#### Model checking - overview

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  - e.g., if posterior would claim that hazardous chemical decreases probability of death
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  - compare predictions to completely new observations
  - cf. relativity theory predictions
- Internal validation
  - posterior predictive checking
  - cross-validation predictive checking

#### Chapter 6

- 6.1 The place of model checking in applied Bayesian statistics
- 6.2 Do the inferences from the model make sense?
- 6.3 Posterior predictive checking
- 6.4 Graphical posterior predictive checks
  - this can be skimmed, see instead the paper Gabry et al. (2019). Visualization in Bayesian workflow https://doi.org/10.1111/rssa.12378
- 6.5 Model checking for the educational testing example

#### Model checking

- demo6\_1: Posterior predictive checking light speed
- demo6\_2: Posterior predictive checking sequential dependence
- demo6\_3: Posterior predictive checking poor test statistic
- https://avehtari.github.io/BDA\_R\_demos/demos\_rstan/brms\_demo.html

#### Simon Newcomb's light of speed experiment in 1882

Newcomb measured (n = 66) the time required for light to travel from his laboratory on the Potomac River to a mirror at the base of the Washington Monument and back, a total distance of 7422 meters.

- Newcomb's speed of light measurements
  - model  $y \sim \text{normal}(\mu, \sigma)$  with prior  $(\mu, \log \sigma) \propto 1$

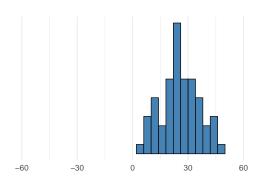
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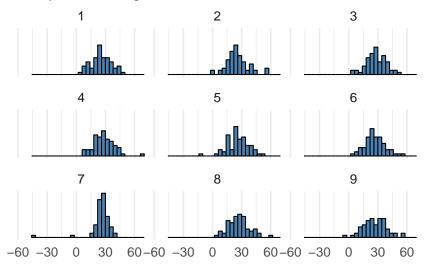
#### Replicates vs. future observation

Predictive ỹ is the next not yet observed possible observation.
 y<sup>rep</sup> refers to replicating the whole experiment (potentially with same values of x) and obtaining as many replicated observations as in the original data.

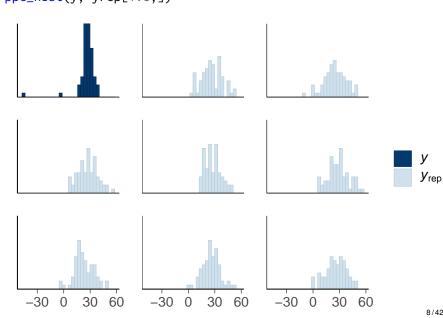
Generate several replicated datasets y<sup>rep</sup>

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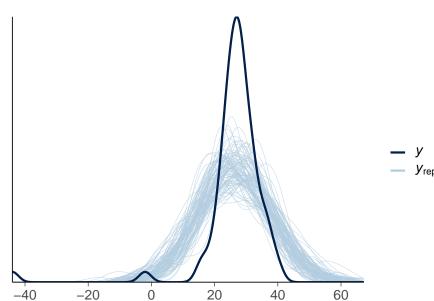
- Generate several replicated datasets y<sup>rep</sup>
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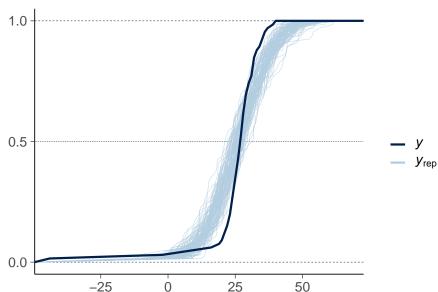
# Posterior predictive checking – bayesplot ppc\_hist(y, yrep[1:8,])



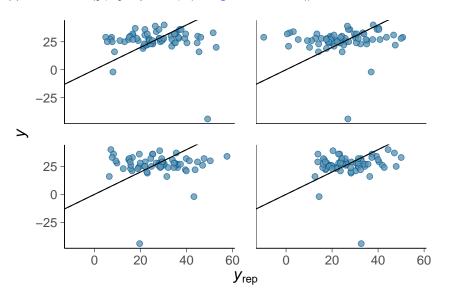
ppc\_dens\_overlay(y, yrep[1:100,])



ppc\_ecdf\_overlay(y, yrep[1:100,])



ppc\_scatter(y, yrep[1:4,]) + geom\_abline()



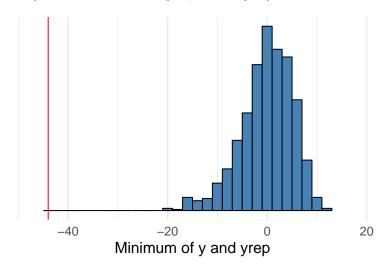
# Posterior predictive checking with test statistic

- Replicated data sets y<sup>rep</sup>
- Test quantity (or discrepancy measure)  $T(y, \theta)$ 
  - summary quantity for the observed data  $T(y, \theta)$
  - summary quantity for a replicated data  $T(y^{rep}, \theta)$
  - can be easier to compare summary quantities than data sets

• Compute test statistic for data  $T(y, \theta) = \min(y)$ 

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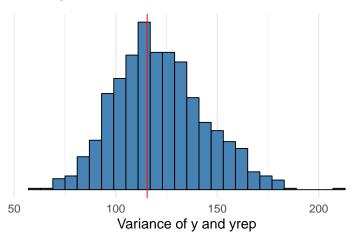
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- Good test statistic is (almost) ancillary (fi: aputunnusluku)
  - ancillary if it depends only on observed data and if its distribution is independent of the parameters of the model

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- Bad test statistic is highly dependent of the parameters
  - e.g. variance for normal model



# Posterior predictive checking

Posterior predictive p-value

$$p = \Pr(T(y^{\text{rep}}, \theta) \ge T(y, \theta)|y)$$
$$= \int \int I_{T(y^{\text{rep}}, \theta) \ge T(y, \theta)} p(y^{\text{rep}}|\theta) p(\theta|y) dy^{\text{rep}} d\theta$$

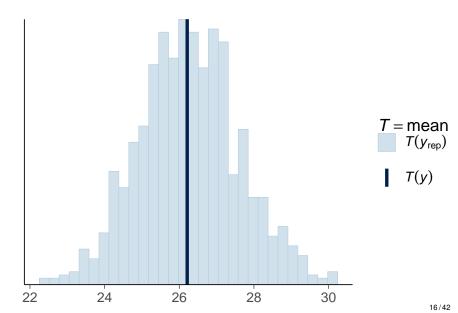
where I is an indicator function

• having  $(y^{\text{rep }(s)}, \theta^{(s)})$  from the posterior predictive distribution, easy to compute

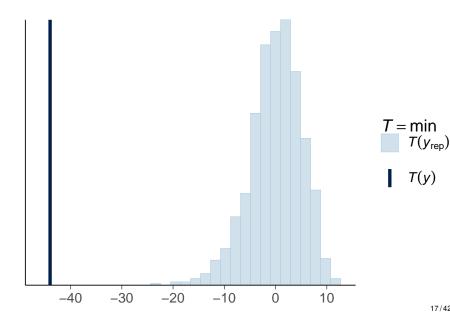
$$T(y^{\text{rep}(s)}, \theta^{(s)}) \ge T(y, \theta^{(s)}), \quad s = 1, \dots, S$$

- Posterior predictive p-value (ppp-value) estimates whether difference between the model and data could arise by chance
- Not commonly used, as
  - not calibrated in case of non-ancillary statistic
  - the distribution of test statistic has more information

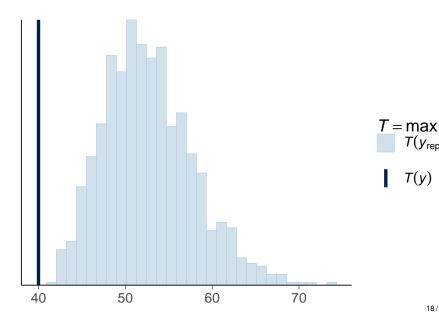
ppc\_stat(y, yrep), the default statistic "mean" is usually bad



ppc\_stat(y, yrep, stat="min")



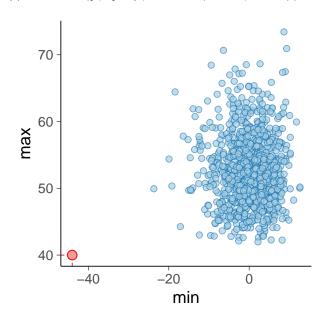
ppc\_stat(y, yrep, stat="max")



 $T(y_{rep})$ 

T(y)

ppc\_stat2d(y, yrep, stat=c("min", "max"))



 $T = (\min, \max)$ 



 $T(y_{\text{rep}})$ 

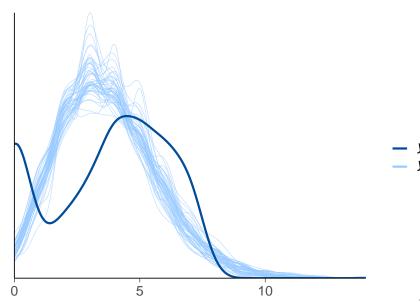
#### Posterior predictive checking - Stan code

demo demos\_rstan/ppc/poisson-ppc.Rmd

```
data {
  int<lower=1> N;
  array[N] int<lower=0> y;
parameters {
  real<lower=0> lambda:
model {
  lambda ~ exponential(0.2);
 v ~ poisson(lambda);
generated quantities {
  real log_lik[N];
  array[N] int y_rep;
  for (n in 1:N) {
    y_rep[n] = poisson_rng(lambda);
    log_lik[n] = poisson_lpmf(y[n] | lambda);
```

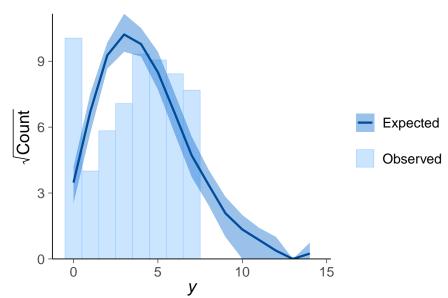
#### PPC for count data - Poisson model

ppc\_dens\_overlay(y, yrep[1:50,])



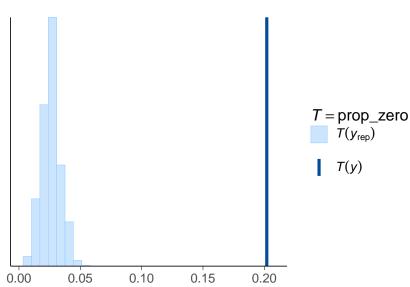
#### PPC for count data - Poisson model

ppc\_rootogram(y, yrep)



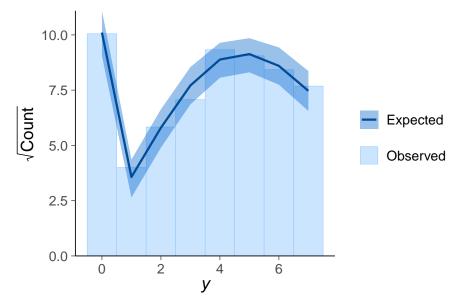
#### PPC for count data - Poisson model

```
prop_zero <- function(x) mean(x == 0)
ppc_stat(y, yrep, stat = "prop_zero")</pre>
```



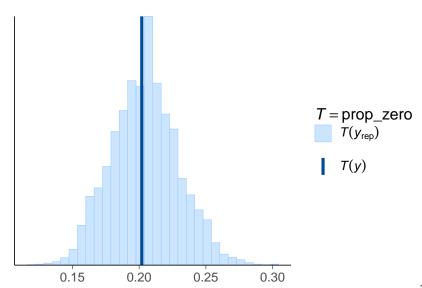
#### PPC for count data - hurdle truncated Poisson model

ppc\_rootogram(y, yrep2)



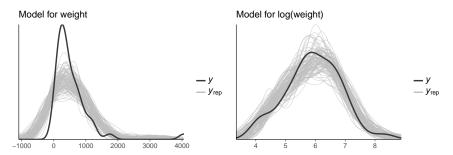
#### PPC for count data - hurdle truncated Poisson model

```
prop_zero <- function(x) mean(x == 0)
ppc_stat(y, yrep2, stat = "prop_zero")</pre>
```



## Posterior predictive checking: Mesquite bushes

#### Positive target: normal vs log-normal model

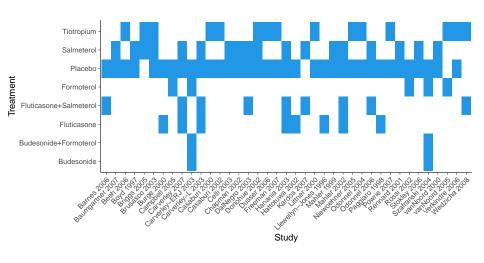


Predicting the yields of mesquite bushes.

Gelman, Hill & Vehtari (2020): Regression and Other Stories, Chapter 11.

#### Meta-analysis

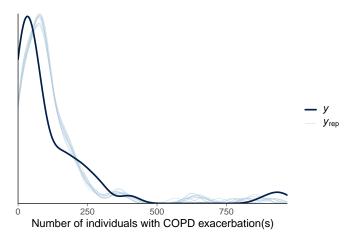
Pharmacologic treatments for chronic obstructive pulmonary disease



## Posterior predictive checking

Pharmacologic treatments for chronic obstructive pulmonary disease

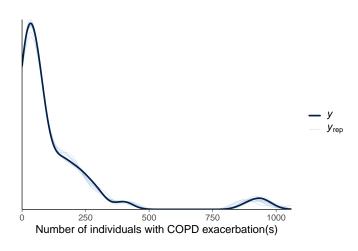
#### Pooled over studies, separate for treatments



## Posterior predictive checking

Pharmacologic treatments for chronic obstructive pulmonary disease

Hirerachical for studies, hierarchical for treatments

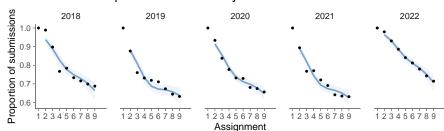


#### Student retention

#### Latent hierarchical linear + spline

```
nstudents | trials(nstudents1) ~
  s(assignment, k=4) + (assignment | year),
  family=binomial()
```

#### Latent functions + posterior uncertainty

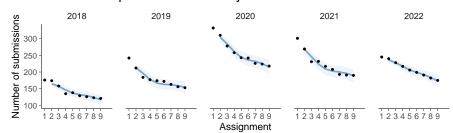


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#### Student retention

1. Latent hierarchical linear model

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```

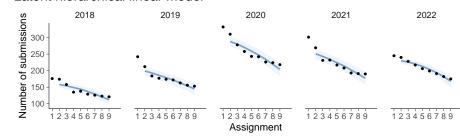
2. Latent spline + hierarchical linear model

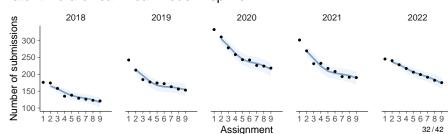
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nstudents | trials(nstudents1) ~
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```

#### Student retention – Posterior predictive distributions

with tidybayes

#### Latent hierarchical linear model

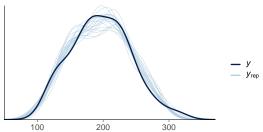


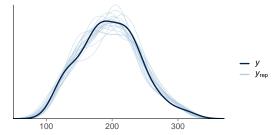


## Student retention – Marginal PPC (brms)

pp\_check(fit, ndraws=100)

#### Latent hierarchical linear model

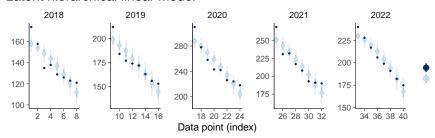


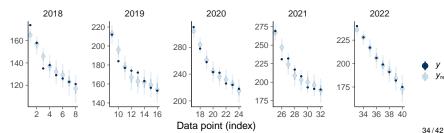


### Student retention – Posterior predictive intervals (brms)

pp\_check(fit, type = "intervals\_grouped", group="year")

#### Latent hierarchical linear model

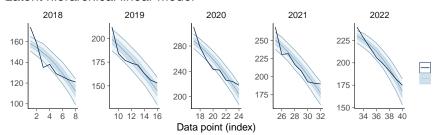


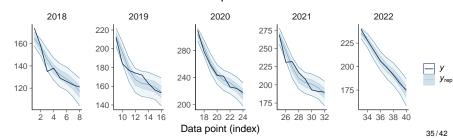


## Student retention – Posterior predictive ribbon (brms)

pp\_check(fit, type = "ribbon\_grouped", group="year")

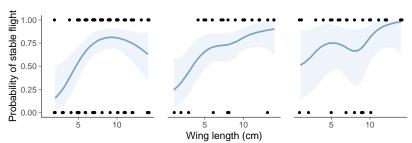
#### Latent hierarchical linear model





## PPC for binary target – Helicopters (brms)

```
stable_flight ~ s(wing_length) + s(wing_length, by = nclips),
family = bernoulli()
```



### PPC for binary target — Helicopters (brms) stable\_flight ~ s(wing\_length) + s(wing\_length, by = nclips), family = bernoulli() Probability of stable flight 1.00 0.75 0.50 0.25 5 10 10 Wing length (cm) pp\_check(fit, ndraws=20) Continuous kernel density estimate for binary target doesn't make $y_{rep}$ sense!

0.00

0.25

0.50

0.75

1.00

#### PPC for binary target — Helicopters (brms) stable\_flight ~ s(wing\_length) + s(wing\_length, by = nclips), family = bernoulli() Probability of stable flight 1.00 0.75 0.50 0.25 5 10 10 Wing length (cm) pp\_check(fit, type="bars") 120 1 Bar plot showing 90 marginal probabilities Count $y_{rep}$ for binary data is al-60 most always useless! 30

-0.5

0.0

0.5

1.0

1.5

#### PPC for binary target - Helicopters (brms) stable\_flight ~ s(wing\_length) + s(wing\_length, by = nclips), family = bernoulli() Probability of stable flight .00 0.75 0.50 0.25 10 10 5 10 Wing length (cm) pp\_check(fit, type="bars\_grouped") 0 60 Count $y_{rep}$ 40 20

0.0 0.5 1.0

-0.5

1.5 -0.5

0.0 0.5 1.0

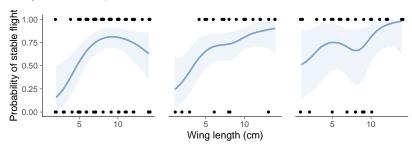
1.5 -0.5

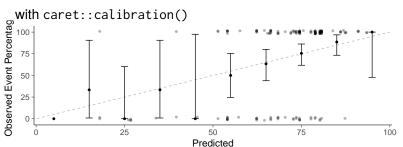
0.5 1.0

0.0

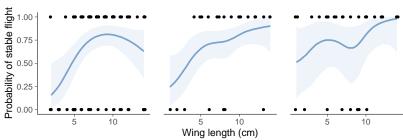
## PPC for binary target - Helicopters (brms) stable\_flight ~ s(wing\_length) + s(wing\_length, by = nclips),

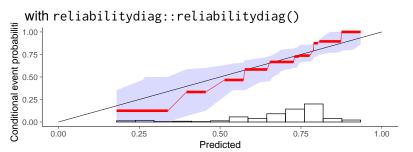
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```





# $\begin{array}{l} \text{PPC for binary target} - \text{Helicopters (brms)} \\ \text{stable\_flight} & \text{s(wing\_length)} + \text{s(wing\_length, by = nclips),} \\ \text{family = bernoulli()} \end{array}$

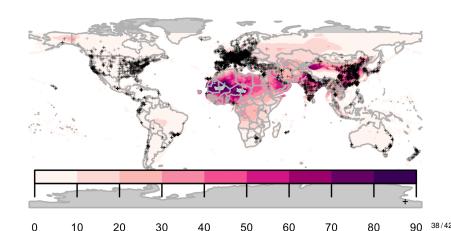




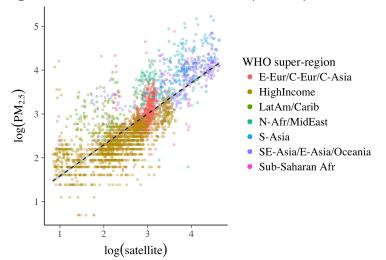
Use this for binary data!

- Example from Gabry, Simpson, Vehtari, Betancourt, and Gelman (2019). Visualization in Bayesian workflow. https://doi.org/10.1111/rssa.12378
- Estimation of human exposure to air pollution from particulate matter measuring less than 2.5 microns in diameter (PM<sub>2.5</sub>)
  - Exposure to PM<sub>2.5</sub> is linked to a number of poor health outcomes and a recent report estimated that PM<sub>2.5</sub> is responsible for three million deaths worldwide each year (Shaddick et al., 2017)
  - In order to estimate the public health effect of ambient  $PM_{2.5}$ , we need a good estimate of the  $PM_{2.5}$  concentration at the same spatial resolution as our population estimates.

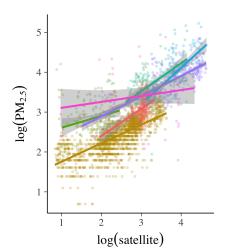
- Direct measurements of PM 2.5 from ground monitors at 2980 locations
- High-resolution satellite data of aerosol optical depth



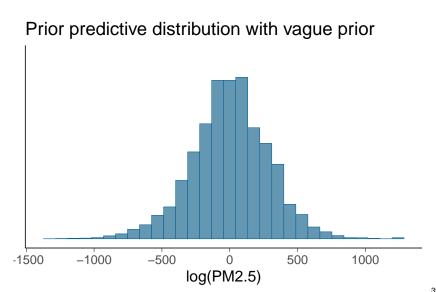
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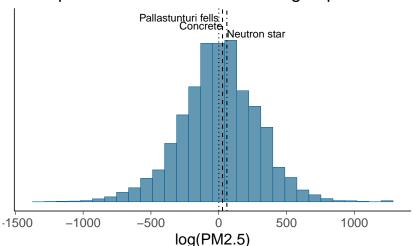


Prior predictive checking



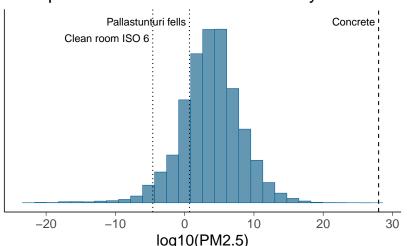
Prior predictive checking

### Prior predictive distribution with vague prior



Prior predictive checking

#### Prior predictive distribution with weakly informative



### Further reading and examples

- Gabry, Simpson, Vehtari, Betancourt, and Gelman (2019).
   Visualization in Bayesian workflow.
   https://doi.org/10.1111/rssa.12378.
- Säilynoja, Johnson, Martin, and Vehtari (2025).
   Recommendations for visual predictive checks in Bayesian workflow.
  - https://teemusailynoja.github.io/visual-predictive-checks/
- Graphical posterior predictive checks using the bayesplot package http://mc-stan.org/bayesplot/articles/graphical-ppcs.html
- brms demos https://avehtari.github.io/BDA\_R\_demos/demos\_ rstan/brms\_demo.html

 How much different choices in model structure and priors affect the results

- How much different choices in model structure and priors affect the results
  - test different models and priors
     priorsense and adjustr packages use importance sampling
     for faster prior sensitivity analysis

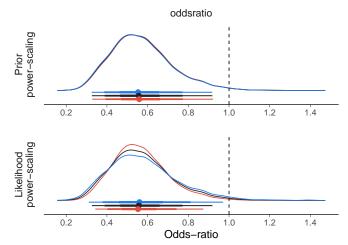
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     priorsense and adjustr packages use importance sampling
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    - Kallioinen, Paananen, Bürkner, and Vehtari (2024). Detecting and diagnosing prior and likelihood sensitivity with power-scaling. https://doi.org/10.1007/s11222-023-10366-5

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  - alternatively combine different models to one model
    - e.g. hierarchical model instead of separate and pooled
    - e.g. *t* distribution contains Gaussian as a special case
  - robust models are good for testing sensitivity to "outliers"
    - e.g. t instead of Gaussian

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  - robust models are good for testing sensitivity to "outliers"
    - e.g. t instead of Gaussian
- Compare sensitivity of essential inference quantities
  - extreme quantiles are more sensitive than means and medians
  - extrapolation is more sensitive than interpolation

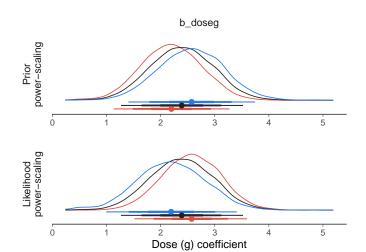
## priorsense — prior and likelihood sensitivity analysis

- Power-scale prior and likelihood separately as  $p(\theta)^{\alpha}$  and  $p(y|\theta)^{\alpha}$
- Beta blockers randomized control-treatment experiment
  - no prior sensitivity
  - likelihood is informative



## priorsense — prior and likelihood sensitivity analysis

- Power-scale prior and likelihood separately as  $p(\theta)^{\alpha}$  and  $p(y|\theta)^{\alpha}$
- Sorafenib Toxicity Binomial model meta analysis
  - prior-data conflict



## priorsense — prior and likelihood sensitivity analysis

- Power-scale prior and likelihood separately as  $p(\theta)^{\alpha}$  and  $p(y|\theta)^{\alpha}$
- Sorafenib Toxicity Binomial model meta analysis
  - prior-data conflict
  - due to accidentally too narrow prior

