Dependency Discovery via Multiscale Graph Correlation

Cencheng Shen

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Collaborators: Joshua T. Vogelstein, Carey E. Priebe, Shangsi Wang, Youjin Lee, Mauro Maggioni, Qing Wang, Alex Badea.

Acknowledgment: NSF DMS, DARPA SIMPLEX.

R package available in CRAN and https://github.com/neurodata/MGC/Matlab code available in https://github.com/neurodata/mgc-matlab

Overview

- 1. Motivation
- 2. Methodology
- 3. Theoretical Properties
- 4. Simulations and Experiments
- 5. Summary

Given paired data $(\mathcal{X}_n, \mathcal{Y}_n) = \{(x_i, y_i) \in \mathbb{R}^p \times \mathbb{R}^q, \text{ for } i = 1, \dots, n\},$

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- How are they related?

X	Y
brain connectivity	creativity / personality
brain shape	health
gene / protein	cancer
social networks	attributes
anything	anything else

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Formal Definition of Independence Testing

$$(x_i, y_i) \stackrel{i.i.d.}{\sim} F_{XY}, \quad i = 1, \dots, n$$

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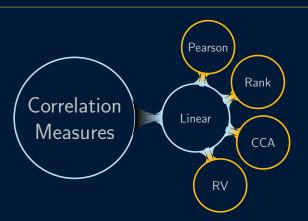
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Without loss of generality, we shall assume ${\cal F}_{XY}$ has finite second moments.

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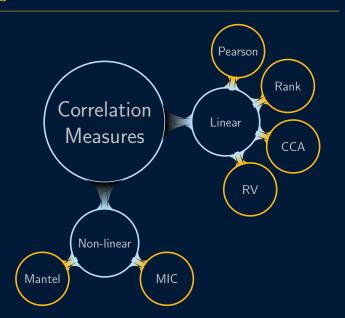


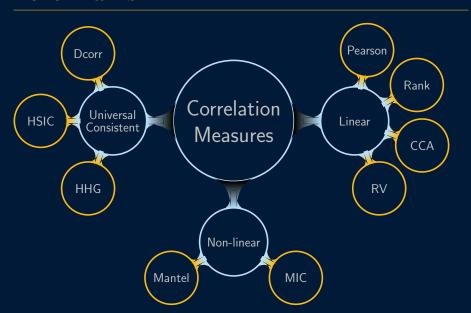




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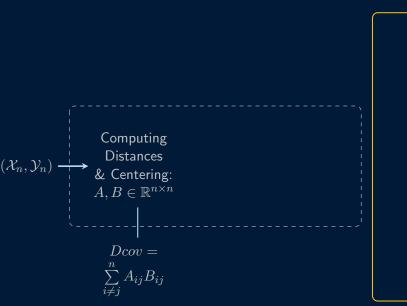
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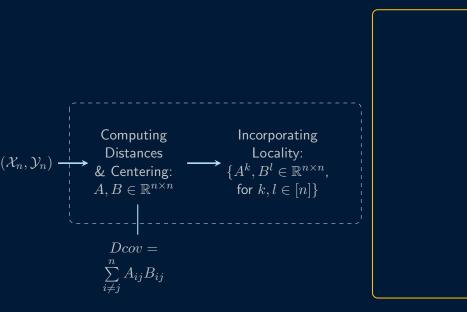
To that end, we propose the **multiscale graph correlation** in [Shen et al.(2018)].

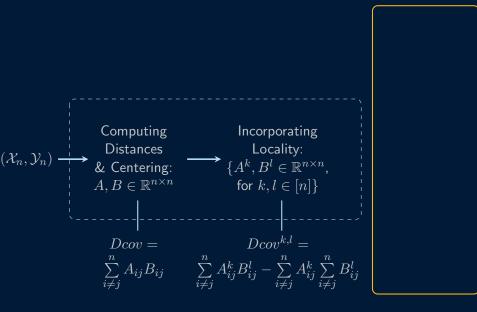
Methodology

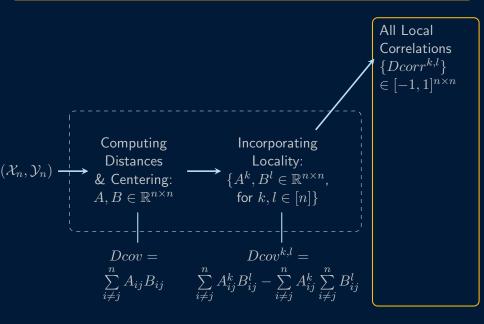


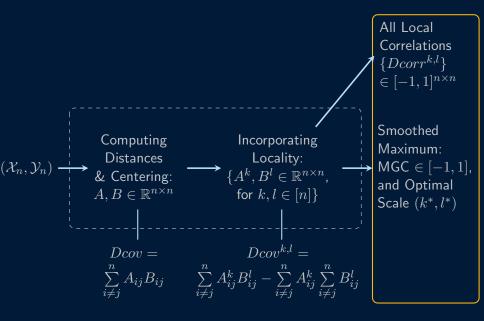


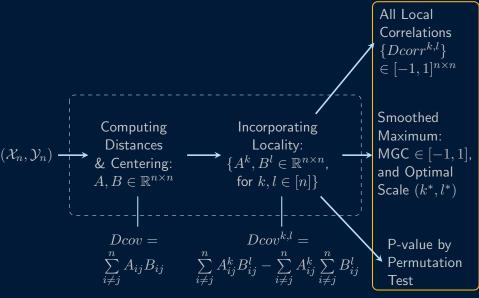












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Computing Distance and Centering

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Centering: Then we center \tilde{A} and \tilde{B} by columns, with the diagonals excluded:

$$A_{ij} = \begin{cases} \tilde{A}_{ij} - \frac{1}{n-1} \sum_{s=1}^{n} \tilde{A}_{sj}, & \text{if } i \neq j, \\ 0, & \text{if } i = j; \end{cases}$$
 (1)

similarly for B.

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Ranking: Define $\{R_{ij}^A\}$ as the "rank" of x_i relative to x_j , that is, $R_{ij}^A=k$ if x_i is the k^{th} closest point (or "neighbor") to x_j , as determined by ranking the set $\{\tilde{A}_{1j}, \tilde{A}_{2j}, \dots, \tilde{A}_{nj}\}$ by ascending order. Similarly define R_{ii}^B for the y's.

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For any $(k,l)\in [n]^2$, define the rank truncated matrices A^k,B^l , and the joint distance matrix C^{kl} as

$$A_{ij}^k = A_{ij} \mathbf{I}(R_{ij}^A \le k),$$

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When ties occur, minimal rank is recommended, e.g., if Y only takes two value, R^B_{ij} takes value in $\{1,2\}$ only. We assume no ties for each of presentation.

A Family of Local Correlations: Let \circ denote the entry-wise product, $\hat{E}(\cdot) = \frac{1}{n(n-1)} \sum_{i \neq j}^n (\cdot)$ denote the diagonal-excluded sample mean of a square matrix, then the sample local covariance, variance, and correlation are defined as:

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$$dCov^{k,l}(\mathcal{X}_n, \mathcal{Y}_n) = \hat{E}(A^k \circ B^{l'}) - \hat{E}(A^k)\hat{E}(B^l),$$

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for k, l = 1, ..., n. If $dVar^k(\mathcal{X}_n) \cdot dVar^l(\mathcal{X}_n) \leq 0$, we set $dCorr^{kl}(\mathcal{X}_n, \mathcal{Y}_n) = 0$ instead.

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There are a maximum of n^2 different local correlations. At k=l=n, $dCorr^{kl}(\mathcal{X}_n,\mathcal{Y}_n)$ equals the "global" distance correlation $dCorr(\mathcal{X}_n,\mathcal{Y}_n)$ by Szekely et al.(2007).

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But directly taking the maximum local correlation

$$\max_{(k,l)\in[n]^2} \{Dcorr^{k,l}(\mathcal{X}_n,\mathcal{Y}_n)\}$$

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Instead, we take a smoothed maximum, by finding a connected region in the local correlation map with significant local correlatons — if such a region exists, use the maximum within the region.

Pick a threshold $\tau \geq 0$ (we choose by an approximate null distribution of Dcorr, which is symmetric beta and converges to 0 as $n \to \infty$), compute the set

 $\{(k,l) \text{ such that } Dcorr^{k,l}(\mathcal{X}_n,\mathcal{Y}_n) > \max\{\tau, Dcorr(\mathcal{X}_n,\mathcal{Y}_n)\}\},\$

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Permutation Test

To get a p-value by MGC for any given data, we utilize the permutation test: randomly permute index the second data set for r times, compute the permuted MGC statistic for each, and compute the percentage that the original MGC is no larger than the permuted MGC statistic, and take it as the p-value.

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This is a standard nonparametric testing procedure employed by Mantel, Dcorr, HHG, HSIC, where the null distribution of the dependency measure cannot be exactly derived.

Distance computation takes $\mathcal{O}(n^2 \max(p,q))$, centering takes $\mathcal{O}(n^2)$, ranking takes $\mathcal{O}(n^2 log(n))$, all local correlations can be iteratively computed in $\mathcal{O}(n^2)$, and the smoothing step takes $\mathcal{O}(n^2)$.

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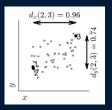
The permutation test takes $\mathcal{O}(n^2 \max(r, p, q, log(n)))$ for r random permutations.

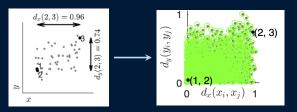
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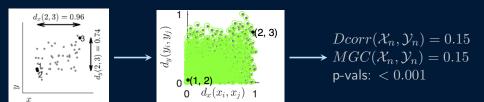
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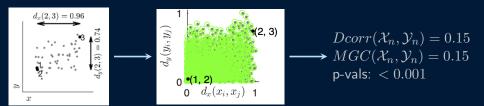
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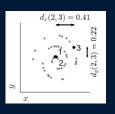
There are a number of ways to speed up the method for big data: faster implementation when p=q=1, null distribution approximation by subsampling and spectral embedding.

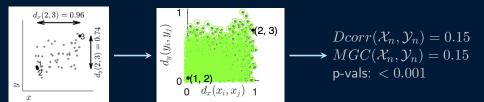


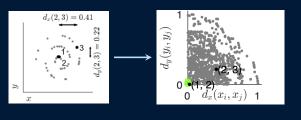


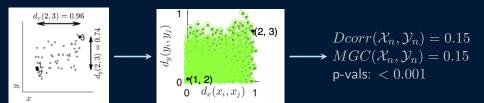














Theoretical Properties

Theorem 1 (Well-behaved Correlation Measure)

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- 3. Non-negative Variance: $c(\mathcal{X}_n, \mathcal{X}_n) \geq 0$ with equality if and only if F_X is non-degenerate.

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- 4. Invariant: $c(\mathcal{X}_n, \mathcal{Y}_n) = c(\{\phi(x_i)\}, \{\delta(y_i)\})$ for any distance-preserving transformations ϕ, δ (i.e., rotation, scaling, translation, reflection).

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- 5. 1-Linear: $c(\mathcal{X}_n,\mathcal{Y}_n)=1$ if and only if F_X is non-degenerate and (X,uY) are dependent via an isometry for some non-zero constant u.

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- 2. Valid Test: Under the permutation test, Sample MGC is a valid test, i.e., it controls the type 1 error level α .
- 3. Consistency: At any type 1 error level α , testing power $\beta(c(\mathcal{X}_n,\mathcal{Y}_n)) \stackrel{n \to \infty}{\to} 1$ against any dependent F_{XY} .

Suppose (X,Y),(X',Y'),(X'',Y''),(X''',Y''') are $\it iid$ as $\it F_{XY}.$

Suppose (X,Y),(X',Y'),(X'',Y''),(X''',Y''') are *iid* as F_{XY} . Let $I(\cdot)$ be the indicator function, define two random variables

$$\mathbf{I}_{X,X'}^{\rho_k} = \mathbf{I}\left(\int_{B(X,\|X'-X\|)} dF_X(u) \le \rho_k\right)$$
$$\mathbf{I}_{Y',Y}^{\rho_l} = \mathbf{I}\left(\int_{B(Y',\|Y'-Y\|)} dF_Y(u) \le \rho_l\right)$$

for $\rho_k, \rho_l \in [0, 1]$.

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for $\rho_k, \rho_l \in [0, 1]$. Further define

$$d_X^{\rho_k} = (\|X - X'\| - \|X - X''\|) \mathbf{I}_{X,X'}^{\rho_k}$$

$$d_{Y'}^{\rho_l} = (\|Y' - Y\| - \|Y' - Y'''\|) \mathbf{I}_{Y',Y}^{\rho_l}$$

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$$I_{Y',Y}^{\rho_l} = I\left(\int_{B(Y',\|Y'-Y\|)} dF_Y(u) \le \rho_l\right)$$

for $\rho_k, \rho_l \in [0, 1]$. Further define

$$d_X^{\rho_k} = (\|X - X'\| - \|X - X''\|) \boldsymbol{I}_{X,X'}^{\rho_k}$$

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The population local covariance can be defined as

$$Dcov^{\rho_k,\rho_l}(X,Y) = E(d_X^{\rho_k} d_{Y'}^{\rho_l}) - E(d_X^{\rho_k}) E(d_{Y'}^{\rho_l}).$$

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Normalizing and taking a smoothed maximum yield population MGC.

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with respect to a non-negative weight function w(t,s) on $(t,s) \in \mathbb{R}^p \times \mathbb{R}^q$. The weight function is defined as:

$$w(t,s) = (c_p c_q |t|^{1+p} |s|^{1+q})^{-1},$$

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Can be similarly adapted to the local correlation.

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The last three properties also hold for any local correlation by $(\rho_k,\rho_l)=(rac{k-1}{n-1},rac{l-1}{n-1}).$

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C. Shen ◀중▶◀불▶ MGC: 24/3

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If the relationship is linear (or with independent noise), the global scale is always optimal and MGC(X,Y) = Dcorr(X,Y).

Conversely, the optimal scale being local, i.e., MGC(X,Y) > Dcorr(X,Y), implies a non-linear relationship.

MGC is applicable to similarity / kernel matrix

Theorem 6 (Transforming kernel to distance)

Given any characteristic kernel function $k(\cdot,\cdot)$, define an induced semi-metric as

$$d(\cdot, \cdot) = 1 - k(\cdot, \cdot) / \max\{k(\cdot, \cdot)\}.$$

Then $d(\cdot, \cdot)$ is of strong negative type, and the resulting MGC is universally consistent.

Namely, given a sample kernel matrices $K_{n\times n}$, one can compute the induced distance matrix by

$$D = J - K / \max_{i, j \in [1, \dots, n]^2} \{K(i, j)\},$$

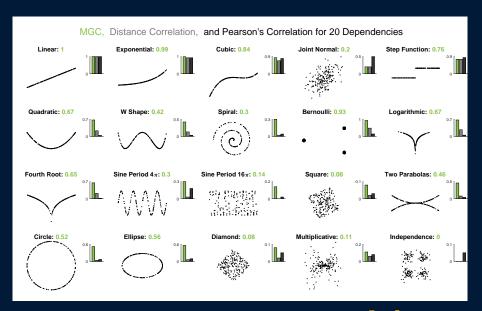
and apply MGC to the induced distance matrices.

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Simulations and Experiments

Visualizations of 20 Simulation Settings

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Evaluation Criterion

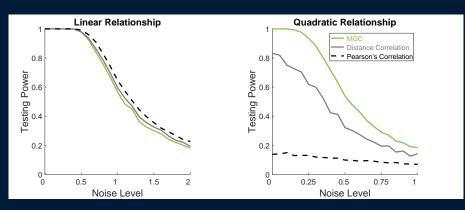
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- Required sample size $N_{\alpha,\beta}(c)$ to achieve a power of β at type 1 error level α using a statistic c.

Testing Power: Linear vs Nonlinear

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$$\begin{split} n &= 30, p = q = 1, \\ X &\sim Uniform(-1,1), \\ \epsilon &\sim Normal(0,noise), \\ Y &= X + \epsilon \text{ and } Y = X^2 + \epsilon. \end{split}$$

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We consider univariate (1D) and multivariate (10D) cases.

Median Size Table

Testing Methods	1D Lin	1D Non-Lin	10D Lin	10D Non-Lin
MGC	50	90	60	165
Dcorr	50	250	60	515
Pearson / RV / CCA	50	>1000	50	>1000
HHG	70	90	100	315
HSIC	70	95	100	400
MIC	120	180	n/a	n/a

C. Shen

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Adjusted for multiple testing, MGC uniquely revealed one particular protein, neurogranin, which is exclusively expressed in brain tissue among normal tissues and has not been linked with any other cancer type.

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If we compare kNN (K=3) leave-one-subject-error of these peptides:

Peptides	False Positives	True Positives
Neurogranin (MGC)	2	5
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They made MGC advantageous in theory and practice.

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 - Any dissimilarity / similarity / kernel matrix.
- 3. Intuitive to understand and efficient to implement in $\mathcal{O}(n^2 log n)$.

References

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