

# Causality in Biomedicine

## Lecture Series: Lecture 3

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# Info

**GitHub:** [https://github.com/avakhamseh/Causality\\_in\\_Biomedicine\\_Lectures](https://github.com/avakhamseh/Causality_in_Biomedicine_Lectures)

**Causal machine learning master class, deadline 11 Feb 2020**

**Tuesday 03 Mar 2020, Time: 10:00-17:30**

Organised by Alan Turing Institute, London

- Average treatment effect:  $\tau = \hat{\mathbb{E}}[\tau^{(i)}] = \hat{\mathbb{E}}[y_1^{(i)} - y_0^{(i)}] = \frac{1}{N} \sum_{i=0}^N \left( y_1^{(i)} - y_0^{(i)} \right)$
- Targeted maximum likelihood estimation for ATE
- Doubly robust estimates: Regression + propensity score
- Machine learning considerations
- Longitudinal data with time-dependent confounding

# So Far ...

- **Matching:** Stratification, Propensity score, IPTW, ...

$$y_1^{(i)}, y_0^{(i)} \perp\!\!\!\perp t^{(i)} \mid b(x)$$

- Estimation of propensity scores directly from the data & algorithms

$$e(x) = p(t = 1 \mid x)$$

- **Sensitivity analysis:** No guarantee that matching leads to balance on variables we did not match for, people who look comparable may differ.

If there is hidden bias, how severe is it:

- Does the conclusion change from statistically significant to not?
- Does it change the direction of effect?

**Notice:** There are **two sources of uncertainty** (confidence interval):

- 1) Due to the causal estimates
- 2) Due to sensitivity analysis

# Overview of the field

## Learning Causality with Data

### Learning Causal Effects

with Unconfoundedness

Regression Adjustment  
Propensity Score

Covariate Balancing

with Unobserved Confounders

IV

Front-door Criterion

RDD

### Learning Causal Relationships

i.i.d. Data

Constraint-based

Score-based

FCMs

non-i.i.d. Data

Constraint-based

FCMs

### Connections to Machine Learning

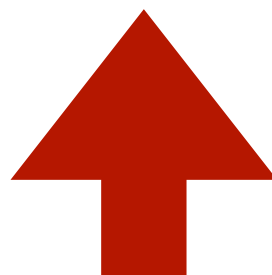
Supervised Learning and SSL

Domain Adaptation

RL

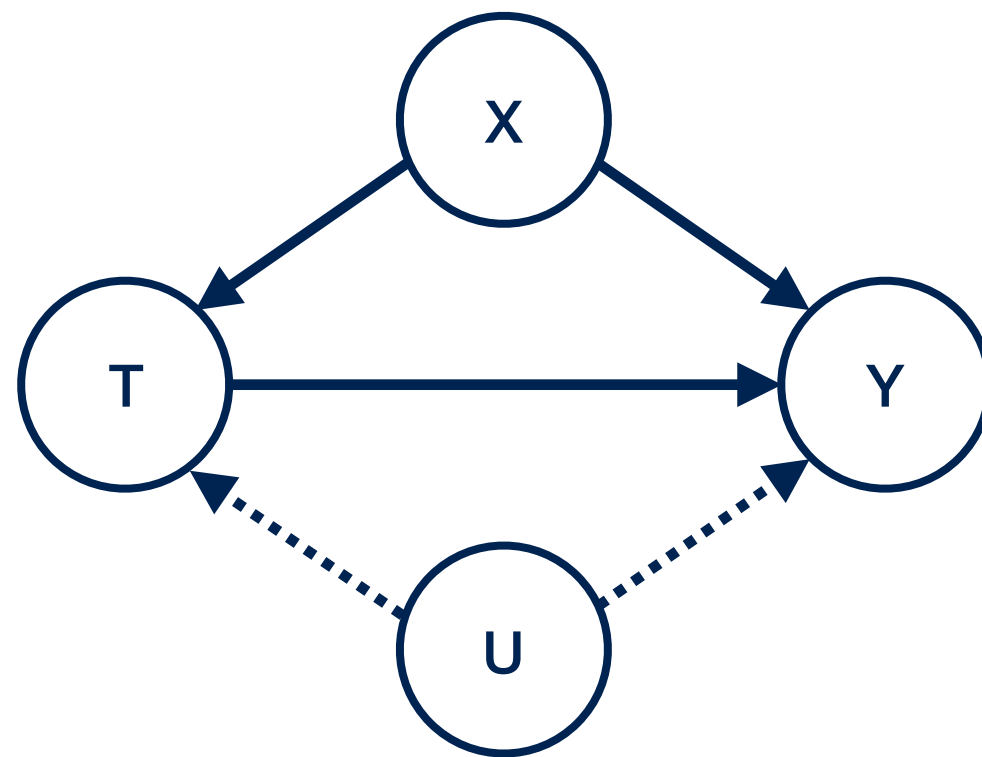
Rubin

Rubin, Pearl



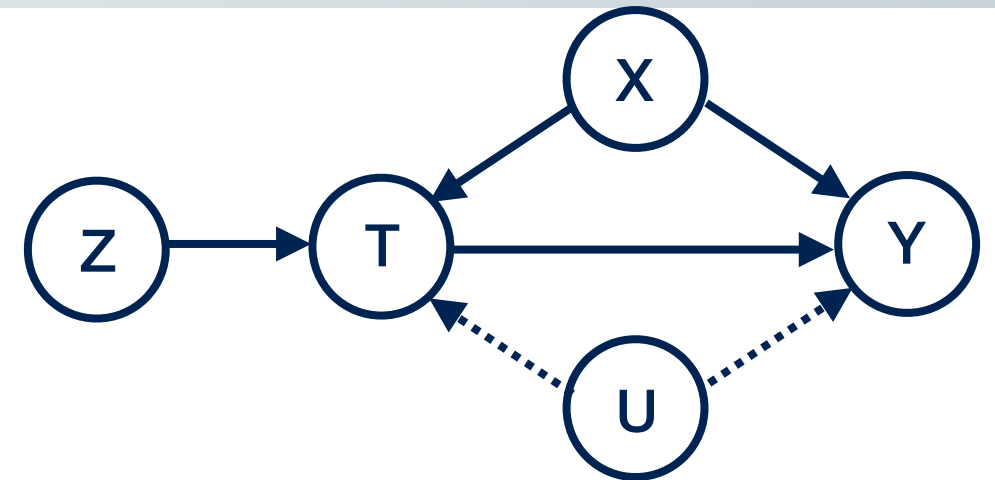
# Instrumental Variable (Originally due to Rubin)

- Unobserved confounders (U), **violates unconfoundedness (ignorability)**, i.e. conditioning on X alone, would not results in a randomised treatment assignment
- Unconfoundedness is fundamentally unverifiable



# Instrumental Variable example

- Example 1:
  - T: smoking during pregnancy
  - Y: birthweight
  - X: parity, mother's age, weight, ...
  - U: Other unmeasured confounders



- Randomise Z (intention-to-treat): either receive encouragement to stop smoking ( $Z=1$ ), or receive usual care ( $Z=0$ )
- Intention-to-treat analysis gives causal effect estimator of encouragement  $z$  on outcome  $y$ :

$$\mathbb{E}(y|z = 1) - \mathbb{E}(y|z = 0)$$

- What can we say about the causal effect of smoking itself?

# Instrumental Variable assumptions

- **SUTVA**: Potential outcomes for each individual  $i$  are unrelated to the treatment status of other individuals:

$$Y^{(i)}(\mathbf{Z}, \mathbf{T}) = Y^{(i)}(Z^{(i)}, T^{(i)}) , \quad |\mathbf{Z}| = |\mathbf{T}| = N \text{ individuals}$$

- Treatment assignment  $Z$  (**associated** with the treatment) is **random**:

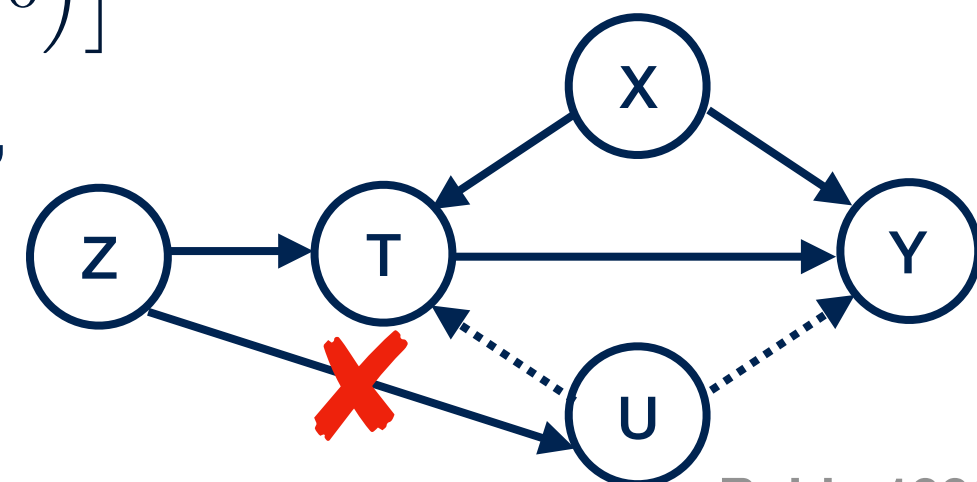
$$P(Z^{(i)} = 0) = P(Z^{(i)} = 1) , \quad \forall i$$

- **Exclusion Restriction**: Any effect of  $Z$  on  $Y$  is via an effect of  $Z$  on  $T$ , i.e.,  $Z$  should not affect  $Y$  when  $T$  is held constant  $\left(Y^{(i)}|_{z=1,t}\right) = \left(Y^{(i)}|_{z=0,t}\right)$

- **Non-zero Average**:  $\mathbb{E} \left[ \left(T^{(i)}|_{z=1}\right) - \left(T^{(i)}|_{z=0}\right) \right]$

- **Monotonicity** (increasing encouragement “dose” increases probability of treatment, no defiers):

$$\left(T^{(i)}|_{z=1}\right) \geq \left(T^{(i)}|_{z=0}\right)$$



# Instrumental Variable: Potential values of T

Population	T z=0	T z=1	Description
Never-takers	0	0	Causal effect of Z on Y is zero, since $\left(T^{(i)} z=1\right) - \left(T^{(i)} z=0\right) = 0$
Compliers	0	1	Treatment received is randomised, $\left(T^{(i)} z=1\right) - \left(T^{(i)} z=0\right) = 1$ <u>causal effect inference</u> : $\left(Y^{(i)} T^{(i)}=1\right) - \left(Y^{(i)} T^{(i)}=0\right)$
Defiers	1	0	Rule out by <b>monotonicity</b> , since $\left(T^{(i)} z=1\right) - \left(T^{(i)} z=0\right) = -1$
Always-takers	1	1	Causal effect of Z on Y is zero, since $\left(T^{(i)} z=1\right) - \left(T^{(i)} z=0\right) = 0$



# Instrumental Variable: The estimand

**Want ATE:**  $\mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \right]$



**“Almost”**

**Will estimate:**

$$\hat{\tau} = \frac{\mathbb{E} \left[ \left( Y^{(i)} | z = 1 \right) - \left( Y^{(i)} | z = 0 \right) \right]}{\mathbb{E} \left[ \left( T^{(i)} | z = 1 \right) - \left( T^{(i)} | z = 0 \right) \right]}$$

# Instrumental Variable: The estimand

$$\mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \right]$$

$$\hat{\tau} = \frac{\mathbb{E} \left[ \left( Y^{(i)} | z = 1 \right) - \left( Y^{(i)} | z = 0 \right) \right]}{\mathbb{E} \left[ \left( T^{(i)} | z = 1 \right) - \left( T^{(i)} | z = 0 \right) \right]}$$

**Derivation:**

$$\begin{aligned} & \left( Y^{(i)} | T^{(i)}(z = 1) \right) - \left( Y^{(i)} | T^{(i)}(z = 0) \right) \quad \text{t is either t=0 or t=1, and exclusion restriction} \\ &= \left[ Y^{(i)} \left( t^{(i)} = 1 \right) \cdot \left( t^{(i)} | z = 1 \right) + Y^{(i)} \left( t^{(i)} = 0 \right) \cdot \left( 1 - \left( t^{(i)} | z = 1 \right) \right) \right] \\ &- \left[ Y^{(i)} \left( t^{(i)} = 1 \right) \cdot \left( t^{(i)} | z = 0 \right) + Y^{(i)} \left( t^{(i)} = 0 \right) \cdot \left( 1 - \left( t^{(i)} | z = 0 \right) \right) \right] \\ &= \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \cdot \left( \left( t^{(i)} | z = 1 \right) - \left( t^{(i)} | z = 0 \right) \right) \end{aligned}$$

Hence, the causal effect of Z on Y for individual i, is the product of the causal effect of Z on T, and, the casual effect of T on Y.

# Instrumental Variable: The estimand

To continue the derivation, we use the fact that:

$$\mathbb{E}[xy] = \int \int xy p(x, y) dx dy = \int dy y p(y) \int dx x p(x|y) = \int dy y p(y) \mathbb{E}[x|y]$$

and write,

$$\begin{aligned} & \mathbb{E} \left[ \left( Y^{(i)} | T^{(i)}(z=1) \right) - \left( Y^{(i)} | T^{(i)}(z=0) \right) \right] \\ &= \mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \cdot \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) \right) \right] \end{aligned} \quad \nearrow \mathbf{0, 1, -1}$$

# Instrumental Variable: The estimand

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and write,

$$\begin{aligned} & \mathbb{E} \left[ \left( Y^{(i)} | T^{(i)}(z=1) \right) - \left( Y^{(i)} | T^{(i)}(z=0) \right) \right] \quad \xrightarrow{\quad} \mathbf{0, 1, -1} \\ &= \mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \cdot \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) \right) \right] \\ &= \mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \mid \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) \right) = 1 \right] \cdot \\ & \quad P \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) = 1 \right) \\ & - \mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \mid \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) \right) = -1 \right] \cdot \\ & \quad P \left( \left( t^{(i)} | z=1 \right) - \left( t^{(i)} | z=0 \right) = -1 \right) \\ & \quad \xleftarrow{\quad} \mathbf{0, by monotonicity} \end{aligned}$$

# Instrumental Variable: The estimand

$$\frac{\mathbb{E} \left[ \left( Y^{(i)} | T^{(i)}(z = 1) \right) - \left( Y^{(i)} | T^{(i)}(z = 0) \right) \right]}{\mathbb{E} \left[ \left( t^{(i)} | z = 1 \right) - \left( t^{(i)} | z = 0 \right) \right]}$$
$$= \mathbb{E} \left[ \left( Y^{(i)} \left( t^{(i)} = 1 \right) - Y^{(i)} \left( t^{(i)} = 0 \right) \right) \mid \left( \left( t^{(i)} | z = 1 \right) - \left( t^{(i)} | z = 0 \right) \right) = 1 \right]$$

i.e. restricting to *compliers*, the average casual effect of Z on Y is proportional to the average causal effect of T on Y.

Rubin 1996

- In this example, Z was randomly assigned as part of the study
- IV can also be randomised in nature (nature randomiser):
  - Mendelian randomisation
  - Quarter of birth

# **Pearl's framework**

## **Graphical models & Do-calculus**

# Causal Inference: DoWhy (a unifying language)

- **Model** a causal inference problem using assumptions, [**Pearl's** Causal Graphical Models]
- **Identify** an expression for the causal effect under these assumptions (“causal estimand”), [**Pearl's** Causal Graphical Models]
- **Estimate** the expression using statistical methods such as matching or instrumental variables, [**Rubin's** Potential Outcomes] ✓
- **Verify** the validity of the estimate using a variety of robustness checks. ✓

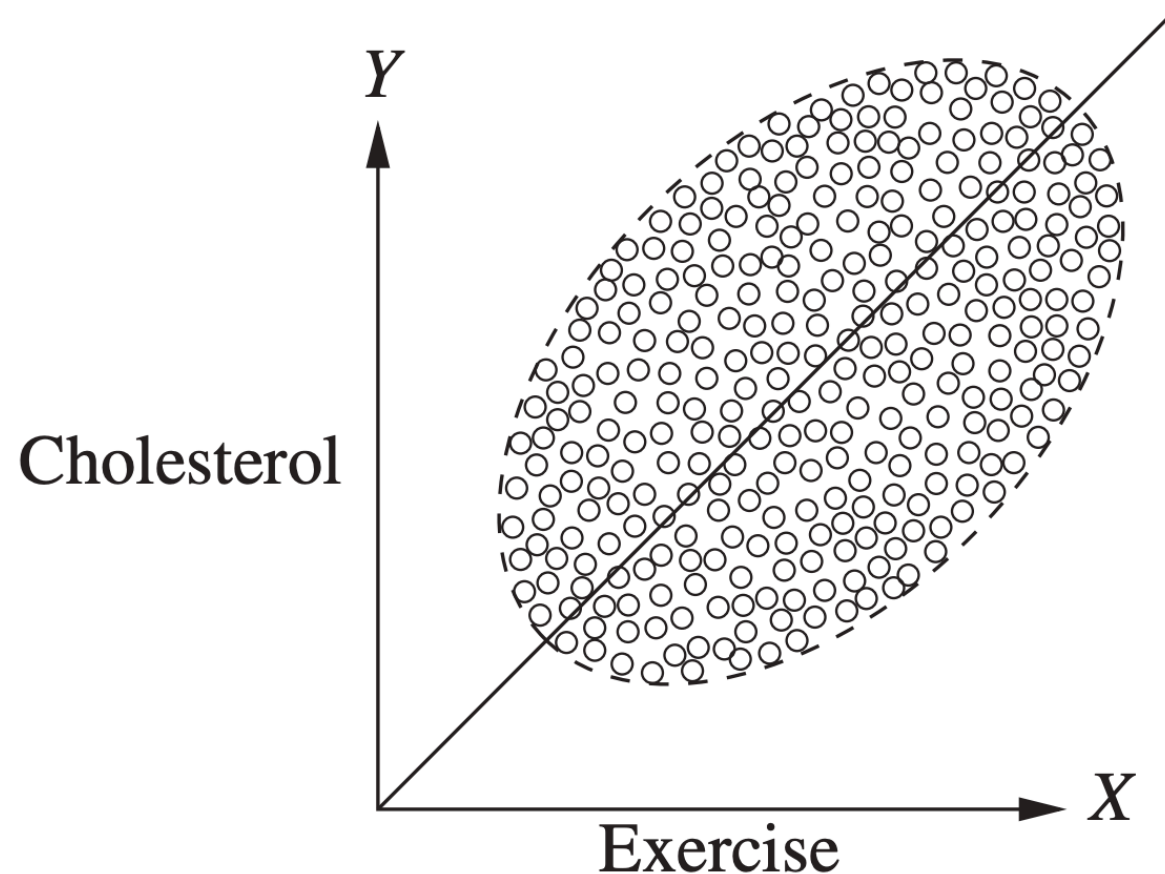
# Pearl's Causal Model

- Ladder of causation:
  - **Association:** What does a symptom tell me about a disease?
  - **Intervention (perturbation):** If I take aspirin will my headache be cured?
  - **Counterfactual:** Was it the aspirin that stopped the headache?  
(alternative versions of past events, strongest causal statements e.g. **physical laws**)
- Aim: To **model** and **identify** the causal estimand
- Causal graphical models + structural equations



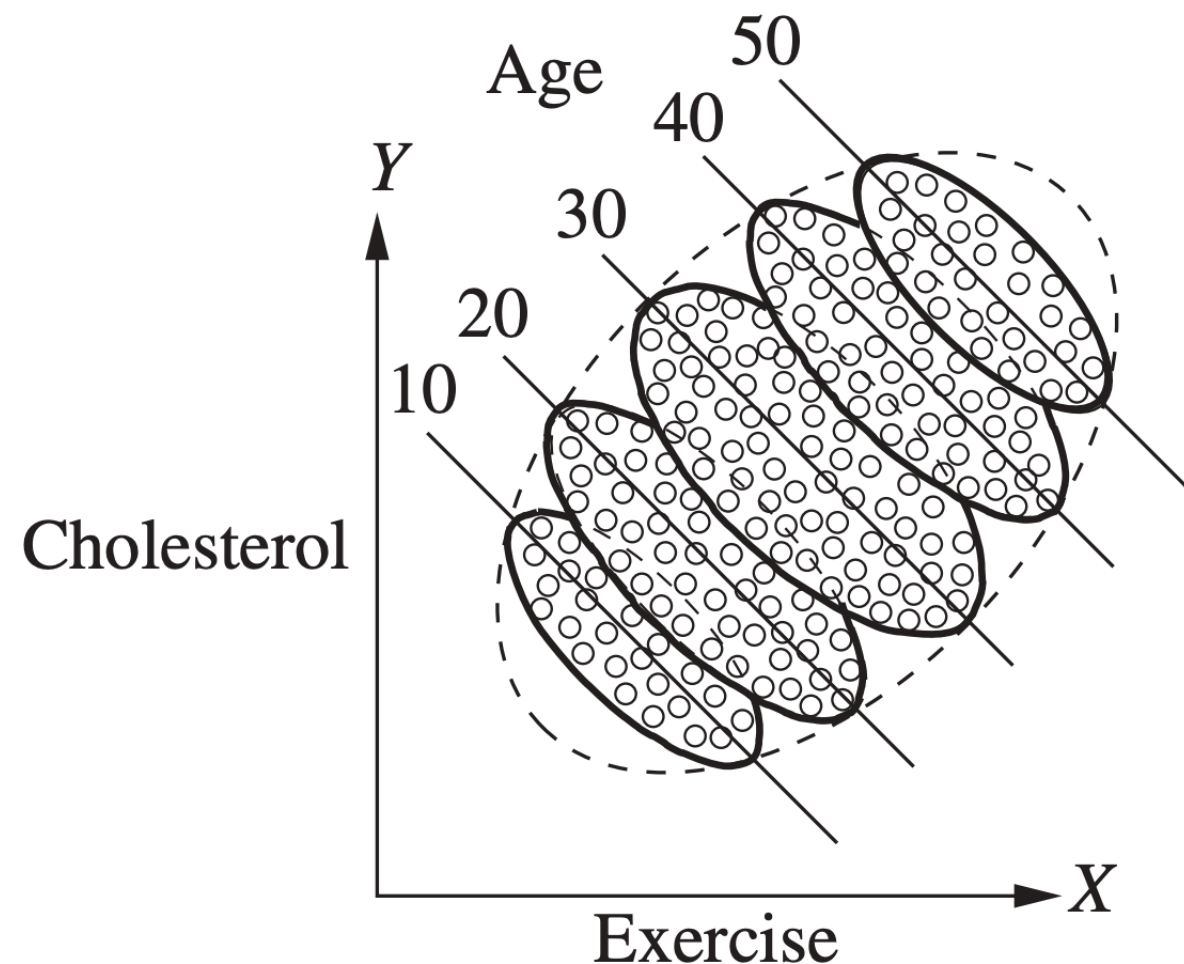
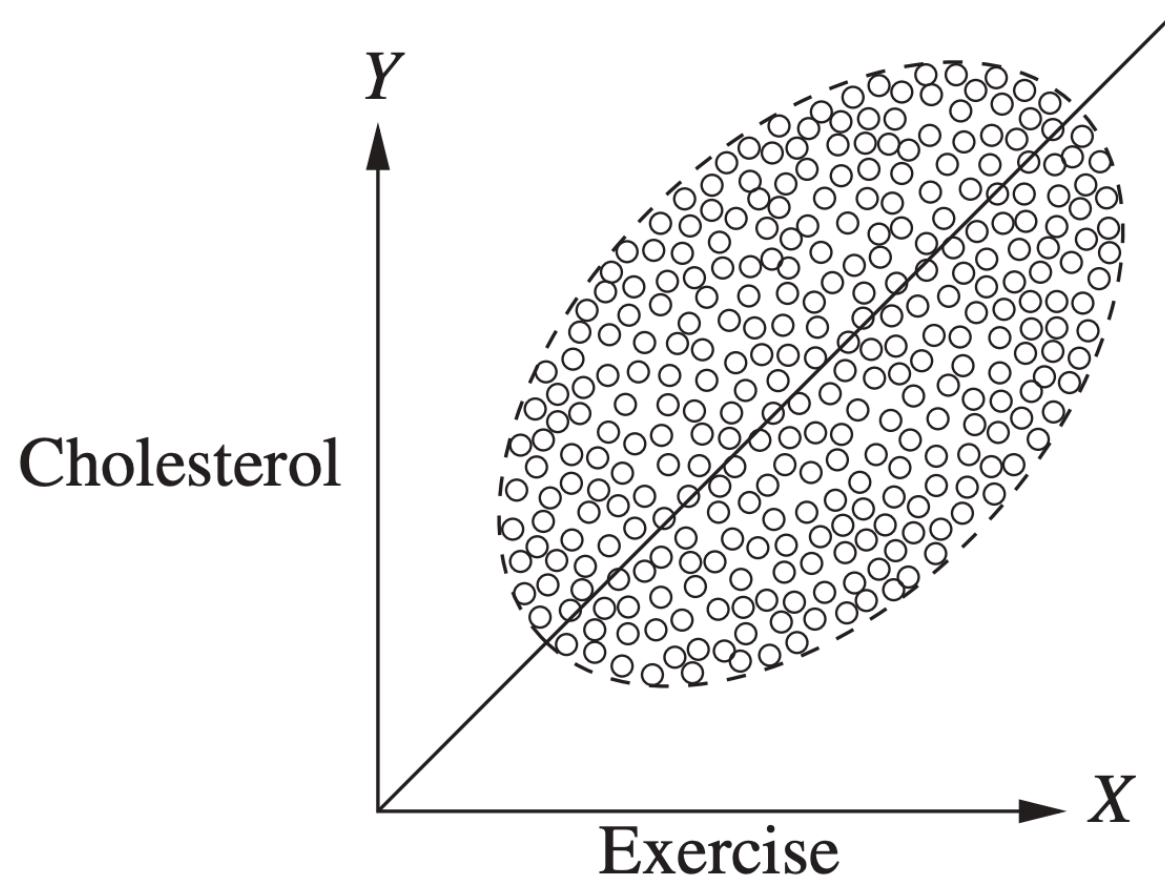
# Simpson's Paradox

- Why concluding causality from purely associational measures, i.e. correlation, can be **very wrong** (not just neutral): “It would have better not to make any statements!”



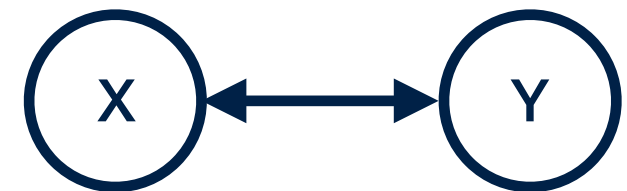
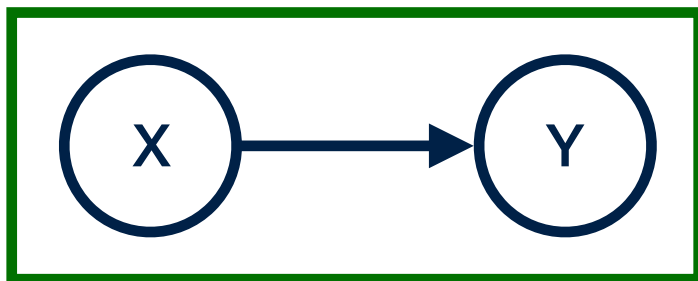
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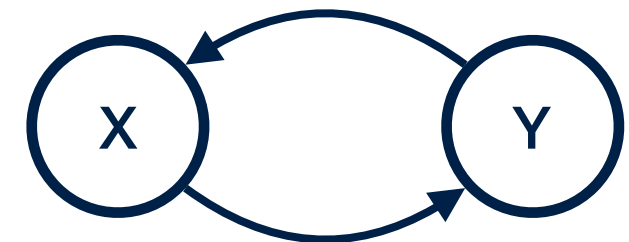


# Causal Graphical Models

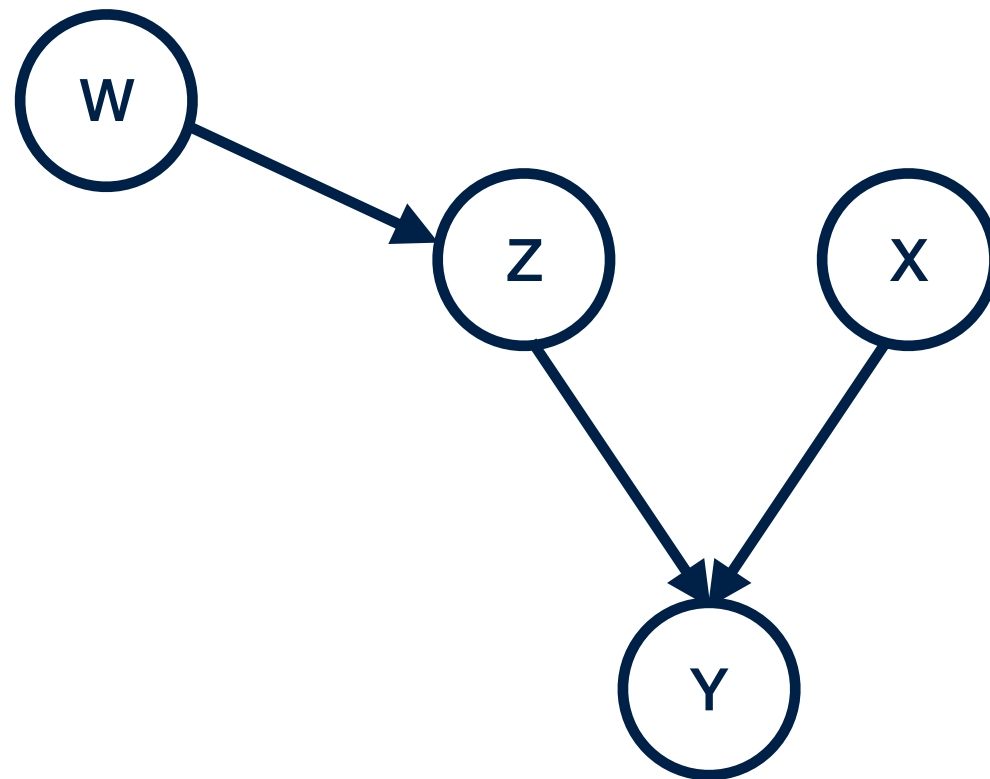
- Diagrammatic representation of probability distributions + **causal info**
- **Graph:** Consists of a set of **vertices**  $V$  (nodes), **edges**  $E$
- $V$  are the variables and  $E$  contains information between the variables
- Graphs can be directed, undirected and bidirectional (confounder?)



- Directed graphs may include directed cycles, i.e., mutual causation/feed-back process.
- A graph with no directed cycles is an **acyclic** graph.



# Directed Acyclic Graphs (DAGs)



Z, X are parents of Y  
Z, X, W are ancestors of Y  
Y has no children  
X has no parents

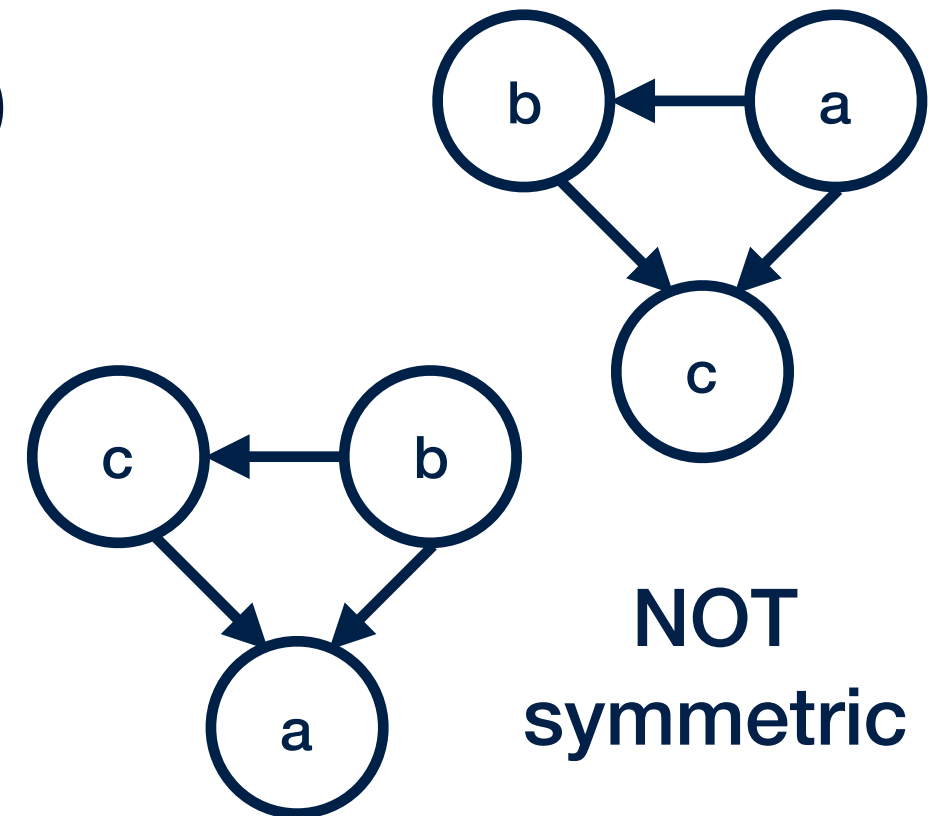
- DAG in which every node has at most one parent is a **tree**
- A tree in which every node has at most one child is a **chain**
- **DAG:**
  - Expresses **model assumptions** explicitly
  - Represent **joint probability** functions
  - Provides **efficient inference** of observations

# DAG contains more info than joint probability

$$p(a, b, c) = p(c|a, b)p(a, b) = p(c|a, b)p(b|a)p(a)$$

$$p(a, b, c) = p(a|b, c)p(b, c) = p(a|b, c)p(c|b)p(b)$$

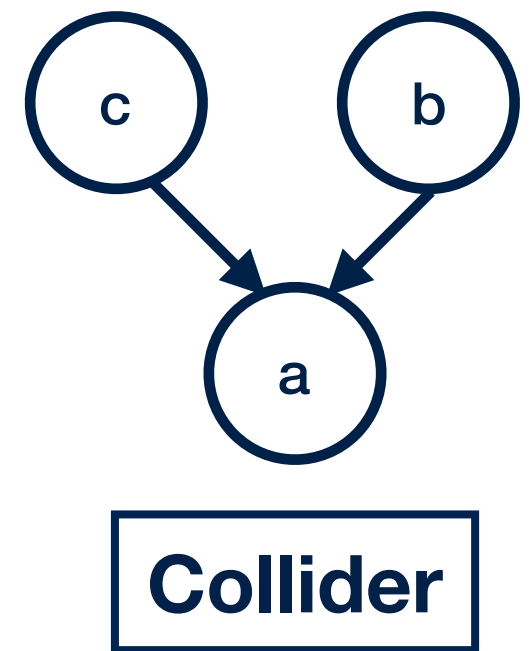
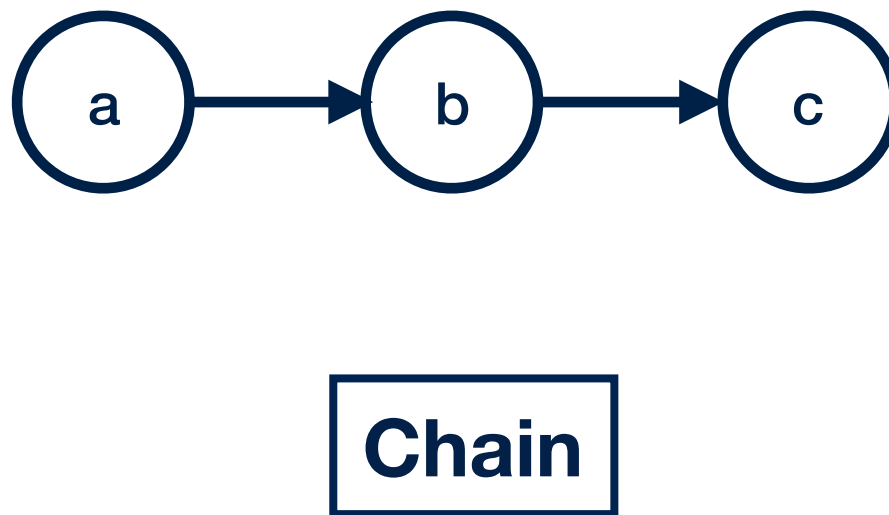
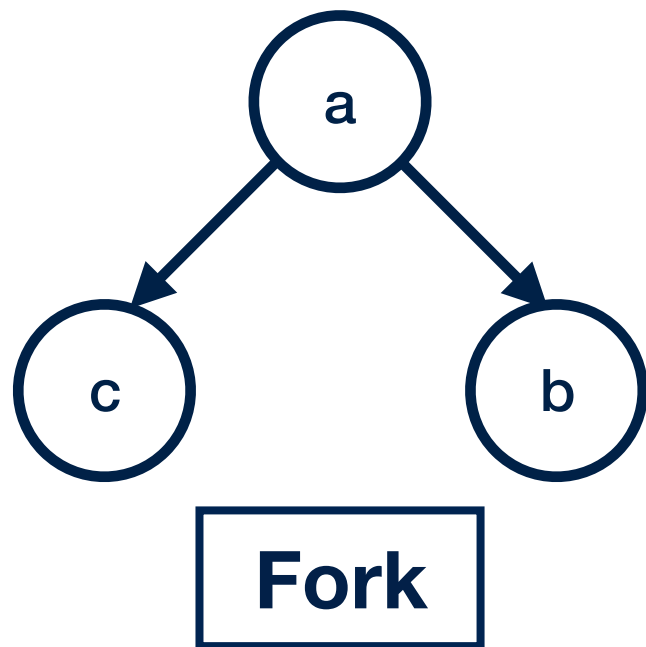
Symmetric  
in a, b, c



- Probabilistic notations are not enough to describe causal aspects
- Using repeated application of Bayes' rule, one can write any joint probability distribution in terms of its marginals
- A graph is **fully connected** if there is a link between every pair of nodes
- The interest lies in the **absence** of a link and link **direction**.

# Next time

- Conditional independence via graphs
- D-separation
- Discuss the 3 main graph structures:



- **Do-calculus** and **identification**

# Causality in Biomedicine

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