

Binomial Models.

3

Let $i = 1, \dots, n$ reference a particular row in our standings dataset (e.g. LA Rams in 2023)

w_i = # of wins for that team (in that season)

A simple model for w_i

$w_i \sim \text{Binomial}(n_i, p_i)$

$n_i = 16$ or 17 depending on the season

p_i = probability team i wins each of its n_i games.

Problems with this model:

- * don't know p_i (need to estimate it)

- * Assumes $P(\text{team } i \text{ beats team } j) = p_i$
not depending on team j 's probability

- * Quality of team i same all season.

But it's a starting point

How do we get p_i ? What should it depend on?

an idea: p_i depends on prop. of wins last year q_i

Try this: $p_i = b_0 + b_1 q_i$

problem is that p_i should be between 0 and 1

This model does not guarantee that.

Thinking ahead: to estimate any parameters, least squares is probably not appropriate because responses are not normal.

Generalized Linear Model:

1. Response distribution

$$Y_i \sim \text{Distribution}(\theta_i)$$

~~response~~ specify a good model for the response.

2. Link function

$$\theta_i = f(a_i)$$

$$a_i = f^{-1}(\theta_i)$$

ensure parameter θ_i falls within valid range. Link function should be invertible.

3. Linear form

$$a_i = b_0 + b_1 x_{i1} + b_2 x_{i2} + \dots$$

Relate parameter to other variables.

For the binomial model:

1. $W_i \overset{\text{ind}}{\sim} \text{Binomial}(n_i, p_i)$

2. $p_i = \frac{\exp(a_i)}{1 + \exp(a_i)} := \text{expit}(a_i)$

$a_i = \log\left(\frac{p_i}{1-p_i}\right) := \text{logit}(p_i)$

3. $a_i = b_0 + b_1 q_i$ or $a_i = b_0 + b_1 \text{logit}(q_i)$ No epsilons!
transforms are allowed.

Looking ahead. For hurricanes:

1. $Y_i \overset{\text{ind}}{\sim} \text{Poisson}(\mu_i)$ (hurricane count in year i)

2. $\mu_i = \exp(a_i)$
 $a_i = \log(\mu_i)$

3. $a_i = b_0 + b_1 x_i$ $x_i = \text{sea surface temp.}$
No epsilons!

Go over the code for NFL data:

Actual wins is more variable than that predicted by the binomial GLM

Could be due to various factors:

Need to inject more variability into the model

Generalized linear mixed model:

1. $W_i \overset{\text{ind}}{\sim} \text{Binomial}(n_i, p_i)$
2. $p_i = \exp(a_i) / (1 + \exp(a_i))$
3. $a_i = b_0 + b_1 \text{logit}(q_i) + \epsilon_i$
 $\epsilon_i \overset{\text{ind}}{\sim} N(0, \sigma^2)$

Basically, we are saying that the proportion of wins is that predicted by prior season, plus ~~an~~ a random effect, which could kick it in either direction.

Beta Binomial Model:

1. $W_i \stackrel{\text{ind}}{\sim} \text{Binomial}(n_i, p_i)$
2. $p_i \sim \text{Beta}(\alpha_i, \beta_i)$

α_i and β_i (or a transformation thereof) may depend on other variables, such as z_i

$$E(p_i) = \frac{\alpha_i}{\alpha_i + \beta_i} = \frac{\cancel{b_0 + b_1 \text{logit}(z_i)}}{\cancel{b_0 + b_1 \text{logit}(z_i)} + 1} = \frac{\exp(a_i)}{1 + \exp(a_i)}$$

$$\phi_i = \frac{1}{\alpha_i + \beta_i + 1} = c_0$$

3. ~~ϕ_i~~ $a_i = b_0 + b_1 \text{logit}(z_i)$

Same ideas as gen. lin. mixed model:

p_i is drawn from a distribution centered on $\text{expit}(b_0 + b_1 \text{logit}(z_i))$

Just a different distribution.

$$E(p) = \frac{\alpha}{\alpha + \beta} \quad \phi = \frac{1}{\alpha + \beta + 1}$$

solve for α and β

$$\frac{1}{\phi} = \alpha + \beta + 1$$

$$\frac{\alpha}{E(p)} = \frac{\alpha + \beta}{\cancel{\alpha + \beta}} \rightarrow \frac{1}{\phi} - 1 = \alpha + \beta$$

$$\frac{\alpha}{E(p)} = \frac{1}{\phi} - 1 \quad \left(\frac{1}{\phi} - 1\right) - E(p)\left(\frac{1}{\phi} - 1\right) = \beta$$

$$\alpha = E(p)\left(\frac{1}{\phi} - 1\right) \quad (1 - E(p))\left(\frac{1}{\phi} - 1\right) = \beta$$

Beta binomial pmf:

$$P(X=k) = \binom{n}{k} \frac{B(k+\alpha, n-k+\beta)}{B(\alpha, \beta)}$$

Maximum Likelihood Estimation

likelihood is the joint density function, evaluated at the data, viewed as a fun. of parameters.

For binomial GLM: data $w_1, \dots, w_n, q_1, \dots, q_n$

$$\begin{aligned} f(w) &= \prod_{i=1}^n \binom{n_i}{w_i} p_i^{w_i} (1-p_i)^{n_i-w_i} \\ &= \prod_{i=1}^n \binom{n_i}{w_i} \text{expit}(a_i)^{w_i} (1 - \text{expit}(a_i))^{n_i-w_i} \\ &= \prod_{i=1}^n \binom{n_i}{w_i} \text{expit}(b_0 + b_1 \text{logit}(q_i))^{w_i} (1 - \text{expit}(b_0 + b_1 \text{logit}(q_i)))^{n_i-w_i} \end{aligned}$$

complicated!

Generally, there won't be a formula for \hat{b}_0, \hat{b}_1 like in regression (i.e. $(X^T X)^{-1} X^T y$)

But we can use numerical procedures to maximize the logarithm of the likelihood.

- Fisher scoring is the preferred method.