q_B+1)}_{g}} \exp \left\{ -\frac{1}{\sigma_B^2} \underbrace{\left( \sum_i \frac{(B_i - \mu)' R (B_i - \mu)}{2} + \frac{1}{r_B} \right)}_{\tilde{r}} \right\} \\
 & = IG(\tilde{g}, \tilde{r})
 \end{aligned}$$

$$[p | \cdot] \propto \left( \prod_{i=1}^n [B_i | \mu, \sigma_B^2, p] \right) [p] \quad (\text{use M-H})$$

$$\begin{aligned}
 & [\sigma^2 | \cdot] \propto \left( \prod_{i=1}^n [y_i | B_i, \sigma^2] \right) [\sigma^2] \\
 & \propto (\sigma^2)^{\frac{n-3}{2}} \exp \left\{ -\frac{1}{\sigma^2} \sum_i \frac{(y_i - x_i B_i)' (y_i - x_i B_i)}{2} \right\} (\sigma^2)^{\frac{-(q+1)}{2}} \exp \left\{ -\frac{1}{\sigma^2 r} \right\} \\
 & \propto (\sigma^2)^{\underbrace{(\frac{n-3}{2} + q+1)}_{\tilde{g}}} \exp \left\{ -\frac{1}{\sigma^2} \underbrace{\left( \sum_i \frac{(y_i - x_i B_i)' (y_i - x_i B_i)}{2} + \frac{1}{r} \right)}_{\tilde{r}} \right\} \\
 & = IB(\tilde{g}, \tilde{r})
 \end{aligned}$$

Generalization to hierarchical model:

1.) Let data-level variance vary  $\sigma_i^2 \sim \text{IG}(q, r)$

2.) Let data-level variance be random:

$$\log \sigma_i \sim N(\mu_0, \sigma_0^2)$$

$$\mu_0 \sim N(\mu_{00}, \sigma_{00}^2)$$

$$\sigma_0^2 \sim \text{IG}(q_0, r_0)$$

3.) Let  $\beta_{0i}, \beta_{1i}$  have diff. variances:

$$\beta_{0i} \sim N(\mu_0, \sigma_0^2) \Rightarrow \rho = 0$$

$$\beta_{1i} \sim N(\mu_1, \sigma_1^2)$$

4.) Let  $\beta_i$  have fully unknown covariance:

$$\Sigma = \begin{bmatrix} \sigma_0^2 & \sigma_0 \sigma_1 \rho \\ \sigma_0 \sigma_1 \rho & \sigma_1^2 \end{bmatrix}$$

recall correlation:

$$\rho = \frac{\sigma_0 \sigma_1 \rho}{\sqrt{\sigma_0^2 \sigma_1^2}}$$

5.) Scalable prior for  $\Sigma$ :  $\Sigma^{-1} \sim \text{Wish}((Sv)^T, v)$   
"wishart" w/ mean:

$$E(\Sigma^{-1}) = (Sv)^T v = S^{-1}$$

b.) GLMAs:  $y_i \sim \text{Pois}(z_i)$   
 $\log(z_i) = x_i \beta_i, \beta_i \sim N(\mu, \Sigma)$

?) More covariates:  $\beta_i \sim N(\mu, \Sigma)$   
 $p \times 1 \quad p \times p$