ME8135 - State Estimation - Assignment 1

Austin Vuong

May 2023

1 Question 1

 \mathbf{x} is a random variable of length K:

$$\mathbf{x} = \mathcal{N}(\mathbf{0}, \mathbf{1}) \tag{1}$$

a) What type of random variable is the following random variable?

$$\mathbf{y} = \mathbf{x}^{\top} \mathbf{x} \tag{2}$$

 \mathbf{y} is a continuous random variable.

b) Calculate the mean and variance of y. First a few notes as helpers:

$$\mathbf{y} = \sum_{i}^{K} x_i^2 = (x_1^2 + \dots + x_K^2)$$

and

$$E[x_i x_i] = 1, E[x_i x_j] = 0$$
 when $i \neq j$

by the trivial case of (2.41), $E[\mathbf{x}\mathbf{x}^{\top}] = \Sigma$.

The mean is:

$$\mu = E[\mathbf{y}] = E[\sum_{i}^{K} x_{i}^{2}]$$

$$= E[x_{1}x_{1}] + \dots + E[x_{K}x_{K}]$$

$$= K$$

The variance is:

$$Var(\mathbf{y}) = E[(\mathbf{y} - \mu)^{2}] = E[(\mathbf{y} - K)^{2}]$$

$$= E[\mathbf{y}^{2} - 2\mathbf{y}K + K^{2}]$$

$$= E[\mathbf{y}^{2}] - 2KE[\mathbf{y}] + K^{2}$$

$$= E[\mathbf{y}^{2}] - K^{2}$$

$$= E[(\sum_{i}^{K} x_{i}^{2})^{2}] - K^{2}$$

$$= E[\sum_{i}^{K} x_{i}^{4} + \sum_{i=1}^{K} \sum_{j \neq i}^{K} x_{i}^{2} x_{j}^{2}] - K^{2}$$

$$= \sum_{i}^{K} E[x_{i}x_{i}x_{i}x_{i}] + \sum_{i=1}^{K} \sum_{j \neq i}^{K} E[x_{i}x_{i}x_{j}x_{j}] - K^{2}$$

$$= \sum_{i}^{K} 3 + \sum_{i=1}^{K} \sum_{j \neq i}^{K} 1 - K^{2}$$

$$= 3K + K(K - 1) - K^{2}$$

$$= 3K - K + K^{2} - K^{2}$$

$$= 2K$$

where the expectations are found using (2.40) and the earlier helper note to obtain:

$$E[x_i x_i x_i x_i] = 3E[x_i x_i] E[x_i x_i] = 3$$

and

$$E[x_i x_i x_j x_j] = E[x_i x_i] E[x_j x_j] + E[x_i x_j] E[x_i x_j] + E[x_i x_j] E[x_i x_j]$$

= 1 + 0 + 0 = 1

c) Using Python, plot the PDF of \mathbf{y} for $K=1,\,2,\,3,\,10,\,100.$

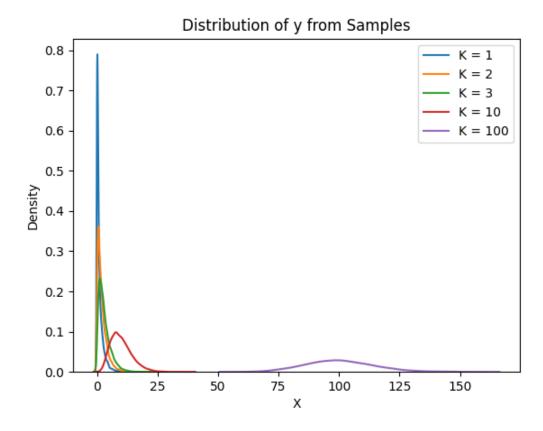


Figure 1: Sampled plot of the PDF of y for K = 1, 2, 3, 10, 100

2 Question 2

 \mathbf{x} is a random variable of length N:

$$\mathbf{x} = \mathcal{N}(\mu, \mathbf{\Sigma}) \tag{3}$$

a) Assume \mathbf{x} is transformed linearly, i.e. $\mathbf{y} = \mathbf{A}\mathbf{x}$, where \mathbf{A} is an $N \times N$ matrix. Calculate the mean and covariance of \mathbf{y} . Show the derivations.

I take the question to mean the covariance matrix. In this case of matrix \mathbf{A} we can apply the expectation operator directly as is done in (2.58).

$$\begin{split} \mu_y = E[\mathbf{y}] = E[\mathbf{A}\mathbf{x}] = \mathbf{A}E[\mathbf{x}] = \mathbf{A}\mu \\ \mathbf{\Sigma}_{yy} = E[(\mathbf{y} - \mu_y)(\mathbf{y} - \mu_y)^\top] = \mathbf{A}E[(\mathbf{x} - \mu)(\mathbf{x} - \mu)^\top]\mathbf{A}^\top = \mathbf{A}\mathbf{\Sigma}\mathbf{A}^\top \end{split}$$

b) Repeat a), when $\mathbf{y} = (\mathbf{A_1} + \mathbf{A_2})$

$$A_1x + A_2x = (A_1 + A_2)x = Ax$$

Now substitute into a).

$$\mu_y = \mathbf{A}\mu = (\mathbf{A_1} + \mathbf{A_2})\mu$$

$$oldsymbol{\Sigma}_{yy} = \mathbf{A} oldsymbol{\Sigma} \mathbf{A}^ op = (\mathbf{A_1} + \mathbf{A_2}) oldsymbol{\Sigma} (\mathbf{A_1} + \mathbf{A_2})^ op$$

c) If **x** is transformed by a nonlinear differentiable function, i.e. $\mathbf{y} = f(\mathbf{x})$, compute the covariance matrix of **y**. Show the derivation.

The Jacobian, **J**, can be used in place of a linear transform in order to use a nonlinear differentiable function. It can be substituted into the covariance derivation from a).

$$\boldsymbol{\Sigma}_{yy} = E[(\mathbf{y} - \mu_y)(\mathbf{y} - \mu_y)^{\top}] = \mathbf{J}E[(\mathbf{x} - \mu)(\mathbf{x} - \mu)^{\top}]\mathbf{J}^{\top} = \mathbf{J}\boldsymbol{\Sigma}\mathbf{J}^{\top}$$

d) Apply c) when

$$\mathbf{x} = \begin{bmatrix} \rho \\ \theta \end{bmatrix}, \mathbf{\Sigma} = \begin{bmatrix} \sigma_{\rho\rho}^2 & \sigma_{\rho\theta}^2 \\ \sigma_{\rho\theta}^2 & \sigma_{\theta\theta}^2 \end{bmatrix}, \text{ and } \mathbf{y} = \begin{bmatrix} \rho \cos \theta \\ \rho \sin \theta \end{bmatrix}$$
(4)

Compute the covariance of y analytically. This models how range-bearing measurements in the polar coordinate frame are converted to a Cartesian coordinate frame.

For brevity we will replace each term with a variable then compute the covariance. The terms can be later substituted back.

$$\begin{split} \boldsymbol{\Sigma}_{yy} &= \mathbf{J} \boldsymbol{\Sigma} \mathbf{J}^{\top} \\ &= \begin{bmatrix} \frac{\partial f_1}{\partial \rho} & \frac{\partial f_1}{\partial \theta} \\ \frac{\partial f_2}{\partial \rho} & \frac{\partial f_2}{\partial \theta} \end{bmatrix} \begin{bmatrix} \sigma_{\rho\rho}^2 & \sigma_{\rho\theta}^2 \\ \sigma_{\rho\theta}^2 & \sigma_{\theta\theta}^2 \end{bmatrix} \begin{bmatrix} \frac{\partial f_1}{\partial \rho} & \frac{\partial f_1}{\partial \theta} \\ \frac{\partial f_2}{\partial \rho} & \frac{\partial f_2}{\partial \theta} \end{bmatrix}^{\top} \\ &= \begin{bmatrix} \cos \theta & -\rho \sin \theta \\ \sin \theta & \rho \cos \theta \end{bmatrix} \begin{bmatrix} \sigma_{\rho\rho}^2 & \sigma_{\rho\theta}^2 \\ \sigma_{\rho\theta}^2 & \sigma_{\theta\theta}^2 \end{bmatrix} \begin{bmatrix} \cos \theta & \sin \theta \\ -\rho \sin \theta & \rho \cos \theta \end{bmatrix} \\ &= \begin{bmatrix} a & b \\ c & d \end{bmatrix} \begin{bmatrix} e & f \\ g & h \end{bmatrix} \begin{bmatrix} a & c \\ b & d \end{bmatrix} \\ &= \begin{bmatrix} ae + bg & af + bh \\ ce + dg & cf + dh \end{bmatrix} \begin{bmatrix} a & c \\ b & d \end{bmatrix} \\ &= \begin{bmatrix} (ae + bg)a + (af + bh)b & (ae + bg)c + (af + bh)d \\ (ce + dg)a + (cf + dh)b & (ce + dg)c + (cf + dh)d \end{bmatrix} \end{split}$$

Where the covariance terms are as follows and are left unsimplified for clarity:

$$(ae + bg)a + (af + bh)b = (\sigma_{\rho\rho}^2 \cos \theta + \sigma_{\rho\theta}^2 \sin \theta)\cos \theta + (\sigma_{\rho\theta}^2 \cos \theta - \sigma_{\theta\theta}^2 \sin \theta)(-\rho \sin \theta)$$

$$(ae + bg)c + (af + bh)d = (\sigma_{\rho\rho}^2 \cos \theta + \sigma_{\rho\theta}^2 \sin \theta)\sin \theta + (\sigma_{\rho\theta}^2 \cos \theta - \sigma_{\theta\theta}^2 \sin \theta)(\rho \cos \theta)$$

$$(ce + dg)a + (cf + dh)b = (\sigma_{\rho\rho}^2 \sin \theta + \sigma_{\rho\theta}^2 \cos \theta)\cos \theta + (\sigma_{\rho\theta}^2 \sin \theta + \sigma_{\theta\theta}^2 \cos \theta)(-\rho \sin \theta)$$

$$(ce + dg)c + (cf + dh)d = (\sigma_{\rho\rho}^2 \sin \theta + \sigma_{\rho\theta}^2 \cos \theta)\sin \theta + (\sigma_{\rho\theta}^2 \sin \theta + \sigma_{\theta\theta}^2 \cos \theta)(\rho \cos \theta)$$

e) Simulate d) using the Monte Carlo simulation, i.e. assume

$$\mathbf{x} = \begin{bmatrix} 1m \\ 0.5^{\circ} \end{bmatrix}, \mathbf{\Sigma} = \begin{bmatrix} 0.01 & 0 \\ 0 & 0.005 \end{bmatrix} \tag{5}$$

Sample 1000 points from this distribution and plot the transformed results on x-y coordinates. Plot the uncertainty ellipse, calculated from part d). Overlay the ellipse on the point samples.

I assume the given \mathbf{x} here as the mean, μ , and apply it to equation 3. I also assume the given variance of θ is in degrees. The result of the latter is a clearer contrast between the inputs. The uncertainty ellipse represents a 95% confidence. The scale of the axes are equal in their respective plots.

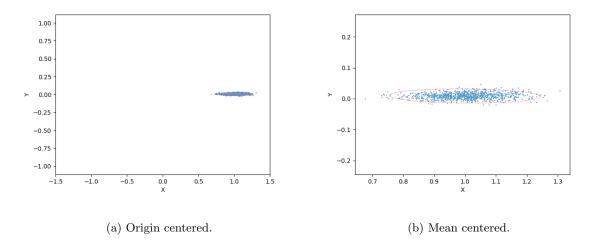


Figure 2: Plots of the transformed distribution.