Learning from Data 6. Classification: Naïve Bayes

© Prakash Ishwar Spring 2017

Classification

- Supervised (preditive) learning: given examples with labels, predict labels for all unseen examples
 - Classification:
 - label = category,
 - $y \in \mathcal{Y} = \{1, \dots, m\}, m = \text{number of classes}$
 - $\ell(\mathbf{x}, y, h) = 1(h(\mathbf{x}) \neq y)$, Risk = $P(Y \neq h(\mathbf{X})) = P(\text{Error})$



x = facial geometry featuresy = gender label

- Generative learning
- d features (components) in each feature vector:

$$\mathbf{x} = (x_1, \dots, x_i, \dots, x_d)^{\top}$$

 Naïve Bayes Assumption: all features are conditionally independent given the class label

$$\forall y, \ p(\mathbf{x}|y,\theta) = p(x_1|y,\theta_1) \cdots p(x_d|y,\theta_d) = \prod_{i=1}^d p(x_i|y,\theta_i)$$
$$\theta = (\theta_0, \theta_1, \dots, \theta_d)^\top, \quad \theta_0 = (p(y=1), \dots, p(y=m))^\top$$

- Resulting MPE decision rule: Naïve Bayes classifier
- Why Naïve? We do not expect all features to be independent even conditional on class label

- Although assumption is false, the resulting classifier works quite well in practice.
- It has only on the order of O(md) parameters to learn → relatively immune to overfitting
 - Example: Suppose that all features are binary, say taking only values +1 or -1, then:
 - In NB model, we need to specify only 1 scalar parameter per feature per class, namely $P(X_i = +1|y), y = 1,...,m$ or md scalar parameters in total
 - But specifying the joint pmf of all d binary features requires $(2^{d}-1)$ scalar parameters per class or $m(2^{d}-1)$ scalar parameters in total

Class-conditional likelihood functions:

$$p(x_i|y,\theta_i), i=1,\ldots,d$$

- may be either completely parametric or completely nonparametric or some components may be modeled parametrically and others non-parametrically
- some or all feature components may be discrete, continuous, or mixed
- Gaussian Naïve Bayes: $\forall i, y, \ p(x_i|y, \theta_i) \sim \mathcal{N} \left(\mu_{iy}, \sigma_{iy}^2\right)$
- Categorical Naïve Bayes: all features are discrete and take only a finite number of possible values

Notation summary:

```
y = \text{class label} \in \{1, \dots, m\}

j = \text{sample index} \in \{1, \dots, n\}

i = \text{feature (component) index} \in \{1, \dots, d\}

\theta_i = \text{parameters for class-conditional pdfs/pmfs of feature } i \text{ across all classes}

\theta_i = \{\theta_{iy}, y = 1, \dots, m\}, \text{ where}

\theta_{iy} = \text{parameters for class-conditional pdf/pmf of feature } i \text{ in class } y, \text{ and}

p(x_i|y,\theta_i) = p(x_i|y,\theta_{iy}) \text{ for all } i,y

\mathcal{D} = \{(\mathbf{x}_1,y_1),\dots,(\mathbf{x}_n,y_n)\} \text{ (training set)}
```

- Training set: $\mathcal{D} = \{(\mathbf{x}_j, y_j), j = 1, ..., n\}$
- Training set for feature *i*:

$$\mathcal{D}_i = \{(x_{ij}, y_j), j = 1, ..., n\}$$

Training set for feature i and class y:

$$\mathcal{D}_{iy} = \{(x_{ij}, y_j), j: y_j = y\}$$

- Key Result: for each feature i and each class y, the ML estimate of θ_{iy} based on \mathcal{D} is the same as the ML estimate of θ_{iy} based on \mathcal{D}_{iy}
- Intuition: features independent in every class ⇒
 parameters of feature distribution can be estimated
 independently for each feature in each class

$$\begin{aligned} Proof: & p(\mathcal{D}|\theta) = \prod_{j=1}^n p(\mathbf{x}_j, y_j|\theta) \\ &= \prod_{j=1}^n \left[p(y_j) \prod_{i=1}^d p(x_{ij}|y_j, \theta_i) \right] & \text{Na\"ive Bayes assumption:} \\ &= \left(\prod_{j=1}^n p(y_j) \right) \prod_{i=1}^d \prod_{j=1}^n p(x_{ij}|y_j, \theta_i) \\ &= \left(\prod_{j=1}^m p(y)^{n_y} \right) \prod_{i=1}^d \prod_{j=1}^m p(x_{ij}|y_j, \theta_i) \\ &= \left(\prod_{y=1}^m p(y)^{n_y} \right) \prod_{i=1}^d \prod_{y=1}^m p(x_{ij}|y_j, \theta_i) & n_y = \# \operatorname{class} y \operatorname{samples} \\ &= \sum_{j=1}^n 1(y_j = y) \end{aligned} \\ \Rightarrow \ln p(\mathcal{D}|\theta) = \sum_{y=1}^m n_y \ln p(y) + \sum_{i=1}^d \sum_{y=1}^m \sum_{j:y_j = y} \ln p(x_{ij}|y, \theta_{iy}) \\ \Rightarrow \frac{1}{n} \ln p(\mathcal{D}|\theta) = \sum_{y=1}^m \left(\frac{n_y}{n} \right) \ln \theta_{0y} + \sum_{i=1}^d \sum_{y=1}^m \left(\frac{n_y}{n} \right) \left[\frac{1}{n_y} \sum_{i:y_i = y} \ln p(x_{ij}|y, \theta_{iy}) \right] \end{aligned}$$

Thus,

$$\widehat{\theta}_{ML}(\mathcal{D}) = \arg \max_{\theta} \frac{1}{n} \ln p(\mathcal{D}|\theta),$$

$$\theta = (\theta_0, \{\theta_{iy}, i = 1, \dots, d, y = 1, \dots, m\})$$

$$\Rightarrow \widehat{\theta}_{0,ML}(\mathcal{D}) = \arg \max_{\theta_0} \left[\sum_{y=1}^{m} \left(\frac{n_y}{n} \right) \ln \theta_{0y} \right],$$

$$\forall i, y, \ \widehat{\theta}_{iy,ML}(\mathcal{D}) = \arg \max_{\theta_{iy}} \left[\frac{1}{n_y} \sum_{j: y_j = y} \ln p(x_{ij}|y, \theta_{iy}) \right]$$

$$= \widehat{\theta}_{iy,ML}(\mathcal{D}_{iy})$$

- All features are categorical:
 - Without loss of generality (w.l.o.g.): $\forall i, x_i \in \{1, \dots, W\}$
- Notation summary:

$$y = \text{class label } \in \{1, \dots, m\}$$

 $j = \text{sample index } \in \{1, \dots, n\}$
 $i = \text{feature (component) index } \in \{1, \dots, d\}$
 $w = \text{feature value } \in \{1, \dots, W\}$

Model:

$$orall y, j, i, w, \ P(X_{ij}=w|Y_j=y) = eta_{w,y,i}$$
, where $orall y, i, w, \ eta_{w,y,i} \geq 0, \ ext{and} \ orall y, i, \sum_{i=1}^{W} eta_{w,y,i} = 1$

Parameters:

$$\theta = \{\underbrace{p(y)}_{\theta_{0y}}, \underbrace{\beta_{w,y,i}, w = 1, \dots, W}_{\theta_{iy} = \{\beta_{w,y,i}, w = 1, \dots, W\}}, y = 1, \dots, m, i = 1, \dots, d\}$$

Total number of scalar parameters = m + mdW

Bayes classifier for 0-1 loss (MPE rule = MAP rule):

$$h_{\text{MPE}}(\mathbf{x}) = \arg \max_{y=1,\dots,m} p(y) p(\mathbf{x}|y,\theta)$$
$$= \arg \max_{y=1,\dots,m} p(y) \prod_{i=1}^{d} \beta_{x_i,y,i}$$

- ML (frequentist) estimate of θ : Let
 - $-n_v$ = number of class y training examples
 - $-n_{w,y,i}$ = number of class y training examples in which the i-th feature value = w
- then, $\sum_{y=1}^{m} n_y = n$ and for all i, $\sum_{w=1}^{W} n_{w,y,i} = n_y$
- Also, let $n_{w,y} = \sum_{i=1}^d n_{w,y,i} = \text{number of times } w$ occurs across all d components in class y examples then, $\sum_{w=1}^W n_{w,y} = dn_y$

ML (frequentist) estimate of θ:

$$\widehat{\theta}_{ML}(\mathcal{D}) = \arg \max_{\theta} \frac{1}{n} \ln p(\mathcal{D}|\theta),$$

$$\theta = (\theta_0, \{\theta_{iy}, i = 1, \dots, d, y = 1, \dots, m\})$$

$$\widehat{\theta}_{0,ML}(\mathcal{D}) = \arg \max_{\theta_0} \left[\sum_{y=1}^m \left(\frac{n_y}{n} \right) \ln \theta_{0y} \right],$$

$$\forall i, y, \ \widehat{\theta}_{iy,ML}(\mathcal{D}) = \arg \max_{\theta_{iy}} \left[\frac{1}{n_y} \sum_{j: y_j = y} \ln p(x_{ij}|y, \theta_{iy}) \right]$$

$$= \arg \max_{\theta_{iy}} \left[\frac{1}{n_y} \sum_{j: y_j = y} \ln \beta_{x_{ij}, y, i} \right]$$

$$= \arg \max_{\theta_{iy}} \left[\sum_{w=1}^W \left(\frac{n_{w,y,i}}{n_y} \right) \ln \beta_{w,y,i} \right]$$

ML (frequentist) estimate of θ solution:

$$\forall y, \ \widehat{\theta}_{0y,ML} = \widehat{p}(y) = \frac{n_y}{n}$$
$$\forall w, y, i, \ \widehat{\beta}_{w,y,i} = \frac{n_{w,y,i}}{n_y}$$

• If for all w,y,i, $\beta_{w,y,i} = \beta_{w,y}$ then,

$$\forall y, \ \widehat{\theta}_{0y,ML} = \widehat{p}(y) = \frac{n_y}{n}$$

$$\forall w, y, i, \ \widehat{\beta}_{w,y,i} = \frac{\sum_{i=1}^d n_{w,y,i}}{\sum_{i=1}^d n_y}$$

$$= \frac{n_{w,y}}{d \cdot n_y}$$

- Solution based on following result
- Result: Let $p_1, ..., p_L$, and $q_1, ..., q_L$, denote two pmfs over L items, i.e., they are non-negative and sum to one. Then,

$$\arg \max_{q_1, \dots, q_L} \sum_{l=1}^{L} p_l \ln q_l = \{ p_l, l = 1, \dots, L \}$$

• Proof: $\forall t > 0$, $\ln t \le t - 1$ with equality, if, and only if (iff), t = 1. Replacing t with q_l/p_l , multiplying by p_l , and summing over all l, we get

$$\sum_{l=1}^{L} p_l \ln \left(\frac{q_l}{p_l} \right) \le \sum_{l=1}^{L} p_l \left(\frac{q_l}{p_l} - 1 \right) = \sum_{l=1}^{L} \left(q_l - p_l \right) = 1 - 1 = 0.$$

Alternatively, this follows from the fact that the KL-divergence D(p||q) is non-negative and is zero if, and only if, the two pmfs p,q are identical

Overfitting problem

- If W >> n, it is quite likely that there is a value w_0 which never occurs in the training set, but occurs in a test example \mathbf{x}_{test}
- For such a w_0 , $n_{w_0,y,i} = \beta_{w_0,y,i} = 0$ for all y, i, and $p(\mathbf{x}_{\text{test}}|y,\hat{\theta}_{\text{ML}}) = 0$, $\forall y$, and the decision reduces to random guessing. This ignores information from values that were seen in both training and test sets
- Solution 1: remove words that were not seen during training and proceed as before. Better than random guessing, but still ignores information in new words
- Solution 2: Regularize estimation of β by incorporating prior beliefs via a pdf $\pi(\beta)$.

Bayesian Naïve Bayes with Dirichlet prior

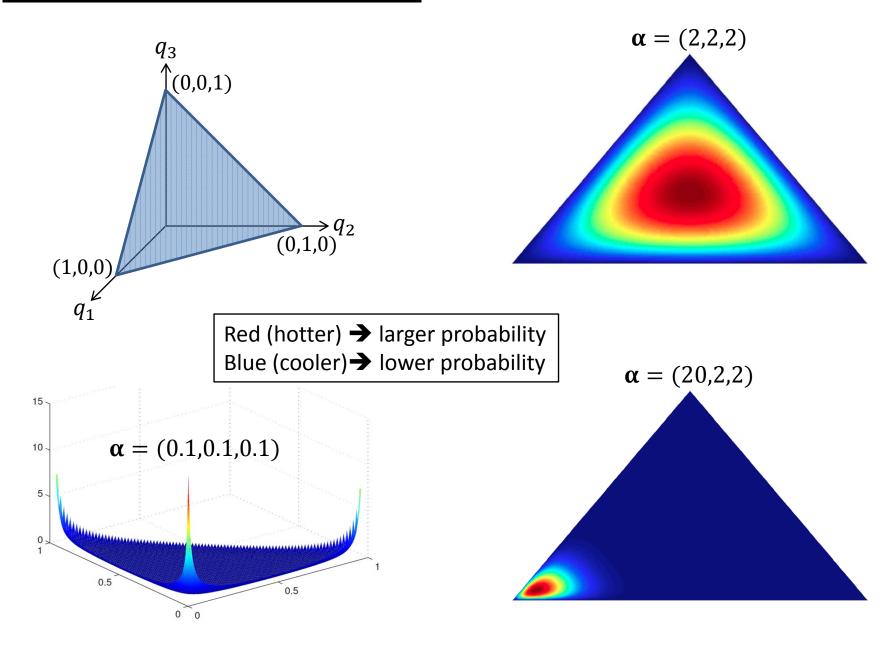
- Dirichlet distribution $Dir(\mathbf{q}|\mathbf{\alpha})$: is a family of continuous multivariate probability distributions parameterized by a W-dimensional vector $\mathbf{\alpha}$ of positive reals called the concentration parameters.
- It is a pdf over all pmfs over W values, i.e., a pdf over the (W-1)-dimensional probability simplex:

$$S_W := \left\{ \mathbf{q} : 0 \le q_i \le 1, i = 1, \dots, W, \sum_{i=1}^W q_i = 1 \right\}$$

$$\operatorname{Dir}(\mathbf{q}|\boldsymbol{\alpha}) := \frac{1(\mathbf{q} \in \mathcal{S}_W)}{Z(\boldsymbol{\alpha})} \prod_{i=1}^W q_i^{\alpha_i - 1}$$

$$Z(\boldsymbol{\alpha}) = \text{normalization constant} = \frac{1}{\Gamma(\sum_{i=1}^{W} \alpha_i)} \prod_{i=1}^{W} \Gamma(\alpha_i)$$

Dirichlet distribution



Dirichlet distribution

- $\alpha_0 := \sum_{i=1}^{W} \alpha_i$ controls how peaked the distribution is (larger \Rightarrow more peaked)
- Dir(1,1,1) is uniform over the probability simplex
- Dir(2,2,2) is a broad distribution centered at (1/3,1/3,1/3)
- Dir(20,20,20) is a narrow distribution centered at (1/3,1/3,1/3)
- If $\alpha_i < 0$ for all i, we get "spikes" at the corners of the probability simplex

Bayesian Naïve Bayes with Dirichlet prior

Dirichlet Prior for β:

$$\pi(\beta) = \prod_{y=1}^{m} \prod_{i=1}^{d} \frac{1}{Z(\boldsymbol{\alpha}_{yi})} \prod_{w=1}^{W} (\beta_{w,y,i})^{\alpha_{w,y,i}-1}$$

• MAP (Bayesian) estimate of θ :

$$\widehat{\theta}_{\text{MAP}} = \arg \max_{\theta} p(\theta | \mathcal{D})
= \arg \max_{\theta} p(\mathcal{D} | \theta) \pi(\beta)
= \arg \max_{\theta} \left[\prod_{y=1}^{m} (p(y))^{n_y} \right] \cdot \left[\prod_{w=1}^{W} \prod_{y=1}^{m} \prod_{i=1}^{d} (\beta_{w,y,i})^{n_{w,y,i} + \alpha_{w,y,i} - 1} \right]$$

Bayesian Naïve Bayes with Dirichlet prior

• MAP (Bayesian) estimate of θ solution:

$$\forall y, \ \widehat{p}(y) = \frac{n_y}{n}$$

$$\forall w, y, i, \ \widehat{\beta}_{w,y,i} = \frac{n_{w,y,i} + \alpha_{w,y,i} - 1}{n_y + \sum_{w=1}^{W} (\alpha_{w,y,i} - 1)}$$

• If for all w,y,i, $\beta_{w,y,i}=\beta_{w,y}$ then taking $\alpha_{w,y,i}=\alpha_{w,y}$ for all w,y,i, we get

$$\forall y, \ \widehat{p}(y) = \frac{n_y}{n}$$

$$\forall w, y, i, \ \widehat{\beta}_{w,y,i} = \frac{n_{w,y} + \alpha_{w,y} - 1}{dn_y + \sum_{w=1}^{W} (\alpha_{w,y} - 1)}$$

Remarks

- alphas can be interpreted as "prior" counts and the MAP solution as updating these prior counts with empirical counts from the likelihood
- If all alphas are equal to 2, then the prior counts are equal to one. This is referred to as add-one smoothing of the ML estimate or Laplace's rule of succession.
- Can also incorporate a separate Dirichlet prior for p(y) in a similar way
- The ML estimates of θ are asymptotically consistent